

Bargaining and Wage Rigidity in a Matching Model for the US*

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Abstract

This paper uses robust econometric methods to assess previous empirical results for the Mortensen and Pissarides (1994) matching model. Assuming all wages are negotiated each period is inconsistent with the history dependence in US wages, even allowing for heterogeneous match productivities, time to build vacancies and credible bargaining. Flexible wages for job changers, with rigid wages for job stayers, allows the model to capture this history dependence and is not inconsistent with parameter calibrations in the literature. Such wage rigidity affects only the timing of wage payments over the duration of matches; conclusions about other characteristics are unaffected by it.

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1 Introduction

The Mortensen and Pissarides (1994) matching model is the basis for most recent discussions of unemployment and vacancies at the macroeconomic level in the US.¹ The empirical implementations in Cole and Rogerson (1999), Yashiv (2000), Shimer (2005), Yashiv (2006) and Hagedorn and Manovskii (2008) assume Nash bargaining of wages in all matches each period. Hall (2005a), Hall (2005b), Hall and Milgrom (2008) retain bargaining of wages in all matches each period but replace Nash by other forms of bargaining. Others, such as Shimer (2004), Gertler and Trigari (2009), Pissarides (2009), Rudanko (2009), Rudanko (2011), Haefke et al. (2013) Kudlyak (2014) and Gertler et al. (2015), argue that introducing some form of wage rigidity enables the matching model to better capture the relationship between unemployment, vacancies and wages in US data. There is also a substantial literature on New Keynesian models with search and matching frictions that has looked at wage persistence (Christoffel et al, 2009, Krause and Lubik, 2007, Krause et al 2008, Trigari, 2009).

In most of these studies, empirical results are based on calibration or full-information estimation, both of which require strong assumptions to pin down the distribution of the data. The studies using calibration rely on specific assumptions about the distribution of shocks. Those using full-information estimation rely on other aspects of the full model being correctly specified. In addition, all empirical results reported in the literature rely on the assumption of strong identification, since none of the papers use methods that are robust to weak identification or weak instruments. Hence the rejection of the spot market model, and evidence on any of the proposed extensions, is conditioned on those assumptions.

The contribution of this paper is to assess the robustness of results in previous studies to dropping those strong assumptions. We use limited-information analysis so that our results do not rely on correct specification of the other aspects of the model.² Moreover, the models we estimate

¹There is also a growing literature applying disaggregated versions of the matching model with heterogeneous firms and employees to micro data. For examples, see Cahuc et al. (2006) and Robin (2011).

²This approach has been used successfully in many other areas of macroeconomics. e.g., see Gali and Gertler (1999) for the New Keynesian Phillips curve.

are forward-looking rational expectations models and this characteristic provides straightforward criteria for determining valid instruments. Furthermore, our estimation is based on the generalized method of moments (GMM) making use of methods of inference that are robust to weak instruments, see Stock and Wright (2000). This innovation relative to the literature on matching in wage determination is important because weak instruments are known to pose a very serious threat to empirical validity across many areas of economics, see, e.g., Stock *et al* (2002).

The paper establishes two main robust findings on the basis of these estimation methods. First, none of the formulations in the aforementioned literature with wages in all matches negotiated each period satisfactorily captures the history dependence in wages. Second, there is a formulation of wage rigidity that does so and, consistent with the micro evidence reported in Pissarides (2009) and reiterated by Haefke et al. (2013), applies only to continuing matches.

This second finding has important implications. If wage rigidity applies to new, as well as continuing, matches, as in Gertler and Trigari (2009), it affects job creation and hence vacancies and unemployment. But, as pointed out by Malcomson (1999, Section 4) and Pissarides (2009), if it applies only to continuing matches, it affects only the timing of wage payments over the duration of a match. So conclusions for vacancies and unemployment drawn from studies without wage rigidity continue to apply.³

The underlying problem for capturing the history dependence of wages with models in which wages in all matches are determined by the Nash bargain in every period is illustrated in Figure 1. This compares the variation in the wage predicted by the calibration in Hagedorn and Manovskii (2008) with the average wage in the data. The residuals are the difference between these. It is apparent from the figure that the residuals are highly persistent (as measured by the autocorrelation function), which is contrary to what one would expect if the model captured adequately the history dependence of wages in the data.

Subsequent contributions have extended the model with all wages negotiated each period to address this problem. Hall and Milgrom (2008) consider a different model of bargaining, *credible*

³Gertler et al. (2015) criticizes this conclusion but does not test for history-dependence in wages, the issue we address in this paper.

bargaining. Mortensen and Nagypál (2007) allow for separation shocks in addition to productivity shocks. Hagedorn and Manovskii (2011) allow for “time to build” in vacancy creation. Hagedorn and Manovskii (2013) add heterogeneity to match productivities and allow “on the job” search by workers. To these can be added the fixed costs (for training, negotiation or administration) incurred *after* matching suggested by Pissarides (2009). In this paper, we construct an empirically implementable model that encompasses all these when all wages are negotiated each period and show that, even with them all combined, the model still does not satisfactorily capture the history dependence of wages in the US.

As an alternative to all wages being negotiated in each period, this paper constructs an empirically implementable model with wage rigidity for continuing matches but wages negotiated afresh for all new matches. The form of wage rigidity is that developed in Gertler and Trigari (2009) but applied only to continuing, not to new, matches. This formulation satisfactorily captures the history dependence of wages in the US, despite having essentially only one additional parameter. Alternative forms of wage rigidity for continuing matches are developed by Thomas and Worrall (1988), Beaudry and DiNardo (1991), Rudanko (2009) and MacLeod and Malcomson (1993). The first three, however, depend on workers being risk averse, which is not part the standard matching model, and MacLeod and Malcomson (1993) is harder to implement empirically with aggregate data. In any case, the purpose here is not to select between different models of wage rigidity for continuing matches but to show that a wage equation with these general characteristics can capture the history dependence of wages. The model we use is an empirically tractable one that suffices for this purpose.

The paper is organized as follows. The next section sets out the theoretical wage equations we use for econometric analysis. That is followed by sections on empirical specifications, the data and estimation results. These are, in turn, followed by a conclusion. Appendix A sets out the full details of the theoretical model from which the wage equations are derived. Appendix B contains further derivations of equations in Appendix A, Appendix C additional information about the data used, and Appendix D supplementary empirical results and robustness checks.

2 Theoretical wage equations

2.1 Wage equations without wage rigidity

The basic framework used here for wage equations without wage rigidity is the matching model of Mortensen and Pissarides (1994) as developed by Hagedorn and Manovskii (2008) for empirical analysis. The model consists of five equations: (1) a value equation for a filled job, (2) a value equation for an unfilled vacancy that is set to zero because free entry is assumed, (3) a value equation for an employed worker, (4) a value equation for an unemployed worker, and (5) a Nash bargaining equation that determines wages. We adapt those equations to encompass the credible bargaining model of Hall and Milgrom (2008) as an alternative to the Nash bargaining model, “time to build” in vacancy creation as suggested in Hagedorn and Manovskii (2011), heterogeneity in match productivities as suggested in Hagedorn and Manovskii (2013) and fixed costs incurred only *after* matching as suggested by Pissarides (2009), together with some minor generalizations of inessential restrictions that there is no reason to require the data to satisfy. The full model is set out in Appendix A. There we derive a wage equation that, when wages in all matches are negotiated each period and with the notation in Table 1, takes the form

$$w_t = \frac{z_t}{1 + \lambda} + \frac{\lambda}{1 + \lambda} \left[p_t + c_t \frac{f_t}{(1 - f_t \kappa) q_t} \right] + \frac{\lambda}{1 - \lambda^2} \frac{\gamma_t - E_t (\delta_{t+1}^\kappa \gamma_{t+1})}{1 - f_t \kappa}. \quad (1)$$

This reduces to exactly the wage equation in Hagedorn and Manovskii (2008) when the parameters incorporating the extensions to the model are set to appropriate values; specifically, $\kappa = 0$ (employment starts at $t + 1$ for new matches at t) and $\gamma_t = 0$ for all t (no cost to making offers so bargaining is Nash). Given appropriate specifications for z_t , c_t and γ_t , (1) is an equation for the average wage that can be estimated from available data.

To interpret (1), start with Nash bargaining ($\gamma_t = 0$ for all t). If employment lasted only a single period, the worker would receive payoff z_t and the firm 0 (because, at the bargaining stage, c_t is a sunk cost) if they do not form a match at t . If they form a match, the parties have productivity p_t to

share between them. So Nash bargaining would result in $w_t = z_t + \beta (p_t - z_t)$, where $\beta \in [0, 1]$ is the bargaining power of the worker. (To encompass the credible bargaining model, β is replaced by $\lambda / (1 + \lambda)$ in (1).) With a continuing match, there is also the future to consider. With Nash bargaining at $t + 1$ as well as at t , the worker's expected future gains are proportional to the firm's expected future gains. Moreover, because of free entry, the firm's expected gains from $t + 1$ on equal the cost c_t of posting a vacancy at t less the period t gains. So, all future payoffs can be written explicitly in terms of variables known at t , including c_t as in (1). See Appendix A for the detailed derivation. The term in (1) including γ_t and γ_{t+1} needs to be added to incorporate credible bargaining. For that model, λ has a different interpretation, see Table 1, but it still affects the other terms in the same way as with Nash bargaining. when individual productivity is not observed

Equation (1) is also consistent with “on the job” search as modelled in Hagedorn and Manovskii (2013), see Appendix A. In their paper, history dependence refers to the dependence of individual wages on past labour market conditions when individual match productivity is not observed. They show that such history dependence can arise through history dependence in individual match productivity induced by “on the job” search. Because equation (1) aggregates over all matches and aggregate productivity is measurable, such history dependence is already captured through our measure of productivity.

For reasons explained above, wage equation (1) depends on the future gains from the relationship being satisfactorily measured in terms of the cost c_t of creating a vacancy at t . It therefore depends on the value equation for an unfilled vacancy being properly specified. If it is not, (1) is mis-specified even if wages in all matches are negotiated every period. In particular, there are some specifications of the time to build a vacancy in Hagedorn and Manovskii (2011) that affect the value of an unfilled vacancy but are not captured in the specification used to derive (1). Thus, our subsequent finding that (1) does not fit the data could result purely from this mis-specification, not because some wages are not negotiated every period.

To rule out this possibility, we derive in Appendix A an alternative wage equation when all wages are bargained each period that, although making use of the free entry condition that the

value of creating a vacancy is zero, does not rely on the specification of the equation for that value. This wage equation has the form

$$(1 - f_t \kappa) (w_t - z_t) - \lambda (p_t - w_t) + E_t \sum_{n=1}^{\infty} [\delta_{t,n}^{\kappa} - \delta_t \delta_{t+1,n-1}^{\kappa} (1 - s_{t+1})] (1 - f_{t+n} \kappa) (w_{t+n} - z_{t+n}) - \frac{\lambda}{1 - \lambda} [\gamma_t - \delta_t E_t (1 - s_{t+1}) \gamma_{t+1}] = 0. \quad (2)$$

where $\delta_{t,n}^{\kappa} = \prod_{i=1}^n \delta_{t+i}^{\kappa}$, with $\delta_{t,0}^{\kappa} = 1$, and $\delta_t^{\kappa} = \delta_{t-1} (1 - s_t - f_{t-1} + \kappa s_t f_{t-1})$. Instead of capitalising a worker's future gains from forming a match in terms of the cost c_t of creating a vacancy, wage equation (2) spells out those future gains explicitly in the terms $w_{t+n} - z_{t+n}$ for $n \geq 1$. It is thus robust to the specification of the equation for the value of creating a vacancy and, in particular, remains valid for any length of time to build a vacancy. It thus encompasses all the specifications in Hagedorn and Manovskii (2011). Estimation of (2) requires the terms under the summation sign to be truncated at some finite horizon.⁴ With δ_t^{κ} strictly less than 1 as implied by the model, however, the approximation error from truncation can be made arbitrarily small for a sufficiently long horizon. Given appropriate specifications for z_t and γ_t , (2) is then an equation in current and future average wages that can be estimated from available data.

