

MAIN

The Oxford Paranoia Defence Behaviours Questionnaire (O-PDQ): assessing paranoia-related safety-seeking behaviours

Sinead Lambe¹ , Sophie Mulhall², Jessica Bird³, Kira Williams^{1,4}, Joanna Mitchell^{1,4} , Mollie Roddan¹, Glory Sokunle^{1,2}, Laina Rosebrock^{1,4} , Daniel Freeman^{1,4} and Felicity Waite^{1,4}

¹Department of Experimental Psychology, University of Oxford, Oxford, UK, ²South London and Maudsley NHS Foundation Trust, London, UK, ³Department of Child and Adolescent Psychiatry, IoPPN, King's College London, London, UK and ⁴Oxford Health NHS Foundation Trust, Oxford, UK

Corresponding author: Sinead Lambe; Email: sinead.lambe@psy.ox.ac.uk

(Received 13 March 2024; revised 27 May 2025; accepted 29 May 2025; first published online 27 August 2025)

Abstract

Background: Defence behaviours – actions carried out to reduce perceived threat – are an important maintenance factor for persecutory delusions. Avoidance of feared situations and subtle in-situation behaviours reduce opportunities for new learning and are erroneously credited for the non-occurrence of harm; hence inaccurate fears are maintained. In contrast, exposure to feared situations whilst dropping defence behaviours – a key technique of cognitive therapy for paranoia – allows the discovery of new information concerning safety, thereby reducing persecutory delusions.

Aim: We aimed to develop for use in research and clinical practice a self-report assessment of paranoia-related defence behaviours.

Method: A 64-item pool was developed from interviews with 106 patients with persecutory delusions, and completed by 53 patients with persecutory delusions, 592 people with elevated paranoia, and 2108 people with low paranoia. Exploratory and confirmatory factor analyses were used to derive the measure. Reliability and validity were assessed.

Results: Two scales were developed: a 12-item avoidance scale and a 20-item in-situation defences scale. The avoidance scale had three factors (indoor spaces, outdoor spaces, and interactions) with an excellent model fit (CFI = 0.98, TLI = 0.97, RMSEA = 0.04, SRMR = 0.027). The in-situation defences scale had a 5-factor model (maintaining safety at home, mitigating risk, staying vigilant, preparing for escape, and keeping a low profile) with a good fit (CFI = 0.95, TLI = 0.94, RMSEA = 0.046, SRMR = 0.039). Both scales demonstrated good internal reliability, test–retest reliability, and construct validity.

Conclusions: The Oxford Paranoia Defence Behaviours Questionnaire is a psychometrically robust scale that can assess a key factor in the maintenance of persecutory delusions.

Keywords: assessment tool; paranoia; psychosis; questionnaire development; safety-seeking behaviours; schizophrenia

Introduction

Cognitive theory conceptualises persecutory delusions as inaccurate threat beliefs about harm from other people (Freeman, 2016; Freeman *et al.*, 2025). The perceived threat leads to anxiety and an attempt to obtain safety through preventative action. Avoidance is the most common preventative action (e.g. avoiding places where it is difficult to escape). When this is not possible,

within-situation strategies such as keeping a low profile or rushing through tasks, are used to reduce risk. In the face of a true threat, safety-seeking behaviours would be adaptive, but when fears are a misperception, as is the case with persecutory delusions, these behaviours prevent the receipt and processing of contradictory information. The non-occurrence of harm is misattributed to the use of the safety-seeking behaviours and thus the threat belief is maintained. In our paranoia treatment we have called these unhelpful safety-seeking behaviours ‘defences’, because the term provides a helpful metaphor for the blocking of new learning and we wish to avoid confusion when truly discovering safety is a clear goal of cognitive therapy, for example, the Feeling Safe programme, a highly efficacious cognitive therapy for persecutory delusions (Freeman *et al.*, 2021a). In this paper we report the development of the first self-report assessment of paranoia-related defences.

The key role of safety-seeking behaviours in the maintenance of unfounded threat beliefs was first conceptualised thirty years ago in relation to panic disorder (Salkovskis, 1991). Several experimental manipulation studies across different anxiety disorders have demonstrated that dropping safety-seeking behaviours during a brief exposure leads to greater reductions in the threat belief and anxiety compared with exposure alone (McManus *et al.*, 2009; Salkovskis *et al.*, 1999; Salkovskis *et al.*, 2003). Reducing the use of safety-seeking behaviours to allow new learning has since become a central mechanism in cognitive treatments for panic disorder and agoraphobia (Clark and Salkovskis, 2009), post-traumatic stress disorder (Ehlers and Clark, 2000), obsessive compulsive disorder (Salkovskis, 1999) and social anxiety (Clark and Wells, 1995). The concept was subsequently applied to persecutory delusions (Freeman *et al.*, 2001; Freeman *et al.*, 2023). The widespread use of such behaviours by patients with persecutory delusions was established (Freeman *et al.*, 2007), and an experimental manipulation study showed that dropping defences leads to large reductions in persecutory delusions (Freeman *et al.*, 2016).

The most widely used measure of defences in relation to persecutory delusions is a semi-structured interview, the Safety Behaviours Questionnaire – Persecutory Beliefs (SBQ) (Freeman *et al.*, 2001). The SBQ assesses defence behaviours across several categories, including avoidance, in-situation, escape, compliance, help-seeking, and aggression. An interview method offers depth, captures idiosyncratic behaviours, and judgement of whether a response is a safety-seeking behaviour or a more adaptive coping strategy. However, interviewer assessments can be more burdensome to administer and are more likely to be influenced by the skill of the interviewer. In contrast, self-report measures can be quicker, easier to use, and ensure a greater degree of standardisation. They are more likely to be used in larger-scale research. Self-report measures of safety-seeking behaviours are widely used in the research and treatment of anxiety disorders, with specific measures for obsessive compulsive disorders (Foa *et al.*, 2002), panic disorder (Clark and Salkovskis, 2009), agoraphobia (Chambless *et al.*, 1985; Chambless *et al.*, 2011), social anxiety (Clark, 2005; Cuming *et al.*, 2009), and post-traumatic stress disorder (Dunmore *et al.*, 2001). Tully and colleagues (2017) developed a scale of common cognitive and behavioural responses to psychosis, of which a 5-item subscale measured defence behaviours. To the best of our knowledge there is no self-report measure of paranoia-related defence behaviours. This study aimed to address this gap by developing such a measure. To ensure the measure was grounded in patient experience we used a comprehensive item pool generated from patient interviews using the SBQ. The final measure was developed through classical test theory analysis in a separate, larger sample.

