

LETTER

Life history adaptations to fluctuating environments: Combined effects of demographic buffering and lability

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Funding information

Norges Forskningsråd, Grant/Award Number: 244404

Editor: Robin Snyder

Abstract

Demographic buffering and lability have been identified as adaptive strategies to optimise fitness in a fluctuating environment. These are not mutually exclusive, however, we lack efficient methods to measure their relative importance for a given life history. Here, we decompose the stochastic growth rate (fitness) into components arising from nonlinear responses and variance–covariance of demographic parameters to an environmental driver, which allows studying joint effects of buffering and lability. We apply this decomposition for 154 animal matrix population models under different scenarios to explore how these main fitness components vary across life histories. Faster-living species appear more responsive to environmental fluctuations, either positively or negatively. They have the highest potential for strong adaptive demographic lability, while demographic buffering is a main strategy in slow-living species. Our decomposition provides a comprehensive framework to study how organisms adapt to variability through buffering and lability, and to predict species responses to climate change.

KEYWORDS

climate change, comparative study, convexity, demographic buffering, demographic lability, environmental variance, matrix population model, nonlinearity, stochasticity, temporal covariance

INTRODUCTION

Understanding life history adaptations to fluctuating environments is increasingly important, as anthropogenic climate change is altering the temporal variability of multiple climatic drivers (IPCC, 2021; Laufkötter et al., 2020; Pendergrass et al., 2017). For instance, while an increased variance in daily and seasonal temperature and precipitation is expected across much of Europe in summer, a decrease is projected in other regions (Huntingford et al., 2013; IPCC, 2021; Kotz et al., 2021; Pendergrass et al., 2017). Fluctuations in abiotic and biotic environmental drivers experienced by organisms may affect their relative fitness and select for specific adaptations to live in variable environments.

Two main processes have been identified as adaptations to environmental variability, optimising fitness: Demographic buffering reduces the variance in demographic parameters (e.g. survival, fertility), thereby minimising the effects of bad environments (Hilde et al., 2020; Morris & Doak, 2004), while demographic lability lets the organisms take advantage of good environments by mounting a large increase in some demographic parameters compared to an average or bad environment, and therefore increasing their mean (Barraquand & Yoccoz, 2013; Jongejans et al., 2010; Koons et al., 2009; see **Box 1** for Glossary). The two processes are not mutually exclusive but can be selected simultaneously, so that different demographic parameters of a given life history can show different responses to an environmental driver. Yet, these processes have often been investigated

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BOX 1 Glossary

Stochastic population growth rate - $\ln(\lambda_s)$: The long-term rate of population growth on a logarithmic scale, a measure of fitness in a stochastic density-independent environment.

Growth rate in the mean environment - $\ln(\lambda_0)$: Dominant eigenvalue of the projection matrix in the mean environment ($z = 0$) $\mathbf{A}(0)$ on a logarithmic scale.

Mean growth rate - $\ln(\bar{\lambda})$: Dominant eigenvalue of the mean projection matrix across variable environments $\bar{\mathbf{A}}$ on a logarithmic scale.

Demographic lability/labile demographic parameter: A labile demographic parameter fluctuates with temporal variation in environmental conditions. The relationship between a labile demographic parameter and the environment (e.g. a key environmental driver) can be convex, concave or linear, so that the average value of this demographic parameter in a variable environment becomes $>$, $<$ or $=$ to the demographic parameter estimated in the mean environment ($z = 0$) respectively. The same definition applies to labile vital rates (e.g. survival, fecundity, transition).

Adaptive demographic lability (demographic lability hypothesis): Selection for demographic parameters to track environmental fluctuations that leads to an overall increased fitness, $\ln(\lambda_s)$. Increase in $\ln(\lambda_s)$ occurs when an increase in the demographic parameter means due to convexity in their responses leads to a shift in the arithmetic mean of annual population growth rates $\ln(\bar{\lambda})$, that overcomes the negative effect of temporal variance in the annual population growth rates (variance–covariance component σ^2). This hypothesis relies on the assumption that the nonlinearity index D (defined below) is positive.

Nonlinearity index (D): D measures the total effect of nonlinearity of demographic parameters in a life history, and is a key component to describe the nonlinearity component of the fitness decomposition (Equation 3). This index corresponds to the sum over all (st)ages of the second derivatives of the demographic parameters (depending on vital rates) in the mean environment ($z = 0$), weighted by the sensitivities of λ_0 to the corresponding demographic parameters (matrix elements). When positive (/negative), D is an indicator of adaptive (/non-adaptive) lability through overall positive (/negative) contributions from convexity (/concavity) of the demographic parameters. Adaptive lability can create a positive overall effect of environmental variability if D is positive and the negative effects of increased variance–covariance of the demographic parameters are not too large (see Equation 3).

Demographic buffering/buffered demographic parameters: Low variance of a demographic parameter in response to temporal variation in the environmental variable z . A more flat relationship between the demographic parameter and the environment z leads to such low parameter variance, and to the mean demographic parameter in the variable environment remaining approximately equal to demographic parameter value in the mean environment ($z = 0$). The same definition applies to buffered vital rates (e.g. survival, fecundity, transition).

Adaptive demographic buffering (demographic buffering hypothesis): The prediction that natural selection should favour a reduction in variance of the demographic parameters with the strongest influence on fitness in the mean environment, reducing the variance–covariance component σ^2 and leading to an overall stable or increased fitness in variable environments. The assumption that $\ln(\bar{\lambda})$ is not affected by environmental variance ($\ln(\lambda_0) \approx \ln(\bar{\lambda})$), is often made for this hypothesis.

separately, and we lack efficient methods to disentangle and predict their relative importance for a given life history and environment. To understand how organisms combine lability and buffering of their demographic parameters to enhance fitness in varying environments, we need a demographic model framework to predict two main fitness components: (1) the effects of nonlinearity in responses of all demographic parameters to an environmental driver, and (2) the effects of variance–covariance of these parameters. While the latter is well described in stochastic demographic theory (Lande et al., 2003), we know much less about the impacts of nonlinearity, representing the potential for adaptation to varying environments through lability.

A key prediction from classical theory for evolutionary bet-hedging and stochastic population growth is that the long-term fitness will be reduced if the temporal variance of fitness is increased (Lewontin & Cohen, 1969). This result is assuming an unstructured population with annual population growth rates that are IID (independently and identically distributed). The fitness is then the logarithm of the geometric mean of these growth rates (Lewontin & Cohen, 1969). In structured populations, the stochastic growth arises from more complex demographic pathways, due to variation in demographic parameters of individuals in different (st)ages. Fluctuations in the (st)age structure introduce autocorrelation in the annual growth

rates (Caswell, 2001). Still, under the assumption of small fluctuations in the demographic parameters, Tuljapurkar (1982) derived an important approximation of long-term growth rate in stage-structured populations, emphasising how the variance in fitness is linked to variances and covariances of demographic parameters in different stages (Equation 1). The key conclusion from this approximation is that temporal variability in demographic parameters and/or positive covariance will have a negative effect on fitness, and should be selected against, in particular for demographic parameters that have a large impact on fitness in the mean environment. Accordingly, the demographic buffering hypothesis predicts that natural selection should favour a reduction in variance of the demographic parameters with the strongest influence on population growth (Boyce et al., 2006; Gaillard & Yoccoz, 2003; Hilde et al., 2020; Pfister, 1998; Tuljapurkar & Orzack, 1980).