2.2 Wage equation with wage rigidity

Gertler and Trigari (2009) use a form of wage rigidity that enables the matching model to account for the cyclical behaviour of wages and labour market activity. Their form of wage rigidity has a fixed probability that the wage for a match, whether new or continuing, is bargained in any one period. But Haefke et al. (2013) find little evidence of wage rigidity for new hires at the microeconomic level, reinforcing the micro evidence surveyed by Pissarides (2009). Here, therefore, we use a model in which wages in continuing matches are subject to wage rigidity of the type analysed by

⁴This approach has been used by, for example, Rudd and Whelan (2006) for studying the new Keynesian Phillips curve.

Gertler and Trigari (2009) but those in new matches are all bargained.

In Appendix A, we derive a wage equation for the average wage in new matches, w_t^* , when the wage for continuing matches is renegotiated with probability $1 - \psi$. We allow for the possibility that wages not renegotiated may be adjusted automatically to inflation by scaling them by the factor $\pi_t^{-\mu}$, where π_t is the ratio of prices at t to prices at $t - 1$ and μ is a parameter to be estimated. Wages here are measured in real terms. Thus, for $\mu = 1$, the unrenegotiated wage is set in nominal terms, for $\mu = 0$ in real terms, with $\mu \in (0, 1)$ interpreted as the proportion of unrenegotiated wages set in nominal terms. The wage equation then takes the form

$$\begin{aligned} (1 - f_t \kappa) (w_t^* - z_t) - \lambda \left[(1 - f_t \kappa) (p_t - w_t^*) + c_t \frac{f_t}{q_t} \right] \\ - \frac{\lambda}{1 - \lambda} [\gamma_t - E_t (\delta_{t+1}^\kappa \gamma_{t+1})] + \psi (1 + \lambda) E_t \left\{ \delta_{t+1}^\kappa \left(\pi_{t+1}^{-\mu} w_t^* \right. \right. \\ \left. \left. - w_{t+1}^* \right) \left[\sum_{i=1}^{\infty} \prod_{j=2}^i (\delta_{t+j-1} (1 - s_{t+j}) \psi \pi_{t+j}^{-\mu}) \right] \right\} = 0, \quad (3) \end{aligned}$$

where, as before, $\delta_t^\kappa = \delta_{t-1} (1 - s_t - f_{t-1} + \kappa s_t f_{t-1})$. When $\psi = 0$, $w_t^* = w_t$ and wage equation (3) reduces to (1). In (1), with Nash bargaining only contemporaneous variables appear. (With credible bargaining, the costs of making offers at $t + 1$ also appear.) Wage rigidity gives rise to forward-looking behaviour because the wage currently negotiated may continue to apply at future dates, which is captured by $\psi \neq 0$. Thus, our specification of wage rigidity involves relaxing the single restriction $\psi = 0$, which turns out to make a big difference empirically.

The data contain the average wage w_t but not the average wage for new matches w_t^* . Appendix A shows that, under the assumptions of the model, the relationship between these is given by

$$w_t^* = \frac{w_t - \pi_t^{-\mu} w_{t-1}}{1 - \psi (1 - s_t) j_{t-1}/j_t} + \pi_t^{-\mu} w_{t-1}. \quad (4)$$

As with (2), estimation of (3) requires the terms under the summation sign on the right-hand side to be truncated at some finite horizon. With $\delta_{t+j-1} (1 - s_{t+j}) \psi$ strictly less than 1, the

approximation error from truncation can be made arbitrarily small for a sufficiently long horizon.

3 Empirical specifications

Empirical implementation of (1)–(3) requires specifications for z_t , γ_t and c_t . In Hagedorn and Manovskii (2008), productivity is detrended and z_t is a constant. The corresponding assumption here is that z_t is proportional to trend productivity, denoted \bar{p}_t .⁵ We make the same assumption for γ_t . These give the specifications

$$z_t = z \bar{p}_t, \quad \gamma_t = \gamma \bar{p}_t, \quad \text{where } z \geq 0. \quad (5)$$

For the vacancy posting cost, Hagedorn and Manovskii (2008) include two components, capital and labour. Capital costs in period t are proportional to productivity in period t and so can be written $c^K p_t$, where c^K is a non-negative constant.⁶ Labour costs in period t are proportional to the cyclical component of productivity, $\tilde{p}_t = p_t / \bar{p}_t$, raised to the power $\xi \in [0, 1]$, which Hagedorn and Manovskii (2008) interpret as the elasticity of the labour cost of those engaged in hiring with respect to productivity, and so can be written $c^W \tilde{p}_t^\xi \bar{p}_t$, where c^W is a non-negative constant. See Hagedorn and Manovskii (2008) for a detailed discussion of the motivation for these formulations. We add to these the post-matching fixed costs suggested in Pissarides (2009), denoted by H_t , which we allow to have both capital and labour components specified in ways corresponding to the vacancy posting costs in Hagedorn and Manovskii (2008), so

$$H_t = H^K p_t + H^W \tilde{p}_t^\xi \bar{p}_t, \quad H^K, H^W \geq 0.$$

⁵Specifying $z_t = z^0 \bar{p}_t + z^1 p_t$ to make it procyclical as in Chodorow-Reich and Karabarbounis (2016) would make z_t collinear with p_t , so it would not affect the fit of the model and the implications for wage rigidity. It would, however, make β unidentified unless additional restrictions are imposed.

⁶Hagedorn and Manovskii (2008) actually assume that capital costs are $c^K p_t / \bar{p}$ and normalize \bar{p} to one, but then HP filter the data, so this is equivalent to assuming that, after detrending the cost by the productivity trend, it is proportional to the business cycle variation in productivity.

Because these costs are incurred only in the event of a match, they are multiplied by the probability of matching q_t in their impact on vacancy creation. Combining all these components for c_t , we get the empirical specification

$$c_t = c^K p_t + c^W \tilde{p}_t^\xi \bar{p}_t + \left(H^K p_t + H^W \tilde{p}_t^\xi \bar{p}_t \right) q_t, \quad c^K, c^W, H^K, H^W \geq 0, \xi \in [0, 1]. \quad (6)$$

For estimation, we normalize the wage equations (1), (2) and (3) by trend productivity, which corresponds to the use of detrended productivity in Hagedorn and Manovskii (2008) and ensures all variables are stationary.⁷

4 Data

We use data on the nonfarm business sector of the USA, mainly from the Bureau of Labor Statistics (BLS) and the OECD. The data are quarterly and cover the period 1951q1 to 2011q4. Our baseline estimation results are for the period up to 2004q4 for comparability with earlier studies. A fuller description of the data is in Appendix C.

We use this data to construct model-consistent data series. The number of new matches at t , m_t , is given by the total number of filled jobs at t , j_t , less the number of continuing matches at t , $(1 - s_t) j_{t-1}$, so

$$m_t = j_t - (1 - s_t) j_{t-1}. \quad (7)$$

The stock of vacancies at the end of period t , after matching takes place, is denoted v_t . Hence, the total number of vacancies available to be filled in period t is $v_t + m_t$. The stock of unemployed workers seeking matches in period t consists of workers who were unemployed in the previous period, $l_{t-1} - j_{t-1}$ (where l_t is the labour force at t), workers who were employed in the previous period but have lost their job, $s_t j_{t-1}$, and net new entrants to the labour force, $\Delta l_t = l_t - l_{t-1}$,

⁷We measure trend productivity using the HP filter with parameter 1600, as in Hagedorn and Manovskii (2008). We detrend wages and productivity by the same productivity trend. The results are robust to the alternative of HP filtering each series separately, see Table 13 in Appendix D.

making $l_t - (1 - s_t) j_{t-1}$ in total. Equivalently, this is given by the stock of unemployed workers at the end of the period, u_t , plus the total matches during the period, $u_t + m_t$. Thus, the probability of filling a vacancy in period t is given by

$$q_t = \frac{m_t}{v_t + m_t} \quad (8)$$

and the job-finding probability for unemployed workers by

$$f_t = \frac{m_t}{u_t + m_t}. \quad (9)$$

Employment j_t and unemployment u_t are constructed and seasonally adjusted by the BLS from the CPS. They correspond to the last month in the quarter in accordance with the model used here. Employment consists of total nonfarm dependent employment (excluding the self-employed). The labour force is the sum of employed and unemployed.

We adopt the practice discussed by Blanchard and Diamond (1990) of constructing a series for separations from the number of short-term unemployed, u_t^s , in our case (because we are using quarterly data) those with spells shorter than 14 weeks. Moreover, if the increase in the labour force all goes through the unemployment pool first, this increase should be subtracted from the short-term unemployed before calculating the separation rate. We adjusted the data for this, though the effect on the calculated series for the separation rate s_t is very small. We also adjusted for direct job-to-job flows using the procedure suggested in Shimer (2005) based on the idea that, on average, a worker losing a job has half a period to find a new one before being recorded as unemployed. Thus, short-term unemployment satisfies

$$u_t^s = \left(1 - \frac{1}{2} f_t\right) (\Delta l_t + s_t j_{t-1}).$$

Use of (7) and (9) to express f_t as $[j_t - (1 - s_t) j_{t-1}] / [l_t - (1 - s_t) j_{t-1}]$ enables us to solve for

a series for s_t that is consistent with the model.⁸ The resulting series is plotted in Figure 2. This series is higher than the monthly separation rate series reported elsewhere (e.g., Shimer (2005, Figure 7)), but it matches the cyclical pattern of the (monthly) series exactly. It illustrates the point made by Mortensen and Nagypál (2007) and by Shimer (2005) that separation rates have not been constant over this period.

Vacancy stocks v_t are measured using the Conference Board Help-Wanted Index (HWI), which is available in quarterly frequency from 1951 to 2008.⁹ The index is converted to total units using the job-openings series from the Job Openings and Labor Turnover Survey (JOLTS), which is available only since December 2000. The HWI is known to contain low frequency fluctuations, such as those resulting from newspaper consolidation in the 1960s and the internet revolution recently, that are unrelated to labour market trends, see Shimer (2005). Following Shimer (2005), we remove the effect of those trends using a low frequency filter, see Appendix C for details. The probability of filling a vacancy q_t is then calculated using data on employment and vacancies via equations (7) and (8). The resulting series is plotted in Figure 3. We also plot on the same graph the corresponding series for v_t derived using the JOLTS data over the period (2001 on) for which it is available. This shows that the two series match very closely (their correlation is 0.9).

Wages and productivity are from the BLS, which provides a measure of the labour share (including non-wage compensation) and output per person in the nonfarm business sector. We adjust for the ratio of marginal to average productivity using the scaling factor 0.679 computed by Hagedorn and Manovskii (2008). Because we use quarterly data, we specify the discount factor as $\delta_t = \frac{1}{1+r_t/4}$, where r_t is the annualized gross real interest rate, which we measure as the quarterly average of daily 3-month Treasury bill interest rates deflated using the implicit price deflator for nonfarm business obtained from the BLS.

⁸Even with the adjustment suggested by Shimer (2005), the measure of separations does not include workers moving directly from jobs to self-employment or to leaving the labour force but it is not clear how to allow for that.

⁹The HWI series based on printed newspaper advertising was replaced by online advertising after 2008. The two series have coexisted since 2005. Barnichon (2010) compiled a composite print and online HWI index that extends to 2011. For comparability with Shimer (2005) and Hagedorn and Manovskii (2008), we use the original HWI series. Robustness checks to alternative vacancy series are reported in Appendix D.

5 Estimation results

Estimation of equations (1), (2) and (3) is performed with GMM (Hansen (1982)). For robustness to weak identification, we use the continuously updated estimator (CUE) proposed by Hansen et al. (1996), and the S test proposed by Stock and Wright (2000). Confidence sets based on the S test can be empty if the identifying restrictions are rejected for all admissible values of the parameters. To check whether this is the case, it suffices to compare the minimum value of the S test statistic to its critical value. This coincides with the J statistic of Hansen (1982) that tests the validity of over-identifying restrictions, but unlike the standard Hansen test, the use of a higher critical value makes it robust to weak identification.¹⁰

Let $\phi_t(\theta)$ denote the expression on the left-hand side of (1), (2) or (3), as appropriate, with the expectations operator removed, where θ is the vector of model parameters. This is a parametric function of observed variables, whose expectation conditional on variables known at t is zero if expectations are rational. The model's testable implications can then be expressed in terms of orthogonality restrictions of the form $E[Z_t\phi_t(\theta)] = 0$, where Z_t is a vector of instruments at t . Rational expectations imply that lagged values of variables are uncorrelated with current and future error terms so, as standard in macroeconomic time-series models with rational expectations, we use lags of the variables in the model as instruments. Given the quarterly frequency of the sample, we use four lags.¹¹ For the estimation, parameter values are constrained to be consistent with the model, specifically $z, \kappa, \psi, \mu, \xi \in [0, 1]$, and $\lambda, c^K, c^W, H^K, H^W \geq 0$.