Method

Participants

We recruited a total of 2753 participants across the full range of severity of paranoia. There were three groups: patients with persecutory delusions who were attending mental health secondary care services, individuals in the general population with elevated paranoia, and individuals in the

Table 1. Demographic and clinical characteristics

	General population (low paranoia) (<i>n</i> = 2108)	General population (elevated paranoia) (<i>n</i> = 592)	Patient group (<i>n</i> = 53)
Age	46.2 (15.7)	43.0 (14.7)	41 (13.6)
Gender	<i>n</i> (%)	<i>n</i> (%)	<i>n</i> (%)
Male	505 (24.0)	180 (30.4)	33 (62.3)
Female	1566 (74.3)	398 (67.2)	20 (37.7)
Non-binary/prefer not to say	37 (1.1)	14 (2.3)	0(0)
Ethnicity			
White (any background)	1991 (94.4)	558 (94.3)	44 (83)
Asian	37 (1.8)	8 (1.4)	2 (3.8)
Black	10 (0.5)	4 (0.7)	4 (7.6)
Mixed or multiple ethnic groups	39 (1.9)	8 (1.4)	3 (5.7)
Other/prefer not to say	31 (1.2)	14 (2.4)	0 (0)
Diagnosis			
Schizophrenia	—	—	22 (41.5)
Schizoaffective disorder	—	—	13 (24.5)
Delusional disorder	—	—	4 (7.5)
Psychosis not otherwise specified	—	—	14 (26.5)

general populations with low paranoia. See Table 1 for a summary of demographic and clinical information about the participants.

Patients with persecutory delusions were participants in the THRIVE Trial (Freeman *et al.*, 2023), a randomised controlled trial of an automated VR therapy. The patients completed the item pool for the new scale at the end of treatment follow-up assessment (8 weeks). The main inclusion criteria for the THRIVE trial were (i) aged 16 years or older, (ii) a diagnosis of schizophrenia or related conditions (e.g. non-affective psychosis), (iii) persistent (at least 3 months) persecutory delusion (as defined by Freeman and Garety, 2000) held with at least 50% conviction, and (iv) who verbally reported feeling threatened when with other people. Participants from the general population were recruited online using social media advertisements. The inclusion criteria were: (i) aged 18 years or older and (ii) resident in the UK. To evaluate test–retest reliability, 343 participants from the general population repeated the questionnaire after a fortnight. The general population group was divided into those with elevated paranoia (*n* = 592) and those with low paranoia (*n* = 2108) using an established cut-off (6 or above) on the Revised Green *et al.* Paranoid Thoughts Scale – Part B (R-GPTS-B) (Freeman *et al.*, 2021b).

Assessments

Item pool for the Oxford Paranoia Defence Behaviours Questionnaire (O-PDQ)

The initial O-PDQ item pool was generated from an analysis of SBQ interview data from the first 106 participants in the Feeling Safe trial (Freeman *et al.*, 2021a). The interviews provided 777 unique descriptions of defence behaviours, which were analysed using inductive content analysis (ICA; Elo and Kyngas, 2007). Behaviours with the same meaning were grouped together (e.g. ‘Rushing through tasks’ and ‘I’d do things as quickly as I can’; see Supplementary material, S2). The goal was to produce a potentially exhaustive list of avoidance and in-situation defence behaviours used in response to persecutory beliefs. ICA was carried out independently by two researchers (S.M. and S.L.) and discrepancies were discussed with another member of the research team (F.W.). ICA identified 22 avoided situations and 48 different in-situation defence behaviours. These were reviewed by five people with lived experience of psychosis for readability, ease of understanding, and completeness. The item pool was also reviewed by our team (F.W., J.B. and D.F.) to ensure face validity and completeness. Six in-situation behaviours (e.g. praying, listening to music) were excluded as they could be helpful coping behaviours rather than defence behaviours (e.g. one may be

praying for the strength to face their fears whilst someone else may be praying so god will stop others harming them), which would be difficult to distinguish in a self-report assessment.

The final O-PDQ item pool had two separate scales: *Avoidance* and *In-situation defences*. These were introduced as strategies that may be used when people feel under threat of harm from others. The avoidance scale item pool included 22 everyday situations. Participants were asked to rate on a scale from 0 (never) to 3 (always) how much over the past two weeks they avoided each situation to prevent themselves being harmed by others. Higher scores indicated greater avoidance. The in-situation defences scale item pool included 42 in-situation behaviours. Participants were asked to rate on a scale from 0 (never) to 3 (always) how often over the past two weeks they had used each of these strategies to try to prevent themselves being harmed by others. Higher scores indicated greater use of in-situation defence behaviours. See Supplementary material (S1) for initial item pool.

Safety Behaviours Questionnaire – Persecutory Beliefs (SBQ)

The SBQ (Freeman *et al.*, 2001) is a semi-structured interview that assesses defence behaviours used in response to persecutory beliefs. It is administered by an interviewer who assesses if a behaviour is being used to prevent harm by others. It assesses seven categories of behaviour: (a) avoidance of situations considered dangerous, (b) in-situation behaviours carried out when threat feels imminent, (c) escape from situations where danger is deemed imminent, (d) compliance with the demands of wishes of the persecutor, (e) getting help from others to reduce the threat, (f) aggression in response to threat and (g) delusional actions which are behaviours aimed at reducing the threat but where there is no logical relation to the achievement of this aim. After a safety behaviour has been elicited, the participant is asked to rate its frequency over the last month on a 4-point scale (1, behaviour definitely occurred on at least one occasion; 2, occurred more than once but not frequently, e.g. not more than five or more times; 3, occurred frequently, e.g. at least five times; 4, present more or less continuously, at least every day). Scores are then summed for each of the subscales. A distinction is made between avoidance, which can be considered a negative behaviour (SBQ avoidance score) and a combined score of the other six subscales, which reflects active attempts to reduce the threat, i.e. a positive behaviour (SBQ in situation defences).

Revised Green et al. Paranoid Thoughts Scale (R-GPTS)

The R-GPTS (Freeman *et al.*, 2021b) is a self-report measure assessing paranoid thinking. In the THRIVE trial it was used to assess paranoia over the last fortnight. The R-GPTS contains two separate scales assessing ideas of reference, R-GPTS-A (8 items), and ideas of persecution, R-GPTS-B (10 items). Items are rated on scale from 0 (not at all) to 4 (totally), with higher scores indicating higher levels of paranoid thinking. The score ranges for the ideas of persecution scale are: average (0–5); elevated (6–10); moderately severe (11–17); severe (18–27); and very severe (28+). The Cronbach's alpha was 0.93 ($n = 1907$) for ideas of reference and was 0.95 ($n = 1859$) for ideas of persecution.

Measure of Common Responses to Psychosis (MCRP)

The MCRP (Tully *et al.*, 2017) is a 15-item self-report measure of cognitive and behavioural responses to psychosis. It has three subscales, each containing five items: conscious self-regulation, threat monitoring and avoidance, and social control and reassurance seeking. Respondents rate how much they engage in each behaviour in response to psychosis on a scale from 1 (never) to 4 (always). Higher scores indicate greater use of behavioural responses to psychosis. Cronbach's alpha was 0.85 ($n = 1774$) for conscious self-regulation, 0.83 ($n = 1774$) for threat monitoring and avoidance, and 0.77 ($n = 1780$) for social control and reassurance seeking.