However, positive effects of environmental variability have also been demonstrated under strong negative covariance among demographic parameters (Colchero et al., 2019; Doak et al., 2005; Tuljapurkar, 1990), negative environmental autocorrelation (Tuljapurkar, 1982) and convex relationships between demographic parameters and the environment. The latter represents a case of adaptive lability as described by Koons et al. (2009). In contrast to adaptive demographic buffering, which optimises fitness by reducing the variance of most influential demographic parameters, lability can be adaptive if the benefit of an increase in the arithmetic mean of the annual growth rates through increased demographic parameter means can overcome the negative effect of increased demographic variance on fitness (Box 1). Nonlinearity in population and demographic parameter responses to environmental drivers may be common in the wild (Barraquand & Yoccoz, 2013; Clark & Luis, 2020; Dahlgren et al., 2011; Drake, 2005; Hansen et al., 2021; Jenouvrier et al., 2012; Lawson et al., 2015; Louthan & Morris, 2021; Mysterud et al., 2001), highlighting the potential importance of lability as an adaptation to environmental variability. However, with structured life histories the combined effects of nonlinearity in different demographic parameters on fitness are challenging to predict (Koons et al., 2009).

Somewhat contrasting predictions have been made as to which demographic parameters should be labile or buffered, and the relative importance of each process for a given life history. Demographic lability has been suggested to affect mainly the demographic parameters with least effect on fitness (Hilde et al., 2020), as a consequence of selection for buffering of more influential demographic parameters. Other studies suggest that lability can be equally important to demographic buffering, and that it can also occur in highly influential demographic parameters (Jongejans et al., 2010; Koons et al., 2009; McDonald et al., 2017). Based on the latter

prediction, recent research suggests that adaptive lability and buffering can be located at the opposite ends of a continuum, encompassing a wide range of demographic strategies (Salguero-Gómez, 2021; Santos et al., 2021). Yet, the extent to which lability among the least or the most influential demographic parameters can be adaptive strategies for coping with varying environments, relative to buffering, remains largely unexplored (e.g. Barraquand & Yoccoz, 2013).

We thus need a more thorough understanding of how the opportunity for selection on demographic buffering and lability depends on major axes of life history variation such as the slow–fast continuum (Gaillard et al., 2016; Salguero-Gómez, Jones, Jongejans, et al., 2016; Stearns, 1992). For instance, populations of fast-living species have been predicted to be more responsive to environmental variability than those of slow-living species, and to be more likely to show adaptive lability in their demographic parameters (Dalglish et al., 2010; Iles et al., 2019; Koons et al., 2009; Morris et al., 2008). According to demographic buffering hypothesis, species towards the slow end of the continuum benefit most from reduced variance in annual survival of the mature stages, while fast-living species gain relatively more from reduction of variance in annual fertility and/or survival of the immature stages (Gaillard & Yoccoz, 2003; Hilde et al., 2020; Rotella et al., 2012). These effects can be predicted from Tuljapurkar's small noise approximation (Tuljapurkar, 1982; Equation 1), but we lack a similar expression to describe the net impact of nonlinearity in different demographic parameters of the same life history. Here, we introduce a new 'nonlinearity index' to predict changes in the arithmetic mean arising from nonlinearity in different demographic parameter responses to an explicit environmental driver. We decompose the stochastic growth rate into contributions from nonlinearity effects and variance–covariance effects. We then apply the decomposition to study how organisms may combine adaptive buffering and lability responses depending on generation time, which closely correlates with the species' position along the slow–fast continuum (Gaillard et al., 2005). We use population models from the COMADRE animal matrix database (Salguero-Gómez, Jones, Archer, et al., 2016) as a starting point for our calculations, representing a broad range of life histories in the mean environment. We then add stochastic environmental variation and perform the decomposition under different scenarios for nonlinearity and covariance among demographic parameters (Figure 1). Our study provides a method to disentangle the effects of buffering and lability for any given life history, and the subsequent analysis addresses two main questions: First, what is the opportunity for positive effects due to adaptive lability to overcome negative impacts through the variance–covariance of demographic parameters, and how does this pattern depend on generation time? Second, are demographic

