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## Epidemiological Studies of Low-dose Ionizing Radiation and Cancer: Summary Bias Assessment and Meta-Analysis

--Manuscript Draft--

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## Response to reviewer comments

We thank editor and reviewers for their comments. Below is a detailed response to all issues raised. In addition, we explain a change we made based on a comment we received on the overview manuscript in this series of manuscripts.

Because the accepted versions of manuscripts will now be published on the Journal site before they undergo copy editing and typesetting by the Journal, we are asking authors to carefully proofread their manuscripts for correct spelling and grammar as part of the revision process.

**Response:** Agreed. We have carefully checked the revised manuscript and made a few corrections which do not affect the content.

This is part of a series of submitted manuscripts which describe a systematic assessment of information from epidemiological studies on low-dose radiation cancer risks. In this manuscript, the authors describe a formal test of hypotheses “that the median excess relative risk per 100 mGy equals 0” (e.g., for leukemia and solid cancers) and a meta-analysis to estimate the risk per 100 mGy – based primarily on minimally biased low-dose epidemiological studies.

Overall, this manuscript is extremely well-written and of excellent quality. The authors have carefully considered the minefield of issues that greatly complicate the drawing of inferences about low-dose radiation risks from epidemiological study results. I have only a few questions/comments:

1) Is there another way to express the basic hypothesis “that the median excess relative risk per 100 mGy equals 0”? Isn’t it essentially whether the risk per unit dose (at doses below 100 mGy) equals 0 (and it doesn’t matter what that unit dose is)? More fundamentally, what are the underlying assumptions for the test of this hypothesis”? Are the authors assuming that the dose response (for specific ages of exposure/attained age?) is linear, but that the slopes can depend on study population?

**Response:** Agreed. We concur with the editor and have changed „excess relative risk per 100 mGy“ into „excess relative risk per unit dose“. The sign test assumes that the observations (i.e., the ERR per unit dose) are independently identically distributed. Normality is not required for this, although it is for the meta-analysis, as we mention on pages 16-17. We do not assume that the dose-response relationship is linear, but that the best linear approximation of the dose-response relationship for all ages of exposure and attained ages is characterized by an ERR per unit dose of zero. Under this assumption, each study provides a realization and we expect about half of the ERRs per unit dose to be above and below zero.

2) One of the reviewer’s comments was that the sign-test approach seemed simplistic “and off the mark.” I don’t necessarily agree, but I recommend that the authors explain why they chose this approach. For example, why was a nonparametric test necessary and, then, why the sign-test versus some other type of nonparametric test?

**Response:** We agree with the Reviewer that the sign test is “simplistic”. We think that the test is intuitive and easily understood by readers with little or no statistical background, which is precisely why we think it was worthwhile to include. As we explain, we „... could not assess the magnitude of the bias in most situations because the required data were not available in the publications.“ (page 17). The sign test is the only test of the median being equal to a hypothesized

value (e.g., zero) which does not take the magnitude of the observations into account but just whether or not they exceed the hypothesized value. Alternative tests take the actual values into account (t-test) or their ranks (Wilcoxon signed rank test). Moreover, the sign test does not make any assumptions about the shape of the distribution, while other tests require normality (t-test) or symmetry (Wilcoxon signed rank test). We added on page 17 after the sentence cited above: „Our summary analysis was therefore limited to the simple sign test since alternative tests are based on the actual values (t-test) or their ranks (Wilcoxon signed rank test).“

3) Overall, I found the presentation of conclusions to be fair/balanced. However, the sentence (see, for example, the 2nd to last sentence of the abstract) that “exclusion of these studies did not change our conclusion, which is that these new epidemiological studies directly support excess cancer risks from low-dose ionizing radiation” might give some pause. Although that is technically true (brushing aside the American Statistical Association’s statement on hypothesis tests?), it is also true that after excluding those studies the p-value for solid cancers increased from about 0.001 to 0.05.

**Response:** Agreed. Throughout this effort we aimed to be in line with the ASA statement on p-values, particularly by providing context for their appropriate interpretation. Before we specify the changes made, we note that due to an error in abstracting, the ERR for solid cancer reported in the study of Rocketdyne workers by Boice et al (2001) was unfortunately included as 0.02 when it should have been -0.02. We regret this oversight and have updated all calculations involving this figure. While the meta-ERR and associated figures changed only slightly, the p-value of the sign test excluding positive studies potentially biased away from zero became non-significant. We changed the corresponding text as follows:

Abstract, results: “For solid cancers, 16 of 22 studies reported positive ERRs per unit dose and we rejected the hypothesis that the median ERR equals zero ( $p=0.026$ ). After exclusion of four positive studies with potential positive bias, 12 of 18 studies reported positive ERRs per unit dose ( $p=0.119$ ).”

Abstract, conclusion: “Our systematic assessments in this monograph showed that these new epidemiological studies are characterized by several limitations, but only a few positive studies were potentially biased away from the null. After exclusion of these studies, the majority of studies still reported positive risk estimates.”

Discussion, last paragraph: “Our systematic assessment in this monograph showed that these epidemiological studies of low-dose radiation and cancer risk are characterized by several limitations, but we found that only a small minority of the studies had biases whose correction could have moved a positive ERR towards the null. After exclusion of these studies, the majority of studies still reported positive risk estimates.”

We still conclude that “that there is now a large body of epidemiological data which supports excess cancer risks from low-dose ionizing radiation” because the sign test is only one piece of evidence we use to assess whether the identified data are consistent with no effect of radiation at low doses. Also, we made the conservative assumption that there is no bias which changed a positive ERR into a negative.

4) With respect to the reviewer’s comment “to say something”, I think the authors have already done that! However, it would be appropriate to discuss future research needs (either here or in the summary paper by Gilbert et al.). For example, for solid cancer, the test of the null hypothesis of no

effect yielded results of “borderline significance”. Can we expect that future analyses such as this one will yield something more definitive, and what can be done to assure that would be the case?

**Response:** We explicitly make recommendations for the presentation of future studies in each of the papers in this monograph, and we describe this in the discussion section of this manuscript: „We have made several recommendations in each of our papers that would facilitate assessments of bias in the risk estimates in the future. For example, we recommend the routine publication of assessments of dose uncertainty and levels of loss to follow-up by exposure and outcome.” (Pages 17/18). We further demonstrate how evidence from sub-studies (e.g., which include data on confounders not available for an entire cohort) can be used to improve risk assessment. In addition, we added text in the discussion about the distribution of doses in the included and excluded studies (based on comments we received about the overview manuscript, see below). There, we recommend pooled analyses of individual subjects with doses below 100 mGy.

5) Please also respond to the other comments of the reviewer.

**Response:** See below.

### Comments from the Reviewers

Please note: All the comments to authors we have received are included below, regardless of the numbering of the reviewers.

Reviewer 1: 1. This paper is a description of earlier chapters in a monograph over-viewing epidemiological studies of ionizing radiation and cancer. It includes summary analyses and meta-analytical estimates. While I do not have the other chapters, this summary is informative.  
2. The summary is affirmative of what is known. Are there conclusions? A need for a new BEIR report or other assessment of risks for policy? Say something!

**Response:** Please see our response above to the fourth point of the editor.

Page 5 "...to be small...": what does this mean?

**Response:** Agreed. We changed the phrase to „because the risks are likely to be small compared to non-radiation risk factors“.

Page 5 "...underpowered...": in relation to?

**Response:** Agreed. We changed the phrase to „studies may have power below 80% (the conventional threshold of adequacy in this respect)“.

Page 7 "We performed a one-sided sign test for the reported ERRs...": The simplest approach seems simplistic and off the mark, given the depth of prior findings.

**Response:** Please see our response to the editor's second point.

Page 8 "A one-side p-value below 0.05...": Why one-sided for negative?

**Response:** We assumed that radiation exposure is either not associated with cancer risk (null hypothesis, median ERR=0) or is positively associated with an increased cancer risk (alternative, median ERR>0). This results in a one-sided test. This is stated in the first sentence of the same paragraph on page 8: „We performed a one-sided sign test for the reported ERRs, separately for solid cancers and leukemia, to evaluate the hypothesis that the median of the ERRs equals zero versus the alternative that the median ERR exceeds zero.“

Page 18 "...large effects are not considered a necessary criterion...": Hill did not describe his 9 points as "criteria"

**Response:** Agreed. We replaced the term „criterion“ by „viewpoint“ (twice) and „aspect“ (once) which are terms used by Bradford Hill in the original publication.

#### **Additional change made in response to comment on other manuscript**

We extended the discussion of the exclusion of studies with a mean cumulative dose exceeding 100 mGy on page 14 by referring to new information in the revised overview manuscript on the proportion of subjects with cumulative doses exceeding 100 mGy per study. The paragraph now reads: "Several studies were ineligible for our review. Some details are presented for studies that had conducted an internal dose-response analysis but were ineligible because they then failed on an additional criterion (Table 3). One reason for exclusion was that the mean cumulative dose exceeded 100 mGy. This led to the exclusion of (borderline) significantly positive studies such as Techa River, Mayak worker, U.S. scoliosis, Chornobyl clean-up workers and Chinese medical workers, and the non-significantly negative study of background exposure in Kerala. While the proportion of subjects with cumulative doses exceeding 100 mGy in these excluded studies ranged between 20% and 80%, it was below 10% among all included studies except for 5 studies for which it ranged between 11% and 22%. Excluding these latter 5 studies from our test whether the median ERR per unit dose equals 0 still resulted in rejection of this hypothesis (solid cancer:  $p=0.021$  for 15 of 20 positive studies, leukemia:  $p=0.011$  for 13 of 16 positive studies). As an alternative, we recommend future analyses of individual subjects exposed to cumulative doses below 100 mGy, as recently done for leukemia after childhood exposure." The text in the manuscript includes references to all mentioned studies.

#### **Additional changes made to increase consistency with other manuscripts in this monograph**

We added reference numbers to studies wherever possible. We also made corrections to reflect that the Korean worker study and the French nuclear worker study presented 90% confidence intervals (not 95%), that the supplements of the UK NRRW study included 95% confidence intervals (which we used), and that we used the estimate for solid cancer in the Taiwanese cohort instead of that for all cancers excl. leukemia (hardly any difference). Because of some of these changes, the meta-analysis results changed very slightly.

# **Epidemiological Studies of Low-dose Ionizing Radiation and Cancer: Summary Bias Assessment and Meta-Analysis**

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## **Abstract**

**Background:** Ionizing radiation is an established carcinogen, but risks from low-dose exposures are controversial. Since the Biological Effects of Ionizing Radiation (BEIR) VII review of the epidemiological data in 2006 many subsequent publications have reported excess cancer risks from low-dose exposures. Our aim was to systematically review these studies to assess the magnitude of the risk, and whether the positive findings could be explained by biases.

**Methods:** Eligible studies had mean cumulative doses <100 mGy, individualized dose estimates, risk estimates and confidence intervals (CI) for the dose-response and were published in 2006-2017. We summarized the evidence for bias (dose error, confounding, outcome ascertainment) and its likely direction for each study. We tested whether the median excess relative risk (ERR) per unit dose equals zero and assessed the impact of excluding positive studies with potential bias away from the null. We performed a meta-analysis to quantify the ERR and assess consistency across studies for all solid cancers and leukemia.

**Results:** Of the 26 eligible studies, 8 concerned environmental, 4 medical and 14 occupational exposure. For solid cancers, 16 of 22 studies reported positive ERRs per unit dose and we rejected the hypothesis that the median ERR equals zero ( $p=0.03$ ). After exclusion of four positive studies with potential positive bias, 12 of 18 studies reported positive ERR per unit dose ( $p=0.12$ ). For leukemia, 17 of 20 studies were positive and we rejected the hypothesis that the median ERR per unit dose equals zero ( $p=0.001$ ), also after exclusion of five positive studies with potential positive bias ( $p=0.02$ ). For adulthood exposure, the meta-ERR/100 mGy was 0.030 (95% CI: -0.02, 0.081) for solid cancers and 0.16 (95% CI: 0.07, 0.25) for leukemia. For childhood exposure, the meta-

ERR/100 mGy for leukemia was 2.84 (95% CI: 0.37, 5.32); there were no eligible studies of all solid cancers.