Agoraphobia Mobility Inventory (AMI)

The AMI (Chambless *et al.*, 1985) consists of 26 items assessing avoidance of situations due to anxiety (i.e. agoraphobia). Items ask about avoidance of places (e.g. theatres, department stores,

museums), transport (e.g. aeroplanes, buses, cars), specific situations (e.g. being home alone, standing in queues), and spaces (e.g. enclosed spaces, high places, open spaces). Items are coded on a 1 (never avoid) to 5 (always avoid) scale. There is also an option to select 'not applicable'; as a result, mean item scores are calculated. Higher mean scores indicate higher levels of agoraphobia. The Cronbach's alpha for the AMI was 0.96 ($n = 1777$).

Brief Fear of Negative Evaluation Scale (BFNE)

The BFNE (Leary, 1983) is a 12-item self-report scale measuring fears of negative evaluation, especially in social settings. Items are rated on a 5-point scale ranging from 0 (not at all characteristic of me) to 4 (extremely characteristic of me). Higher scores indicate a greater fear of being negatively evaluated by others. The Cronbach's alpha for the BFNE was 0.74 ($n = 1719$).

Patient Health Questionnaire-9 (PHQ-9)

The PHQ-9 (Kroenke *et al.*, 2001) consists of nine items assessing symptoms of depression over the past two weeks. Items are rated on a 0 (not at all) to 3 (nearly every day) scale. Higher scores indicate higher levels of depression. The Cronbach's alpha for the PHQ-9 was 0.93 ($n = 1731$).

Generalised Anxiety Disorder-7 (GAD-7)

The GAD-7 (Spitzer *et al.*, 2006) is a 7-item scale assessing symptoms of generalised anxiety over the past two weeks. Response options range from 0 (not at all) to 3 (nearly every day). Higher scores indicate higher levels of generalised anxiety. The Cronbach's alpha for the GAD-7 was 0.94 ($n = 1920$).

Oxford Agoraphobic Avoidance Scale (O-AS)

The O-AS (Lambe *et al.*, 2023) is an 8-item self-report measure of agoraphobic avoidance of everyday activities. Respondents are asked to rate whether they could complete the activity: 0 (yes – I can do this now) or 1 (no – I would get too anxious); and how anxious they would feel doing the task, ranging from 0 (no distress) to 10 (extreme distress). Higher avoidance and distress scores indicate higher levels of agoraphobia symptoms. The Cronbach's alpha for O-AS avoidance was 0.83 ($n = 1980$), and 0.94 ($n = 1614$) for O-AS distress.

Procedure

All participants completed the O-PDQ item pool and the R-GPTS. Patients within the THRIVE trial also completed SBQ administered by a research assistant. Participants in the general population also completed the MCRP, AMI, PHQ-9, GAD-7 and BFNE.

Statistical analysis

All analyses were conducted in R, version 3.6.1 (R Core Team, 2013). Only participants with complete O-PDQ item pool responses were included. Factor analysis was appropriate for the derivation sample for both the avoidance scale and in-situation defences scales, as Bartlett's test of sphericity was significant (avoidance: $\chi^2 = 5794.0$, d.f. = 231, $p < 0.001$; in-situation defences: $\chi^2 = 9213.007$, d.f. = 741, $p < 0.001$) and the Kaiser-Meyer-Olkin (KMO) test of sampling adequacy was excellent (avoidance: KMO = 0.94; in-situation defences: KMO = 0.94).

Deriving the O-PDQ

A combined paranoia sample was formed with the participants from the general population with elevated paranoia and the patients with persecutory delusions ($n = 645$). A randomly selected 80%

of this combined paranoia group was used as a derivation sample ($n = 516$) to extract items. This ensured that items selection and model development was conducted with the intended users (i.e. those with paranoia) of the measure. A confirmatory sample ($n = 2237$) was then formed from the remaining 20% of the combined paranoia group and the participants from the general population with low paranoia.

Exploratory factor analysis (EFA) was conducted with the derivation sample ($n = 516$) in the 'psych' package (Revelle, 2020). Items that were highly correlated with other items ($r \geq 0.8$) were deleted prior to the EFA to avoid issues of multi-collinearity. EFA was estimated using principal axis factoring to account for non-normality in the data (Costello and Osborne, 2005) and oblique rotation. Parallel analysis and inspection of scree plots were used to determine the number of factors to extract. Items were deleted if they were theoretically inconsistent or redundant, did not load onto any factor, or had cross-loadings above 0.30 on multiple factors.

The model fit of the final set of items was examined using confirmatory factor analysis (CFA) in the confirmation sample ($n = 2237$). CFA was conducted in the 'lavaan' package (Rosseel, 2012) using the robust maximum likelihood (MLR) estimator due to non-normality in the data. Good model fit was determined using recommended thresholds of 0.90 (good) and 0.95 (excellent) on the comparative fit index (CFI) and the Tucker-Lewis index (TLI), and < 0.10 and < 0.06 on the root mean square error of approximation (RMSEA) and the standardised root mean square residual (SRMR), respectively (Bentler and Bonett, 1980).

Test properties

The internal reliability of the avoidance scale and the in-situation defences scale was determined using Cronbach's alpha. To examine test-retest reliability, a two-way, mixed effects, absolute agreement interclass correlation coefficient (ICC) was conducted. Validity was examined using simple correlations. Construct validity (convergent and divergent) was assessed in several ways. Convergent validity was tested using the MCRP threat monitoring and avoidance scale, as this scale measures similar defence behaviours in psychosis. The avoidance scale was validated against established measures of avoidance – the AMI and O-AS avoidance scale – while the in-situation defence behaviours scale was examined in relation to the O-AS distress score, as higher situational distress would likely be associated with increased use of in-situation defence behaviours. In the patient group the O-PDQ was compared with the gold standard for measuring defence behaviours in paranoia, the SBQ. The avoidance scale was compared with the number and frequency of avoided places and the distress scale was compared with both the number and frequency of defence behaviours identified by the SBQ. Construct validity was further examined by testing the O-PDQ's relationship with paranoia severity using the R-GPTS ideas of reference and persecution scales, as the cognitive model suggests safety behaviours are a response to paranoid thinking.

An analysis of variance (ANOVA) was used to assess differences in O-PDQ scores across the different groups (general population-low paranoia; general population-elevated paranoia, and patients with persecutory delusions). Discriminant validity was tested using the MCRP seeking social support/reassurance scale and self-regulation scale to test whether the O-PDQ differentiates between defence behaviours and other behavioural responses to psychosis. The BFNE was used to demonstrate that the O-PDQ captures paranoia-specific responses rather than general social anxiety-related behaviours. Relationships with GAD-7 and PHQ-9 were examined as anxiety and depression frequently co-occur with paranoia and negative affect in relation to paranoia often drives the use of defence behaviours (Freeman *et al.*, 2007). The ability of the O-PDQ to detect defence behaviours was assessed against the SBQ by descriptively comparing both the number of defence behaviours endorsed on the in-situation defence behaviour scale with the number of defence behaviours identified through interview on the SBQ. The number of places avoided identified on the O-PDQ was also compared with the number of places avoided in the avoidance section of the SBQ.