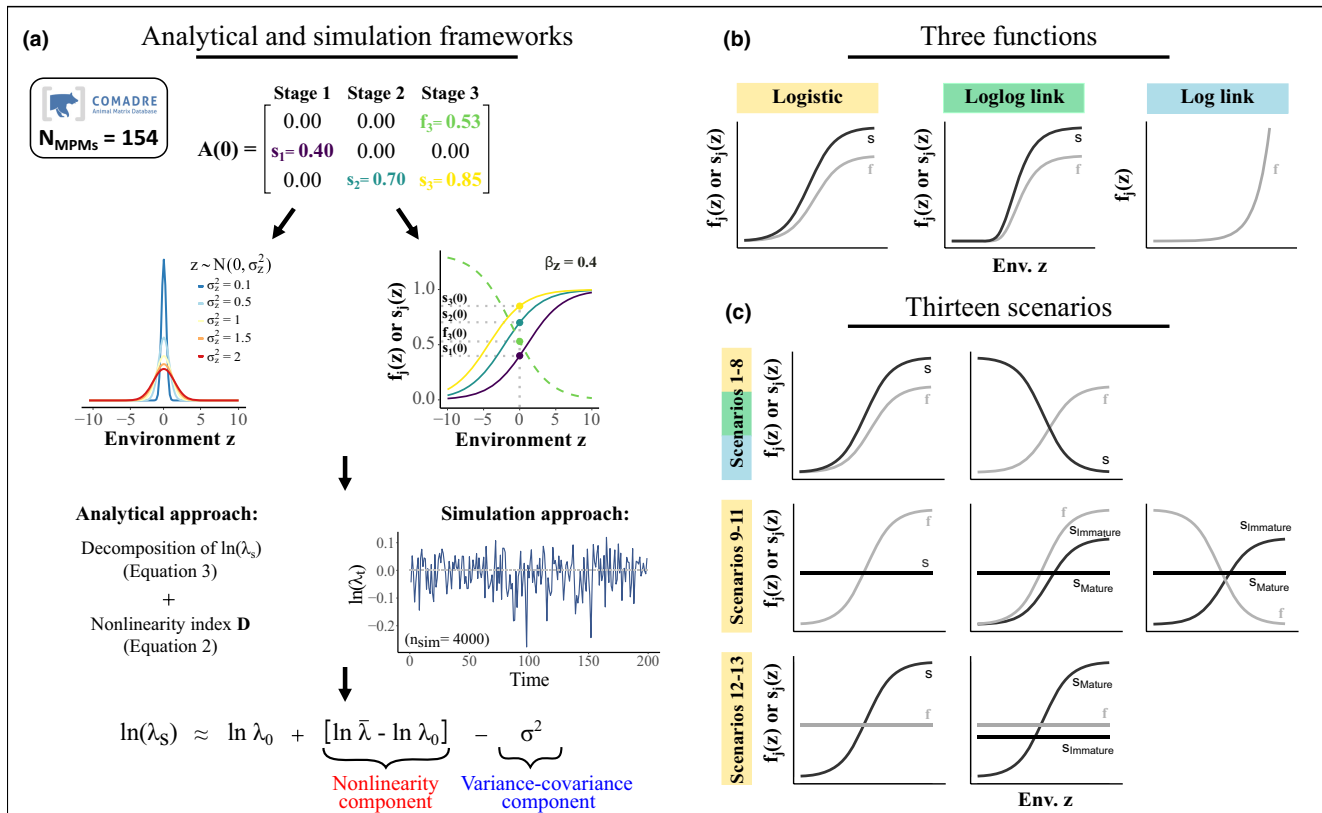


FIGURE 1 Framework used to study the effects of environmental variability on fitness (stochastic growth rate $\ln(\lambda_s)$). (a) Our calculations define demographic parameters as nonlinear functions of the environmental driver z (see methods), where $A(0)$ (from our selected, standardised COMADRE models, $N_{tot} = 154$) defines the values of (st)age-specific survival rates $s_j(0)$ and fertilities $f_j(0)$ in the mean environment ($z = 0$). Different levels of environmental variance levels σ_z^2 and environmental strength $|\beta_z|$ of z on the demographic parameters were considered. In the analytical approach, $\ln(\lambda_s)$ was calculated and decomposed into main components capturing nonlinearity and variance–covariance effects following Equation (3). The accuracy of this decomposition was tested using simulations (Supporting Information S4). (b) Two or three different link functions were considered for survival $s_j(z)$ and fertility $f_j(z)$ respectively. (c) Scenarios 1–8: Four combinations were examined including logistic functions for all parameters, loglog link functions for all parameters and two combinations of exponential fertilities $f_j(z)$ (log link) with logistic or loglog link function for $s_j(z)$. Positive or negative covariance between survival and fertility was tested for each combination, assuming positive covariance between $s_j(z)$, and between $f_j(z)$. Scenarios 9–11: Scenarios of forced buffering considering demographic lability in the fertility coefficients and survival rates of the immature stages ($S_{immature}$). Scenarios 12–13: Scenarios of forced buffering assuming demographic lability in all survival rates $s_j(z)$ or in only the mature stages (S_{mature}). Logistic functions were used to define lability while the other rates were held constant and fixed to the values reported in the standardised COMADRE projection matrix.

parameters that show adaptive lability typically the least or most influential demographic parameters for fitness?

MATERIAL AND METHODS

To explore fitness responses to environmental variability along the slow–fast continuum, we decomposed the long-term stochastic growth rate $\ln(\lambda_s)$, a measure of fitness (Caswell, 2001; Lande et al., 2003; Tuljapurkar, 1990), into main components capturing effects of nonlinearity in demographic parameters as a function of an environmental driver z , and effects of variance–covariance among the parameters. Our approach builds on Tuljapurkar's approximation which assumes linear relationships between demographic parameters and an IID environmental variable (Tuljapurkar, 1990):

$$\ln(\lambda_s) \approx \ln \lambda_0 - \sigma^2, \text{ with } \sigma^2 = \frac{\text{Var}(\lambda_t)}{2\lambda_0^2}. \quad (1)$$

Here λ_0 is the arithmetic growth rate in the mean environment, which is assumed equal to the mean arithmetic growth rate $\bar{\lambda}$ (ignoring nonlinear responses), while $\text{Var}(\lambda_t)$ is the variance in annual population growth caused by temporal variance and covariance in the demographic parameters. We show in the next section that including nonlinear effects of the environment on demographic parameters mainly affects $\ln(\lambda_s)$ through the mean arithmetic growth rate λ , but also through the variance–covariance term. After defining main components of the stochastic growth rate, we perform a theoretical exploration of how the different components will vary across generation time, using different scenarios regarding nonlinear functions for survival and fertility (Figure 1a–c). We also confront

hypotheses about demographic lability, through scenarios that specifically consider effects of nonlinearity in the demographic parameters of immature or mature individuals only, keeping other parameters constant ('forced buffering' scenarios, Figure 1c). All simulations and calculations were performed in R, version 4.0.3 (R Core Team, 2020). R code is provided in Supporting Information S7.

Decomposing the stochastic growth rate with nonlinear effects

We assume that the environment at each time step is described by a stochastic variable z (IID), with mean 0 and variance σ_z^2 . Population growth from one time step to the next is given by $\mathbf{n}_{t+1} = \mathbf{A}(z_t)\mathbf{n}_t$, where \mathbf{n}_t is the vector containing the number of individuals in each stage at time t , and $\mathbf{A}(z)$ is the population projection matrix. Elements of $\mathbf{A}(z)$ are the demographic parameters describing survival, fertility and transitions as functions of z . To derive the stochastic growth rate, we approximate this projection matrix using $\mathbf{A}(z) \approx \mathbf{A}(0) + \frac{\sigma_z}{2}\mathbf{A}''(0) + \varepsilon$, where ε is the matrix describing the noise terms with mean elements 0 (i.e. stochastic deviations from the expectations defined by the other two terms on the right side of the equation), $\mathbf{A}(0)$ is the projection matrix of the mean environment (with asymptotic growth rate λ_0) and $\mathbf{A}''(0) = \mathbf{A}''(z)|_{z=0}$ contains the second derivatives of elements of $\mathbf{A}(z)$. Using this second derivative matrix, the reproductive value vector \mathbf{v} and the stable stage structure \mathbf{u} calculated from the matrix $\mathbf{A}(0)$, we define a nonlinearity index (Supporting Information S3).

$$D = \mathbf{v}\mathbf{A}''(0)\mathbf{u} = \sum_{ij} \frac{\partial \lambda_0}{\partial A_{ij}(0)} A''_{ij}(0), \quad (2)$$

which measures the overall degree of nonlinearity in the life history defined by $\mathbf{A}(z)$. A positive D indicates adaptive lability. A matrix element (i.e. demographic parameter) with strong convex curvature may still have a low impact on D if the corresponding sensitivity of λ_0 to that element is low, and vice versa.