Conclusions: Our systematic assessments in this monograph showed that these new epidemiological studies are characterized by several limitations, but only a few positive studies were potentially biased away from the null. After exclusion of these studies, the majority of studies still reported positive risk estimates. We therefore conclude that these new epidemiological studies directly support excess cancer risks from low-dose ionizing radiation. Furthermore, the magnitude of the cancer risks from these low-doses radiation exposures was statistically compatible with the radiation dose-related cancer risks of the atomic bomb survivors.

## **Introduction**

The evidence for cancer risks provided by epidemiological studies of low-dose ionizing radiation exposure is of key relevance to radiation protection since a large fraction of the population is exposed to low doses of ionizing radiation from diagnostic medical procedures or occupationally, in addition to natural background radiation. Careful and sophisticated interpretation of the results from these studies is required, however, because the risks are likely to be small compared to those associated with non-radiation risk factors, studies may have power below 80% (the conventional threshold of adequacy in this respect), dose estimation may be limited and/or retrospective and studies may suffer from biases typical of observational studies such as confounding.

The last major US review of the epidemiological and experimental evidence for cancer risks from low-dose exposures (which we denote as  $<100$  mGy) in 2006 concluded that “the available scientific evidence is consistent with a linear dose-response relationship between ionizing radiation and the development of cancer in humans” (1). This conclusion was largely based on studies of populations exposed to higher doses combined with experimental data. Subsequent to 2006, several new epidemiological studies of populations exposed primarily to low doses have been published and several existing studies have reported new results from extended follow-up. Most of these new publications report excess cancer risks from low dose radiation exposures. The aim of this monograph is to systematically evaluate whether there is direct human evidence of excess cancer risks from low-dose ( $<100$  mGy) radiation exposure, and, if so, what the magnitude of the risk is and whether the positive findings could be explained by biases.

Here, we provide a synthesis from our in-depth systematic assessments of the methodology for the eligible studies published during 2006-2017 that we evaluated, and the associated potential for the risk estimates to be biased due to dose error, confounding, selection bias, or outcome

misclassification (2-5). We assess for each study the direction of the biases from any of these sources. As a general evaluation of whether the studies support cancer risks from low-dose ionizing radiation we conducted a sign test for whether the median of the excess relative risks (ERRs) equals zero, and then calculated the impact of excluding the positive studies identified as being biased away from the null. Finally, to quantify the magnitude of the estimated risks, we conduct a meta-analysis for all solid cancers and for leukemia from childhood or adulthood exposure to low-dose ionizing radiation.

## **Methods**

We included epidemiological studies published since the BEIR VII report in 2006 (1) and before 2018. Studies were eligible for inclusion if they were based on human populations exposed to low dose, predominantly low linear energy transfer (LET), radiation (mean cumulative dose < 100 mGy). We required individualized dose estimates for the study participants, and that the publications provided risk estimates and confidence intervals (CI) for the dose-response for cumulative radiation dose. For a full description of the eligible studies see the overview paper in this monograph (6). In brief, we used the results of the methodological assessments conducted in this monograph (dosimetry (2), confounding (5), outcome (4)) to derive, for each aspect, an assessment of the potential for bias in the risk estimate and the direction of the bias.

Summarizing the evaluations of the different study aspects was as follows. We assessed the strengths and weaknesses of dosimetry systems with respect to the directness, complexity, and completeness of the dosimetry, the dosimetric uncertainty and the validity of dose estimates (2). This process identified studies with a known or suspected bias in dose estimates and the likely direction of the bias in the risk estimate.

In assessing the evidence for confounding and selection bias, we summarized methods to control confounding and assessed the likelihood of uncontrolled confounding as well as its direction (5). This assessment was based on available data from the eligible studies and related publications including some examples of quantitative bias assessment to examine the potential magnitude of the bias.

The outcome evaluation paper in this monograph reviewed the possible impact of differential outcome ascertainment across radiation dose levels (4). The evaluation also considered loss to follow-up, under- or over-ascertainment of cancer outcomes, misclassification of outcomes, and changing classifications over time. The main objective was to identify studies whose outcome ascertainment was differential regarding exposure level and hence to bias the relative risk estimate. We then performed a summary of the assessments of different biases for each study and carefully considered both the direction of the observed effect and the direction of the bias. For our general question of whether the studies overall support excess cancer risks (as opposed to the question of the magnitude of the risk) our priority was to identify the positive studies with bias in the positive direction or bias of uncertain direction. If there were several biases acting in different directions, or it was not possible to determine the potential direction, we classified these studies as potentially biased away from the null. Since the magnitude of bias is difficult to determine from the published reports, we conservatively counted any study with a negative ERR estimate as negative regardless of any potential for bias. We also indicated whether the estimated power (if available) was low ( $<50\%$ ) or reasonable ( $\geq 50\%$ ) (3).

We performed a one-sided sign test for the reported ERRs, separately for solid cancers and for leukemia, to evaluate the hypothesis that the median of the ERRs per unit dose equals zero versus the alternative that the median ERR per unit dose exceeds zero (7). The sign test excluded the

INWORKS study (8, 9) because it is a pooled analysis of UK, French and U.S. nuclear workers, which were included separately. To assess the impact of the studies identified as potentially biased we then repeated the sign test after excluding the studies where bias adjustment could move a positive ERR towards the null. A one-sided p-value below 0.05 was considered statistically significant.

Finally, we conducted a meta-analysis of the published ERR estimates at 100 mGy to quantify the magnitude of the risk and to assess the consistency across studies for both all solid cancers and leukemia. Here we excluded the INWORKS study (8, 9) because of overlap as described above and the US radiologic technologists (USRT) studies (10-12) because only site-specific ERRs were reported. All other studies were included regardless of potential bias because as described above we generally could not quantify the magnitude of the bias and hence the magnitude of bias-corrected ERRs, and because excluding the subset of positive studies with positive biases would bias the summary risk estimate away from the null. We generated standard errors of the ERR at 100 mGy based on the upper (UL) and lower (LL) limit of the CI as  $(UL-LL)/2*1.96$  or  $(UL-LL)/2*1.645$ , depending on whether 95% or 90% CI were reported, respectively. We acknowledge that this is an approximation that may not be fully adequate given the skewed dose distributions and small numbers of cases in some of the studies. The meta-ERR estimate was derived from a random effects model using the iterative method of Paule and Mandel (13) as outlined in DerSimonian and Kacker (14). We also computed Cochran's Q, which is the weighted sum of squared differences between individual study effects and the pooled effect across studies, to test homogeneity, as well as the  $I^2$  statistic (variance due to heterogeneity (15)). If homogeneity was rejected, we excluded the studies with the largest signed contribution to the Q statistic. Meta-ERRs were calculated separately for studies on adult solid cancers, adult leukemia and childhood

leukemia. We did not calculate meta-ERR estimates for childhood solid cancers because only site-specific ERRs were reported.

## **Results**

Of the 26 eligible studies, we found that three studies had a known or suspected bias in dose estimates that could bias the risk estimate away from the null due to possible recall or selection bias (Chornobyl residents (16), Chornobyl liquidators (17), Ukrainian Chornobyl liquidators (18), Tables 1 and 2) and one study that was likely biased towards the null (Three Mile Island, leukemia) (19). The direction of potential dose bias was uncertain in the USRT study on breast cancer (12). Among the three case-control studies of leukemia in Chornobyl liquidators (17, 18) and residents (16) individual dose estimates relied extensively on information obtained by interview after case ascertainment. In two of these studies there was evidence from the manuscripts that the risk estimate was reduced after exclusions of a study center (Chornobyl residential childhood leukemia study) (16) and proxy respondents (Ukrainian liquidators leukemia study) (18).

Our evaluation of confounding and selection bias identified several sources that could bias the risk estimates: clinical indication for studies of medical diagnostic exposures during childhood, lifestyle factors for environmental, medical, and occupational studies of adult cancers and for occupational studies, other workplace exposures and healthy worker survivor bias. In addition to the impact of a potential dose error considered above (2), the Chornobyl residential case-control study (16) may also have suffered from control selection bias that may have upwardly biased the risk estimate. We assessed that a potential for uncontrolled confounding biasing the risk estimate towards the null from healthy worker survivor bias was especially likely for the Korean workers (29) and Japanese nuclear workers (33) studies, which did not adjust for socio-economic status



(SES) or duration of employment, and the German nuclear workers study (35), which did not adjust for birth cohort and SES (Tables 1 and 2). For the solid cancer results in the Canadian (34) and German nuclear workers study (35) and the leukemia findings in the Japanese nuclear workers study (33), bias adjustment would move the ERR towards the null. For the Japanese nuclear workers, the direction of the bias was uncertain, given that smoking may be a positive confounder in this cohort (40). Therefore, we could not draw a definitive conclusion on the impact of bias adjustment with the available data.

Four studies may have had cancer ascertainment possibly differential by radiation exposure, which could have biased the risk estimate. These include the Japanese nuclear workers (33) with bias towards the null through early loss to follow-up during periods with higher radiation exposure; the Chinese background study (20) with possible bias away from the null because of a higher loss to follow-up in regions with lower background radiation exposure compared to regions with high background radiation exposure; the cardiovascular imaging study (25), possibly biased away from the null because those undergoing diagnostic/therapeutic procedures with imaging were more likely to have cancer outcomes detected; and the Korean radiation workers population (29), where medical surveillance was required for radiation workers but not for the general population or for the worker comparison group (manufacture of motor vehicles) (Tables 1 and 2).

When considering the three potential biases for the 22 studies of solid cancers, there were three studies with negative ERR estimates where adjustment would likely move the ERR towards the null (Three Mile Island (19), Canadian (34) and German nuclear workers (35)) and four studies with positive ERR estimates where it was uncertain whether adjustment would move the ERR towards or away from the null (Canadian cardiac imaging (25), Korean (29) and Japanese workers (33) and USRT breast cancer (12), Table 1). For leukemia there was one study with a negative

ERR where adjustment would likely have moved the ERR towards the null (Japanese workers (33)), three studies with a positive ERR where adjustment would likely have moved the ERR towards the null (Chornobyl residents (16), Chornobyl liquidators (17) and Ukrainian Chornobyl liquidators (18)), and two positive studies where the direction of the bias was uncertain (Chinese background (20) and Korean workers (29), Table 2).

For solid cancers 16 of 22 studies reported a positive ERR, leading to rejection of the hypothesis that the median ERR per unit dose equals zero ( $p=0.03$ ) from the sign test (Supplementary Table 1). After exclusion of the four studies for which bias adjustment could move a positive ERR towards null (Canadian cardiac imaging (25), Korean workers (29), Japanese nuclear workers (33) and USRT breast cancer (12)), 12 of the remaining 18 studies were positive ( $p=0.12$ ). For leukemia, 17 of 20 studies were positive and we rejected the hypothesis that the median ERR per unit dose equals zero ( $p=0.001$ ) (Supplementary Table 2). This conclusion was not changed by the exclusion of five studies (Chornobyl residents (16), Chinese background (20), Korean workers (29), Chornobyl liquidators (17) and Ukrainian Chornobyl liquidators (18)) for which bias adjustment could move a positive ERR towards null ( $p=0.02$ ).

Power to reject the null for these studies was evaluated under A-bomb survivor-based alternative hypotheses. For all cancers except leukemia, studies with a statistically significantly elevated ERR (at the 5% level) had reasonable power ( $\geq 50\%$ ), while studies with statistically non-significant ERR estimates had low power ( $< 50\%$ ). The only exception was the basal cell carcinoma analysis within the USRT study (11), with a statistically non-significant negative ERR estimate in the presence of reasonable power (Table 1). For leukemia, the pattern was less clear. Although the power of studies with statistically non-significant ERR estimates was generally low, three studies with statistically significant ERR estimates were estimated to have low power (UKNRRW (30),

UK pediatric CT (27) and Taiwanese residents (24)) and two studies to have reasonable power (Great Britain [GB] background (21) and INWORKS (8)) (Table 2).