All calculations were performed using Ox, see Doornik (2007).

5.1 Results with wages bargained in all matches every period

We first present estimates for the model with wages bargained in all matches each period. Table 2 reports results for wage equation (1).

¹⁰See, for example, Mavroeidis et al. (2014, Section A.2.6).

¹¹Because results may become unreliable when the number of instruments is large, see Andrews and Stock (2007), we avoid using a larger number of instruments, but we find that our results are robust to different sets of instruments, see Appendix D.

Column 1 of Table 2 reports estimates for the specification in Hagedorn and Manovskii (2008) with κ , the fraction of newly matched jobs that become active in less than one period, and the cost parameters H^K and H^W set to 0. (The parameters γ and λ do not appear in this specification.) A “period” in Hagedorn and Manovskii (2008) is one week. Here each period is a quarter, so this specification corresponds to a longer lag between the decision to create a vacancy and the possibility of the job becoming productive, as suggested by Hagedorn and Manovskii (2011). (A shorter lag corresponds to $\kappa > 0$, which we allow for in column 2.) The point estimate of z , the value of non-work activity, is essentially identical to the calibrated value 0.936 in Hagedorn and Manovskii (2011) for the corresponding productivity series and the other parameters have large standard errors — reassuring evidence that, when applied to the same formulation, our estimation procedure yields results not inconsistent with their calibration procedure. But the over-identifying restrictions of this model (implied by errors uncorrelated with the lagged instruments we use) are overwhelmingly rejected by the Hansen test, even that using the more conservative projection p value that is robust to weak identification. Column 2 reports estimates with κ restricted only to $[0, 1]$. The point estimate of κ is zero and the value of the GMM objective function indicates that this generalization of the timing does not improve fit. Column 3 of Table 2 reports estimates allowing for the fixed costs H^K and H^W in the spirit of Pissarides (2009). This generalization does not significantly improve the fit of the model either. For both these extensions, the Hansen test continues to overwhelmingly reject the over-identifying restrictions. The standard errors are large, especially so for the cost parameters. But these standard errors are unreliable for constructing confidence intervals because there is no guarantee that the assumptions underlying standard t tests are satisfied.

An alternative to Nash bargaining is the credible bargaining of Hall and Milgrom (2008). Equation (1) with γ unrestricted and $\lambda \in [0, 1)$ corresponds to our formulation of credible bargaining for application to quarterly data. Column 4 of Table 2 reports estimates for this model. Allowing γ to be non-zero makes little difference to the fit, as measured by the value of the GMM objective function. It also leaves the over-identifying restrictions rejected just as overwhelmingly by the

Hansen test.

Rejection by the Hansen test does not provide information about which aspects of the model fail to fit the data. One way to assess this informally is by regressing the residuals of the model on the instruments. This reveals that the instrument that drives the rejection is w_{t-1} : its coefficient is close to one, while the coefficients on all other instruments are very close to zero. This supports the message from Figure 1 that it is the history dependence of wages that the model fails to capture.

Wage equation (1) is derived making use of the equation for the value of a vacancy. Rejection of (1) might, therefore, be the result of misspecification of that equation and not a rejection of wage bargaining for every job in every period. We check for this by estimating wage equation (2), which is derived without any assumption about the form of the equation for the value of a vacancy. It is consistent with any length of time to build a vacancy of the type considered in Hagedorn and Manovskii (2011). Table 3 reports the results of estimating that.

For the results in Table 3, the infinite sum in (2) is truncated at 13 quarters, which makes use of all the data available for the variables in that sum while keeping the estimation sample the same as in Table 2, for comparability with Hagedorn and Manovskii (2008). The point estimates are, however, very insensitive to the truncation length. Column 1 gives results for Nash bargaining, column 2 for credible bargaining. Because wage equation (2) does not include c_t , it does not yield estimates of the cost parameters. But the important point here is that, despite these specifications not depending on the form of the equation for the value of a vacancy, both Nash and credible bargaining formulations continue to be resoundingly rejected. The implication is that no alternative specification of the value of creating a vacancy will enable either bargaining model with wages for all jobs bargained in every period to capture the history dependence of wages in the data.

The most important point from these results is that, for all the specifications with wages negotiated in all matches in every period, the Hansen tests resoundingly reject the over-identifying restrictions with a p value of 1% or lower even using the most conservative critical values robust to weak identification, implying that the instruments are correlated with the residuals. Another implication is that the confidence intervals discussed above that are robust to weak identification

are completely empty even at the 99% level for every parameter, despite the large standard errors. There just does not exist any set of economically feasible parameter values that enable these models to capture in a statistically satisfactory way the pattern of history dependence in the data. Because the formulation allows for heterogeneous match productivity and on the job search as in Hagedorn and Manovskii (2013), this is strong evidence that these extensions of the basic matching model, either alone or together, are insufficient to enable the model with the wage for each job bargained in every period to capture the pattern of wages in the data. Moreover, the results in Table 3 imply that no alternative formulation of the value of creating a vacancy, whether to incorporate time to build as in Hagedorn and Manovskii (2011) or anything else, can overcome this.

The implication is that something more is required to enable the model to fit the history dependence of wages. It need not be wage rigidity. But in the next section we show that the addition of wage rigidity of the form modelled in Section 2.2 enables the model to do so.

5.2 Results with wage rigidity

Wage equation (3) can be used to test the Nash and credible bargaining specifications in the presence of wage rigidity of the form in Section 2.2. The results of estimating that specification, with w_t^* specified by (4) and the infinite sum truncated at 28 quarters, are reported in columns 1–5 of Table 4. (Truncation at 28 quarters makes use of all the data available up to the end of 2011 for the variables in the infinite sum while leaving the estimation sample at 1952q2–2004q4 as used by Hagedorn and Manovskii (2008). Table 12 in Appendix D gives results for other truncation lengths. Beyond 8 quarters, the point estimates are completely insensitive to the truncation length.) The results in columns 1–4 are directly comparable to the corresponding columns of Table 2. The most important finding is that, unlike the specifications without wage rigidity, these specifications all comfortably pass the Hansen tests of over-identifying restrictions at conventional levels of significance. Addition of the single extra parameter η is the reason for this. Thus for the model with wage rigidity, unlike for that without wage rigidity, there exist sets of parameter values that satisfactorily capture the history dependence of wages in US data.

The cost parameters c^K, c^W, ξ, H^K and H^W are not precisely estimated from the aggregate time-series data we use for columns 1-4 of Table 4. (This is shown formally by the confidence intervals robust to weak identification that we report below). Calibration studies typically use other data to determine values of cost parameters. In column 5, we report estimates of β and z for the Nash bargaining model (together with the wage rigidity parameters) when the cost parameters are calibrated following the procedure in Hagedorn and Manovskii (2008). Fixing the cost parameters in this way hardly affects the fit of the model (compare with column 3, the unrestricted specification) and this specification still comfortably passes the Hansen tests. Thus wage rigidity allows the model to capture the history dependence of wages with values of the cost parameters consistent with those commonly used in calibrated versions of the matching model.

Because standard errors are unreliable for constructing confidence intervals in the present context, we construct confidence intervals that are robust to weak identification using the S test of Stock and Wright (2000) in order to see how large the set of statistically acceptable parameter values is. Unlike in the models in which wages are bargained for every match in every period, confidence intervals for standard significance levels constructed in this way are not empty. Confidence intervals at the 95% and 90% level for the specifications that correspond to columns 3 and 4 in Table 4 are reported in the first two columns of Table 5, respectively. (The final column of Table 5 is discussed below). Confidence intervals for parameters not reported in Table 5 comprise the entire parameter space; they are completely uninformative. In the case of the cost parameters c^K, c^W, ξ, H^K, H^W , and γ , this is explained by the fact that they are unidentified when $\lambda = 0$ because then these parameters drop out of equation (3) and the restriction $\lambda = 0$ (which also corresponds to $\beta = 0$) is acceptable at the 10% level in both specifications.

It may help with interpreting the confidence intervals in Table 5 to explain how they are constructed. Consider the confidence interval for the parameter ψ (the parameter that determines the degree of wage rigidity). For each value of ψ , say ψ_0 , in the range of economically feasible values (in this case $[0, 1]$), we check whether there are any values of the remaining parameters such that the model's identifying restrictions are statistically acceptable at the desired level of significance.

Specifically, we compare the value of the S statistic, minimized over all parameters subject to the restriction $\psi = \psi_0$, to the appropriate 95% or 90% quantile of the χ^2 distribution with degrees of freedom equal to the number of identifying restrictions. The confidence interval contains all the values ψ_0 for which this test accepts. This procedure is repeated over a grid of values from 0 to 1, with increment .01 (so the confidence intervals are correct to two decimal places). Thus, the top row of confidence intervals for ψ in Table 5 shows that there is no set of values for the parameters other than ψ that enables the Nash bargaining specification to pass the S test at the appropriate significance level for a value of ψ less than 0.15 or greater than 0.90. But there is a set that does so for each intermediate value of ψ . For this parameter, all the confidence intervals exclude zero, confirming the results of the previous section that, without wage rigidity, the model is statistically unacceptable. They also exclude one, that is, completely rigid wages in continuing matches. However, the confidence intervals for μ cover the entire parameter space, indicating that this parameter is not identified, so the data is not sufficiently informative to distinguish between nominal and real wage rigidity.

For z , the value of non-work activity as a proportion of productivity, the confidence intervals in the first two columns of Table 5 are sufficiently wide to contain the values reported in earlier studies that used calibration. In the Nash bargaining specification, the confidence intervals for β are also wide, including values from zero to more than 0.8. The corresponding parameter λ in the credible bargaining specification has the confidence intervals that contain all the values from 0 to 0.999 that we searched over. (For $\lambda = 1$, the model is not defined.) In particular, they contain the calibrated value 0.995 in Hall and Milgrom (2008). There is thus very considerable latitude to choose parameter values based on other data sources used to calibrate matching models that will enable the model to capture the history dependence of wages if allowance is made for wage rigidity.

The column “Nash Barg calibr.” reports confidence intervals for ψ , β and z in the Nash bargaining specification with wage rigidity when $\mu = 0$ (real wage rigidity) and all other parameters are fixed at the calibrated values in column 5 of Table 4. The confidence intervals for these parameters

are very much smaller than in the other columns, so the use of additional sources of information to calibrate the cost parameters substantially reduces the uncertainty surrounding the point estimates. These confidence intervals have β below 0.25, and the value of non-market activity z no lower than 0.88, and neither inconsistent with their values in Hagedorn and Manovskii (2008) and Hagedorn and Manovskii (2011). Such values of non-market activity are, however, higher than Mortensen and Nagypál (2007) and Hall and Milgrom (2008) regard as plausible (Hall and Milgrom (2008) suggest a calibrated value of 0.71), so this remains an important puzzle that cannot be addressed just by allowing for wage rigidity of the form used here.

The key message from these results is that there is a form of wage rigidity that enables the model to capture in a statistically satisfactory way the history dependence of wages in US data. The time series data we use are not themselves sufficiently informative to tie down the parameters of the model tightly. But there is plenty of scope for determining parameter values from other data sources that will capture that history dependence when wage rigidity is included in the model. That is in contrast to the model with wages for all matches bargained every period, for which the time series data are sufficiently informative to rule out *any* set of economically feasible parameter values that does this. These conclusions remain robust to several variations in the data and specification, see Appendix D.