Applying a Taylor approximation to the mean change of the logarithm of the total reproductive value $V_t = \sum_j n_{j,t} v_j$, we show in Supporting Information S3 that the stochastic growth rate is given by

$$\ln(\lambda_s) = E[\ln V_{t+1} - \ln V_t | V_t] \approx \ln \lambda_0 + \underbrace{\frac{\sigma_z^2}{2\lambda_0} D \left(1 - \frac{\sigma_z^2}{4\lambda_0} D\right)}_{\ln \bar{\lambda} - \ln \lambda_0} - \underbrace{\frac{\sigma_z^2}{2\lambda_0^2} \left(B + \frac{\sigma_z^2}{2} C\right)}_{\sigma^2}, \quad (3)$$

where D is the nonlinearity index defined above, $B = \sum_{ij} \frac{\partial \lambda_0}{\partial A_{ij}(0)} \frac{\partial \lambda_0}{\partial A_{kl}(0)} A'_{ij}(0) A'_{kl}(0)$ (where $\mathbf{A}'(0) = \mathbf{A}'(z)|_{z=0}$ is the matrix of first derivatives), and $C = \sum_{ij} \frac{\partial \lambda_0}{\partial A_{ij}(0)} \frac{\partial \lambda_0}{\partial A_{kl}(0)} A''_{ij}(0) A''_{kl}(0)$. The stochastic growth rate is

thus decomposed into the growth rate in the mean environment, $\ln \lambda_0$, plus two additive terms describing changes mainly due to nonlinearity ($\ln \bar{\lambda} - \ln \lambda_0$), and changes mainly due to variance–covariance (σ^2) of demographic parameters in a stochastic environment. The first term can be positive or negative, depending on the nonlinearity index D , and can be further decomposed into effects of survival and fertility coefficients (code in Supporting Information S7). The second term corresponds largely to the variance–covariance term in the approximation of Tuljapurkar (1982), except that here there is also a small effect of nonlinearity through C . However, the effect of nonlinearity on the second term is very small compared to the effect of nonlinearity on the mean, therefore we refer to the first term as the nonlinearity component and second term as the variance–covariance component. In the Supporting Information S4 we demonstrate the accuracy of this approximation using simulations.

Applying the decomposition

To explore life history variation in the main components of the stochastic growth rate, we used age- and stage-structured matrix population models (MPMs) from the COMADRE Animal Matrix Database (v.4.20.5; Salguero-Gómez, Jones, Archer, et al., 2016) as a starting point, considering different scenarios for effects of the environment z on the demographic parameters. Each MPM includes a projection matrix that depends on the (st)age-specific fertilities, transitions and survival rates for a given time interval (see Figure 1). We let this projection matrix represent the matrix in the mean environment, $\mathbf{A}(0)$. We selected MPMs from unmanipulated and free-ranging populations, considering only 'mean' matrices (i.e. one matrix per population) with annual time steps. Before the analysis we standardised all MPMs to have $\lambda_0 = 1$ by dividing each matrix element by λ_1 calculated from the original matrix (see Supporting Information S1 for complete description of selection criteria). One hundred fifty-four MPMs were selected, describing two amphibian, 35 bird, 22 bony fish, three insect, 61 mammal and 31 reptile populations, belonging to 107 species. Generation time was calculated as the mean age of parents at the stable (st)age distribution (Bienvenu & Legendre, 2015) and ranged from 1.1 to 265.6 years.

Nonlinear relationships

We added environmental effects to the survival and fertility coefficients. Since some models were stage

structured, we first separated out the two matrices containing these coefficients: Each stage structured projection matrix can be written as $\mathbf{A} = \mathbf{GS} + \mathbf{QB}$ (Vindenes et al., 2021). Here \mathbf{G} and \mathbf{Q} are matrices describing the stage transition rates of individuals and new offspring, respectively, assumed constant in our analysis. The matrix \mathbf{B} contains the stage-specific fertility coefficients $f_j(z)$ on the diagonal and zeroes elsewhere, while the matrix \mathbf{S} contains stage-specific survival rates $s_j(z)$ on the diagonal and zeroes elsewhere. For each MPM, we chose a link function for the survival rates $s_j(z)$ (logistic or loglog link) and a link function for the fertility coefficients $f_j(z)$ (logistic, loglog or log link), defining the relationship of $\mathbf{A}(z)$ to the environmental driver z . For each scenario we defined different link functions (Figure 1b,c), where $s_j(0)$ and $f_j(0)$ corresponded to the values from the standardised MPM in COMADRE. For instance, with a loglog link function, the survival rates are defined as $s_j(z) = \exp(-\exp(-\beta_0 - \beta_{z_S} z))$, and the parameter β_0 is defined as $\beta_0 = -\ln(-\ln(s_j(0)))$. The parameter β_{z_S} defines the strength of the environmental effect on $s_j(z)$, and affects the curvature and variance of survival probability in stage j (Figure 1a; Figure S6 shows survival and fertility coefficients for different β_{z_S} and β_{z_F} values). Fertility coefficients are defined in a similar way, but here we also defined a maximum $\text{Max}F = M \times f_j(0)$ with $M = 2.5$ (results for different values of M are shown in Supporting Information S5), so that the fertility in the mean environment was set as a proportion of the maximum fertility. The values in the mean environment $s_j(0)$ and $f_j(0)$, defined by the standardised MPM, affect the second derivatives of the link functions (Figure 1a; Figure S2). A complete description including equations for all link functions and their derivatives is provided in Supporting Information S2 and S7.

To limit the number of scenarios we made the simplifying assumption that survival rates of different (st)ages all have the same value of β_{z_S} , and similarly all fertility coefficients have the same β_{z_F} . This means that there is always positive covariance among survival rates of different (st)ages and among fertilities of different (st)ages (but curvature and variance vary according to the demographic parameter values in the mean environment; Figure S2), while covariance between survival and fertility is controlled in our scenarios by the sign of β_{z_S} and β_{z_F} . The assumption of positive covariance is biologically relevant for populations where individuals of different (st)ages live in the same environment, and where survival of different stages and reproduction of different stages are affected in the same direction by a key driver. A range of other scenarios are also possible but not considered here, such as no covariance among demographic parameters.