Results

Deriving the O-PDQ

Avoidance

During EFA, 10 of the original 22 avoidance items were removed, including two items with low communalities (<0.3) (*Staying at home alone*; *Staying at home with others*), two items with extremely high loadings (>0.96) (*Busy places*; *Restaurants*), three items that did not load on any factor (>0.3) (*Staying at home alone*; *Enclosed spaces*; *Unfamiliar places*), one item that was cross-loading (*Meeting people or social gatherings*), and two items that were deleted for theoretical reasons or redundant content (*Supermarkets*; *Being far from home*). With the remaining 12 items, a three-factor structure consisting of indoor public spaces (5 items), outdoor spaces (4 items), and interactions (3 items) explained 51% of the variance (see Table 2 for factor loadings). Correlations

Table 2. EFA factor loadings for O-PDQ avoidance and in-situation defences scale

		EFA loadings	CFA loadings	CFA loadings
Factor	Items	Derivation group	Derivation group	Confirmatory group
Avoidance				
Indoor spaces	My local shop	0.777	0.750	0.742
	Shopping centres	0.814	0.784	0.811
	Using public transport (e.g. bus, train)	0.451	0.726	0.731
	Cafés	0.649	0.720	0.813
	Gyms	0.565	0.647	0.709
Outdoor spaces	Walking on the street	0.786	0.759	0.813
	Parks	0.728	0.768	0.796
	Open spaces	0.695	0.744	0.754
Interactions	People in authority (e.g. the police)	0.685	0.693	0.700
	My neighbours	0.604	0.671	0.708
	GP surgery or health centre	0.690	0.668	0.721
	My workplace or place of education	0.472	0.553	0.667
Defences				
Maintaining safety at home	I keep the curtains closed so I can't be watched	0.562	0.658	0.627
	I repeatedly check all the locks and windows	0.542	0.630	0.614
	I check if the items in my home have been interfered with	0.547	0.580	0.635
	I stay away from the windows so I can't be watched	0.727	0.754	0.728
Mitigating risk	I only go out at certain times of the day that I think are safer	0.788	0.739	0.779
	I carefully plan my routes to avoid danger	0.683	0.859	0.864
	I keep someone with me for protection	0.409	0.540	0.533
Preparing for escape	I have an escape plan ready	0.813	0.731	0.769
	I wear clothes that I can easily run away in	0.597	0.725	0.656
Staying vigilant	I plan for the dangers that might exist	0.418	0.798	0.808
	I check for smells that might indicate danger	0.424	0.574	0.723
	I position myself so that I have a clear view of everyone	0.467	0.727	0.736
	I watch out for signs of danger	0.746	0.782	0.807
Keeping a low profile	I listen out for any sounds that might indicate danger	0.823	0.799	0.803
	I watch out for threatening body language	0.679	0.714	0.663
	I keep my head down	0.637	0.720	0.545
	I don't stay anywhere for too long	0.586	0.793	0.713
	I try to keep a low profile	0.834	0.800	0.821
	I try to do what I need to do as quickly as possible	0.651	0.799	0.835
	I avoid making eye contact to prevent others from harming me	0.304	0.614	0.763

between the factors were $r=0.78$ (indoor and outdoor spaces), $r=0.70$ (indoor spaces and interactions), and $r=0.54$ (outdoor spaces and interactions).

CFA in the derivation sample ($n=516$) showed the three-factor model had an excellent model fit ($\chi^2=102.97$, d.f. = 51, $p<0.001$, CFI = 0.97, TLI = 0.97, RMSEA = 0.037, SRMR = 0.031). Similarly, in the confirmatory sample ($n=2237$), CFA demonstrated an excellent model fit ($\chi^2=355.83$, d.f. = 51, $p<0.001$, CFI = 0.98, TLI = 0.97, RMSEA = 0.04, SRMR = 0.027).

In-situation defences

From the 42-item pool, two items were deleted due to high correlations ($r>0.80$) (*I scan the area around me for any disturbances; I scan faces for signs of danger*). During the EFA, nine items were deleted due to low communalities (<0.30) (*I do not answer the front door; I call the police and legal professionals so that they can stop people targeting me; I keep track of suspicious things and events by taking pictures, videos, and audio recordings; I only eat or drink things I have prepared or opened myself, in case it has been interfered with; I wear things that hide my face, e.g. sunglasses, a cap; I carry something that I can use to protect myself; I make sure to avoid any cameras so they can't track me; I confront others before they can attack me; I try to be nice and agreeable to prevent others from harming me*); three items were deleted due to not loading on any factor (*I try to alter my physical appearance so I'm less of a target; I avoid doing things that make me feel vulnerable to attack (e.g. showering, sleeping); I try to be unpredictable so I can't be followed*); two items were deleted as they loaded with another item on a two-item factor (*I try to look tough to discourage people from attacking me; I avoid using technological devices in case they have been bugged or interfered with*); one item was deleted due to cross loading (*I position myself near an exit or escape route*); and five items were deleted for theoretical reasons such as high endorsement by both elevated and low paranoia groups, not clearly a defence behaviour, duplication of another item, or long and poorly worded (*I do not answer the phone; I repeatedly check, update, and/or review the security; I try to be ready for a fight; If I see any sign of danger I quickly leave the situation, even if I have not finished what I was doing; I make sure I'm visible so others can protect me*). With the remaining 20 items, a 5-factor structure was identified that explained 54% of the variance: safety at home, mitigating risk, preparing for escape, staying vigilant, and keeping a low profile. Correlations between the factors were $r=0.43$ – 0.66 .

CFA in the derivation sample ($n=516$) showed the five-factor model had a good model fit ($\chi^2=355.59$, d.f. = 160, $p<0.001$, CFI = 0.95, TLI = 0.94, RMSEA = 0.049, SRMR = 0.048). Similarly, in the confirmatory sample ($n=2237$), CFA demonstrated a good model fit ($\chi^2=160.05$, d.f. = 160, $p<0.001$, CFI = 0.95, TLI = 0.94, RMSEA = 0.046, SRMR = 0.039).

Scale reliability

Internal consistency was excellent for both the avoidance scale ($\alpha=0.92$, $n=2753$), and in-situation defences scale ($\alpha=0.94$, $n=2753$). A total score can be calculated for the avoidance scale by summing all 12 items. The avoidance scale has one 'n/a' option for item 12 ('My workplace or place of education'). If selected, this item is scored with the mean item score, calculated from the other 11 items (item 12 = sum of items 1–11 divided by 11). Similarly, a total score can be calculated for the in-situation scale by summing all 20 items. Test–retest reliability after two weeks in the 343 participants from the general population was good for both avoidance (ICC = 0.81, 95% CI = 0.77–0.85) and in-situation defences (ICC = 0.89, 95% CI = 0.86–0.91).