5.3 Implications of the results

One implication of our results is that matching models with fully flexible wages of the type Hagedorn and Manovskii argue in a series of papers captures the behaviour of vacancies and unemployment do not capture the history dependence of wages. The alternative proposed by Gertler and Trigari (2009) with the same degree of wage rigidity in all matches has been criticised by Pissarides (2009) as inconsistent with the micro evidence that wages of job changers (new matches) are substantially more flexible than those of job stayers (continuing matches). A second implication of our results is that a wage equation consistent with this micro evidence, with wage rigidity only in continuing matches and flexible wages for all new matches, can capture the history dependence of

wages.

This second implication is important. As Malcomson (1999, Section 4) and Pissarides (2009) point out, wage rigidity that applies only to continuing matches, unlike wage rigidity that applies to all matches as in Gertler and Trigari (2009), has implications for unemployment and vacancies no different from fully flexible wages, provided it does not result in inefficient separations. Wage negotiation for new matches takes account of the wage rigidity and sets an initial wage such that the expected present value of wages over the duration of the match is unaffected by that wage rigidity. The wage rigidity thus results merely in an intertemporal redistribution of that expected present value and so has no effect on the incentives for vacancy creation. Inefficient separations do not occur in the model used here because all separations are assumed to be for exogenous reasons, though even with endogenous separations they can be avoided by renegotiation. There is thus no need to use wage equations that have implications for unemployment and vacancies different from fully flexible wages to capture the history dependence of wages.

Other messages that come across strongly from our results are the following. The credible bargaining model of Hall and Milgrom (2008) with all wage negotiated each period fares no better than Nash bargaining in capturing the history dependence of wages. With our formulation of wage rigidity that applies only to continuing matches, with the wage negotiated for all new matches, the lower bound of the 95% confidence interval for the proportion of wages not negotiated each quarter is comfortably above zero, and this conclusion is robust to weak identification. The time-series data we use are, however, not sufficiently informative to enable us to identify whether the rigidity should be modelled in nominal or in real wages.

Moreover, among the sets of parameter values that capture the history dependence, there is plenty of scope for selecting a set that is consistent with the other empirical evidence typically used for calibration of matching models. In particular, those sets include the calibrated values in Hagedorn and Manovskii (2008), Hall and Milgrom (2008), Pissarides (2009), Hagedorn and Manovskii (2011), Haefke et al. (2013) and many other papers in the literature. Thus our results are not inconsistent with the findings of other studies concerning those parameters, as long as

allowance is made for wage rigidity in continuing matches. Disagreements between those studies about appropriate values for the parameters need to be settled by other evidence.

6 Conclusion

In this paper, we have investigated econometric wage equations for a matching model of the US that are robust both to dropping the strong assumptions typically required for calibration and for full-information estimation methods and to weak instruments. We establish two main findings. First, we show that none of the formulations in the literature with wages in all matches negotiated each period robustly satisfies the natural criterion of adequately capturing the history dependence of wages. Second, we provide a formulation of wage rigidity that does so and, consistent with the micro evidence reported in Pissarides (2009) and Haefke et al. (2013), applies only to continuing matches, with wages negotiated for all new matches.

We reach this conclusion by nesting the Nash bargaining model and the credible bargaining model of Hall and Milgrom (2008) within a common over-arching framework of which each is a special case. The framework allows for heterogeneous match productivities and “on the job” search as modelled in Hagedorn and Manovskii (2013) and gives rise to a wage equation that can be estimated in a way that allows for the time to build in vacancy creation in Hagedorn and Manovskii (2011) of any length. It enables us to apply statistical tests that are robust to weak instruments to investigate which models are statistically acceptable restrictions of the over-arching framework and, in particular, capture the history dependence of wages. Our statistically acceptable specification includes a parameter that allows for wage rigidity in continuing, but not new, matches. Only with this parameter strictly positive can the model capture the history dependence of wages. But this form of wage rigidity has implications for unemployment and vacancies no different from fully flexible wages. There is no need to use a wage equation with wage rigidity for new matches, with its markedly different implications for unemployment and vacancies, to capture the history dependence of wages.

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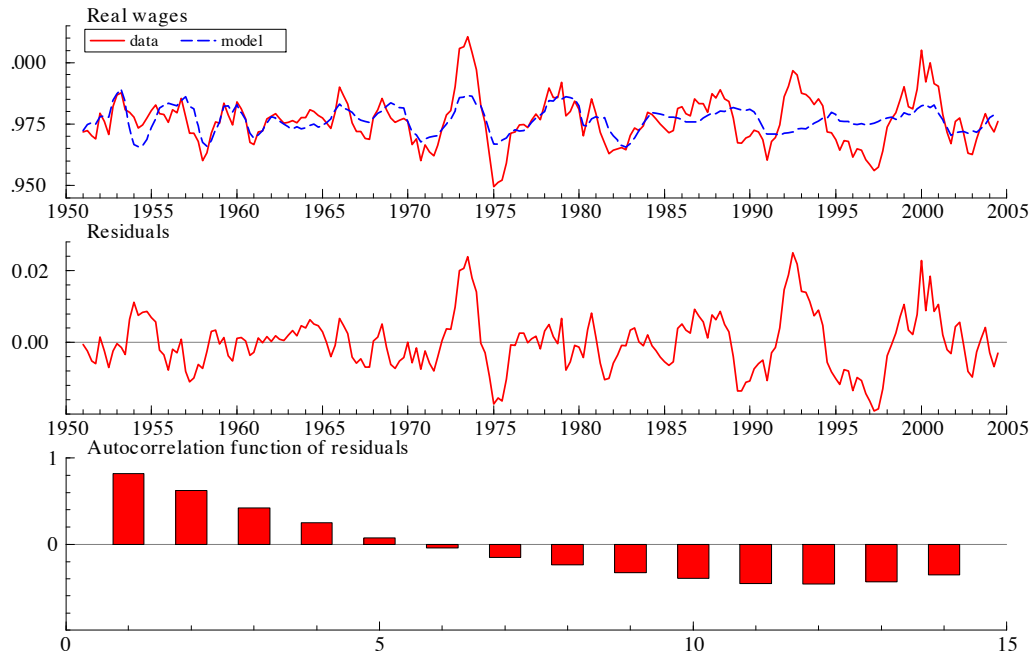


Figure 1: Top panel: real wages and corresponding fitted values for the calibration in Hagedorn and Manovskii (2008). Middle panel: residuals = differences between data and model. Bottom panel: autocorrelation function of residuals. Data: Hagedorn and Manovskii (2008).

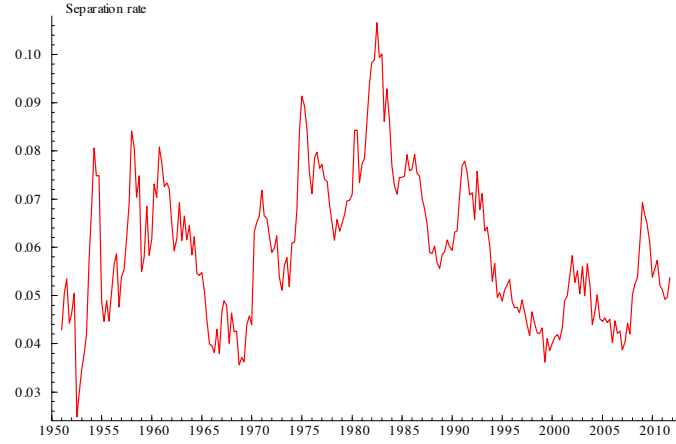


Figure 2: The separation rate s_t computed from $u_t^s = \left(1 - \frac{1}{2} \frac{j_t - (1-s_t)j_{t-1}}{l_t - (1-s_t)l_{t-1}}\right) (\Delta l_t + s_t l_{t-1})$ using employment, unemployment and short-term unemployment data from the BLS.

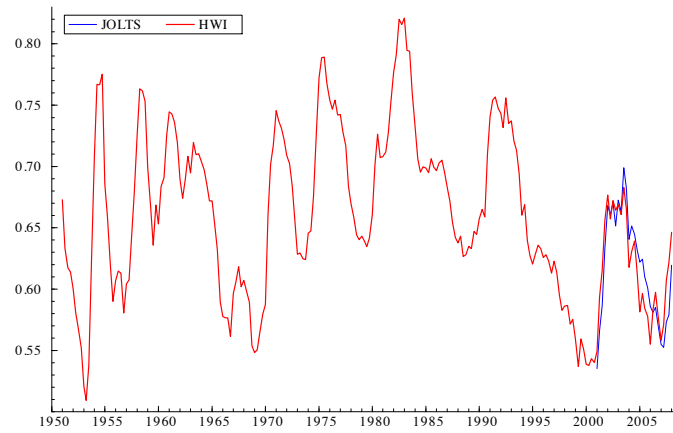


Figure 3: The vacancy-filling probability q_t using employment data from the BLS and vacancy data from the Conference Board HWI and from JOLTS.

Notation	Description
p_t	average productivity in matches at t
w_t	average wage in matches at t
δ_t	discount factor applied to $t + 1$ at t
s_t	match separation probability at t
c_t	vacancy posting cost at t
γ_t	difference in cost to firm and worker of making offers at t with credible bargaining
q_t	probability of filling a vacancy at t
f_t	probability of an unemployed worker finding employment at t
z_t	value of non-work activity at t
λ	Nash bargain: $\beta / (1 - \beta)$, for β the bargaining power of a worker
	credible bargain: probability of negotiation breakdown
κ	probability that employment starts at t for new match at t
δ_t^κ	$\delta_{t-1} (1 - s_t - f_{t-1} + \kappa s_t f_{t-1})$

Table 1: Notation for all wages bargained each period

Parameter	1	2	3	4
β	0.036 (0.891)	0.036 (0.843)	0.033 (2.568)	— (—)
z	0.937 (0.060)	0.937 (0.056)	0.887 (0.308)	0.880 (0.411)
c^K	0.440 (1803)	0.440 (953)	0.006 (751)	0.201 (579)
c^W	0.267 (1817)	0.270 (966)	0.399 (777)	0.254 (627)
ξ	0.382 (4292)	0.383 (2290)	0.006 (1891)	0.000 (2388)
κ	— (—)	0.000 (0.390)	0.069 (0.382)	0.046 (0.425)
H^K	— (—)	— (—)	2.279 (245)	1.753 (1100)
H^W	— (—)	— (—)	0.157 (243)	0.599 (1313)
γ	— (—)	— (—)	— (—)	0.716 (86.3)
λ	— (—)	— (—)	— (—)	0.031 (3.632)
GMM objective	34.459	34.457	33.395	33.315
Hansen test p value	0.001	0.001	0.000	0.000
Hansen test proj. p value	0.007	0.007	0.010	0.010

Table 2: Estimates of models without wage rigidity, eq. (1).

Notes: In all models, z is the value of non-work activity, c^K , c^W are capital and labour vacancy posting costs, ξ is the elasticity of the labour cost of those engaged in hiring with respect to productivity, κ is the fraction of matched jobs that become active within the quarter, and H^K , H^W are Pissarides (2009) fixed costs. For the Nash bargaining model (columns 1–3), β is the workers' bargaining weight. For the credible bargaining model (column 4), γ is the difference between firms' and workers' costs of making offers, and λ is the probability negotiations do not break down between offers. Estimation method is CUE-GMM with Newey-West weight matrix with prewhitening over the sample 1952q2-2004q4, with a constant and four lags of w_t , p_t , f_t and q_t as instruments. Standard errors in parentheses. The Hansen test p value is computed from a $\chi^2(N - \nu)$ distribution where N is the number of instruments and ν is the number of parameters estimated in the interior of the parameter space. The Hansen test proj. p value is derived from a $\chi^2(N)$ distribution.

Parameter	1	2
β	0.111 (0.030)	—
z	0.865 (0.018)	0.933 (0.019)
κ	0.000 (0.356)	1.000 (0.242)
γ	0.000 (0.000)	−0.926 (0.219)
λ	—	0.092 (0.023)
GMM objective	37.899	36.486
Hansen test p value	0.001	0.001
Hansen test proj. p value	0.003	0.004

Table 3: Estimates of models without wage rigidity independent of the value of creating a vacancy, eq. (2), with the infinite sum truncated at 13 quarters.