For all solid cancers following adulthood exposure, homogeneity was not rejected only after exclusion of the Canadian cardiac imaging study (25) ( $p=0.20$ ), which introduced statistically significant heterogeneity due to the very small standard deviation in relation to the size of the ERR. Based on the remaining 15 studies, the meta-ERR at 100 mGy was 0.030 (95% CI: -0.020, 0.081) with 18% of the variability explained by heterogeneity (Figure 1; Table 4). For leukemia following adulthood exposure ( $n=14$  studies), the meta-ERR at 100 mGy was 0.16 (95% CI: 0.07, 0.25) with no indication of heterogeneity ( $p=0.98$ ) (Table 4; Figure 2). Homogeneity was not rejected for 6 studies of leukemia after childhood exposure ( $p=0.27$ ) with a meta-ERR at 100 mGy of 2.84 (95% CI: 0.37, 5.32) (Table 4; Figure 3).

## **Discussion**

This summary report combines the results of our detailed assessments of potential biases in risk estimates for the 26 eligible human studies on low dose radiation exposure and cancer risk. Most of the studies reported positive ERRs; 16 of 22 studies of solid cancers and 17 of 20 studies of leukemia. After a systematic assessment of methodological issues that may be associated with the potential for bias in the risk estimate, we concluded that only a small subset of the positive studies had biases where adjustment would move the risk estimate towards the null. The sign test rejected the hypothesis of no radiation effect for solid cancers and leukemia; for leukemia, this is true even after exclusion of the positive studies that were possibly biased away from the null ( $n=5$ ). For solid cancer, the evidence was borderline after exclusion of positive studies that were possibly biased away from the null ( $n=4$ ). Finally, a meta-analysis of the published risk estimates yielded

statistically significantly elevated risks for leukemia separately among children and among adults, and a borderline statistically significantly elevated risk for solid cancers among adults.

For solid cancers following adulthood exposure, the summary risk estimate from our meta-analysis of 0.030 was very similar to the recent estimate from the Life Span Study for males of 0.027 per 100 mGy but lower than the estimate for females of 0.064 (41). As most of the studies of solid cancers following adulthood exposure were from nuclear workers, the comparison with the males from the Life Span Study is probably most appropriate. For leukemia our summary risk estimate for adulthood exposure of 0.16 per 100 mGy is double the most recent risk estimate from the Life Span Study of 0.08 per 100 mGy for males and females combined, although statistically compatible with the 95% CI (0.003, 0.19) (42). Our estimate is very similar to the meta-estimate of 0.19 (95% CI: 0.07, 0.32) based on 10 studies of protracted exposure to low-dose radiation (43), which included earlier follow-up of cohorts in several studies also included in this report (9, 17, 36).

Several studies were ineligible for our review (6). Some details are presented for studies that had conducted an internal dose-response analysis but were ineligible because they then failed on an additional criterion (Table 3). One reason for exclusion was that the mean cumulative dose exceeded 100 mGy. This led to the exclusion of (borderline) statistically significantly positive studies such as Techa River (46), Mayak worker (47), US scoliosis (44), Chornobyl liquidators (48) and Chinese medical workers (49), and the statistically non-significantly negative study of background exposure in Kerala (45). While the proportion of subjects with cumulative doses exceeding 100 mGy in these excluded studies ranged between 20% and 80%, it was below 10% among all included studies except for 5 studies (16-18, 20, 23) for which it ranged between 11% and 22% (6). Excluding these latter 5 studies from our test whether the median ERR per unit dose

equals 0 still resulted in rejection of this hypothesis (solid cancer:  $p=0.02$  for 15 of 20 positive studies, leukemia:  $p=0.01$  for 13 of 16 positive studies). As an alternative, we recommend future analyses of individual subjects exposed to cumulative doses below 100 mGy, as recently done for leukemia after childhood exposure (57). The seven studies excluded because they only published risk estimates for categories of dose were mostly null (50-54). It is possible that the non-increasing categorical risks were the reason for not presenting risks per continuous dose, which could be a form of reporting bias. Finally, two background radiation studies used dose rate instead of cumulative dose (55,56). Both were largely null.

A recent review conducted by the National Council on Radiation Protection and Measurements (NCRP) (58) assessed the different, albeit related question, of whether the recent epidemiological data from 29 low-dose and low-dose rate studies support the linear-no-threshold (LNT) model for radiation protection purposes. There were considerable differences between the studies included in the NCRP review and ours; only 12 of the studies were in both reviews. One reason for the difference in eligible studies is that we restricted our review to low-dose studies, defined as mean cumulative dose  $<100$  mGy. The approach to reviewing the evidence was also different. NCRP adopted the traditional approach to assessing study quality and excluded several studies due to their being classified as low quality. As we have shown in this monograph, there is not always a direct relationship between study quality and bias. Some of the studies that were classified as high quality could still have been subject to bias, e.g., the Japanese nuclear workers study (33), and those judged as of poor quality are not necessarily biased, e.g., the GB background study (21) and the Taiwanese residents study (24). We showed formally that mostly these methodological issues and errors are unlikely to have resulted in biases that would have impacted the interpretation of the study results. Nevertheless, after their exclusions the NCRP committee still concluded that

“twenty studies [of the 29 reviewed, with the committee designating four studies as inconclusive] (80%) provided some support for the LNT model, including five studies (20%) providing strong support and four (16%) providing moderate support” (58). As mentioned previously, the BEIR VII (1) and other earlier reviews of evidence for cancer risks from low doses mostly depended on epidemiological studies of higher dose exposures and then invoked the linear no-threshold assumption from experimental data to support the conclusion that low doses are likely to cause cancer. Although we cannot rule out the possibility that the risks are influenced by higher doses, our results derived from studies with a mean cumulative dose of <100 mGy apply to populations that are mostly exposed to low doses and therefore directly address the question of whether there is epidemiological evidence for cancer risks from low-dose exposures. Despite the different approaches and different studies included, the two previous reviews (1, 58) and the pooled analysis of leukemia after childhood exposure to doses below 100 mGy (57) are all in general agreement with our result that there is evidence of cancer risks from low-dose ionizing radiation. Our summary risk estimates from the meta-analysis are broadly consistent with the Life Span Study of atomic bomb survivors (41, 42).

The observation that most of the study findings were positive ( $ERR > 0$ ) raises the question of whether there could be publication bias. As these epidemiological studies required extensive effort and all had the primary (and usually the only) aim of evaluating whether low-dose ionizing radiation causes cancer, the likelihood that a null or statistically non-significant result would not be published is likely quite low in our view. Furthermore, the field of radiation epidemiology is relatively small and we are not aware of recent studies or updates that were not published, although a few were published or submitted for publication after the study period.

Confidence intervals and study power need to be considered in interpreting study findings. The confidence intervals for ERRs from all eligible studies included positive values; for the positive estimated ERRs, the upper confidence limits were often several times the estimate. Thus, even when the null hypothesis could not be rejected, findings were compatible with positive effects. It should also be noted that a statistically significant estimate in a low power study may reflect a false positive finding and suggest that the estimated ERR is biased. However, one should be cautious about the assessments of statistical power reported here (Table 1 and 2) (and see also Gilbert et al (3)) as they are based on summary data reported by the study and several assumptions, most importantly that the ERR from the Life Span Study can be transported to other populations.

To our knowledge, this monograph provides the first systematic assessment of the impact of methodological issues and errors on the risk estimate in the studies of low-dose radiation exposure eligible for our analysis. More generally, we used a novel approach in this systematic review to formally assess biases based on published data combined with epidemiological and statistical theory. Traditionally, systematic reviews classify the quality of a study but without formally considering whether the quality of information translates into a bias. For example, low quality dosimetry does not automatically result in bias, e.g., the GB background study (21). In addition, the direction of the bias, and, if possible, the magnitude of the potential bias needs to be assessed. Without these further considerations exclusions based on quality or potential bias could result in substantial loss of information. Such an approach has been recently recommended over other approaches, such as use of a “risk of bias” checklist (59).

There were, however, several limitations to our review. Primarily because we were working only with published data, there were several instances where we had insufficient information to assess the direction of the bias. As examples, this occurred in assessing the direction of confounding if

the relationship between the exposure and the confounder was not published, in the lack of quantitative estimates about the completeness of follow-up, and in absence of information about the completeness and accuracy of vital statistics and cancer registry data for studies using linkage of the cohort with such databases. We also could not assess the magnitude of the bias in most situations because the required data were not available in the publications. Our summary analysis was therefore limited to the simple sign test, which does not take account of the size or precision of the estimated ERRs, since alternative tests are based on the actual values (t-test) or their ranks (Wilcoxon signed rank test). We have made several recommendations in each of our papers that would facilitate assessments of bias in the risk estimates in the future. For example, we recommend the routine publication of assessments of dose uncertainty and levels of loss to follow-up by exposure and outcome. Some of the data to inform our assessments came from sub-studies and we assumed that these applied to the full study population. For example, there was no evidence that smoking was related to occupational radiation exposure in a case-control study of leukemia in U.S. nuclear workers (60). As smoking data were not available for the full cohort, we assumed the findings from this sample were generalizable. In addition, the sign test requires that the distribution of the estimated ERRs per unit dose (under the assumption that the true ERRs are zero) has a median of zero, although it does not depend on the ERRs following approximately normal distributions. Our assessment of whether a study was positive or null as input for the sign test did not account for the width of the confidence intervals. Many of these studies had wide confidence intervals, which overlapped zero and therefore one could argue that bias adjustment which moves a statistically non-significant positive ERR towards the null does not change the conclusion. Results on all solid cancers combine a diverse group of cancer sites. This heterogeneity has to be taken into account when comparing results with those for leukemia. Finally, the meta-analyses



included all studies, even those identified as potentially biased and were based on the assumption that the estimated ERRs from the individual studies were approximately normally distributed. With the highly skewed dose distributions and small numbers of cases in many of these studies, these approximations may not be accurate (3). Thus, tests and confidence intervals could be distorted, especially since the meta-analyses are based on the Wald method. Furthermore, our assessment necessarily had a subjective component.

Age at exposure and time since exposure are important effect modifiers of cancer risk from ionizing radiation exposure, and the eligible studies were heterogeneous in these respects. A pooled analysis with individual subject data is necessary, therefore, to quantify the risk accounting for these effect modifiers. The childhood thyroid pooling project included here (28) and the more recently published pooled analysis of leukemia after childhood radiation exposure (57) provide the most reliable estimates of these cancer risks that account for these modifiers. Similarly, the INWORKS study (9) reviewed here provides a reliable, and we judge minimally biased, estimate of the solid cancer and leukemia risks from adulthood occupational radiation exposures incorporating age and time since exposure (8, 9).

It is well established that moderate to high doses of ionizing radiation cause cancer. As discussed in the paper by Gilbert et al (3), the available low dose human data also now satisfy most of Bradford Hill's viewpoints on causality: consistency, temporality, biological gradient, plausibility and coherence (61). The first of Bradford Hill's viewpoints, strength of effect or, in other words, magnitude of association, is usually not met in studies of cancer risk after low-dose radiation exposure. Large effects are an important indicator of causality since they are less likely to be entirely due to bias. However, large effects are not considered a necessary aspect and indeed Bradford Hill explicitly cautions against dismissal of weak observed associations. Our systematic

assessment in this monograph showed that these epidemiological studies of low-dose radiation and cancer risk are characterized by several limitations, but we found that only a small minority of the studies had biases whose correction could have moved a positive ERR towards the null. After exclusion of these studies, the majority of studies still reported positive risk estimates. We therefore conclude that there is now a large body of epidemiological data which supports excess cancer risks from low-dose ionizing radiation and the magnitude of the excess relative cancer risk from these low dose studies is statistically compatible with the atomic bomb survivors.

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Figure 1: Meta-analysis of the excess relative risk (ERR) at 100 mGy for all solid cancers after adulthood radiation exposure. The size of the ERR symbol is proportional to the inverse variance of the study-specific ERR.

Figure 2: Meta-analysis of the excess relative risk (ERR) at 100 mGy for leukemia after adulthood radiation exposure. The size of the ERR symbol is proportional to the inverse variance of the study-specific ERR.

Figure 3: Meta-analysis of the excess relative risk (ERR) at 100 mGy for leukemia after childhood radiation exposure. The size of the ERR symbol is proportional to the inverse variance of the study-specific ERR.