Scenarios

We decomposed the stochastic growth rate under 13 scenarios (Figure 1c) varying (1) the type of link function defining $s_j(z)$ and $f_j(z)$, (2) the covariance between survival and fertility; negative or positive (scenarios 1–8), and by applying (3) special cases of forced buffering, turning off the effect of z for certain demographic parameters (thus nonlinearity, variance and covariance of demographic parameters were affected; scenarios 9–13). In the first eight scenarios, effects of z were added to survival and fertility of all stages as described above. Four combinations of link functions were tested, including logistic functions for all parameters, loglog link functions for all parameters and two combinations of log-link function for $f_j(z)$ with logistic or loglog link functions for $s_j(z)$. Each of these four combinations was tested using positive or negative covariance between survival and fertility (Figure 1). In the scenarios of demographic lability with forced buffering, mature stages were defined as all stages equal to or larger than the stage with first non-zero fertility, and immature stages as all stages preceding this stage. Either $s_{\text{mature}}(z)$ or all $s_j(z)$ (scenarios 9–11), or all $f_j(z)$ and $s_{\text{immature}}(z)$ (scenarios 12–13) were held constant and equal to their value in the mean environment as reported in the standardised COMADRE MPM. We used logistic functions for the other demographic parameters (Figure 1c). These scenarios reflect different assumptions of demographic lability and buffering within the least or the most influential demographic parameters on population growth, assessed qualitatively depending on the position of the populations along the slow–fast continuum of life histories (Gaillard & Yoccoz, 2003; Sæther & Bakke, 2000; Stearns, 1989). Survival of immature stages and fertility coefficients are assumed to show a higher contribution to population growth in fast-living species, while survival rates of the mature stages are assumed to be more influential for slow life histories.

Decomposition

For each population in each scenario, we calculated and decomposed the stochastic growth rate $\ln(\lambda_s)$ following Equation (3). Since all the MPMs from COMADRE were standardised so that $\ln(\lambda_0) = 0$, the stochastic growth rate is a sum of the nonlinearity and the variance–covariance components. The sign of the stochastic growth rate directly reflects whether the fitness effects of environmental variance (σ_z^2) are positive or negative in that population and scenario. All calculations shown in the main text use the value $\sigma_z^2 = 1$, and altering this value only affects the magnitude of the effects (Supporting Information S4). In our analyses, $|\beta_{z_F}|$ and $|\beta_{z_S}|$ were both set to 0.4 (Figure 1a; results for other values shown in Supporting Information S4).

RESULTS

Combined effects of nonlinearity and variance–covariance among demographic parameters

In all scenarios, life histories with short to intermediate generation times (<10 years) showed consistently stronger fitness responses to environmental variability than slow life histories (Figures 2 and 3). Whether these responses are positive or negative, strongly depends on the combined impacts of covariance structure between the (st)age-specific survival rates and fertility coefficients and their curvatures.

Positive effects of lability on the mean fitness $\ln(\bar{\lambda})$ were found mainly among the fast-living species, and positive effects occurred through both survival and fertility (Figures 2 and 3, right panel). The nonlinearity index D correlated almost perfectly with this nonlinearity component (i.e.

$\ln \bar{\lambda} - \ln \lambda_0$ in Equation 3; Spearman coefficient >0.999 and 0.928 in all scenarios without and with bony fish MPMs respectively), suggesting that this is a reliable indicator of adaptive lability. However, as fitness $\ln(\lambda_s)$ also depends on the variance–covariance structure of the demographic parameters, this must also be taken into account.

With positive covariance between survival and reproduction, $\ln(\lambda_s)$ was consistently reduced compared to the mean environment, regardless of the type of link functions used (e.g. Figure 2a,b). In these scenarios, positive nonlinearity components still occurred, but were not sufficient to overcome the negative variance–covariance component. In contrast, populations of fast-living species could show an overall positive fitness $\ln(\lambda_s)$ if survival and fertility covaried negatively (Figures 2 and 3), although less frequent when $s_j(z)$ and $f_j(z)$ were defined by loglog link functions (Figure 2d). Positive effects were stronger when

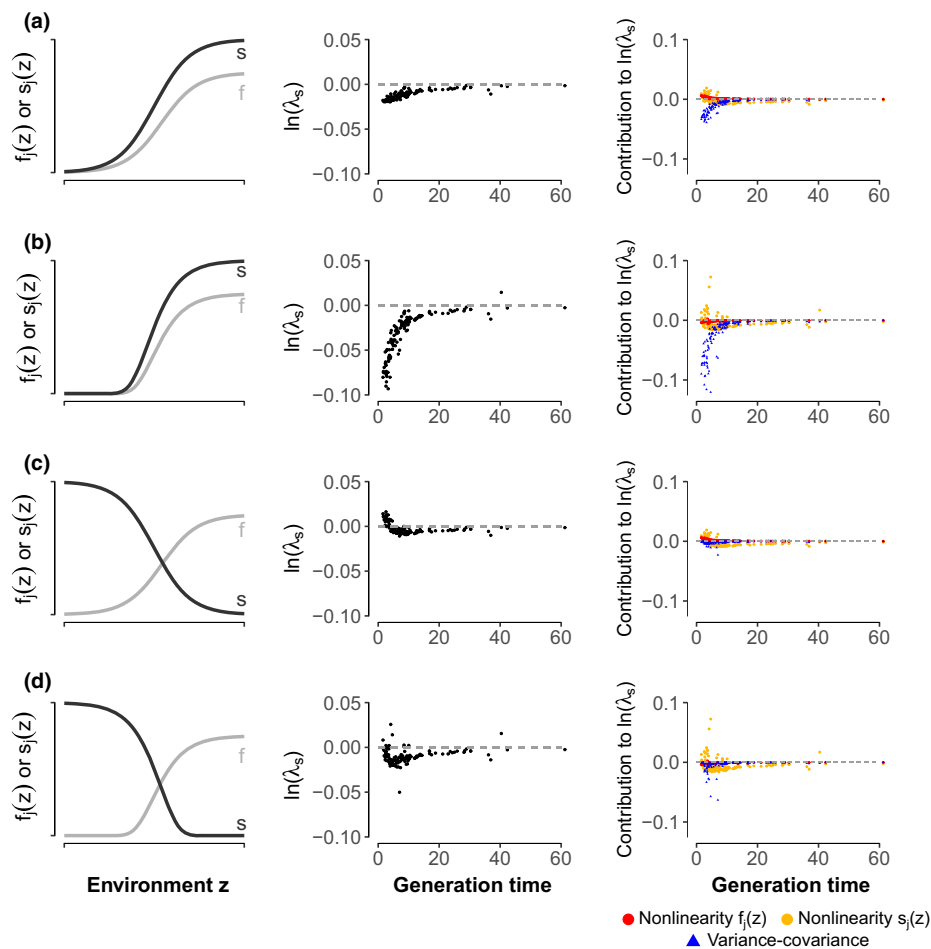


FIGURE 2 Mid panels: Stochastic growth rate (fitness) $\ln(\lambda_s)$ across generation time, under four scenarios of covariance and link functions of the demographic parameters. Left panels: Illustration of scenarios, with black and grey lines corresponding to the (st)age-specific survival rates $s_j(z)$ and fertility coefficients $f_j(z)$ respectively (functions varied for each stage depending on $s_j(0)$ and $f_j(0)$; only one function is shown for survival and fertility here). We assumed positive covariance between survival rates of different (st)ages and between the fertilities of different (st)ages. For each scenario and for each population, positive (panels a, b) or negative (panels c, d) covariance between $f_j(z)$ and $s_j(z)$ were considered, treating $f_j(z)$ and $s_j(z)$ as logistic functions (panels a, c) or loglog link functions (panels b, d) of the environment z . Right panels: Decomposition of $\ln(\lambda_s)$ into main components capturing variance–covariance effects (blue triangles) and lability effects generated by nonlinear responses of $f_j(z)$ (red circles) and $s_j(z)$ (orange circles). Results for bony fish populations and populations with generation time >62 years are not shown ($N_{\text{MPMs}} = 129$; see Figure S14 for all MPMs).