Notes: In both models, z is the value of non-work activity and κ is the fraction of matched jobs that become active within the quarter. For the Nash bargaining model (column 1), β is the workers' bargaining weight. For the credible bargaining model (column 2), γ is the difference between firms' and workers' costs of making offers, and λ is the probability negotiations do not break down between offers. Estimation method is CUE-GMM with Newey-West weight matrix with prewhitening over the sample 1952q2-2004q4, with a constant and four lags of w_t , p_t , f_t and q_t as instruments. Standard errors in parentheses. The Hansen test p value is computed from a $\chi^2(N - v)$ distribution where N is the number of instruments and v is the number of parameters estimated in the interior of the parameter space. The Hansen test proj. p value is derived from a $\chi^2(N)$ distribution.

Parameter	1	2	3	4	5
ψ	0.743 (0.027)	0.606 (0.069)	0.606 (0.073)	0.631 (0.084)	0.742 (0.026)
μ	0.123 (0.228)	0.023 (0.314)	0.023 (0.366)	0.000 (0.454)	0.130 (0.218)
β	0.090 (3.303)	0.077 (2.230)	0.077 (5.154)	— (—)	0.060 (0.045)
z	0.946 (0.198)	0.943 (0.137)	0.943 (0.340)	1.000 (0.125)	0.947 (0.013)
c^K	0.001 (1931.862)	0.007 (606.350)	0.007 (876.154)	0.000 (121.400)	0.194 (—)
c^W	0.149 (1935.586)	0.046 (607.306)	0.046 (879.118)	0.083 (130.336)	0.045 (—)
ξ	0.293 (9394.965)	0.085 (12360.664)	0.085 (17529.835)	0.000 (1169.534)	0.449 (—)
κ	— (—)	1.000 (0.334)	1.000 (0.359)	1.000 (0.395)	0.000 (—)
H^K	— (—)	— (—)	0.000 (101.850)	0.011 (1200.852)	0.000 (—)
H^W	— (—)	— (—)	0.000 (102.460)	1.145 (1317.720)	0.000 (—)
γ	— (—)	— (—)	— (—)	−1.376 (176.249)	0.000 (—)
λ	— (—)	— (—)	— (—)	0.131 (16.2)	— (—)
GMM objective	10.594	10.485	10.485	8.923	10.615
Hansen test p value	0.390	0.399	0.399	0.629	0.643
Hansen test proj. p value	0.877	0.882	0.882	0.943	0.876

Table 4: Estimates of models with wage rigidity, eq. (3), with w_t^* specified by (4) and the infinite sum truncated at 28 quarters.

Notes: In all models, ψ is the proportion of wages not negotiated in a quarter, μ is the proportion of wages not negotiated that are set in nominal terms, z is the value of non-work activity, c^K, c^W are capital and labour vacancy posting costs, ξ is the elasticity of the labour cost of those engaged in hiring with respect to productivity, κ is the fraction of matched jobs that become active within the quarter, and H^K, H^W are Pissarides (2009) fixed costs. For the Nash bargaining model (columns 1–3 and 5), β is the workers’ bargaining weight. For the credible bargaining model (column 4), γ is the difference between firms’ and workers’ costs of making offers, and λ is the probability negotiations do not break down between offers. In column 5, the cost parameters are calibrated using the approach of Hagedorn and Manovskii (2008). Estimation method is CUE-GMM with Newey-West weight matrix with prewhitening over the sample 1952q2–2004q4, with a constant and four lags of w_t, p_t, f_t and q_t as instruments. Standard errors in parentheses. The Hansen test p value is computed from a $\chi^2(N - \nu)$ distribution where N is the number of instruments and ν is the number of parameters estimated in the interior of the parameter space. The Hansen test proj. p value is derived from a $\chi^2(N)$ distribution.

Parameter		Nash Barg.	Cred. Barg.	Nash Barg. calibr.
ψ	95%	[0.15, 0.90]	[0.25, 0.90]	[0.62, 0.90]
	90%	[0.21, 0.88]	[0.25, 0.88]	[0.64, 0.87]
β	95%	[0.00, 0.94]	—	[0.00, 0.24]
	90%	[0.00, 0.84]	—	[0.00, 0.22]
z	95%	[0.25, 1.00]	[0.11, 1.00]	[0.88, 0.97]
	90%	[0.65, 1.00]	[0.20, 1.00]	[0.89, 0.97]
λ	95%	—	[0.00, 0.999]	
	90%	—	[0.00, 0.999]	

Table 5: Confidence intervals based on the S test of Stock and Wright (2000) for the specifications in columns 3, 4 and 5 of Table 4.

Notes: Confidence intervals reported only for parameters for which they do not comprise the entire admissible parameter range. In the first two columns all the remaining parameters are unrestricted. In the column “Nash Barg. calibr.”, $\mu = 0$ and the rest of the parameters are fixed at the calibrated values given in column 5 of Table 4.

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Appendix A Full theoretical model

A.1 Basic framework

The path of output in a match is determined by a random draw at the time the match is formed but may change over time (because, for example, of a general increase in productivity) at a rate common to all matches. The distribution of match productivity is such that it is always worthwhile to form a match when a vacant job and an unemployed worker meet. Match productivity in match k at time t is denoted p_t^k . In the basic model, separations occur only for exogenous reasons and at the same rate for all matches — there is no on the job search. Thus the distribution of productivity in actual matches is the same as the distribution of the productivity of potential matches. Denote by $J_{\tau,t}^k$ the expected present value of current and future profits at t to a firm from having a filled job in match k whose wage was most recently negotiated at $\tau \leq t$. This equals (i) output p_t^k net of wage costs $w_{\tau,t}^k$ for period t , plus (ii) the expected present value of profits $\tilde{J}_{\tau,t+1}^k$ from period $t+1$ on (when taking account of the possibility that the wage is renegotiated at $t+1$), discounted by the discount factor δ_t and the probability $(1 - s_{t+1})$ that the relationship is not ended before production at $t+1$ because the match is destroyed for exogenous reasons, plus (iii) the expected payoff V_{t+1} (if non-negative) of going back into the market for another employee if the match is destroyed. (A new match results in a new productivity draw and negotiation of a new wage, so V_{t+1} does not depend on k .) Thus

$$J_{\tau,t}^k = p_t^k - w_{\tau,t}^k + \delta_t E_t \left\{ (1 - s_{t+1}) \tilde{J}_{\tau,t+1}^k + s_{t+1} \max[0, V_{t+1}] \right\}, \text{ for all } k, t \geq \tau, \quad (\text{A.1})$$

where E_t is the expectation operator conditional on information available at t . Hagedorn and Manovskii (2008) and Hall and Milgrom (2008) assume s_{t+1} constant for all t . Here we allow for separation shocks in view of the importance Mortensen and Nagypál (2007) attribute to these.

In Hagedorn and Manovskii (2008), a new match at t results in employment starting at $t + 1$ and thus expected future profit $\delta_t E_t \tilde{J}_{t,t+1}$, where $\tilde{J}_{\tau,t}$ (with no superscript and $t > \tau$) is the average, over the distribution of productivities of matches that negotiated the wage most recently at $\tau < t$, of firms' payoffs from such matches from t on. For the empirical work, we are restricted to quarterly data, for which a one-period delay between matching and employment starting may seem implausibly long. Hagedorn and Manovskii (2011), however, argue for a time to build between the decision to create a vacancy and the possibility of a job becoming productive, for which they find 3 months appropriate. With the quarterly data used here, that corresponds to a one-period time to build. To avoid being overly prescriptive, we allow for a one-period delay with probability $1 - \kappa$ and zero delay with probability κ . The expected present value V_t of creating a vacancy at t for which productivity is not yet determined and the wage is to be negotiated is then

$$V_t = -c_t + q_t \left[\kappa J_{t,t} + (1 - \kappa) \delta_t E_t \tilde{J}_{t,t+1} \right] + (1 - q_t) \delta_t E_t \max [0, V_{t+1}], \text{ for all } t, \quad (\text{A.2})$$

where c_t is the vacancy posting cost that must be incurred at the start of period t to create a vacancy, q_t is the probability the vacancy is matched with a worker, $J_{t,t}$ (with no superscript) is the expectation at t of a firm's payoff from t on from a new match at t before match productivity is known and the wage bargained, and $\tilde{J}_{\tau,t}$ (with no superscript and $t > \tau$) is the average, over the distribution of productivities of matches that negotiated wages most recently at $\tau < t$, of firms' payoffs from t on from such matches. (If wages are not negotiated every period, (A.2) implies that a match at t with employment starting at $t + 1$ has its match productivity drawn, and its wage negotiated, at t . Match productivity then changes at the same rate between t and $t + 1$, and the wage has the same possibility for renegotiation at $t + 1$, as in continuing matches. If wages are negotiated every period, $\tilde{J}_{t,t+1} = J_{t+1,t+1}$ so no such assumption is implied.) Free entry of vacancies when (as in the data) new vacancies are created at each date implies

$$V_t = 0, \text{ for all } t. \quad (\text{A.3})$$

For a worker in match k at t whose wage was most recently negotiated at $\tau \leq t$, the expected present value of employment $W_{\tau,t}^k$ is given by

$$W_{\tau,t}^k = w_{\tau,t}^k + \delta_t E_t \left[(1 - s_{t+1}) \tilde{W}_{\tau,t+1}^k + s_{t+1} U_{t+1} \right], \text{ for all } k, t \geq \tau, \quad (\text{A.4})$$

where $\tilde{W}_{\tau,t+1}^k$ is the expected present value of employment from period $t + 1$ on (when taking account of the possibility that the wage is renegotiated at $t + 1$) and U_{t+1} is the expected present value of starting period $t + 1$ unemployed, an event that happens with the probability s_{t+1} that the job comes to an end for exogenous reasons. The probability that a worker unemployed at t finds a job in the matching process at t is denoted by f_t , the value of non-work activity (including any unemployment benefit) by z_t . The expected present value U_t of seeking a match at t is then

$$U_t = f_t \left[\kappa W_{t,t} + (1 - \kappa) \left(z_t + \delta_t E_t \tilde{W}_{t,t+1} \right) \right] + (1 - f_t) (z_t + \delta_t E_t U_{t+1}), \text{ for all } t, \quad (\text{A.5})$$

where $W_{t,t}$ (with no superscript) is the expectation at t of a worker's payoff from t on from a new match at t before match productivity is known and the wage bargained, and $\tilde{W}_{\tau,t}$ (with no superscript and $t > \tau$) is the average, over the distribution of productivities of matches that negotiated the wage most recently at $\tau < t$, of workers' payoffs from t on from such matches. The right-hand side of (A.5) can be interpreted as follows. With probability f_t , the worker is hired at t and receives expected future utility $\left[\kappa W_{t,t} + (1 - \kappa) \left(z_t + \delta_t E_t \tilde{W}_{t,t+1} \right) \right]$ from being matched. With probability $1 - f_t$ the worker is not hired at t and receives utility z_t for period t plus the expected utility from starting period $t + 1$ unmatched.

Pissarides (2009) has argued for an extension of matching models to allow for fixed costs incurred after matching has occurred, such as training, negotiation, or one-off administrative costs. For our purposes, this specification corresponds to replacing the vacancy posting cost term c_t in (A.2) by the composite cost $c_t + q_t H_t$, where H_t is the fixed cost incurred by the firm after meeting a new worker but before the wage is agreed (compare with Pissarides (2009, p. 1364)). To save on notation, we retain the representation in (A.2) for developing the theory and make the appropriate

adjustment for the empirical analysis.

This model nests the case of homogeneous productivity and no wage rigidity by setting $p_t^k = p_t$, $w_{\tau,t}^k = w_t$, $J_{\tau,t}^k = \tilde{J}_{\tau,t}^k = \tilde{J}_{\tau,t} = J_{t,t}$ and $W_{\tau,t}^k = \tilde{W}_{\tau,t}^k = \tilde{W}_{\tau,t} = W_{t,t}$ for all k, τ and $t \geq \tau$.