**Table 1: Assessment of bias from several sources for studies of solid cancers**

		ERR per unit of dose bias						
Study name	Reference	ERR at 100 mGy	95% CI	Dose error	Confounding/selection bias	Outcome misclassification	Could bias adjustment move ERR towards null?*	Estimated power†
<b>Environmental</b>								
Three Mile Island	Han et al. 2011 (19)	-1	(-6 to 3)	↓			Yes. Adjustment possibly moves ERR towards null.‡	NC
Chinese background	Tao et al. 2012 (20)	-0.101	(-0.253 to 0.095)			↕	Uncertain. Adjustment could move ERR towards or away from null.	Low
GB background	Kendall et al. 2013 (21)	2	( -2.0 to 6.0)					NC
Swiss background	Spycher et al. 2015 (22)	2.8	(0.8 to 4.8)					NC
Techa River	Davis et al. 2015 (23)	0.077	(0.013 to 0.150)					Reasonable
Taiwanese residents	Hsieh et al. 2017 (24)	0.04	(0.01 to 0.08)‡					Low
<b>Medical</b>								
Canadian cardiac imaging	Eisenberg et al. 2011 (25)	0.3	(0.2 to 0.4)		↕	↑	Uncertain. Adjustment could move ERR towards or away from null.	NC
French Pediatric CT (brain tumors)	Journey et al. 2016 (26)	0.7	(-0.1 to 1.0)					Low
UK Pediatric CT (brain tumors)	Berrington et al. 2016 (27)	1.2	(0.4 to 3.1)					Reasonable
PIRATES (thyroid cancer)	Lubin et al. 2017 (28)	0.96	(0.37 to 1.70)					Reasonable

# Occupational

Korean workers	Ahn et al. 2008 (29)	0.72	(-0.5 to 2.1)‡	↓	↑	Uncertain. Adjustment could move ERR towards or away from null.	NC
UKNRRW	Muirhead et al. 2009 (30)	0.03	(0 to 0.056)	↓		No. Adjustment would move ERR away from null.	Reasonable
Korean nuclear workers	Jeong et al. 2010 (31)	0.21	(-0.19 to 0.9)	↓		No. Adjustment would move ERR away from null.	Low
Rocketdyne workers	Boice et al. 2011 (32)	0.02	(-0.18 to 0.17)	↓		No. Adjustment would move ERR away from null.	Low
Japanese workers	Akiba et al. 2012 (33)	0.13	(-0.03 to 0.30)	↕	↓	Uncertain. Adjustment could move ERR towards or away from null.	Low
Canadian nuclear workers§	Zablotska et al. 2014 (34)	-0.12	(<-0.15 to 0.24)	↓		Yes. Adjustment possibly moves ERR towards null.¶	NC
German nuclear workers	Merzenich et al. 2014 (35)	-0.1	(-0.4 to 0.1)	↓		Yes. Adjustment possibly moves ERR towards null.¶	NC
US nuclear workers	Schubauer-Berigan et al. 2015 (36)	0.01	(-0.02 to 0.05)	↓		No. Adjustment would move ERR away from null.	Low
INWORKS	Richardson et al. 2015 (9)	0.047	(0.018 to 0.079)‡	↓		No. Adjustment would move ERR away from null.	Reasonable
USRT (breast cancer)	Preston et al. 2016 (12)	0.07	(-0.005 to 0.19)	↕		Uncertain. Adjustment could move ERR towards or away from null.	Low
USRT (brain cancer)	Kitahara et al. 2017 (10)	0.1	(<-0.3 to 1.5)				Low
USRT (skin cancer)	Lee et al. 2015 (11)	-0.001	(-0.04 to 0.05)				Reasonable
French nuclear workers	Leuraud et al. 2017 (37)	0.04	(-0.04 to 0.13)‡	↓		No. Adjustment would move ERR away from null.	Low

\*Reflects an assessment of the presence of bias of the ERR and its likely direction, but not statistical significance

‡Low: <50%, reasonable: ≥50% based on LSS ERR and published dose distributions.

‡90%CI

§The Canadian Study is restricted to the cohort excluding early AECL workers.

||If bias is sufficiently large and additive, adjustment could increase the ERR above null. Since the magnitude of bias is difficult to determine from the published reports, we conservatively assume that adjustment of bias will not increase the ERR to positive.

NC = Not calculated because dose distributions needed for power calculations were not available.

↓ Bias in the negative direction

↑ Bias in the positive direction

↑↓ Bias of unclear direction

**Table 2: Assessment of bias from several sources for studies of leukemia**

Study name	Reference	ERR at 100 mGy	95% CI	Dose error	ERR per unit of dose bias		Could bias adjustment move ERR towards null?†	Estimated power*
					Confounding/ selection bias	Outcome misclassification		
<b>Environmental</b>								
Chornobyl residents	Davis et al. 2006 (16)	3.2	(0.9 to 8.4)	↑	↑		Yes. Exclusion of subgroup with potential recall bias reduced risk to null.	NC
Three Mile Island	Han et al. 2011 (19)	19	(−3 to 45)	↓			No. Adjustment would move ERR away from null.	NC
Chinese background	Tao et al. 2012 (20)	1.068	(<0 to inf)			↕	Uncertain. Adjustment could move ERR towards or away from null.	Low
GB background	Kendall et al. 2013 (21)	12	(3.0 to 22.0)					Reasonable
Swiss background	Spycher et al. 2015 (22)	3.6	(−0.3 to 7.7)					Low
Finnish background	Nikkila et al. 2016 (38)	−3	(−11 to 6)					NC
Taiwanese residents	Hsieh et al. 2017 (24)	0.15	(0.03 to 0.24)‡					Low
<b>Medical</b>								
French Pediatric CT	Journey et al. 2016 (26)	1.6	(−2.3 to 2.7)					Low
UK Pediatric CT	Berrington et al. 2016 (27)	3	(0.3 to 10.9)					Low
<b>Occupational</b>								
Korean workers	Ahn et al. 2008 (29)	1.68	(−3.4 to 14.9)‡		↓	↑	Uncertain. Adjustment could move ERR towards or away from null.	NC
Chornobyl liquidators	Kesminiene et al 2008 (17)	0.5	(−0.38 to 5.70)‡	↑			Yes. Adjustment possibly moves ERR to null.	Low



UKNRRW Rocketdyne workers	Muirhead et al. 2009 (30)	0.18	(-0.006 to 0.50)						Low
	Boice et al. 2011 (32)	0.06	(-0.50 to 1.23)						Low
Japanese workers	Akiba et al. 2012 (33)	-0.19	(-0.61 to 0.86)		↓		↓	Yes. Adjustment possibly moves ERR towards null. <sup>ll</sup>	Low
Ukrainian Chornobyl liquidators	Zablotska et al. 2013 (18)	0.221	(0.005 to 0.761)	↑				Yes. Adjustment possibly moves ERR towards null.	NC
Canadian nuclear workers§	Zablotska et al. 2014 (34)	1.44	(<-0.15 to 14.6)						NC
German nuclear workers	Merzenich et al. 2014 (35)	0.4	(-0.3 to 1.1)		↓			No. Adjustment would move ERR away from null.	NC
US nuclear workers	Schubauer-Berigan et al. 2015 (36)	0.17	(-0.02 to 0.47)						Reasonable
INWORKS	Richardson et al. 2015 (9)	0.3	(0.12 to 0.52)‡						Reasonable
US atomic veterans	Caldwell et al. 2016 (39)	-0.5	(-14 to 4)						Low
French nuclear workers	Leuraud et al. 2017 (37)	0.35	(<0 to 1.6)‡		↓			No. Adjustment would move ERR away from null.	Low

\*Low: <50%, reasonable: ≥50%.

†Reflects an assessment of the presence of bias of the ERR and its likely direction, but not statistical significance

‡90%CI

§The Canadian Study is restricted to the cohort excluding early AECL workers.

llIf bias is sufficiently large and additive, adjustment could increase the ERR above null. Since the magnitude of bias is difficult to determine from the published reports, we conservatively assume that adjustment of bias will not increase the ERR to positive.

NC = Not calculated because dose distributions needed for power calculations were not available.

↓ Bias in the negative direction

↑ Bias in the positive direction

**Table 3: Studies with an internal dose-response analysis that were ineligible for only one reason, reasons for exclusion, and summary of findings**

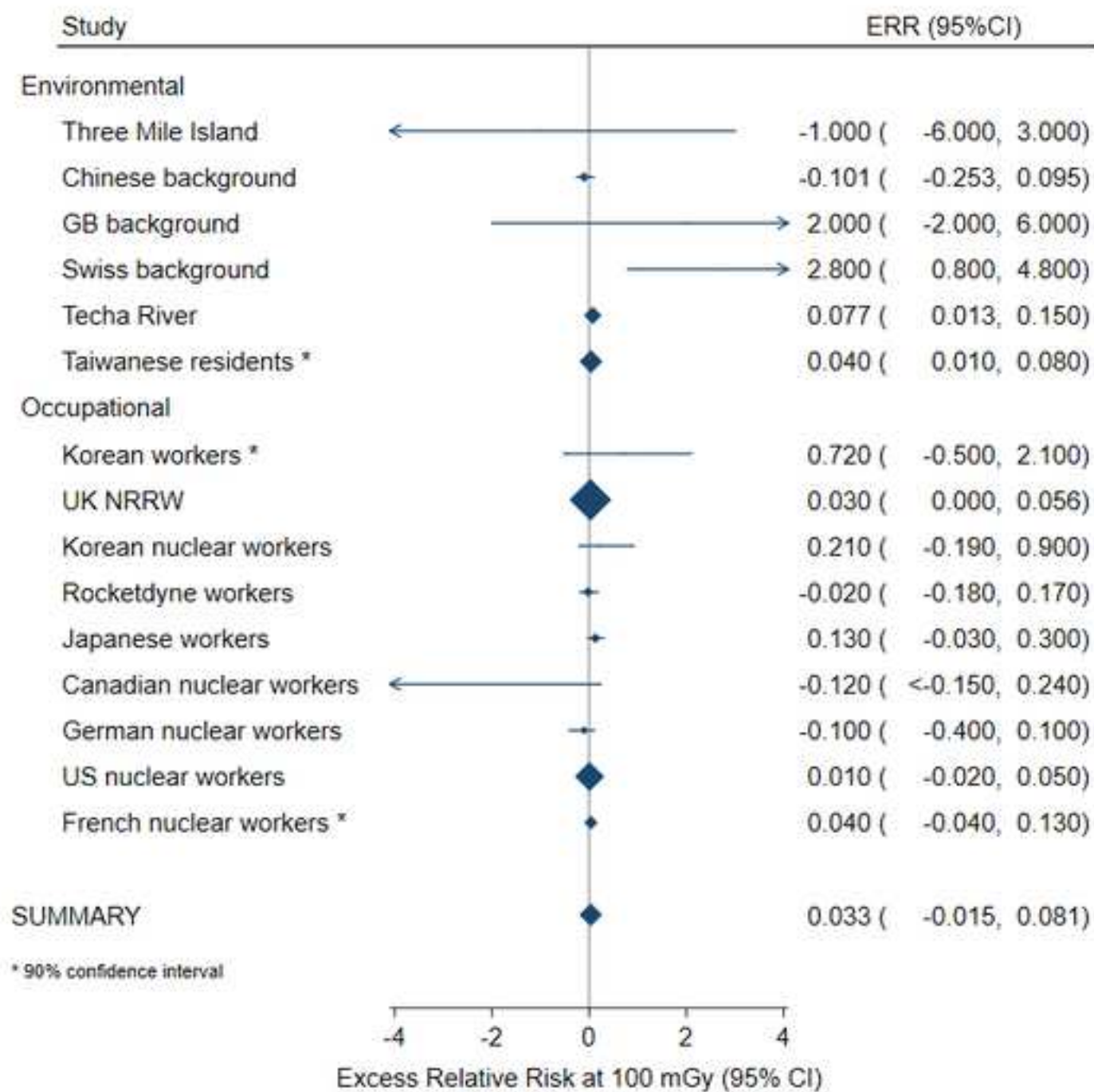
Population	1st author	Publication year	Reason for exclusion	Findings for solid cancers and leukemia (or primary cancer site)
Breast cancer in US scoliosis	Ronckers (44)	2008	Mean cumulative dose = 120 mGy	Borderline statistically significant positive dose-response relationship for breast cancer incidence (p-trend=0.06).
Kerala background	Nair (45)	2009	Mean cumulative dose = 161 mGy	Statistically non-significant negative dose-response for incidence of all solid cancers excluding leukemia (p-trend>0.5).
Techa river	Krestinina (46)	2013	Mean cumulative dose = 410 mGy	Statistically significant positive dose-response for incidence of all leukemias (p-trend<0.001).
Mayak workers	Sokolnikov (47)	2015	Mean cumulative dose = 354 mGy	Statistically significant positive dose-response for mortality of all solid cancers excluding lung, liver and bone cancers (p-trend=0.01).
Chornobyl clean-up workers	Kashcheev (48)	2015	Mean cumulative dose = 132 mGy	Statistically significant positive dose-response for total cancer incidence (p-trend=0.03) and mortality (p-trend=0.05).
Chinese medical workers	Sun (49)	2016	Mean cumulative dose = 250 mGy	Statistically significant positive dose-response for incidence of all solid cancers (p-trend=0.002).
US Shipyard workers	Matanoski (50)	2008	Categorical risk estimates	Statistically non-significant increased risk of leukemia mortality in highest dose category. No trend tests presented.
Australian nuclear test	Gun (51)	2008	Categorical risk estimates	No increased risk of solid cancer incidence or leukemia incidence across dose-categories (p-trend>0.05).
French biology researchers	Guseva (52)	2008	Categorical risk estimates	Statistically significant increasing trend for all cancer deaths across dose-categories (p=0.03 5-year lag).
Finnish reindeer herders	Kurtio (53)	2010	Categorical risk estimates	No overall increased risk of cancer incidence across dose-categories (p-trend=0.28), but statistically significant increased risk with dose for exposure <age 15 years (p-trend=0.003).
Childhood X-rays	Hammer (54)	2009	Categorical risk estimates	No increased risk of solid cancer incidence (p-trend=0.32) or leukemia incidence (p-trend=0.26) across dose-categories.
French background (Geocap case-control)	Demoury (55)	2017	Risk for dose rate not cumulative dose	Acute leukemia incidence not related to background radiation exposure (p-trend>0.05).
German background	Spix (56)	2017	Risk for dose rate not cumulative dose	Statistically non-significant positive increased risk of lymphoid leukemia incidence (p-trend=0.54).