Scale validity

Mean scores and correlations between the O-PDQ avoidance and in-situation defences scales and measures of responses to psychosis, agoraphobic avoidance, paranoia, depression, and generalised anxiety are shown in Table 3.

Table 3. Descriptive statistics and bivariate correlations between the O-PDQ scales and other measures in the general population and patient group

			O-PDQ			
			O-PDQ avoidance scale		in-situation defences scale	
					<i>r</i>	<i>p</i>
	<i>n</i>	<i>Mean (SD)</i>	<i>r</i>	<i>p</i>		
Total group						
O-PDQ avoidance	2753	8.49 (7.83)	—	—	—	—
O-PDQ in-situation defences	2753	17.95 (12.28)	0.76	<0.001	—	—
Ideas of reference (R-GPTS)	1960	7.41 (7.98)	0.57	<0.001	0.63	<0.001
Ideas of persecution (R-GPTS)	1912	6.45 (9.64)	0.47	<0.001	0.56	<0.001
General population						
O-PDQ avoidance	2700	8.43 (7.77)	—	—	—	—
O-PDQ in-situation defences	2700	17.05 (12.18)	0.68	<0.001	—	—
Agoraphobia Mobility Inventory (AMI)	1884	1.97 (0.88)	0.79	<0.001	0.71	<0.001
Ideas of reference (R-GPTS)	1907	7.32 (7.97)	0.57	<0.001	0.60	<0.001
Ideas of persecution (R-GPTS)	1859	6.19 (9.51)	0.46	<0.001	0.55	<0.001
MCRP-Self regulation	1774	11.65 (3.73)	0.31	<0.001	0.44	<0.001
MCRP-Threat monitoring and avoidance	1774	10.25 (3.70)	0.71	<0.001	0.79	<0.001
MCRP-Social support	1780	8.59 (3.03)	0.28	<0.001	0.35	<0.001
O-AS Avoidance	1980	.89 (1.67)	0.62	<0.001	0.55	<0.001
O-AS Distress	1614	18.14 (19.72)	0.74	<0.001	0.73	<0.001
Brief Fears of Negative Evaluation (BFNE)	1719	20.69 (8.55)	0.44	<0.001	0.50	<0.001
Generalised anxiety (GAD7)	1920	9.51 (6.51)	0.61	<0.001	0.60	<0.001
Depression (PHQ-9)	1731	11.56 (7.80)	0.61	<0.001	0.58	<0.001
Patient group						
O-PDQ avoidance	53	11.41 (9.97)	—	—	—	—
O-PDQ in-situation defences	53	27.28 (13.54)	0.74	<0.001	—	—
Ideas of reference (R-GPTS)	53	10.91 (7.44)	0.51	<0.001	0.49	<0.001
Ideas of persecution (R-GPTS)	53	15.42 (9.89)	0.50	<0.001	0.43	<0.001
SBQ number of avoided places	51	2.94 (2.94)	0.75	<0.001	0.52	<0.001
SBQ frequency of avoidance	51	6.75 (7.21)	0.70	<0.001	0.53	<0.001
SBQ number of behaviours	51	5.06 (4.00)	0.66	<0.001	0.54	<0.001
SBQ frequency of behaviours	51	11.80 (9.86)	0.61	<0.001	0.55	<0.001

The O-PDQ showed good construct validity. In terms of convergent validity, both scales showed strong correlations with the MCRP threat monitoring and avoidance scale. Both O-PDQ scales showed moderate to strong correlations with established avoidance measures (AMI and O-AS avoidance scale), with the avoidance scale showing slightly stronger correlations compared with the in-situation defences scale. Both scales were strongly correlated with the O-AS distress scale. In the patient group, the O-PDQ avoidance scale showed a strong correlation with the both the number of avoided places and the frequency of avoidance reported on the SBQ. The distress scale showed a moderate correlation with both the number of defence behaviours reported and the frequency of defence behaviours reported on the SBQ. Supporting the theoretical relationship with paranoia, both O-PDQ scales showed moderate correlations with the R-GPTS ideas of reference and ideas of persecution scales. Further supporting construct validity, there was a main effect of group on both the avoidance scale ($F_{2,2750} = 138.8$, $p < .001$, $\eta^2 = 0.09$, 95% CI [.07, .11]) and the in-situation defences scale ($F_{2,2750} = 247.4$, $p < .001$, $\eta^2 = 0.15$, 95% CI [.13, .18]). Games-Howell *post-hoc* comparisons showed that the general population-low paranoia group (in-situation defences $M = 14.6$, $SD = 10.9$; avoidance $M = 7.2$, $SD = 7.2$) scored significantly lower on both scales compared with both the general population-elevated paranoia group (in-situation defences: $M = 25.8$, $SD = 12.4$, $p < 0.001$; avoidance: $M = 12.9$, $SD = 8.1$; $p < 0.001$) and the patient group (in-situation defences: $M = 27.3$, $SD = 13.5$, $p < 0.001$; avoidance: $M = 11.5$, $SD = 10.0$, $p < 0.01$). The general population-elevated paranoia and patient group did not significantly differ on the in-situation defences ($p = 0.72$) or avoidance ($p = .586$) scales. To examine this further, paranoia levels across the three groups were compared. Again, there was a significant main effect of group

on paranoia ($F_{2,1909} = 1696.5$, $p < 0.001$, $\eta^2 = .61$, 95% CI [.58, .63]). Games-Howell *post-hoc* comparisons showed that the general population-low paranoia group ($M = 0.9$; $SD = 1.4$) scored significantly lower than both the general population-elevated paranoia group ($M = 17.4$; $SD = 9.8$; $p < 0.001$) and the patient group ($M = 15.4$; $SD = 9.9$; $p < 0.001$). As with the O-PDQ scales, the general population-elevated paranoia and patient groups did not significantly differ ($p = 0.34$).

The discriminant validity of both O-PDQ scales was supported by weak correlations with other behavioural responses to psychosis (MCRP social support/reassurance and self-regulation scales). While correlations with the BFNE were moderate, they were slightly weaker than with paranoia. Moderate correlations with generalised anxiety and depression were observed.

Fifty-one patients with persecutory delusions completed both the O-PDQ and SBQ. The average number of defence behaviours endorsed to any degree on the O-PDQ was 14.1 ($SD = 4.8$) and endorsed as *often* or *always* was 8.7 ($SD = 5.8$) compared with an average of 2.16 ($SD = 1.8$) defence behaviours identified using the SBQ. The average number of places avoided to any degree on the O-PDQ was 5.7 ($SD = 4.3$) and avoided *often* or *always* was 3.6 ($SD = 3.6$) compared with an average of 2.9 ($SD = 2.9$) identified by the SBQ.