A.2 Wage determination in the basic model

A.2.1 Nash bargaining

Most of the literature on matching models follows Mortensen and Pissarides (1994) in modelling wage determination by the Nash bargaining solution. The generalized form of that solution for the wage negotiated for match k at t is

$$W_{t,t}^k - U_t = \frac{\beta}{1 - \beta} (J_{t,t}^k - V_t), \text{ for } \beta \in [0, 1), \quad (\text{A.6})$$

where $\beta / (1 - \beta)$ is the bargaining power of workers relative to that of firms. With the free entry condition $V_t = 0$, (A.6) reduces to the formulation in Hagedorn and Manovskii (2008)

$$W_t^k - U_t = \frac{\beta}{1 - \beta} J_t^k, \text{ for } \beta \in [0, 1). \quad (\text{A.7})$$

A.2.2 Credible bargaining

Hall and Milgrom (2008) develop an alternative to the standard Nash bargain in which there is positive probability, here denoted α , that negotiations break down irrevocably each time a new offer is made. They also include a cost to making an offer that we here denote by γ_t^f for each offer the firm makes in period t and γ_t^w for each offer the worker makes in period t . (The specification in Hall and Milgrom (2008) corresponds to $\gamma_t^w = 0$ but, for reasons that will become apparent, it is useful to allow $\gamma_t^w > 0$.) If negotiations break down, the parties search for alternative matches.

In Hall and Milgrom (2008), the parties alternate in making offers, starting with the firm, with at most one offer made each period. Hall and Milgrom (2008) envisage each period as corresponding to a day. With the data available, we are constrained to having each period correspond to a quarter,

so the assumption of at most one offer per period seems implausible. For this reason, we generalize the model to allow offers at fixed intervals that may be less than a whole period. Consider an offer from the firm in match k at time η ($0 \leq \eta < 1$) between t and $t + 1$ that would yield the worker present value payoff $W_{t+\eta,t+\eta}^k$. The worker will accept that offer if $W_{t+\eta,t+\eta}^k$ is at least as great as the payoff from rejecting the offer, having negotiations break down with probability α and receiving payoff $U_{t+\eta}$ of seeking an alternative match, but otherwise making a counter-offer resulting in expected present value payoff denoted $\hat{W}_{t+\eta}^k$. Recognizing this, the firm will make the lowest offer satisfying that requirement, which gives the indifference condition

$$W_{t+\eta,t+\eta}^k = \alpha U_{t+\eta} + (1 - \alpha) \hat{W}_{t+\eta}^k, \quad \eta \in [0, 1). \quad (\text{A.8})$$

Symmetrically, the firm will accept an offer with present value payoff $J_{t+\eta}^k$ made by the worker at time η ($0 \leq \eta < 1$) between t and $t + 1$ if $J_{t+\eta}^k$ is at least as great as the payoff from rejecting the offer, having negotiations break down with probability α and receiving payoff $V_{t+\eta}$ of seeking an alternative match, but otherwise making a counter-offer resulting in expected present value payoff denoted $\hat{J}_{t+\eta}^k$. Recognizing this, the worker will make the lowest offer satisfying that requirement, which gives the indifference condition

$$J_{t+\eta}^k = \alpha V_{t+\eta} + (1 - \alpha) \hat{J}_{t+\eta}^k, \quad \eta \in [0, 1). \quad (\text{A.9})$$

In Hall and Milgrom (2008), the firm makes the first offer and in equilibrium that offer is always accepted, so the bargained outcome corresponds to $W_{t+\eta,t+\eta}^k$ for $\eta = 0$.

In the specification in Hall and Milgrom (2008) with only one offer per period,

$$\hat{J}_t^k = \delta_t E_t \left(-\gamma_{t+1}^f + J_{t+1,t+1}^k \right); \quad \hat{W}_t^k = z_t + \delta_t E_t \left(-\gamma_{t+1}^w + W_{t+1}^k \right), \quad (\text{A.10})$$

where W_{t+1}^k is the payoff to the worker from making an offer at $t + 1$. The alternative we consider here, which seems more appropriate with periods of a quarter, is to let the time interval between

offers go to zero. Then

$$\hat{J}_t^k = -\gamma_t^f + J_{t,t}^k; \quad \hat{W}_t^k = -\gamma_t^w + W_{t,t}^k. \quad (\text{A.11})$$

In that case, the indifference conditions (A.8) and (A.9) with $\eta = 0$ can be solved to give the following sharing rule as an alternative to (A.6):

$$W_{t,t}^k - U_t = (1 - \alpha) (J_{t,t}^k - V_t) + \frac{1 - \alpha}{\alpha} \gamma_t, \quad (\text{A.12})$$

where $\gamma_t = (1 - \alpha) \gamma_t^f - \gamma_t^w$. Note that γ_t^f and γ_t^w cannot be separately identified from (A.12). But permitting $\gamma_t^w > 0$ allows the model to be consistent with an estimated $\gamma_t < 0$.

A.2.3 Nesting Nash and credible bargaining

The Nash and credible bargaining outcomes (A.6) and (A.12) are special cases of the more general formulation

$$W_{t,t}^k - U_t = \lambda (J_{t,t}^k - V_t) + \frac{\lambda}{1 - \lambda} \gamma_t, \quad (\text{A.13})$$

with the models satisfying the restrictions

$$\text{Nash bargaining (A.6) : } \lambda = \frac{\beta}{1 - \beta} \in [0, \infty); \gamma_t = 0;$$

$$\text{Credible bargaining (A.12) : } \lambda = 1 - \alpha \in [0, 1). \quad (\text{A.14})$$

Averaged over all match productivities, (A.13) becomes

$$W_{t,t} - U_t = \lambda (J_{t,t} - V_t) + \frac{\lambda}{1 - \lambda} \gamma_t, \quad (\text{A.15})$$

where we have made use of the linearity property of expectations.¹²

¹²This follows by Tonelli's theorem, see Billingsley (1995, Theorem 18.3).

A.2.4 Wages bargained every period

With all wages bargained every period, the average wage w_t is just the average of the wages for each individual match k given by (A.13). Because, from (A.1) and (A.4), $J_{t,t}^k$ and $W_{t,t}^k$ are linear in $w_{t,t}^k$, the average wage is given by (A.15). In this case, (A.15) can be combined with (A.2), (A.3) and (A.5) to yield wage equation (1) for the average wage. For Nash bargaining, $\lambda = \beta / (1 - \beta)$ and $\gamma_t = 0$ so (1) can be written

$$w_t = \beta p_t + (1 - \beta) z_t + \beta c_t \frac{f_t}{q_t (1 - f_t \kappa)}. \quad (\text{A.16})$$

When $\kappa = 0$, (A.16) is exactly the wage equation in Hagedorn and Manovskii (2008).

A.2.5 Wage equation independent of vacancy creation equation

A wage equation that is independent of the vacancy creation equation can be derived as follows. From the free entry condition (A.3), $V_t = 0$ for all t . With the wage for every job bargained in every period, $\tilde{J}_{\tau,t} = J_{t,t}$ for $\tau \leq t$. With those specifications, (A.1) averaged over all match productivities becomes

$$J_{t,t} = p_t - w_t + \delta_t E_t (1 - s_{t+1}) J_{t+1,t+1}, \text{ for all } t. \quad (\text{A.17})$$

Also, with the wage for every job bargained in every period, $\tilde{W}_{\tau,t} = W_{t,t}$ for $\tau \leq t$. Then (A.4) averaged over all match productivities and (A.5) can be jointly solved forward to write

$$W_{t,t} - U_t = E_t \sum_{n=0}^{\infty} \delta_{t,n}^{\kappa} (1 - f_{t+n} \kappa) (w_{t+n} - z_{t+n}), \quad (\text{A.18})$$

where $\delta_{t,n}^{\kappa} = \prod_{i=1}^n \delta_{t+i}^{\kappa}$, with $\delta_{t,0}^{\kappa} = 1$, and $\delta_t^{\kappa} = \delta_{t-1} (1 - s_t - f_{t-1} + \kappa s_t f_{t-1})$. These two conditions can be used with (A.15) to yield the wage equation (2).

A.3 Extensions to the basic model

Two important generalizations of the basic model in the literature are to on the job search and to wages that are not negotiated every period.

A.3.1 On the job search

In the Hagedorn and Manovskii (2013) model of on the job search, wage determination in match k takes, up to a log-linear approximation, the form (see their equation (1))

$$w_{t,t}^k = (p_t^k)^\zeta (\vartheta_t)^\rho, \quad \text{for all } k, t, \quad (\text{A.19})$$

where ϑ_t is a business cycle indicator that incorporates labour market tightness. For empirical purposes, Hagedorn and Manovskii (2013, eq. (35)) normalize $\zeta = 1$ because it is not identified separately from the standard deviation of the distribution of productivities. With $\zeta = 1$ and averaging over k , (A.19) corresponds to (A.16) with

$$(\vartheta_t)^\rho = (1 - \beta) \frac{z_t}{p_t} + \beta \left[1 + \frac{c_t}{p_t q_t (1 - f_t \kappa)} \frac{f_t}{p_t} \right].$$

A.3.2 Wage rigidity

Following Pissarides (2009), we model wage rigidity as applying only to continuing matches, with wages for new matches all negotiated. The form of wage rigidity is that developed by Gertler and Trigari (2009) but applied only to continuing, not new, matches. Persistence takes the form of a fixed probability $1 - \psi$ that a firm renegotiates its wage in any period. In the absence of such renegotiation, the wage remains the same as in the previous period. Thus the wage at t for a match with wage most recently negotiated at $\tau \leq t - 1$ is

$$\begin{aligned} & w_{t,t}^k, & \text{with probability } 1 - \psi, \\ & \pi_t^{-\mu} w_{\tau,t-1}^k, & \text{with probability } \psi, \end{aligned} \quad , \quad \psi, \mu \in [0, 1], \quad (\text{A.20})$$

where π_t is the ratio of prices at t to prices at $t-1$, which we incorporate to allow for the possibility that the previous period's wage may be adjusted automatically in response to inflation, with μ a parameter to be estimated. With wage rigidity of the form in (A.20),

$$\tilde{J}_{\tau,t}^k = (1 - \psi) J_{t,t}^k + \psi J_{\tau,t}^k = J_{t,t}^k - \psi (J_{t,t}^k - J_{\tau,t}^k) \quad (\text{A.21})$$

$$\tilde{W}_{\tau,t}^k = (1 - \psi) W_{t,t}^k + \psi W_{\tau,t}^k = W_{t,t}^k - \psi (W_{t,t}^k - W_{\tau,t}^k). \quad (\text{A.22})$$

(Recall that $\tilde{J}_{\tau,t}^k$ and $\tilde{W}_{\tau,t}^k$ refer to matches with wage negotiated at τ but not renegotiated before t .)

With wages negotiated in all new matches, (A.1), (A.2), (A.4) and (A.5) continue to apply. Manipulation of these conditions gives

$$\begin{aligned} J_{t+1,t+1}^k - J_{t,t+1}^k \\ = - (W_{t+1,t+1}^k - W_{t,t+1}^k) \end{aligned} \quad (\text{A.23})$$

$$= \left(\pi_{t+1}^{-\mu} w_{t,t}^k - w_{t+1,t+1}^k \right) E_{t+1} \left[\sum_{i=1}^{\infty} \prod_{j=2}^i \left(\delta_{t+j-1} (1 - s_{t+j}) \psi \pi_{t+j}^{-\mu} \right) \right], \quad (\text{A.24})$$

with the convention that the product term equals 1 for $i < j$. For a newly-bargained wage, (A.13) applies. Moreover, with the formulation in (A.20), the distribution of match productivities for new bargains is the same as that for all vacancies so the average newly-bargained wage is still given by (A.15).