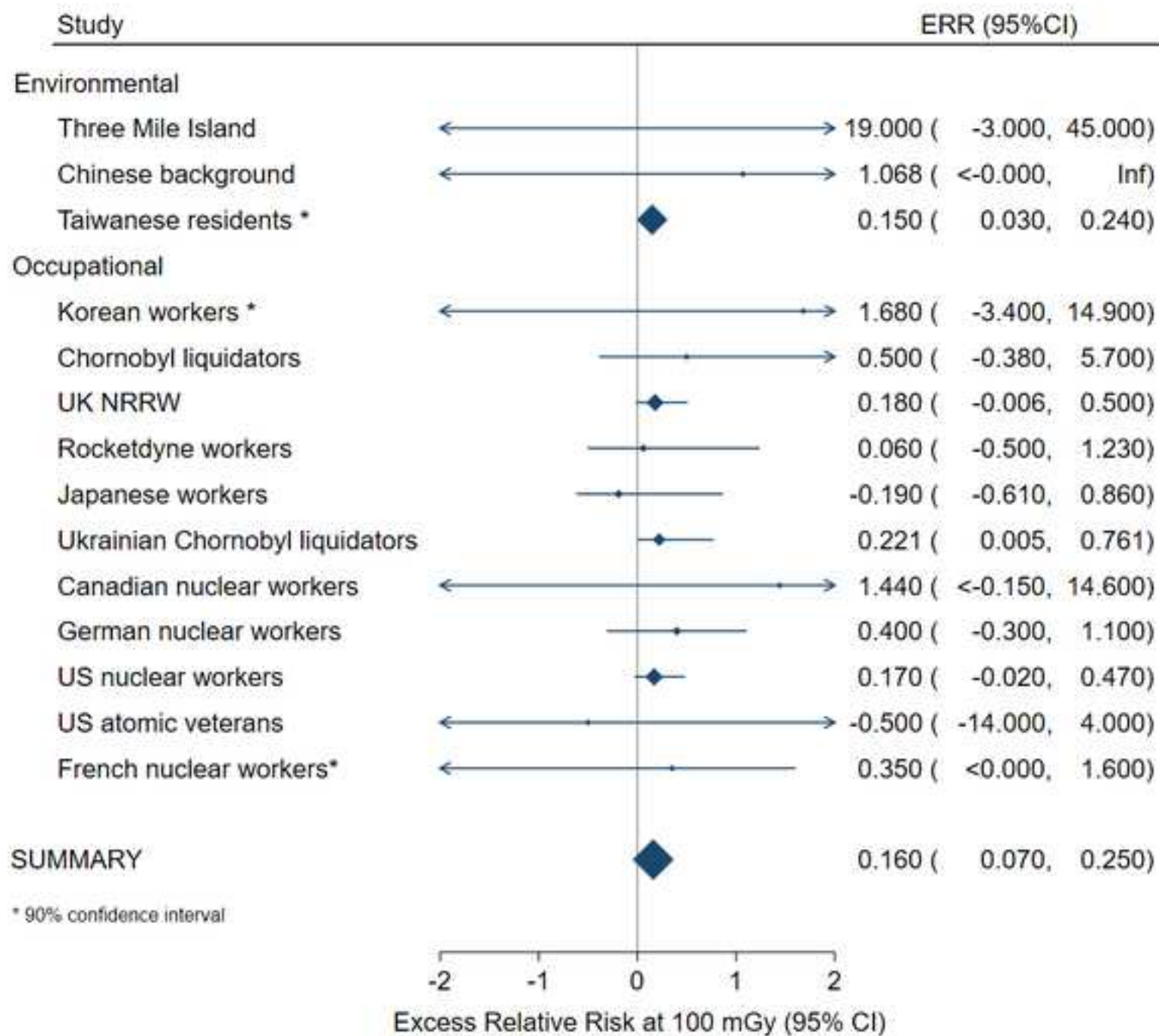
**Table 4: Meta-analysis of excess relative risks (ERR) per 100mGy for all solid cancers and leukemia**

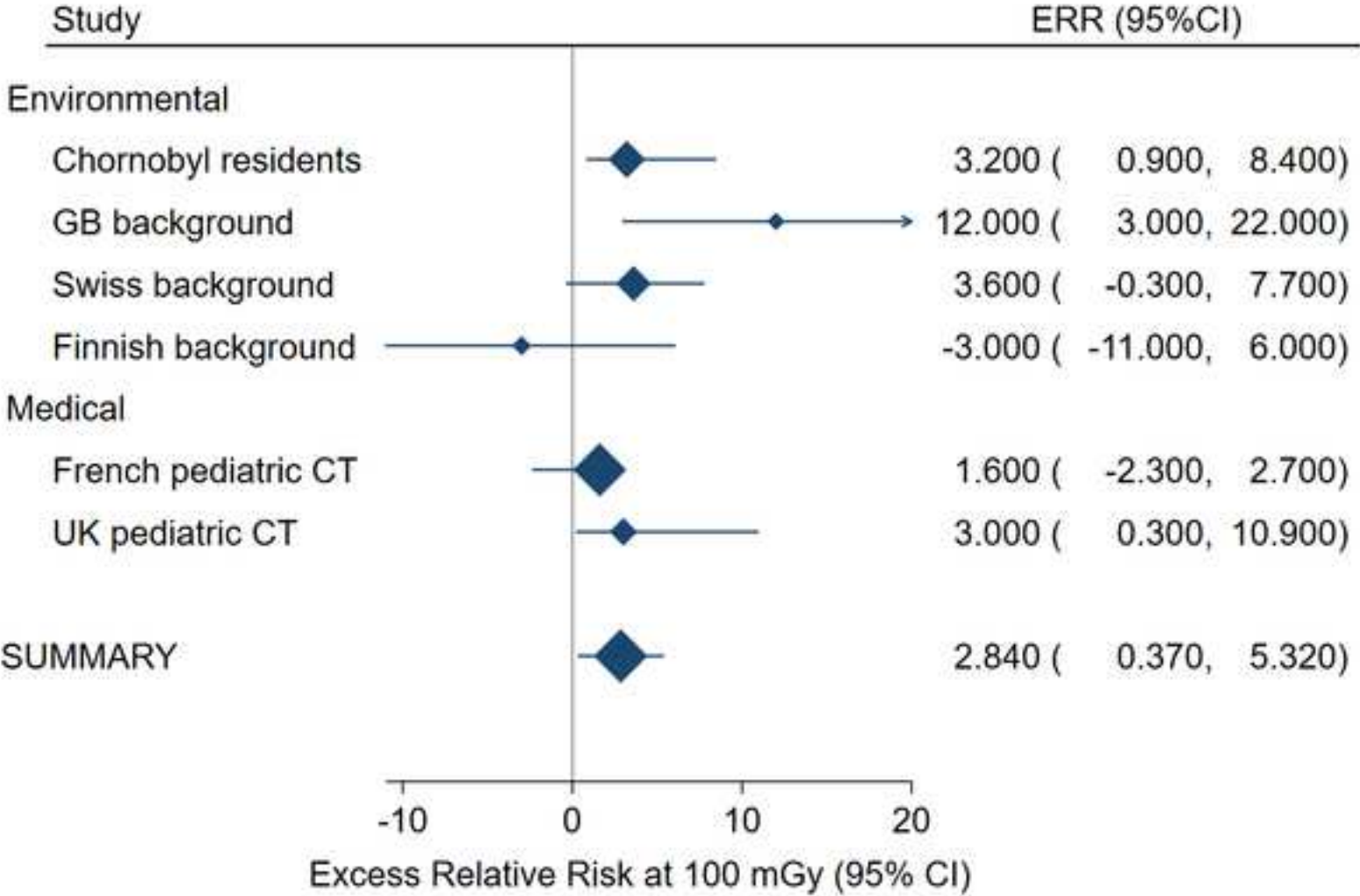
Outcome	Studies	ERR/100 mGy (95% CI)	p-value	Cochran's Q (p-value)	I <sup>2</sup>
Adult solid cancer	16*	0.061 (-0.024, 0.145)	0.08	45.13 (<0.001)	0.67
Adult solid cancer excluding the Canadian cardiovascular imaging study due to heterogeneity	15*	0.033 (-0.015, 0.081)	0.11	17.90 (0.21)	0.22
Adult leukemia	14†	0.160 (0.070, 0.250)	<0.001	4.12 (0.99)	NA
Childhood leukemia	6	2.840 (0.370, 5.320)	0.01	6.40 (0.27)	0.22

\*Excluding INWORKS and site-specific results from USRT

†Excluding INWORKS









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pl\\_tables.docx](#)



# **Epidemiological Studies of Low-dose Ionizing Radiation and Cancer: Summary Bias**

## **Assessment and Meta-Analysis**

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## Abstract

**Background:** Ionizing radiation is an established carcinogen, but risks from low-dose exposures are controversial. Since the Biological Effects of Ionizing Radiation (BEIR) VII review of the epidemiological data in 2006 many subsequent publications have reported excess cancer risks from low-dose exposures. Our aim was to systematically review these studies to assess the magnitude of the risk, and whether the positive findings could be explained by biases.

**Methods:** Eligible studies had mean cumulative doses <100 mGy, individualized dose estimates, risk estimates and confidence intervals (CI) for the dose-response and were published in 2006-2017. We summarized the evidence for bias (dose error, confounding, outcome ascertainment) and its likely direction for each study. We tested whether the median excess relative risk (ERR) per unit dose equals zero and assessed the impact of excluding positive studies with potential bias away from the null. We performed a meta-analysis to quantify the ERR and assess consistency across studies for all solid cancers and leukemia.