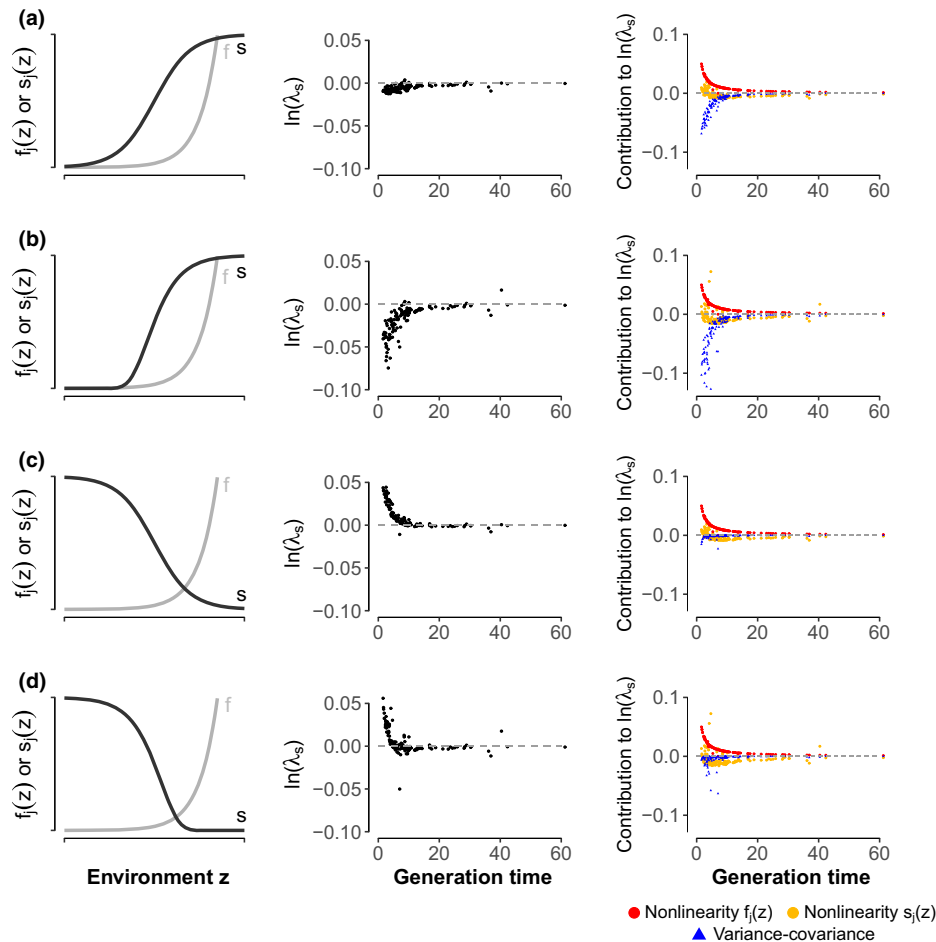


FIGURE 3 Mid panels: Stochastic growth rate (fitness) $\ln(\lambda_s)$ across generation time, considering positive (panels a, b) or negative (panels c, d) covariance between (st)age-specific survival rates $s_j(z)$ and fertilities $f_j(z)$, treating $s_j(z)$ as logistic (panels a, c) or loglog (panels b, d) link functions of the environment z and $f_j(z)$ as log link functions. We assumed positive covariance between survival rates of different (st)ages and between the fertilities of different (st)ages. See Figure 2 for explanation of left and right panels. Results for bony fish populations and populations with generation time >62 years are not presented ($N_{\text{MPMs}} = 129$; see Figure S15 for all MPMs).

we used a log-link function for the fertility coefficients, so that they increased exponentially with the environmental driver z leading to strong convexity (Figure 3c,d). For bony fish MPMs, the signs of the nonlinearity and variance-covariance components were the same as for the other MPMs, but the magnitude was stronger. Here the underlying models from COMADRE showed very high fertility coefficients and low survival rates, yielding extremely high variance in demographic parameters. Under scenarios using loglog link functions for $s_j(z)$ and/or $f_j(z)$, the small noise assumption behind our decomposition of $\ln(\lambda_s)$ was violated to a degree where the approximation broke down for these MPMs (Supporting Information S6).

Demographic lability with forced buffering

In these scenarios, some survival probabilities and fertility coefficients were kept constant and buffered, while others were allowed to vary. The identity of labile demographic parameters, together with the position of the species

along the slow-fast continuum, affected each fitness component and their combined impact on fitness (Figure 4). When lability in all survival rates $s_j(z)$ or in only the mature stages $s_{\text{mature}}(z)$ was combined with a constant fertility (Figure 4a,b), only the fastest-living species showed a positive $\ln(\lambda_s)$. This positive fitness resulted from a positive nonlinearity effect of survival rates and a low negative variance-covariance effect, reflecting buffering. When lability in fertility $f_j(z)$ and survival rates of the immature stages $s_{\text{immature}}(z)$ was combined with constant survival rates in all stages or mature stages, positive values of $\ln(\lambda_s)$ were also detected, especially when immature survival rates and reproduction covaried negatively (Figure 4c-e).

In contrast, for intermediate and slow-living species, labile survival rates of the reproductive stages $s_{\text{mature}}(z)$ combined with constant fertility $f_j(z)$ and constant survival of immature stages $s_{\text{immature}}(z)$ (Figure 4b) always produced negative nonlinearity components, and very small negative variance-covariance components, leading to an overall negative $\ln(\lambda_s)$. The scenarios of lability in fertility coefficients combined with constant (st)age-specific