Discussion

In this study we aimed to develop the first self-report measure of defence behaviours used in response to paranoia. Items were derived from interviews with over 100 patients with persecutory delusions and tested with over 2700 people across the spectrum of paranoia. The resulting Oxford Paranoia Defence Behaviours Questionnaire (O-PDQ) has an excellent model fit, test-retest reliability, and construct validity. Good convergent validity was evidenced through high correlations with established measures of in-situation defences and avoidance. Both scales showed a moderate correlation with paranoia. The moderate, rather than strong, correlation with paranoid thinking reflects the distinct but related nature of these constructs. While defence behaviours are employed in response to paranoid thoughts, individuals may employ other strategies such as seeking support or using adaptive coping techniques. Discriminant validity was evidenced through minimal correlations with other behavioural responses to psychosis and a slightly weaker correlation with social anxiety compared with paranoia. Stronger correlations with anxiety and depression were found compared with paranoia, a pattern previously reported by Tully and colleagues (2017). This supports a close link between defence behaviours and emotional distress. Further validation of the measure comes from a study with a representative sample of 10,000 UK adults that tested 22 cognitive and social factors potentially associated with paranoia: use of in-situation defence behaviours as assessed by the O-PDQ was the factor most strongly associated with paranoia (Freeman and Loe, 2023).

The factor structure of the O-PDQ provides a framework for understanding avoidance and in-situation defence behaviours in paranoia. The avoidance scale has a three-factor structure – (i) avoidance of indoor or enclosed spaces, (ii) avoidance of outdoor or exposed spaces, and (iii) avoidance of interactions, which cuts across indoor and outdoor spaces and instead relates to the people that might be present and the expectation for proximity and interaction (e.g. a workplace or a GP appointment). The in-situation defence behaviour scale had a five-factor structure: (a) securing the home from danger, (b) mitigating the risk, for example by having someone with you or taking a safer route, (c) staying vigilant for signs of threat, (d) preparing for escape in response to signs of threat, and (e) keeping a low profile when out so not to come to the attention of the persecutors. This framework helps conceptualise the different types of avoidance and in-situation defence behaviours.

The avoidance and in-situation behaviours measured by the O-PDQ may maintain paranoia through multiple mechanisms. The non-occurrence of harm may be misattributed to the use of the safety-seeking behaviours (e.g. I wasn't attacked because I had someone with me) and they

prevent opportunities for positive restorative experiences that would allow the development of new more helpful beliefs (e.g. avoiding eye contact and rushing reduces the chances of having a positive interaction with a stranger and strengthening of the belief ‘most people are friendly and helpful’). Defences (e.g. hypervigilance) also shift attentional resources to signs of threat, hampering the receipt and processing of new information or the identification of true signs of safety. Hypervigilance can also increase the processing of cues that activate fears and confirm misinterpretations. Furthermore, acting like a belief is real may make it feel real, possibly through the activation of a fear response or creation of corroborating memories, similar to the concept of fantasy elaboration in grandiose delusions described by Isham *et al.* (2021). For example, having the memory of running home, terrified, because you perceived someone following you, would provide powerful reinforcement for the belief that you are in danger. In addition to directly maintaining paranoia, avoidance and in-situation defences serve to limit and isolate patients with deleterious effects on mental wellbeing. This may also partly account for the moderate correlation between defences and depression found in this study.

It is therefore not surprising that elimination, or at least reduction, of defence behaviours is a key feature of CBT for psychosis approaches (Freeman *et al.*, 2021a; Morrison, 2017). Clinicians need to be able to identify defence behaviours and distinguish them from adaptive coping behaviours. The O-PDQ identifies a number of subtle behaviours (e.g. wearing clothes that one can easily run away in; listening out for any sounds that might indicate danger) that may be easily overlooked in therapy. A long-acknowledged challenge is distinguishing coping and safety-seeking behaviours, particularly when they can be topographically similar (Salkovskis, 1991). In other words, a behaviour may be a defence behaviour in one case (praying for god to prevent you being harmed) but be a positive coping behaviour in another (praying because your faith gives you the strength to face your fears). This requires an assessment of both the intention of the behaviour (e.g. having someone with you to prevent you being attacked versus to provide distraction from voices) and the context it is used in (e.g. only going out at certain times of the day might be appropriate if you live in a dangerous area with high levels of crime). Self-report measures offer less flexibility, compared with interview, in assessing the functions of behaviours and wider context in which they occur. However, the O-PDQ can be a helpful clinical tool prompting a more detailed discussion around the behaviours identified and the reasons for their use. In addition, it can help frame these behaviours as a normal and understandable response to persecutory beliefs. Defence behaviours are often habitual and occur without much conscious awareness. The O-PDQ provides prompts to help people identify their defences; in contrast to interview measures, which depend more on the skill of the interviewer and the interviewee’s awareness, recognition, and recall.

There are a number of limitations. The general population group was recruited online and will not have been representative of the general population. Socio-economic status was not assessed, so it is unclear whether individuals from lower socio-economic backgrounds were represented. This may have influenced responses to avoidance items related to resource-dependent activities (e.g. gym use). Both the clinical and general population group were predominantly white and therefore the measure would benefit from further validation with ethnic minority groups. The clinical group in this study is relatively small and test–retest reliability was only assessed in the general population group. Although paranoid ideation occurs on a spectrum in the general population (Bebbington *et al.*, 2013), the O-PDQ would benefit from further testing in a larger clinical group. Furthermore, an assessment of sensitivity to clinically important change, which was not assessed in the current study, would increase its utility for monitoring progress in therapy. Nevertheless, the O-PDQ demonstrates robust psychometric properties and fills an important gap in the measurement of paranoia-related defence behaviours. The standardised format and ease of administration may facilitate greater attention on defensive behaviours – the key maintaining factor in paranoia – across both research and clinical settings.

Supplementary material. To view supplementary material for this article, please visit <https://doi.org/10.1017/S1352465825100994>

Data availability statement. Data are available upon reasonable request. Deidentified participant data will be available in anonymised form from the corresponding author (SL) on reasonable request (including a study outline), subject to university approval.

Acknowledgements. We would like to thank the participants who took part in this study and the lived experience advisory group. We would also like to thank the three NHS trusts that participated: Oxford Health NHS Foundation Trust, Berkshire Healthcare Foundation Trust NHS, and Central and North West London NHS Foundation Trust.

Author contributions. **Sinead Lambe:** Conceptualization (lead), Formal analysis (lead), Methodology (lead), Writing - original draft (lead); **Sophie Mulhall:** Conceptualization (supporting), Formal analysis (supporting), Writing - review & editing (equal); **Jessica Bird:** Formal analysis (equal), Methodology (equal), Writing - review & editing (equal); **Kira Williams:** Data curation (supporting), Investigation (supporting), Writing - review & editing (equal); **Joanna Mitchell:** Data curation (equal), Investigation (equal), Writing - review & editing (equal); **Mollie Roddan:** Investigation (equal), Writing - review & editing (equal); **Glory Sokunle:** Investigation (equal), Writing - review & editing (equal); **Laina Rosebrock:** Conceptualization (equal), Investigation (equal), Methodology (equal), Writing - review & editing (equal); **Daniel Freeman:** Conceptualization (equal), Formal analysis (equal), Funding acquisition (equal), Methodology (equal), Supervision (equal), Writing - original draft (equal); **Felicity Waite:** Conceptualization (equal), Methodology (equal), Supervision (equal), Validation (equal), Writing - review & editing (equal).