**Results:** Of the 26 eligible studies, 8 concerned environmental, 4 medical and 14 occupational exposure. For solid cancers, 16 of 22 studies reported positive ERRs per unit dose and we rejected the hypothesis that the median ERR equals zero ( $p=0.03$ ). After exclusion of four positive studies with potential positive bias, 12 of 18 studies reported positive ERR per unit dose ( $p=0.12$ ). For leukemia, 17 of 20 studies were positive and we rejected the hypothesis that the median ERR per unit dose equals zero ( $p=0.001$ ), also after exclusion of five positive studies with potential positive bias ( $p=0.02$ ). For adulthood exposure, the meta-ERR/100 mGy was 0.030 (95% CI: -0.02, 0.081) for solid cancers and 0.16 (95% CI: 0.07, 0.25) for leukemia. For childhood exposure, the meta-

ERR/100 mGy for leukemia was 2.84 (95% CI: 0.37, 5.32); there were no eligible studies of all solid cancers.

Conclusions: Our systematic assessments in this monograph showed that these new epidemiological studies are characterized by several limitations, but only a few positive studies were potentially biased away from the null. After exclusion of these studies, the majority of studies still reported positive risk estimates. We therefore conclude that these new epidemiological studies directly support excess cancer risks from low-dose ionizing radiation. Furthermore, the magnitude of the cancer risks from these low-doses radiation exposures was statistically compatible with the radiation dose-related cancer risks of the atomic bomb survivors.

## Introduction

The evidence for cancer risks provided by epidemiological studies of low-dose ionizing radiation exposure is of key relevance to radiation protection since a large fraction of the population is exposed to low doses of ionizing radiation from diagnostic medical procedures or occupationally, in addition to natural background radiation. Careful and sophisticated interpretation of the results from these studies is required, however, because the risks are likely to be small compared to those associated with non-radiation risk factors, studies may have power below 80% (the conventional threshold of adequacy in this respect), dose estimation may be limited and/or retrospective and studies may suffer from biases typical of observational studies such as confounding.

The last major US review of the epidemiological and experimental evidence for cancer risks from low-dose exposures (which we denote as  $<100$  mGy) in 2006 concluded that “the available scientific evidence is consistent with a linear dose-response relationship between ionizing radiation and the development of cancer in humans” (1). This conclusion was largely based on studies of populations exposed to higher doses combined with experimental data. Subsequent to 2006, several new epidemiological studies of populations exposed primarily to low doses have been published and several existing studies have reported new results from extended follow-up. Most of these new publications report excess cancer risks from low dose radiation exposures. The aim of this monograph is to systematically evaluate whether there is direct human evidence of excess cancer risks from low-dose ( $<100$  mGy) radiation exposure, and, if so, what the magnitude of the risk is and whether the positive findings could be explained by biases.

Here, we provide a synthesis from our in-depth systematic assessments of the methodology for the eligible studies published during 2006-2017 that we evaluated, and the associated potential for the risk estimates to be biased due to dose error, confounding, selection bias, or outcome

misclassification (2-5). We assess for each study the direction of the biases from any of these sources. As a general evaluation of whether the studies support cancer risks from low-dose ionizing radiation we conducted a sign test for whether the median of the excess relative risks (ERRs) equals zero, and then calculated the impact of excluding the positive studies identified as being biased away from the null. Finally, to quantify the magnitude of the estimated risks, we conduct a meta-analysis for all solid cancers and for leukemia from childhood or adulthood exposure to low-dose ionizing radiation.

## **Methods**

We included epidemiological studies published since the BEIR VII report in 2006 (1) and before 2018. Studies were eligible for inclusion if they were based on human populations exposed to low dose, predominantly low linear energy transfer (LET), radiation (mean cumulative dose < 100 mGy). We required individualized dose estimates for the study participants, and that the publications provided risk estimates and confidence intervals (CI) for the dose-response for cumulative radiation dose. For a full description of the eligible studies see the overview paper in this monograph (6). In brief, we used the results of the methodological assessments conducted in this monograph (dosimetry (2), confounding (5), outcome (4)) to derive, for each aspect, an assessment of the potential for bias in the risk estimate and the direction of the bias.

Summarizing the evaluations of the different study aspects was as follows. We assessed the strengths and weaknesses of dosimetry systems with respect to the directness, complexity, and completeness of the dosimetry, the dosimetric uncertainty and the validity of dose estimates (2). This process identified studies with a known or suspected bias in dose estimates and the likely direction of the bias in the risk estimate.

In assessing the evidence for confounding and selection bias, we summarized methods to control confounding and assessed the likelihood of uncontrolled confounding as well as its direction (5). This assessment was based on available data from the eligible studies and related publications including some examples of quantitative bias assessment to examine the potential magnitude of the bias.

The outcome evaluation paper in this monograph reviewed the possible impact of differential outcome ascertainment across radiation dose levels (4). The evaluation also considered loss to follow-up, under- or over-ascertainment of cancer outcomes, misclassification of outcomes, and changing classifications over time. The main objective was to identify studies whose outcome ascertainment was differential regarding exposure level and hence to bias the relative risk estimate. We then performed a summary of the assessments of different biases for each study and carefully considered both the direction of the observed effect and the direction of the bias. For our general question of whether the studies overall support excess cancer risks (as opposed to the question of the magnitude of the risk) our priority was to identify the positive studies with bias in the positive direction or bias of uncertain direction. If there were several biases acting in different directions, or it was not possible to determine the potential direction, we classified these studies as potentially biased away from the null. Since the magnitude of bias is difficult to determine from the published reports, we conservatively counted any study with a negative ERR estimate as negative regardless of any potential for bias. We also indicated whether the estimated power (if available) was low (<50%) or reasonable ( $\geq 50\%$ ) (3).

We performed a one-sided sign test for the reported ERRs, separately for solid cancers and for leukemia, to evaluate the hypothesis that the median of the ERRs per unit dose equals zero versus the alternative that the median ERR per unit dose exceeds zero (7). The sign test excluded the

INWORKS study (8, 9) because it is a pooled analysis of UK, French and U.S. nuclear workers, which were included separately. To assess the impact of the studies identified as potentially biased we then repeated the sign test after excluding the studies where bias adjustment could move a positive ERR towards the null. A one-sided p-value below 0.05 was considered statistically significant.

Finally, we conducted a meta-analysis of the published ERR estimates at 100 mGy to quantify the magnitude of the risk and to assess the consistency across studies for both all solid cancers and leukemia. Here we excluded the INWORKS study (8, 9) because of overlap as described above and the US radiologic technologists (USRT) studies (10-12) because only site-specific ERRs were reported. All other studies were included regardless of potential bias because as described above we generally could not quantify the magnitude of the bias and hence the magnitude of bias-corrected ERRs, and because excluding the subset of positive studies with positive biases would bias the summary risk estimate away from the null. We generated standard errors of the ERR at 100 mGy based on the upper (UL) and lower (LL) limit of the CI as  $(UL-LL)/2*1.96$  or  $(UL-LL)/2*1.645$ , depending on whether 95% or 90% CI were reported, respectively. We acknowledge that this is an approximation that may not be fully adequate given the skewed dose distributions and small numbers of cases in some of the studies. The meta-ERR estimate was derived from a random effects model using the iterative method of Paule and Mandel (13) as outlined in DerSimonian and Kacker (14). We also computed Cochran's Q, which is the weighted sum of squared differences between individual study effects and the pooled effect across studies, to test homogeneity, as well as the  $I^2$  statistic (variance due to heterogeneity (15)). If homogeneity was rejected, we excluded the studies with the largest signed contribution to the Q statistic. Meta-ERRs were calculated separately for studies on adult solid cancers, adult leukemia and childhood

leukemia. We did not calculate meta-ERR estimates for childhood solid cancers because only site-specific ERRs were reported.

## **Results**

Of the 26 eligible studies, we found that three studies had a known or suspected bias in dose estimates that could bias the risk estimate away from the null due to possible recall or selection bias (Chornobyl residents (16), Chornobyl liquidators (17), Ukrainian Chornobyl liquidators (18), Tables 1 and 2) and one study that was likely biased towards the null (Three Mile Island, leukemia) (19). The direction of potential dose bias was uncertain in the USRT study on breast cancer (12). Among the three case-control studies of leukemia in Chornobyl liquidators (17, 18) and residents (16) individual dose estimates relied extensively on information obtained by interview after case ascertainment. In two of these studies there was evidence from the manuscripts that the risk estimate was reduced after exclusions of a study center (Chornobyl residential childhood leukemia study) (16) and proxy respondents (Ukrainian liquidators leukemia study) (18).

Our evaluation of confounding and selection bias identified several sources that could bias the risk estimates: clinical indication for studies of medical diagnostic exposures during childhood, lifestyle factors for environmental, medical, and occupational studies of adult cancers and for occupational studies, other workplace exposures and healthy worker survivor bias. In addition to the impact of a potential dose error considered above (2), the Chornobyl residential case-control study (16) may also have suffered from control selection bias that may have upwardly biased the risk estimate. We assessed that a potential for uncontrolled confounding biasing the risk estimate towards the null from healthy worker survivor bias was especially likely for the Korean workers (29) and Japanese nuclear workers (33) studies, which did not adjust for socio-economic status



(SES) or duration of employment, and the German nuclear workers study (35), which did not adjust for birth cohort and SES (Tables 1 and 2). For the solid cancer results in the Canadian (34) and German nuclear workers study (35) and the leukemia findings in the Japanese nuclear workers study (33), bias adjustment would move the ERR towards the null. For the Japanese nuclear workers, the direction of the bias was uncertain, given that smoking may be a positive confounder in this cohort (40). Therefore, we could not draw a definitive conclusion on the impact of bias adjustment with the available data.

Four studies may have had cancer ascertainment possibly differential by radiation exposure, which could have biased the risk estimate. These include the Japanese nuclear workers (33) with bias towards the null through early loss to follow-up during periods with higher radiation exposure; the Chinese background study (20) with possible bias away from the null because of a higher loss to follow-up in regions with lower background radiation exposure compared to regions with high background radiation exposure; the cardiovascular imaging study (25), possibly biased away from the null because those undergoing diagnostic/therapeutic procedures with imaging were more likely to have cancer outcomes detected; and the Korean radiation workers population (29), where medical surveillance was required for radiation workers but not for the general population or for the worker comparison group (manufacture of motor vehicles) (Tables 1 and 2).

When considering the three potential biases for the 22 studies of solid cancers, there were three studies with negative ERR estimates where adjustment would likely move the ERR towards the null (Three Mile Island (19), Canadian (34) and German nuclear workers (35)) and four studies with positive ERR estimates where it was uncertain whether adjustment would move the ERR towards or away from the null (Canadian cardiac imaging (25), Korean (29) and Japanese workers (33) and USRT breast cancer (12), Table 1). For leukemia there was one study with a negative

ERR where adjustment would likely have moved the ERR towards the null (Japanese workers (33)), three studies with a positive ERR where adjustment would likely have moved the ERR towards the null (Chornobyl residents (16), Chornobyl liquidators (17) and Ukrainian Chornobyl liquidators (18)), and two positive studies where the direction of the bias was uncertain (Chinese background (20) and Korean workers (29), Table 2).

For solid cancers 16 of 22 studies reported a positive ERR, leading to rejection of the hypothesis that the median ERR per unit dose equals zero ( $p=0.03$ ) from the sign test (Supplementary Table 1). After exclusion of the four studies for which bias adjustment could move a positive ERR towards null (Canadian cardiac imaging (25), Korean workers (29), Japanese nuclear workers (33) and USRT breast cancer (12)), 12 of the remaining 18 studies were positive ( $p=0.12$ ). For leukemia, 17 of 20 studies were positive and we rejected the hypothesis that the median ERR per unit dose equals zero ( $p=0.001$ ) (Supplementary Table 2). This conclusion was not changed by the exclusion of five studies (Chornobyl residents (16), Chinese background (20), Korean workers (29), Chornobyl liquidators (17) and Ukrainian Chornobyl liquidators (18)) for which bias adjustment could move a positive ERR towards null ( $p=0.02$ ).

Power to reject the null for these studies was evaluated under A-bomb survivor-based alternative hypotheses. For all cancers except leukemia, studies with a statistically significantly elevated ERR (at the 5% level) had reasonable power ( $\geq 50\%$ ), while studies with statistically non-significant ERR estimates had low power ( $< 50\%$ ). The only exception was the basal cell carcinoma analysis within the USRT study (11), with a statistically non-significant negative ERR estimate in the presence of reasonable power (Table 1). For leukemia, the pattern was less clear. Although the power of studies with statistically non-significant ERR estimates was generally low, three studies with statistically significant ERR estimates were estimated to have low power (UKNRRW (30),

UK pediatric CT (27) and Taiwanese residents (24)) and two studies to have reasonable power (Great Britain [GB] background (21) and INWORKS (8)) (Table 2).