A.4 Empirical wage equation

Let w_t^* denote the average newly-negotiated wage at t (that is, the average of $w_{t,t}^k$ for matches k negotiating wages at t). Then, from (A.1) for $\tau = t$ averaged over all match productivities and with $V_{t+1} \geq 0$ for all t (as it must be when the free entry condition (A.3) holds for all t),

$$J_{t,t} = p_t - w_t^* + \delta_t E_t \left[(1 - s_{t+1}) \tilde{J}_{t,t+1} + s_{t+1} V_{t+1} \right], \text{ for all } t. \quad (\text{A.25})$$

This can be combined with (A.2) to give

$$\begin{aligned} J_{t,t} - V_t &= (1 - q_t \kappa) (p_t - w_t^*) + c_t \\ &\quad + \delta_t E_t \left[(1 - s_{t+1} - q_t + q_t \kappa s_{t+1}) (\tilde{J}_{t,t+1} - V_{t+1}) \right]. \end{aligned} \quad (\text{A.26})$$

Use of (A.21) in (A.26) gives

$$\begin{aligned} J_{t,t} - V_t &= (1 - q_t \kappa) (p_t - w_t^*) + c_t \\ &\quad + \delta_t E_t \left[(1 - s_{t+1} - q_t + q_t \kappa s_{t+1}) (J_{t+1,t+1} - V_{t+1} - \psi (J_{t+1,t+1} - J_{t,t+1})) \right]. \end{aligned} \quad (\text{A.27})$$

Similarly, from (A.4) for $\tau = t$ averaged over all match productivities,

$$W_{t,t} = w_t^* + \delta_t E_t \left[(1 - s_{t+1}) \tilde{W}_{t,t+1} + s_{t+1} U_{t+1} \right], \text{ for all } t. \quad (\text{A.28})$$

This can be combined with (A.5) to give

$$\begin{aligned} W_{t,t} - U_t &= (1 - f_t \kappa) (w_t^* - z_t) \\ &\quad + \delta_t E_t \left[(1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) (\tilde{W}_{t,t+1} - U_{t+1}) \right]. \end{aligned} \quad (\text{A.29})$$

Use of (A.22) in (A.29) gives

$$\begin{aligned} W_{t,t} - U_t &= (1 - f_t \kappa) (w_t^* - z_t) + \delta_t E_t \left\{ (1 - s_{t+1} - f_t \right. \\ &\quad \left. + \kappa f_t s_{t+1}) [W_{t+1,t+1} - U_{t+1} - \psi (W_{t+1,t+1} - W_{t,t+1})] \right\}. \end{aligned}$$

Use of (A.15) forwarded one period and (A.23) in this allows it to be written

$$\begin{aligned}
W_{t,t} - U_t &= (1 - f_t \kappa) (w_t^* - z_t) + \delta_t E_t \left\{ (1 - s_{t+1} - f_t \right. \\
&\quad \left. + \kappa f_t s_{t+1}) \left[\lambda (J_{t+1,t+1} - V_{t+1}) + \frac{\lambda}{1 - \lambda} \gamma_{t+1} + \psi (J_{t+1,t+1} - J_{t,t+1}) \right] \right\}. \tag{A.30}
\end{aligned}$$

Use of (A.27), (A.30) and the free entry condition (A.3) in (A.15) can, with some manipulation (see Appendix B), be used to derive the wage equation (3) for newly formed matches in the text.

Empirical implementation of (3) requires a data series for the average newly-negotiated wage w_t^* . Under the assumptions of the model, continuing matches for which wages are renegotiated have a random sample of match productivities. Thus, a series for the average newly-negotiated wage consistent with the model can be derived using an approximation suggested by Gertler and Trigari (2009) based on the law of large numbers. Let j_t denote employment at t , which equals the number of filled jobs. The total number of matches negotiating their wage at t comprises new matches created at t $m_t = j_t - (1 - s_t) j_{t-1}$, plus continuing matches that renegotiate the wage at t , $(1 - \psi) (1 - s_t) j_{t-1}$. Hence the total wage bill at t , with w_t the average wage in all matches at t , is

$$\begin{aligned}
w_t j_t &= \{ [j_t - (1 - s_t) j_{t-1}] + (1 - \psi) (1 - s_t) j_{t-1} \} w_t^* + \psi (1 - s_t) j_{t-1} \pi_t^{-\mu} w_{t-1} \\
&= [j_t - \psi (1 - s_t) j_{t-1}] w_t^* + \psi (1 - s_t) j_{t-1} \pi_t^{-\mu} w_{t-1}.
\end{aligned}$$

So the average newly-negotiated wage is given by (4).

Appendix B Derivations of equations

Table 6 gives definitions of variables in the theoretical models, Table 7 definitions of parameters.

Variable	Description
p_t^k	productivity in match k at t
p_t	average productivity in matches at t
\bar{p}_t	trend productivity at t
\tilde{p}_t	p_t / \bar{p}_t
$w_{\tau,t}^k$	wage in match k at t negotiated at τ
w_t	average wage in matches at t
w_t^*	average of wages newly-negotiated at t
δ_t	discount factor applied to $t + 1$ at t
s_t	match separation probability at t
c_t	vacancy posting cost at t
γ_t^f	cost to firm of making an offer at t with credible bargaining
γ_t^w	cost to worker of making an offer at t with credible bargaining
γ_t	$(1 - \alpha) \gamma_t^f - \gamma_t^w$
H_t	cost of starting employment at t
q_t	probability of filling a vacancy at t
f_t	probability of finding employment at t
z_t	value of non-work activity at t
π_t	ratio of prices at t to prices at $t - 1$
$J_{\tau,t}^k$	payoff to firm at t from match k with wage negotiated at τ
$J_{\tau,t}$	average payoff to firms at t from matches with wage negotiated at τ
$\tilde{J}_{\tau,t}^k$	expected $J_{\tau,t}^k$ before possible renegotiation at t
$\tilde{J}_{\tau,t}$	expected $J_{\tau,t}$ before possible renegotiation at t
$W_{\tau,t}^k$	payoff to worker at t from match k with wage negotiated at τ
$W_{\tau,t}$	average payoff to workers at t from matches with wage negotiated at τ
$\tilde{W}_{\tau,t}^k$	expected $W_{\tau,t}^k$ before possible renegotiation at t
$\tilde{W}_{\tau,t}$	expected $W_{\tau,t}$ before possible renegotiation at t
V_t	value of unfilled vacancy at t
U_t	payoff to worker from seeking match at t
δ_t^κ	$\delta_{t-1} (1 - s_t - f_{t-1} + \kappa s_t f_{t-1})$

Table 6: Variables in theoretical models

Parameter	Definition
κ	probability that employment starts at t for new match at t
β	worker bargaining power in Nash bargain
α	probability of negotiation breakdown with credible bargaining
λ	$\beta / (1 - \beta)$ with Nash bargaining; $1 - \alpha$ with credible bargaining
ψ	degree of wage rigidity (equals 0 if none)
μ	proportion of unrenegotiated wages set in nominal terms

Table 7: Parameters in theoretical models

B.1 Derivation of equation (A.12)

With the specification in (A.11), (A.8) and (A.9) with $\eta = 0$ become

$$W_{t,t}^k = \alpha U_t + (1 - \alpha) (-\gamma_t^w + W_t'^k) \quad (\text{A.31})$$

$$J_t'^k = \alpha V_t + (1 - \alpha) (-\gamma_t^f + J_{t,t}^k). \quad (\text{A.32})$$

The terms in $W_t'^k$ and $J_t'^k$ in (A.31) and (A.32) can be eliminated in the following way. Multiply (A.32) by $1 - \alpha$ and subtract it from (A.31) to get

$$W_{t,t}^k - (1 - \alpha) J_t'^k = \alpha U_t + (1 - \alpha) (-\gamma_t^w + W_t'^k) - \alpha (1 - \alpha) V_t - (1 - \alpha)^2 (-\gamma_t^f + J_{t,t}^k).$$

Next note that $W_t'^k + J_t'^k = W_{t,t}^k + J_{t,t}^k$ necessarily, so this can be written

$$W_{t,t}^k = \alpha U_t - (1 - \alpha) \gamma_t^w + (1 - \alpha) (W_{t,t}^k + J_{t,t}^k) - \alpha (1 - \alpha) V_t - (1 - \alpha)^2 (-\gamma_t^f + J_{t,t}^k)$$

or, subtracting $W_{t,t}^k$ from both sides

$$0 = \alpha U_t - (1 - \alpha) \gamma_t^w - \alpha W_{t,t}^k + (1 - \alpha) (1 - (1 - \alpha)) J_{t,t}^k - \alpha (1 - \alpha) V_t + (1 - \alpha)^2 \gamma_t^f.$$

This can be rewritten as (A.12).

B.2 Derivation of equation (1)

Equation (1) is a special case of (3) with $\psi = 0$ and $w_t^* = w_t$, as is appropriate when all wages are bargained every period.

B.3 Derivation of equation (A.18)

Equation (A.18) follows from solving forward equation (A.29) with, as appropriate when the wage for all jobs is bargained every period, $w_t^* = w_t$ and $E_t \tilde{W}_{t,t+1} = E_t W_{t+1,t+1}$.

B.4 Derivation of equation (2)

Multiply (A.17) through by λ and use (A.15) to substitute for $J_{t,t}$ and $J_{t+1,t+1}$ to get

$$W_{t,t} - U_t - \frac{\lambda}{1-\lambda} \gamma_t = \lambda (p_t - w_t) + \delta_t E_t \left[(1 - s_{t+1}) \left(W_{t+1,t+1} - U_{t+1} - \frac{\lambda}{1-\lambda} \gamma_{t+1} \right) \right]. \quad (\text{A.33})$$

Use of (A.18) to substitute for $W_{t,t} - U_t$ and $W_{t+1,t+1} - U_{t+1}$ in (A.33) yields (2).

B.5 Derivation of equations (A.23) and (A.24)

Let $\tilde{V}_t^k = \max [0, V_t^k]$ for notational simplicity. From (A.1) for $\tau = t$ and (A.21),

$$\begin{aligned} J_{t,t+1}^k &= p_{t+1}^k - \pi_{t+1}^{-\mu} w_{t,t}^k + \delta_{t+1} E_{t+1} \left[(1 - s_{t+2}) \tilde{J}_{t,t+2}^k + s_{t+2} \tilde{V}_{t+2}^k \right] \\ &= p_{t+1}^k - \pi_{t+1}^{-\mu} w_{t,t}^k \\ &\quad + \delta_{t+1} E_{t+1} \left\{ (1 - s_{t+2}) \left[(1 - \psi) J_{t+2,t+2}^k + \psi J_{t,t+2}^k \right] + s_{t+2} \tilde{V}_{t+2}^k \right\}. \end{aligned} \quad (\text{A.34})$$

Moreover, from (A.1) with the wage negotiated at $t + 1$ and (A.21),

$$\begin{aligned} J_{t+1,t+1}^k &= p_{t+1}^k - w_{t+1,t+1}^k + \delta_{t+1} E_{t+1} \left[(1 - s_{t+2}) \tilde{J}_{t+1,t+2}^k + s_{t+2} \tilde{V}_{t+2}^k \right] \\ &= p_{t+1}^k - w_{t+1,t+1}^k \\ &\quad + \delta_{t+1} E_{t+1} \left\{ (1 - s_{t+2}) \left[(1 - \psi) J_{t+2,t+2}^k + \psi J_{t+1,t+2}^k \right] + s_{t+2} \tilde{V}_{t+2}^k \right\}. \end{aligned}$$

Hence,

$$J_{t+1,t+1}^k - J_{t,t+1}^k = \pi_{t+1}^{-\mu} w_{t,t}^k - w_{t+1,t+1}^k + \delta_{t+1} E_{t+1} \left[(1 - s_{t+2}) \psi (J_{t+1,t+2}^k - J_{t,t+2}^k) \right]. \quad (\text{A.35})$$