Financial support. The project was supported by a grant from the Medical Research Council Developmental Pathway Funding Scheme (MR/P02629X/1), an NIHR Senior Investigator (NIHR204013) award to D.F., and the NIHR Oxford Health Biomedical Research Centre. S.L. is supported by an NIHR Doctoral Fellowship (NIHR301483). F.W. is supported by a Wellcome Trust Doctoral Fellowship (102176/B/13/Z). D.F. is an NIHR Senior Investigator. The views expressed are those of the authors and not necessarily those of the NIHR or the Department of Health and Social Care.

Competing interests. The authors declare none.

Ethical standards. All participants provided informed consent. Ethical approval was obtained from the Medical Sciences Inter-Divisional Research Ethics Committee at the University of Oxford for the general population group and the NHS Health Research Authority (HRA) South Central – Oxford B Research Ethics Committee (REC reference: 18/SC/0316) for the patient group.

References

- Bebbington, P. E., McBride, O., Steel, C., Kuipers, E., Radovanovic, M., Brugha, T., Jenkins, R., Meltzer, H. I., & Freeman, D. (2013). The structure of paranoia in the general population. *British Journal of Psychiatry*, 202, 419–427. <https://doi.org/10.1192/bjp.bp.112.119032>
- Bentler, P. M., & Bonett, D. G. (1980). Significance tests and goodness-of-fit in the analysis of covariance structures. *Psychological Bulletin*, 88, 588–606. <https://doi.org/10.1037/0033-2909.88.3.588>
- Chambless, D. L., Caputo, G. C., Jasin, S. E., Gracely, E. J., & Williams, C. (1985). The Mobility Inventory for Agoraphobia. *Behaviour Research and Therapy*, 23, 35–44. [https://doi.org/10.1016/0005-7967\(85\)90140-8](https://doi.org/10.1016/0005-7967(85)90140-8)
- Chambless, D. L., Sharpless, B. A., Rodriguez, D., McCarthy, K. S., Milrod, B. L., Khalsa, S. R., & Barber, J. P. (2011). Psychometric properties of the mobility inventory for agoraphobia: convergent, discriminant, and criterion-related validity. *Behavior Therapy*, 42, 689–699. doi: [10.1016/j.beth.2011.03.001](https://doi.org/10.1016/j.beth.2011.03.001)
- Clark, D. M. (2005) *Three Questionnaires for Measuring Central Constructs in the Cognitive Model of Social Phobia*. King's College London, London.
- Clark, D. M., & Salkovskis, P. M. (2009). *Panic Disorder: Manual for Improving Access to Psychological Therapies (IAPT) High Intensity Therapists*. Available at: https://oxcadatresources.com/wp-content/uploads/2018/06/Cognitive-Therapy-for-Panic-Disorder_IAPT-Manual.pdf. Retrieved 21/10/2022
- Clark, D. M., & Wells, A. (1995). A cognitive model of social phobia. In Heimberg, R. G., Liebowitz, M. R., Hope, D. A., & Schneier, F. (eds), *Social Phobia: Diagnosis, Assessment, and Treatment* (pp. 69–93). New York: Guilford Press.
- Costello, A. B., & Osborne, J. W. (2005) Best practices in exploratory factor analysis: four recommendations for getting the most from your analysis. *Practical Assessment, Research and Evaluation*, 10, 1–9. doi: <https://doi.org/10.7275/jyj1-4868>
- Cuming, S., Rapee, R. M., Kemp, N., Abbott, M. J., Peters, L., & Gaston, J. E. (2009). A self-report measure of subtle avoidance and safety behaviors relevant to social anxiety: development and psychometric properties. *Journal of Anxiety Disorders*, 23, 879–883. doi: [10.1016/j.janxdis.2009.05.002](https://doi.org/10.1016/j.janxdis.2009.05.002)