For all solid cancers following adulthood exposure, homogeneity was not rejected only after exclusion of the Canadian cardiac imaging study (25) ( $p=0.20$ ), which introduced statistically significant heterogeneity due to the very small standard deviation in relation to the size of the ERR. Based on the remaining 15 studies, the meta-ERR at 100 mGy was 0.030 (95% CI: -0.020, 0.081) with 18% of the variability explained by heterogeneity (Figure 1; Table 4). For leukemia following adulthood exposure ( $n=14$  studies), the meta-ERR at 100 mGy was 0.16 (95% CI: 0.07, 0.25) with no indication of heterogeneity ( $p=0.98$ ) (Table 4; Figure 2). Homogeneity was not rejected for 6 studies of leukemia after childhood exposure ( $p=0.27$ ) with a meta-ERR at 100 mGy of 2.84 (95% CI: 0.37, 5.32) (Table 4; Figure 3).

## Discussion

This summary report combines the results of our detailed assessments of potential biases in risk estimates for the 26 eligible human studies on low dose radiation exposure and cancer risk. Most of the studies reported positive ERRs; 16 of 22 studies of solid cancers and 17 of 20 studies of leukemia. After a systematic assessment of methodological issues that may be associated with the potential for bias in the risk estimate, we concluded that only a small subset of the positive studies had biases where adjustment would move the risk estimate towards the null. The sign test rejected the hypothesis of no radiation effect for solid cancers and leukemia; for leukemia, this is true even after exclusion of the positive studies that were possibly biased away from the null ( $n=5$ ). For solid cancer, the evidence was borderline after exclusion of positive studies that were possibly biased away from the null ( $n=4$ ). Finally, a meta-analysis of the published risk estimates yielded

statistically significantly elevated risks for leukemia separately among children and among adults, and a borderline statistically significantly elevated risk for solid cancers among adults.

For solid cancers following adulthood exposure, the summary risk estimate from our meta-analysis of 0.030 was very similar to the recent estimate from the Life Span Study for males of 0.027 per 100 mGy but lower than the estimate for females of 0.064 (41). As most of the studies of solid cancers following adulthood exposure were from nuclear workers, the comparison with the males from the Life Span Study is probably most appropriate. For leukemia our summary risk estimate for adulthood exposure of 0.16 per 100 mGy is double the most recent risk estimate from the Life Span Study of 0.08 per 100 mGy for males and females combined, although statistically compatible with the 95% CI (0.003, 0.19) (42). Our estimate is very similar to the meta-estimate of 0.19 (95% CI: 0.07, 0.32) based on 10 studies of protracted exposure to low-dose radiation (43), which included earlier follow-up of cohorts in several studies also included in this report (9, 17, 36).

Several studies were ineligible for our review (6). Some details are presented for studies that had conducted an internal dose-response analysis but were ineligible because they then failed on an additional criterion (Table 3). One reason for exclusion was that the mean cumulative dose exceeded 100 mGy. This led to the exclusion of (borderline) statistically significantly positive studies such as Techa River (46), Mayak worker (47), US scoliosis (44), Chornobyl liquidators (48) and Chinese medical workers (49), and the statistically non-significantly negative study of background exposure in Kerala (45). While the proportion of subjects with cumulative doses exceeding 100 mGy in these excluded studies ranged between 20% and 80%, it was below 10% among all included studies except for 5 studies (16-18, 20, 23) for which it ranged between 11% and 22% (6). Excluding these latter 5 studies from our test whether the median ERR per unit dose

equals 0 still resulted in rejection of this hypothesis (solid cancer:  $p=0.02$  for 15 of 20 positive studies, leukemia:  $p=0.01$  for 13 of 16 positive studies). As an alternative, we recommend future analyses of individual subjects exposed to cumulative doses below 100 mGy, as recently done for leukemia after childhood exposure (57). The seven studies excluded because they only published risk estimates for categories of dose were mostly null (50-54). It is possible that the non-increasing categorical risks were the reason for not presenting risks per continuous dose, which could be a form of reporting bias. Finally, two background radiation studies used dose rate instead of cumulative dose (55,56). Both were largely null.

A recent review conducted by the National Council on Radiation Protection and Measurements (NCRP) (58) assessed the different, albeit related question, of whether the recent epidemiological data from 29 low-dose and low-dose rate studies support the linear-no-threshold (LNT) model for radiation protection purposes. There were considerable differences between the studies included in the NCRP review and ours; only 12 of the studies were in both reviews. One reason for the difference in eligible studies is that we restricted our review to low-dose studies, defined as mean cumulative dose <100 mGy. The approach to reviewing the evidence was also different. NCRP adopted the traditional approach to assessing study quality and excluded several studies due to their being classified as low quality. As we have shown in this monograph, there is not always a direct relationship between study quality and bias. Some of the studies that were classified as high quality could still have been subject to bias, e.g., the Japanese nuclear workers study (33), and those judged as of poor quality are not necessarily biased, e.g., the GB background study (21) and the Taiwanese residents study (24). We showed formally that mostly these methodological issues and errors are unlikely to have resulted in biases that would have impacted the interpretation of the study results. Nevertheless, after their exclusions the NCRP committee still concluded that

“twenty studies [of the 29 reviewed, with the committee designating four studies as inconclusive] (80%) provided some support for the LNT model, including five studies (20%) providing strong support and four (16%) providing moderate support” (58). As mentioned previously, the BEIR VII (1) and other earlier reviews of evidence for cancer risks from low doses mostly depended on epidemiological studies of higher dose exposures and then invoked the linear no-threshold assumption from experimental data to support the conclusion that low doses are likely to cause cancer. Although we cannot rule out the possibility that the risks are influenced by higher doses, our results derived from studies with a mean cumulative dose of <100 mGy apply to populations that are mostly exposed to low doses and therefore directly address the question of whether there is epidemiological evidence for cancer risks from low-dose exposures. Despite the different approaches and different studies included, the two previous reviews (1, 58) and the pooled analysis of leukemia after childhood exposure to doses below 100 mGy (57) are all in general agreement with our result that there is evidence of cancer risks from low-dose ionizing radiation. Our summary risk estimates from the meta-analysis are broadly consistent with the Life Span Study of atomic bomb survivors (41, 42).

The observation that most of the study findings were positive ( $ERR > 0$ ) raises the question of whether there could be publication bias. As these epidemiological studies required extensive effort and all had the primary (and usually the only) aim of evaluating whether low-dose ionizing radiation causes cancer, the likelihood that a null or statistically non-significant result would not be published is likely quite low in our view. Furthermore, the field of radiation epidemiology is relatively small and we are not aware of recent studies or updates that were not published, although a few were published or submitted for publication after the study period.

Confidence intervals and study power need to be considered in interpreting study findings. The confidence intervals for ERRs from all eligible studies included positive values; for the positive estimated ERRs, the upper confidence limits were often several times the estimate. Thus, even when the null hypothesis could not be rejected, findings were compatible with positive effects. It should also be noted that a statistically significant estimate in a low power study may reflect a false positive finding and suggest that the estimated ERR is biased. However, one should be cautious about the assessments of statistical power reported here (Table 1 and 2) (and see also Gilbert et al (3)) as they are based on summary data reported by the study and several assumptions, most importantly that the ERR from the Life Span Study can be transported to other populations.

To our knowledge, this monograph provides the first systematic assessment of the impact of methodological issues and errors on the risk estimate in the studies of low-dose radiation exposure eligible for our analysis. More generally, we used a novel approach in this systematic review to formally assess biases based on published data combined with epidemiological and statistical theory. Traditionally, systematic reviews classify the quality of a study but without formally considering whether the quality of information translates into a bias. For example, low quality dosimetry does not automatically result in bias, e.g., the GB background study (21). In addition, the direction of the bias, and, if possible, the magnitude of the potential bias needs to be assessed. Without these further considerations exclusions based on quality or potential bias could result in substantial loss of information. Such an approach has been recently recommended over other approaches, such as use of a “risk of bias” checklist (59).

There were, however, several limitations to our review. Primarily because we were working only with published data, there were several instances where we had insufficient information to assess the direction of the bias. As examples, this occurred in assessing the direction of confounding if

the relationship between the exposure and the confounder was not published, in the lack of quantitative estimates about the completeness of follow-up, and in absence of information about the completeness and accuracy of vital statistics and cancer registry data for studies using linkage of the cohort with such databases. We also could not assess the magnitude of the bias in most situations because the required data were not available in the publications. Our summary analysis was therefore limited to the simple sign test, which does not take account of the size or precision of the estimated ERRs, since alternative tests are based on the actual values (t-test) or their ranks (Wilcoxon signed rank test). We have made several recommendations in each of our papers that would facilitate assessments of bias in the risk estimates in the future. For example, we recommend the routine publication of assessments of dose uncertainty and levels of loss to follow-up by exposure and outcome. Some of the data to inform our assessments came from sub-studies and we assumed that these applied to the full study population. For example, there was no evidence that smoking was related to occupational radiation exposure in a case-control study of leukemia in U.S. nuclear workers (60). As smoking data were not available for the full cohort, we assumed the findings from this sample were generalizable. In addition, the sign test requires that the distribution of the estimated ERRs per unit dose (under the assumption that the true ERRs are zero) has a median of zero, although it does not depend on the ERRs following approximately normal distributions. Our assessment of whether a study was positive or null as input for the sign test did not account for the width of the confidence intervals. Many of these studies had wide confidence intervals, which overlapped zero and therefore one could argue that bias adjustment which moves a statistically non-significant positive ERR towards the null does not change the conclusion. Results on all solid cancers combine a diverse group of cancer sites. This heterogeneity has to be taken into account when comparing results with those for leukemia. Finally, the meta-analyses



included all studies, even those identified as potentially biased and were based on the assumption that the estimated ERRs from the individual studies were approximately normally distributed. With the highly skewed dose distributions and small numbers of cases in many of these studies, these approximations may not be accurate (3). Thus, tests and confidence intervals could be distorted, especially since the meta-analyses are based on the Wald method. Furthermore, our assessment necessarily had a subjective component.

Age at exposure and time since exposure are important effect modifiers of cancer risk from ionizing radiation exposure, and the eligible studies were heterogeneous in these respects. A pooled analysis with individual subject data is necessary, therefore, to quantify the risk accounting for these effect modifiers. The childhood thyroid pooling project included here (28) and the more recently published pooled analysis of leukemia after childhood radiation exposure (57) provide the most reliable estimates of these cancer risks that account for these modifiers. Similarly, the INWORKS study (9) reviewed here provides a reliable, and we judge minimally biased, estimate of the solid cancer and leukemia risks from adulthood occupational radiation exposures incorporating age and time since exposure (8, 9).

It is well established that moderate to high doses of ionizing radiation cause cancer. As discussed in the paper by Gilbert et al (3), the available low dose human data also now satisfy most of Bradford Hill's viewpoints on causality: consistency, temporality, biological gradient, plausibility and coherence (61). The first of Bradford Hill's viewpoints, strength of effect or, in other words, magnitude of association, is usually not met in studies of cancer risk after low-dose radiation exposure. Large effects are an important indicator of causality since they are less likely to be entirely due to bias. However, large effects are not considered a necessary aspect and indeed Bradford Hill explicitly cautions against dismissal of weak observed associations. Our systematic

assessment in this monograph showed that these epidemiological studies of low-dose radiation and cancer risk are characterized by several limitations, but we found that only a small minority of the studies had biases whose correction could have moved a positive ERR towards the null. After exclusion of these studies, the majority of studies still reported positive risk estimates. We therefore conclude that there is now a large body of epidemiological data which supports excess cancer risks from low-dose ionizing radiation and the magnitude of the excess relative cancer risk from these low dose studies is statistically compatible with the atomic bomb survivors.

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Figure 1: Meta-analysis of the excess relative risk (ERR) at 100 mGy for all solid cancers after adulthood radiation exposure. The size of the ERR symbol is proportional to the inverse variance of the study-specific ERR.