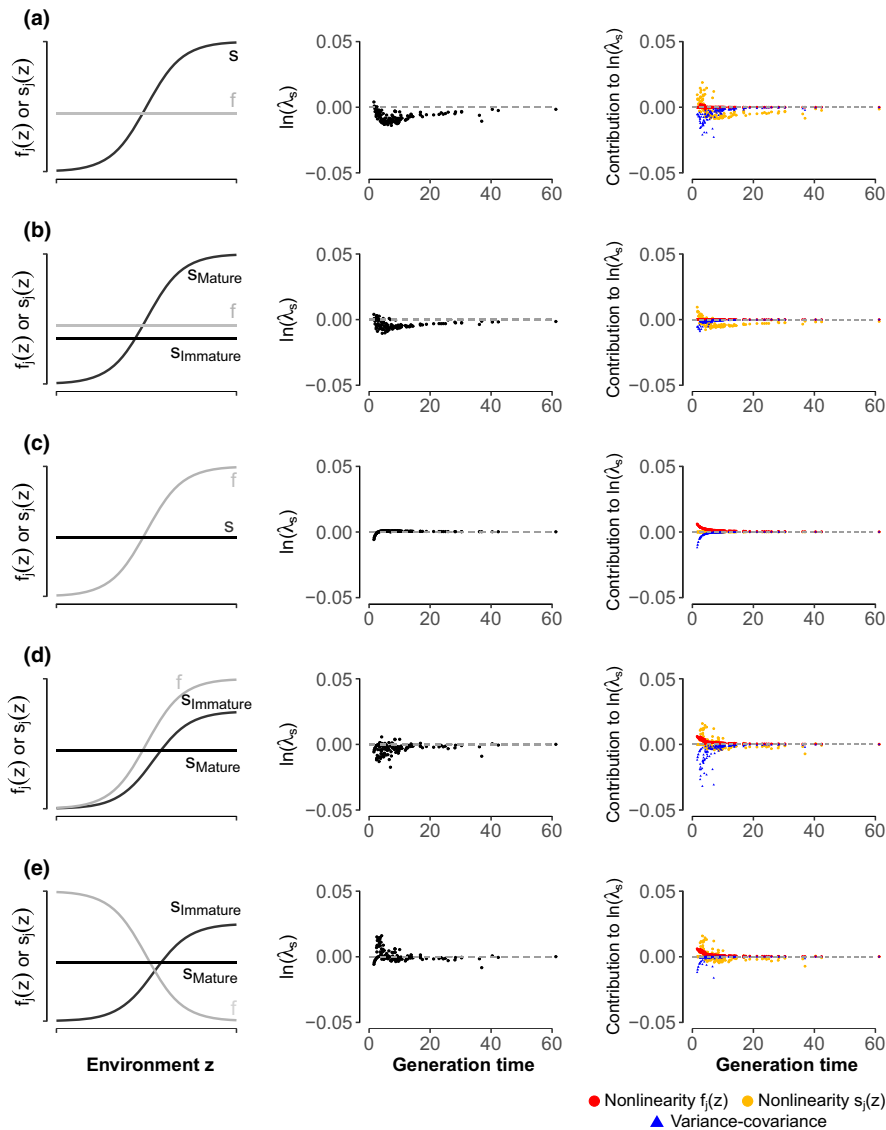


FIGURE 4 Results from scenarios of forced buffering assuming demographic liability only in (a) (st)age-specific survival rates, (b) survival rates of the mature stages only, (c) (st)age-specific fertilities and (d, e) fertilities and survival rates of the immature stages. For each scenario, the long term fitness $\ln(\lambda_s)$ and its main components reflecting variance–covariance effects (blue triangles) and liability effects due to nonlinearity of $f_i(z)$ (red circles) and $s_i(z)$ (orange circles) are plotted against generation time (mid and right panels; see Figure S16 for all MPMs). See Figure 2 for explanation of left panel.

survival rates or in only the mature stages (Figure 4c–e) showed a weak negative variance–covariance component while the nonlinearity component was zero or slightly positive, leading to overall fitness $\ln(\lambda_s)$ having values close to zero. In other words, constant (st)age-specific survival rates associated with labile fertility coefficients have a stabilising effect on $\ln(\lambda_s)$ of slow life histories (generation time >10 years; Figure 4c vs. Figure 4a,b).

DISCUSSION

This study emphasises the importance of considering explicit links between environmental drivers and demographic parameters to understand the effects of

environmental variability on fitness, as these links allow effects on nonlinearity to be quantified. We extended Tuljapurkar's approximation of the stochastic growth rate to incorporate effects of nonlinearity in demographic parameters. We also defined a nonlinearity index to measure the overall nonlinearity in a given life history, reflecting the potential for positive fitness effects of environmental variability. Our decomposition of the stochastic growth rate into nonlinearity and variance–covariance components creates a new framework to study their joint impacts on fitness, expanding earlier theory focusing mainly on buffering through the variance–covariance component. Applying this decomposition across a range of scenarios and life histories, we identified the faster-living species as the most responsive

to environmental fluctuations, both through the nonlinearity and variance–covariance components. Positive fitness values were only found when positive nonlinearity components were combined with negative covariance between survival and fertility, leading to a smaller negative variance–covariance component. In scenarios with some demographic parameters being constant (forced buffering), lability in both the least and the most expected influential demographic parameters was found to benefit fitness to some extent, but mainly for short-lived species. Our decomposition provides a step forward in our understanding of potential adaptations to environmental variability in a wide range of life histories, and stresses the importance of characterising both nonlinearity and covariance structure of demographic parameters with respect to key environmental drivers. Our framework is also useful for predicting population responses to increased variability under global change.

Lability and buffering in fast versus slow life histories

Several studies have shown evidence that populations located at the fast end of the slow–fast continuum are more sensitive to changes in the different components of climate change. These populations tend to respond more strongly to changes in climate drivers (e.g. Compagnoni et al., 2021), to environmental variability (e.g. Dalglish et al., 2010; Drake, 2005; Koons et al., 2009; Morris et al., 2008, but see Le Coeur et al., 2021; Santos et al., 2021), to shifts in temporal autocorrelation in the environment (e.g. Paniw et al., 2018), and to shifts in the correlation structure of demographic parameters (Iles et al., 2019). In line with these previous studies, we found that populations of faster-living species have larger absolute values of both nonlinearity and variance–covariance components of fitness in a stochastic environment compared to those of slow-living species. On one hand, fast-living species are more vulnerable to environmental fluctuations due to higher negative variance–covariance components, as reported in previous studies (e.g. Dalglish et al., 2010; Morris et al., 2008). On the other hand, they have the largest potential for adaptive lability through convex demographic responses. Our results show that a positive nonlinearity component can overcome the negative variance–covariance and lead to increased fitness especially when there is a negative correlation between fertility and survival (we assumed a perfect correlation in our analysis). We found that the nonlinearity index D is a reliable predictor of the nonlinearity component of the stochastic growth rate (Equation 3).

A majority of studies have focused on effects of the variance–covariance component alone, without explicit reference to the underlying environmental drivers, even though other studies (Drake, 2005; Henden et al., 2008; Koons et al., 2009) highlighted the potentially critical

importance of including such links. Our results support this conclusion, and show that the total impact of environmental fluctuations on the fitness of structured populations may be either positive or negative if nonlinear demographic responses are present (Equation 3). The net effect depends on the variance–covariance component, the shape of the environment–demographic parameter relationship, and the strength of the nonlinearity in demographic parameters. Evidence of convex relationships between demographic parameters or underlying vital rates and key environmental drivers is still limited for natural populations, due to data limitation or a priori linear assumptions in the statistical models. Our study highlights the need for empirical research to determine more systematically the shape and curvature of demographic parameter responses to accurately predict fitness responses to environmental variance. Quantifying the relationships between environmental drivers and all demographic parameters remains, however, a statistical challenge for wild populations (e.g. separating link functions; Gill, 2001) and requires long-term monitoring data (see Lee, 2017 for an alternative method to study nonlinearity in the growth rate response to an environmental driver with discrete levels). This highlights the need to continue and increase the ongoing collection of demographic data.