- Dunmore, E., Clark, D. M., & Ehlers, A. (2001). A prospective investigation of the role of cognitive factors in persistent posttraumatic stress disorder (PTSD) after physical or sexual assault. *Behaviour Research and Therapy*, 39, 1063–1084. doi: [10.1016/s0005-7967\(00\)00088-7](https://doi.org/10.1016/s0005-7967(00)00088-7).
- Ehlers, A., & Clark, D. M. (2000). A cognitive model of posttraumatic stress disorder. *Behaviour Research and Therapy*, 38, 319–345. doi: [https://doi.org/10.1016/s0005-7967\(99\)00123-0](https://doi.org/10.1016/s0005-7967(99)00123-0)
- Elo, S., & Kyngas, H. (2007). The qualitative content analysis process. *Journal of Advanced Nursing*, 62, 107–115. doi: [10.1111/j.1365-2648.2007.04569.x](https://doi.org/10.1111/j.1365-2648.2007.04569.x)
- Foa, E. B., Huppert, J. D., Leiberg, S., Langner, R., Kichic, R., Hajcak, G., & Salkovskis, P. M. (2002). The Obsessive-Compulsive Inventory: development and validation of a short version. *Psychological Assessment*, 14, 485.
- Freeman, D. (2016). Persecutory delusions: a cognitive perspective on understanding and treatment. *The Lancet Psychiatry*, 3, 685–692. doi: [https://doi.org/10.1016/S2215-0366\(16\)00066-3](https://doi.org/10.1016/S2215-0366(16)00066-3)
- Freeman, D., Bradley, J., Antley, A., Bourke, E., DeWeever, N., Evans, N., Černis, E., Sheaves, B., Waite, F., Dunn, G., & Clark, D. M. (2016). Virtual reality in the treatment of persecutory delusions: randomised controlled experimental study testing how to reduce delusional conviction. *British Journal of Psychiatry*, 209, 62–67. doi: <https://doi.org/10.1192/bjp.bp.115.176438>
- Freeman, D., Emsley, R., Diamond, R., Collett, N., Bold, E., Chadwick, E., Isham, S., Bird, J., Edwards, D., Kingdon, D., Fitzpatrick, R., Kabir, T., Waite, F., on behalf of the Oxford Cognitive Approaches to Psychosis Trial Study Group (2021a). Comparison of a theoretically driven cognitive therapy (the Feeling Safe Programme) with befriending for the treatment of persistent persecutory delusions: a parallel, single-blind, randomised controlled trial. *Lancet Psychiatry*, 8, 696–707. doi: [https://doi.org/10.1016/S2215-0366\(21\)00158-9](https://doi.org/10.1016/S2215-0366(21)00158-9)
- Freeman, D., & Garety, P. A. (2000). Comments on the content of persecutory delusions: does the definition need clarification? *British Journal of Clinical Psychology*, 39, 407–414. doi: <https://doi.org/10.1348/014466500163400>
- Freeman, D., Garety, P. A., & Kuipers, E. (2001). Persecutory delusions: developing the understanding of belief maintenance and emotional distress. *Psychological Medicine*, 31, 1293–1306. doi: <https://doi.org/10.1017/S003329170100455x>
- Freeman, D., Garety, P. A., Kuipers, E., Fowler, D., Bebbington, P. E., & Dunn, G. (2007). Acting on persecutory delusions: the importance of safety seeking. *Behaviour Research and Therapy*, 45, 89–99. doi: <https://doi.org/10.1016/j.brat.2006.01.014>
- Freeman, D., Isham, L., & Waite, F. (2025). A counterweight model for understanding and treating persecutory delusions. *Psychological Medicine*, 55, e141.
- Freeman, D., Lister, R., Waite, F., Galal, U., Yu, L.M., Lambe, S., Beckley, A., Bold, E., Jenner, L., Diamond R., Kirkham, M., Twivey, E., Causier, C., Carr, L., Saidel, S., Day, R., Beacco, A., Rovira, A., Ivins, A., Nah, R., Slater, M., Clark, D. & Rosebrock, L. (2023). Automated virtual reality cognitive therapy versus virtual reality mental relaxation therapy for the treatment of persistent persecutory delusions in patients with psychosis (THRIVE): a parallel-group, single-blind, randomised controlled trial in England with mediation analyses. *The Lancet Psychiatry*, 10, 836–847. doi: [10.1016/S2215-0366\(23\)00257-2](https://doi.org/10.1016/S2215-0366(23)00257-2)
- Freeman, D., & Loe, B. S. (2023). Explaining paranoia: cognitive and social processes in the occurrence of extreme mistrust. *BMJ Ment Health*, 26. doi: <https://doi.org/10.1136/bmjment-2023-300880>
- Freeman, D., Loe, B. S., Kingdon, D., Startup, H., Molodynski, A., Rosebrock, L., Brown, P., Sheaves, B., Waite, F. & Bird, J. C. (2021b). The revised Green *et al.* Paranoid Thoughts Scale (R-GPTS): psychometric properties, severity ranges, and clinical cut-offs. *Psychological Medicine*, 51, 244–253. doi: <https://doi.org/10.1017/S0033291719003155>
- Isham, L., Griffith, L., Boylan, A. M., Hicks, A., Wilson, N., Byrne, R., Sheaves, B., Bentall, R. P., & Freeman, D. (2021). Understanding, treating, and renaming grandiose delusions: a qualitative study. *Psychology and Psychotherapy: Theory, Research and Practice*, 94, 119–140. doi: <https://doi.org/10.1111/papt.12260>
- Kroenke, K., Spitzer, R. L., & Williams, J. B. (2001). The PHQ-9: validity of a brief depression severity measure. *Journal of General Internal Medicine*, 16, 606–613. doi: <https://doi.org/10.1046/j.1525-1497.2001.016009606.x>
- Lambe, S., Bird, J. C., Loe, B. S., Rosebrock, L., Kabir, T., Petit, A., Mulhall, S., Jenner, L., Aynsworth, C., Murphy, E., Jones, J., Powling, R., Chapman, K., Dudley, R., Morrison, A., Regan, E. O., Yu, L.-M., Clark, D.M., Waite, F., & Freeman, D. (2023). The Oxford Agoraphobic Avoidance Scale. *Psychological Medicine*, 53, 1233–1243. doi: [10.1017/S0033291721002713](https://doi.org/10.1017/S0033291721002713)
- Leary, M. R. (1983). A brief version of the Fear of Negative Evaluation Scale. *Personality and Social Psychology Bulletin*, 9, 371–375. doi: <https://doi.org/10.1177/0146167283093007>
- McManus, F., Clark, D. M., Grey, N., Wild, J., Hirsch, C., Fennell, M., . . . & Manley, J. (2009). A demonstration of the efficacy of two of the components of cognitive therapy for social phobia. *Journal of Anxiety Disorders*, 23, 496–503. doi: <https://doi.org/10.1016/j.janxdis.2008.10.010>
- Morrison, A. P. (2017). A manualised treatment protocol to guide delivery of evidence-based cognitive therapy for people with distressing psychosis: Learning from clinical trials. *Psychosis*, 9, 271–281. doi: <https://doi.org/10.1080/17522439.2017.1295098>
- R Core Team (2013). R: A Language and Environment for Statistical Computing. R Foundation for Statistical Computing, Vienna. <http://www.R-project.org/>
- Revelle, W. (2020). Psych: Procedures for Personality and Psychological Research. North western University, Evanston, <https://CRAN.R-project.org/package=psych>. R package version 2.0.8

- Rosseel, Y. (2012). lavaan: an R package for structural equation modeling. *Journal of Statistical Software*, 48, article 1. doi: <https://doi.org/10.18637/jss.v048.i02>
- Salkovskis, P. M. (1991). The importance of behaviour in the maintenance of anxiety and panic: a cognitive account. *Behavioural and Cognitive Psychotherapy*, 19, 6–19. doi: <https://doi.org/10.1017/S0141347300011472>
- Salkovskis, P. M. (1999). Understanding and treating obsessive-compulsive disorder. *Behaviour Research and Therapy*, 37, S29–S52. doi: [https://doi.org/10.1016/S0005-7967\(99\)00049-2](https://doi.org/10.1016/S0005-7967(99)00049-2)
- Salkovskis, P. M., Clark, D. M., Hackmann, A., Wells, & Gelder, M. G. (1999). An experimental investigation of the role of safety-seeking behaviours in the maintenance of panic disorder with agoraphobia. *Behaviour Research and Therapy*, 37, 559–574. doi: [10.1016/s0005-7967\(98\)00153-3](https://doi.org/10.1016/s0005-7967(98)00153-3).
- Salkovskis, P. M., Thorpe, S. J., Wahl, K., Wroe, A. L., & Forrester, E. (2003). Neutralizing increases discomfort associated with obsessional thoughts: an experimental study with obsessional patients. *Journal of Abnormal Psychology*, 112, 709–715. doi: [10.1037/0021-843X.112.4.709](https://doi.org/10.1037/0021-843X.112.4.709)
- Spitzer, R. L., Kroenke, K., Williams, J. B., & Löwe, B. (2006). A brief measure for assessing generalized anxiety disorder: the GAD-7. *Archives of Internal Medicine*, 166, 1092–1097. doi: <https://doi.org/10.1001/archinte.166.10.1092>
- Tully, S., Wells, A., Pyle, M., Hudson, J., Gumley, A., Kingdon, D., Schwannauer, M., Turkington, D., & Morrison, A. P. (2017). Measuring common responses to psychosis: assessing the psychometric properties of a new measure. *Schizophrenia Research*, 181, 131–136. doi: [10.1016/j.schres.2016.10.01](https://doi.org/10.1016/j.schres.2016.10.01)