Figure 2: Meta-analysis of the excess relative risk (ERR) at 100 mGy for leukemia after adulthood radiation exposure. The size of the ERR symbol is proportional to the inverse variance of the study-specific ERR.

Figure 3: Meta-analysis of the excess relative risk (ERR) at 100 mGy for leukemia after childhood radiation exposure. The size of the ERR symbol is proportional to the inverse variance of the study-specific ERR.

**Table 1: Assessment of bias from several sources for studies of solid cancers**

		ERR per unit of dose bias						
Study name	Reference	ERR at 100 mGy	95% CI	Dose error	Confounding/selection bias	Outcome misclassification	Could bias adjustment move ERR towards null?*	Estimated power†
<b>Environmental</b>								
Three Mile Island	Han et al. 2011 (19)	-1	(-6 to 3)	↓			Yes. Adjustment possibly moves ERR towards null.‡	NC
Chinese background	Tao et al. 2012 (20)	-0.101	(-0.253 to 0.095)			↕	Uncertain. Adjustment could move ERR towards or away from null.	Low
GB background	Kendall et al. 2013 (21)	2	( -2.0 to 6.0)					NC
Swiss background	Spycher et al. 2015 (22)	2.8	(0.8 to 4.8)					NC
Techa River	Davis et al. 2015 (23)	0.077	(0.013 to 0.150)					Reasonable
Taiwanese residents	Hsieh et al. 2017 (24)	0.04	(0.01 to 0.08)‡					Low
<b>Medical</b>								
Canadian cardiac imaging	Eisenberg et al. 2011 (25)	0.3	(0.2 to 0.4)		↕	↑	Uncertain. Adjustment could move ERR towards or away from null.	NC
French Pediatric CT (brain tumors)	Journey et al. 2016 (26)	0.7	(-0.1 to 1.0)					Low
UK Pediatric CT (brain tumors)	Berrington et al. 2016 (27)	1.2	(0.4 to 3.1)					Reasonable
PIRATES (thyroid cancer)	Lubin et al. 2017 (28)	0.96	(0.37 to 1.70)					Reasonable

# Occupational

Korean workers	Ahn et al. 2008 (29)	0.72	(-0.5 to 2.1)‡	↓	↑	Uncertain. Adjustment could move ERR towards or away from null.	NC
UKNRRW	Muirhead et al. 2009 (30)	0.03	(0 to 0.056)	↓		No. Adjustment would move ERR away from null.	Reasonable
Korean nuclear workers	Jeong et al. 2010 (31)	0.21	(-0.19 to 0.9)	↓		No. Adjustment would move ERR away from null.	Low
Rocketdyne workers	Boice et al. 2011 (32)	0.02	(-0.18 to 0.17)	↓		No. Adjustment would move ERR away from null.	Low
Japanese workers	Akiba et al. 2012 (33)	0.13	(-0.03 to 0.30)	↕	↓	Uncertain. Adjustment could move ERR towards or away from null.	Low
Canadian nuclear workers§	Zablotska et al. 2014 (34)	-0.12	(<-0.15 to 0.24)	↓		Yes. Adjustment possibly moves ERR towards null.¶	NC
German nuclear workers	Merzenich et al. 2014 (35)	-0.1	(-0.4 to 0.1)	↓		Yes. Adjustment possibly moves ERR towards null.¶	NC
US nuclear workers	Schubauer-Berigan et al. 2015 (36)	0.01	(-0.02 to 0.05)	↓		No. Adjustment would move ERR away from null.	Low
INWORKS	Richardson et al. 2015 (9)	0.047	(0.018 to 0.079)‡	↓		No. Adjustment would move ERR away from null.	Reasonable
USRT (breast cancer)	Preston et al. 2016 (12)	0.07	(-0.005 to 0.19)	↕		Uncertain. Adjustment could move ERR towards or away from null.	Low
USRT (brain cancer)	Kitahara et al. 2017 (10)	0.1	(<-0.3 to 1.5)				Low
USRT (skin cancer)	Lee et al. 2015 (11)	-0.001	(-0.04 to 0.05)				Reasonable
French nuclear workers	Leuraud et al. 2017 (37)	0.04	(-0.04 to 0.13)‡	↓		No. Adjustment would move ERR away from null.	Low

\*Reflects an assessment of the presence of bias of the ERR and its likely direction, but not statistical significance

‡Low: <50%, reasonable: ≥50% based on LSS ERR and published dose distributions.

‡90%CI

§The Canadian Study is restricted to the cohort excluding early AECL workers.

||If bias is sufficiently large and additive, adjustment could increase the ERR above null. Since the magnitude of bias is difficult to determine from the published reports, we conservatively assume that adjustment of bias will not increase the ERR to positive.

NC = Not calculated because dose distributions needed for power calculations were not available.

↓ Bias in the negative direction

↑ Bias in the positive direction

↑↓ Bias of unclear direction

**Table 2: Assessment of bias from several sources for studies of leukemia**

Study name	Reference	ERR at 100 mGy	95% CI	Dose error	ERR per unit of dose bias		Could bias adjustment move ERR towards null?†	Estimated power*
					Confounding/ selection bias	Outcome misclassification		
<b>Environmental</b>								
Chornobyl residents	Davis et al. 2006 (16)	3.2	(0.9 to 8.4)	↑	↑		Yes. Exclusion of subgroup with potential recall bias reduced risk to null.	NC
Three Mile Island	Han et al. 2011 (19)	19	(−3 to 45)	↓			No. Adjustment would move ERR away from null.	NC
Chinese background	Tao et al. 2012 (20)	1.068	(<0 to inf)			↓	Uncertain. Adjustment could move ERR towards or away from null.	Low
GB background	Kendall et al. 2013 (21)	12	(3.0 to 22.0)					Reasonable
Swiss background	Spycher et al. 2015 (22)	3.6	(−0.3 to 7.7)					Low
Finnish background	Nikkila et al. 2016 (38)	−3	(−11 to 6)					NC
Taiwanese residents	Hsieh et al. 2017 (24)	0.15	(0.03 to 0.24)‡					Low
<b>Medical</b>								
French Pediatric CT	Journey et al. 2016 (26)	1.6	(−2.3 to 2.7)					Low
UK Pediatric CT	Berrington et al. 2016 (27)	3	(0.3 to 10.9)					Low
<b>Occupational</b>								
Korean workers	Ahn et al. 2008 (29)	1.68	(−3.4 to 14.9)‡		↓	↑	Uncertain. Adjustment could move ERR towards or away from null.	NC
Chornobyl liquidators	Kesminiene et al 2008 (17)	0.5	(−0.38 to 5.70)‡	↑			Yes. Adjustment possibly moves ERR to null.	Low



UKNRRW	Muirhead et al. 2009 (30)	0.18	(-0.006 to 0.50)					Low
Rocketdyne workers	Boice et al. 2011 (32)	0.06	(-0.50 to 1.23)					Low
Japanese workers	Akiba et al. 2012 (33)	-0.19	(-0.61 to 0.86)		↓	↓	Yes. Adjustment possibly moves ERR towards null. <sup>ll</sup>	Low
Ukrainian Chornobyl liquidators	Zablotska et al. 2013 (18)	0.221	(0.005 to 0.761)	↑			Yes. Adjustment possibly moves ERR towards null.	NC
Canadian nuclear workers§	Zablotska et al. 2014 (34)	1.44	(<-0.15 to 14.6)					NC
German nuclear workers	Merzenich et al. 2014 (35)	0.4	(-0.3 to 1.1)		↓		No. Adjustment would move ERR away from null.	NC
US nuclear workers	Schubauer-Berigan et al. 2015 (36)	0.17	(-0.02 to 0.47)					Reasonable
INWORKS	Richardson et al. 2015 (9)	0.3	(0.12 to 0.52)‡					Reasonable
US atomic veterans	Caldwell et al. 2016 (39)	-0.5	(-14 to 4)					Low
French nuclear workers	Leuraud et al. 2017 (37)	0.35	(<0 to 1.6)‡		↓		No. Adjustment would move ERR away from null.	Low

\*Low: <50%, reasonable: ≥50%.

†Reflects an assessment of the presence of bias of the ERR and its likely direction, but not statistical significance

‡90%CI

§The Canadian Study is restricted to the cohort excluding early AECL workers.

llIf bias is sufficiently large and additive, adjustment could increase the ERR above null. Since the magnitude of bias is difficult to determine from the published reports, we conservatively assume that adjustment of bias will not increase the ERR to positive.

NC = Not calculated because dose distributions needed for power calculations were not available.

↓ Bias in the negative direction

↑ Bias in the positive direction

**Table 3: Studies with an internal dose-response analysis that were ineligible for only one reason, reasons for exclusion, and summary of findings**

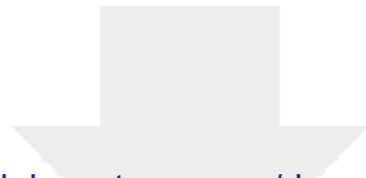
Population	1st author	Publication year	Reason for exclusion	Findings for solid cancers and leukemia (or primary cancer site)
Breast cancer in US scoliosis	Ronckers (44)	2008	Mean cumulative dose = 120 mGy	Borderline statistically significant positive dose-response relationship for breast cancer incidence (p-trend=0.06).
Kerala background	Nair (45)	2009	Mean cumulative dose = 161 mGy	Statistically non-significant negative dose-response for incidence of all solid cancers excluding leukemia (p-trend>0.5).
Techa river	Krestinina (46)	2013	Mean cumulative dose = 410 mGy	Statistically significant positive dose-response for incidence of all leukemias (p-trend<0.001).
Mayak workers	Sokolnikov (47)	2015	Mean cumulative dose = 354 mGy	Statistically significant positive dose-response for mortality of all solid cancers excluding lung, liver and bone cancers (p-trend=0.01).
Chornobyl clean-up workers	Kashcheev (48)	2015	Mean cumulative dose = 132 mGy	Statistically significant positive dose-response for total cancer incidence (p-trend=0.03) and mortality (p-trend=0.05).
Chinese medical workers	Sun (49)	2016	Mean cumulative dose = 250 mGy	Statistically significant positive dose-response for incidence of all solid cancers (p-trend=0.002).
US Shipyard workers	Matanoski (50)	2008	Categorical risk estimates	Statistically non-significant increased risk of leukemia mortality in highest dose category. No trend tests presented.
Australian nuclear test	Gun (51)	2008	Categorical risk estimates	No increased risk of solid cancer incidence or leukemia incidence across dose-categories (p-trend>0.05).
French biology researchers	Guseva (52)	2008	Categorical risk estimates	Statistically significant increasing trend for all cancer deaths across dose-categories (p=0.03 5-year lag).
Finnish reindeer herders	Kurtio (53)	2010	Categorical risk estimates	No overall increased risk of cancer incidence across dose-categories (p-trend=0.28), but statistically significant increased risk with dose for exposure <age 15 years (p-trend=0.003).
Childhood X-rays	Hammer (54)	2009	Categorical risk estimates	No increased risk of solid cancer incidence (p-trend=0.32) or leukemia incidence (p-trend=0.26) across dose-categories.
French background (Geocap case-control)	Demoury (55)	2017	Risk for dose rate not cumulative dose	Acute leukemia incidence not related to background radiation exposure (p-trend>0.05).
German background	Spix (56)	2017	Risk for dose rate not cumulative dose	Statistically non-significant positive increased risk of lymphoid leukemia incidence (p-trend=0.54).

**Table 4: Meta-analysis of excess relative risks (ERR) per 100mGy for all solid cancers and leukemia**

Outcome	Studies	ERR/100 mGy (95% CI)	p-value	Cochran's Q (p-value)	I <sup>2</sup>
Adult solid cancer	16*	0.061 (-0.024, 0.145)	0.08	45.13 (<0.001)	0.67
Adult solid cancer excluding the Canadian cardiovascular imaging study due to heterogeneity	15*	0.033 (-0.015, 0.081)	0.11	17.90 (0.21)	0.22
Adult leukemia	14†	0.160 (0.070, 0.250)	<0.001	4.12 (0.99)	NA
Childhood leukemia	6	2.840 (0.370, 5.320)	0.01	6.40 (0.27)	0.22

\*Excluding INWORKS and site-specific results from USRT

†Excluding INWORKS



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