The decomposition of the stochastic growth rate considers nonlinearity and variance–covariance of demographic parameters, which in turn are functions of underlying vital rates. For instance, fertility depends on both fecundity and survival of offspring or parents, depending on the census of the matrix model. Studies applying the method for specific empirical systems should carefully consider how the demographic parameters depend on lower-level parameters as functions of environmental drivers. Our qualitative conclusions on demographic buffering and lability across generation time are general, but quantitative differences are likely present for instance for models based on pre- versus post-reproductive census, when environmental effects arise through lower-level parameters. This presents an interesting area for future research using the decomposition.

Role of temporal covariance between (st)age-specific demographic parameters

While negative covariance between demographic parameters could arise from life history trade-offs (Stearns, 1989) or opposite responses to the same environmental driver, positive covariances between these parameters are just as likely to occur in a population. Previous theoretical work has shown that positive covariance enhances the variance in population growth while negative covariance reduces it (Tuljapurkar, 1982, 1990). Our results are in line with this result, showing reduced negative variance–covariance component when survival and fertility covaried negatively compared to positively.

Interestingly, there is no general consensus on the degree to which positive or negative covariance in demographic parameters are more common in the wild, nor if the sign, magnitude or type of (st)age-specific demographic parameters involved correlate with the position of a species along the fast–slow continuum (but see a recent comparative study, Fay et al., 2022). From empirical studies, positive covariances have been reported predominantly in long-lived species (e.g. Dahlgren et al., 2016; Rotella et al., 2012; van de Pol et al., 2010) with substantial (e.g. Coulson et al., 2005) or weak (e.g. Altwegg et al., 2007; Compagnoni et al., 2016; Johnson et al., 2010) effects on fitness. In contrast, negative covariances were less often detected (Fay et al., 2022), with often small effects on $\ln(\lambda_s)$. To our knowledge, relatively few studies have specifically addressed this question among species towards the fast end of the continuum.

In our scenarios, we assumed a perfect, positive temporal covariance between (st)age-specific survival rates and between (st)age-specific fertilities, respectively, but positive or negative covariances between survival and fertility. While these assumptions on the direction of covariance between stages and type of demographic parameters are plausible, they are strong in terms of magnitude, and a main environmental driver will never explain all of the (co)variance in demographic parameters. Our results may therefore overestimate the magnitude of the variance–covariance component in the decomposition, compared to wild populations where correlations are likely not perfect. Even though we assumed perfect correlation, we found that variance–covariance had negligible effects on fitness of slow-living populations, reflecting a large degree of buffering in these species. For fast-living species, covariance had contrasting effects on the fitness components. These effects were strengthened in scenarios where link functions implied more asymmetric relationships between demographic parameters and environmental driver.

Demographic lability and buffering of different demographic parameters

The set of scenarios combining lability in some demographic parameters with forced buffering in others, yielded insights into possible demographic strategies along the slow–fast continuum. While different predictions have been made as to which demographic parameters should be selected for lability (Hilde et al., 2020; McDonald et al., 2017), we found that demographic lability in the demographic parameters assumed to be the most ($f_j(z)$ and/or $s_{immature}(z)$) or least ($s_{mature}(z)$) important to fitness, could both lead to enhanced fitness in many fast-living life histories due to positive nonlinearity components and reduced variance–covariance components. However, such positive effects on fitness were stronger and more prevalent with lability in both survival

of the immature stages and fertility (the most influential parameters in fast life histories). In contrast, for slow-living life histories, lability in the survival rates of mature stages, believed to have the highest impact on fitness led to negative effects on fitness due to negative nonlinearity components. Selection for a reduction in variance in (and in positive covariance between) the demographic parameters that contribute the most to fitness, combined with other parameters varying more freely, as stated by the demographic buffering hypothesis, seems likely for slow-living species, at least those with a similar animal life history as in our analysis.

The least and the most influential demographic parameters in our scenarios were qualitatively assigned based on expectations from the demographic buffering and life history theories (Gaillard & Yoccoz, 2003; Sæther & Bakke, 2000; Stearns, 1989). This simple categorisation, while accurate for some life histories, may be different for other populations with the same generation time. Further insights would require quantitatively differentiating the least and most influential (st)age-specific demographic parameters and underlying vital rates in a population based on elasticities of the growth rate in the mean environment.

In conclusion, this study provides a comprehensive framework for assessing the contributions of demographic lability and buffering on fitness of any given population. Positive effects of environmental fluctuations on fitness are only possible to detect if we account for the impacts of nonlinear relationships between demographic parameters and environmental drivers. Our decomposition of the stochastic growth rate into components of nonlinearity and variance–covariance provides a tool to quantify their relative impacts in different life histories and scenarios, and is easily applicable for other study systems and scenarios not considered here. Across the slow–fast continuum of animal life histories, faster-living species have the largest potential for using demographic lability as an adaptive response to variability, while demographic buffering is a main adaptive response in slow-living species. These findings have important implications for predicting population and species responses to changes in environmental fluctuations under climate change and other anthropogenic impacts.

AUTHOR CONTRIBUTIONS

Yngvild Vindenes and Christie Le Coeur designed the study with inputs from Roberto Salguero-Gómez and Nigel G. Yoccoz; Christie Le Coeur performed modelling work and analyses; Yngvild Vindenes did the analytical derivations; Christie Le Coeur wrote the first draft of the manuscript and all authors contributed substantially to revisions.

ACKNOWLEDGEMENTS

Dave Koons and two anonymous reviewers provided helpful comments that considerably improved the

manuscript. This study was supported by the Research Council of Norway (project ECOVAR, grant no. 244404).

FUNDING INFORMATION

Norges Forskningsråd, Grant/Award Number: 244404

PEER REVIEW

The peer review history for this article is available at <https://publons.com/publon/10.1111/ele.14071>.

DATA AVAILABILITY STATEMENT

The data supporting the results and R code used to decompose the stochastic growth rate and to conduct stochastic perturbations are archived on FigShare (DOI: [10.6084/m9.figshare.20065973](https://doi.org/10.6084/m9.figshare.20065973)). Matrix population models are available at www.compadre-db.org.

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SUPPORTING INFORMATION

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How to cite this article: Le Coeur, C., Yoccoz, N.G., Salguero-Gómez, R. & Vindenes, Y. (2022) Life history adaptations to fluctuating environments: Combined effects of demographic buffering and lability. *Ecology Letters*, 25, 2107–2119. Available from: <https://doi.org/10.1111/ele.14071>