Furthermore, again from (A.1) and (A.21),

$$\begin{aligned}
J_{t,t+2}^k &= p_{t+2}^k - \pi_{t+2}^{-\mu} \pi_{t+1}^{-\mu} w_{t,t}^k + \delta_{t+2} E_{t+2} \left[(1 - s_{t+3}) \tilde{J}_{t,t+3}^k + s_{t+3} \tilde{V}_{t+3}^k \right] \\
&= p_{t+2}^k - \pi_{t+2}^{-\mu} \pi_{t+1}^{-\mu} w_{t,t}^k \\
&\quad + \delta_{t+2} E_{t+2} \left\{ (1 - s_{t+3}) \left[(1 - \psi) J_{t+3,t+3}^k + \psi J_{t,t+3}^k \right] + s_{t+3} \tilde{V}_{t+3}^k \right\}
\end{aligned}$$

so, with the use of (A.34) forwarded one period,

$$\begin{aligned}
&J_{t+1,t+2}^k - J_{t,t+2}^k \\
&= p_{t+2}^k - \pi_{t+2}^{-\mu} w_{t+1,t+1}^k + \delta_{t+2} E_{t+2} \left\{ (1 - s_{t+3}) \left[(1 - \psi) J_{t+3,t+3}^k + \psi J_{t+1,t+3}^k \right] + s_{t+3} \tilde{V}_{t+3}^k \right\} \\
&\quad - p_{t+2}^k + \pi_{t+2}^{-\mu} \pi_{t+1}^{-\mu} w_{t,t}^k - \delta_{t+2} E_{t+2} \left\{ (1 - s_{t+3}) \left[(1 - \psi) J_{t+3,t+3}^k + \psi J_{t,t+3}^k \right] + s_{t+3} \tilde{V}_{t+3}^k \right\} \\
&= \pi_{t+2}^{-\mu} \left(\pi_{t+1}^{-\mu} w_{t,t}^k - w_{t+1,t+1}^k \right) + \delta_{t+2} E_{t+2} \left[(1 - s_{t+3}) \psi (J_{t+1,t+3}^k - J_{t,t+3}^k) \right].
\end{aligned}$$

Use of this in (A.35) gives

$$\begin{aligned}
J_{t+1,t+1}^k - J_{t,t+1}^k &= \left(\pi_{t+1}^{-\mu} w_{t,t}^k - w_{t+1,t+1}^k \right) \left[1 + \delta_{t+1} E_{t+1} (1 - s_{t+2}) \psi \pi_{t+2}^{-\mu} \right] \\
&\quad + \delta_{t+1} E_{t+1} \delta_{t+2} \left[(1 - s_{t+2}) \psi (1 - s_{t+3}) \psi (J_{t+1,t+3}^k - J_{t,t+3}^k) \right].
\end{aligned}$$

Proceeding recursively in this way gives

$$J_{t+1,t+1}^k - J_{t,t+1}^k = \left(\pi_{t+1}^{-\mu} w_{t,t}^k - w_{t+1,t+1}^k \right) E_{t+1} \left[1 + \sum_{i=2}^{\infty} \prod_{j=2}^i \left(\delta_{t+j-1} (1 - s_{t+j}) \psi \pi_{t+j}^{-\mu} \right) \right].$$

With the convention that the product term equals 1 for $i < j$, this can be written as (A.24).

From (A.4) for $\tau = t$ and (A.22),

$$\begin{aligned}
W_{t,t+1}^k &= \pi_{t+1}^{-\mu} w_{t,t}^k + \delta_{t+1} E_{t+1} \left[(1 - s_{t+2}) \tilde{W}_{t,t+2}^k + s_{t+2} U_{t+2} \right] \\
&= \pi_{t+1}^{-\mu} w_{t,t}^k \\
&\quad + \delta_{t+1} E_{t+1} \left\{ (1 - s_{t+2}) \left[(1 - \psi) W_{t+2,t+2}^k + \psi W_{t,t+2}^k \right] + s_{t+2} U_{t+2} \right\}. \tag{A.36}
\end{aligned}$$

Moreover, from (A.4) with the wage negotiated at $t + 1$ and (A.22),

$$\begin{aligned}
W_{t+1,t+1}^k &= w_{t+1,t+1}^k + \delta_{t+1} E_{t+1} \left[(1 - s_{t+2}) \tilde{W}_{t+1,t+2}^k + s_{t+2} U_{t+2} \right] \\
&= w_{t+1,t+1}^k + \delta_{t+1} E_{t+1} \left\{ (1 - s_{t+2}) \left[(1 - \psi) W_{t+2,t+2}^k + \psi W_{t+1,t+2}^k \right] + s_{t+2} U_{t+2} \right\}.
\end{aligned}$$

Hence

$$W_{t+1,t+1}^k - W_{t,t+1}^k = w_{t+1,t+1}^k - \pi_{t+1}^{-\mu} w_{t,t}^k + \delta_{t+1} E_{t+1} \left[(1 - s_{t+2}) \psi (W_{t+1,t+2}^k - W_{t,t+2}^k) \right]. \tag{A.37}$$

Furthermore, again from (A.4) and (A.22),

$$\begin{aligned}
W_{t,t+2}^k &= \pi_{t+2}^{-\mu} \pi_{t+1}^{-\mu} w_{t,t}^k + \delta_{t+2} E_{t+2} \left[(1 - s_{t+3}) \tilde{W}_{t,t+3}^k + s_{t+3} U_{t+3} \right] \\
&= \pi_{t+2}^{-\mu} \pi_{t+1}^{-\mu} w_{t,t}^k + \delta_{t+2} E_{t+2} \left\{ (1 - s_{t+3}) \left[(1 - \psi) W_{t+3,t+3}^k + \psi W_{t,t+3}^k \right] + s_{t+3} U_{t+3} \right\}
\end{aligned}$$

so, with the use of (A.36) forwarded one period,

$$\begin{aligned}
&W_{t+1,t+2}^k - W_{t,t+2}^k \\
&= \pi_{t+2}^{-\mu} w_{t+1,t+1}^k + \delta_{t+2} E_{t+2} \left\{ (1 - s_{t+3}) \left[(1 - \psi) W_{t+3,t+3}^k + \psi W_{t+1,t+3}^k \right] + s_{t+3} U_{t+3} \right\} \\
&\quad - \pi_{t+2}^{-\mu} \pi_{t+1}^{-\mu} w_{t,t}^k - \delta_{t+2} E_{t+2} \left\{ (1 - s_{t+3}) \left[(1 - \psi) W_{t+3,t+3}^k + \psi W_{t,t+3}^k \right] + s_{t+3} U_{t+3} \right\} \\
&= \pi_{t+2}^{-\mu} \left(w_{t+1,t+1}^k - \pi_{t+1}^{-\mu} w_{t,t}^k \right) + \delta_{t+2} E_{t+2} \left[(1 - s_{t+3}) \psi (W_{t+1,t+3}^k - W_{t,t+3}^k) \right].
\end{aligned}$$

Use of this in (A.37) gives

$$\begin{aligned} W_{t+1,t+1}^k - W_{t,t+1}^k &= \left(w_{t+1,t+1}^k - \pi_{t+1}^{-\mu} w_{t,t}^k \right) \left[1 + \delta_{t+1} E_{t+1} (1 - s_{t+2}) \psi \pi_{t+2}^{-\mu} \right] \\ &\quad + \delta_{t+1} E_{t+1} \delta_{t+2} \left[(1 - s_{t+2}) \psi (1 - s_{t+3}) \psi (W_{t+1,t+3}^k - W_{t,t+3}^k) \right]. \end{aligned}$$

Proceeding recursively in this way gives

$$\begin{aligned} W_{t+1,t+1}^k - W_{t,t+1}^k &= \left(w_{t+1,t+1}^k - \pi_{t+1}^{-\mu} w_{t,t}^k \right) E_{t+1} \left[1 + \sum_{i=2}^{\infty} \prod_{j=2}^i \left(\delta_{t+j-1} (1 - s_{t+j}) \psi \pi_{t+j}^{-\mu} \right) \right] \\ &= \left(w_{t+1,t+1}^k - \pi_{t+1}^{-\mu} w_{t,t}^k \right) E_{t+1} \left[\sum_{i=1}^{\infty} \prod_{j=2}^i \left(\delta_{t+j-1} (1 - s_{t+j}) \psi \pi_{t+j}^{-\mu} \right) \right], \end{aligned} \tag{A.38}$$

again with the convention that the product term equals 1 for $i < j$. (A.23) then follows directly from (A.24).

B.6 Derivation of equation (A.26)

From (A.25) and (A.2) with $V_{t+1} \geq 0$ for all t ,

$$\begin{aligned}
& J_{t,t} - V_t \\
&= (1 - q_t \kappa) J_{t,t} + c_t - \left[q_t (1 - \kappa) \delta_t E_t \tilde{J}_{t,t+1} + (1 - q_t) \delta_t E_t V_{t+1} \right] \\
&= (1 - q_t \kappa) \left\{ p_t - w_t^* + \delta_t E_t \left[(1 - s_{t+1}) \tilde{J}_{t,t+1} + s_{t+1} V_{t+1} \right] \right\} \\
&\quad + c_t - \left[q_t (1 - \kappa) \delta_t E_t \tilde{J}_{t,t+1} + (1 - q_t) \delta_t E_t V_{t+1} \right] \\
&= (1 - q_t \kappa) (p_t - w_t^*) + c_t + \delta_t E_t \left\{ (1 - q_t \kappa) \left[(1 - s_{t+1}) \tilde{J}_{t,t+1} + s_{t+1} V_{t+1} \right] \right. \\
&\quad \left. - \left[q_t (1 - \kappa) \tilde{J}_{t,t+1} + (1 - q_t) V_{t+1} \right] \right\} \\
&= (1 - q_t \kappa) (p_t - w_t^*) + c_t + \delta_t E_t \left\{ \left[(1 - q_t \kappa) (1 - s_{t+1}) - q_t (1 - \kappa) \right] \tilde{J}_{t,t+1} \right. \\
&\quad \left. + \left[(1 - q_t \kappa) s_{t+1} - (1 - q_t) \right] V_{t+1} \right\} \\
&= (1 - q_t \kappa) (p_t - w_t^*) + c_t + \delta_t E_t \left\{ \left[1 - s_{t+1} - q_t \kappa + q_t \kappa s_{t+1} - q_t + q_t \kappa \right] \tilde{J}_{t,t+1} \right. \\
&\quad \left. + \left[s_{t+1} - q_t \kappa s_{t+1} - 1 + q_t \right] V_{t+1} \right\}.
\end{aligned}$$

This can be re-written as (A.26).

B.7 Derivation of equation (A.29)

From (A.28) and (A.5),

$$\begin{aligned}
& W_{t,t} - U_t \\
&= (1 - f_t \kappa) W_{t,t} - \left[f_t (1 - \kappa) \left(z_t + \delta_t E_t \tilde{W}_{t,t+1} \right) + (1 - f_t) (z_t + \delta_t E_t U_{t+1}) \right] \\
&= (1 - f_t \kappa) \left\{ w_t^* + \delta_t E_t \left[(1 - s_{t+1}) \tilde{W}_{t,t+1} + s_{t+1} U_{t+1} \right] \right\} \\
&\quad - \left[f_t (1 - \kappa) \delta_t E_t \tilde{W}_{t,t+1} + (1 - f_t) \delta_t E_t U_{t+1} + f_t (1 - \kappa) z_t + (1 - f_t) z_t \right] \\
&= (1 - f_t \kappa) (w_t^* - z_t) + \delta_t E_t \left\{ (1 - f_t \kappa) \left[(1 - s_{t+1}) \tilde{W}_{t,t+1} + s_{t+1} U_{t+1} \right] \right. \\
&\quad \left. - \left[f_t (1 - \kappa) \tilde{W}_{t,t+1} + (1 - f_t) U_{t+1} \right] \right\} \\
&= (1 - f_t \kappa) (w_t^* - z_t) + \delta_t E_t \left\{ \left[(1 - f_t \kappa) (1 - s_{t+1}) - f_t (1 - \kappa) \right] \tilde{W}_{t,t+1} \right. \\
&\quad \left. + \left[(1 - f_t \kappa) s_{t+1} - (1 - f_t) \right] U_{t+1} \right\} \\
&= (1 - f_t \kappa) (w_t^* - z_t) + \delta_t E_t \left\{ \left[1 - s_{t+1} - f_t \kappa + f_t \kappa s_{t+1} - f_t + f_t \kappa \right] \tilde{W}_{t,t+1} \right. \\
&\quad \left. + \left[s_{t+1} - f_t \kappa s_{t+1} - 1 + f_t \right] U_{t+1} \right\}.
\end{aligned}$$

This can be re-written as (A.29).

B.8 Derivation of equation (3)

Use of (A.30) in (A.15) yields

$$\begin{aligned}
& (1 - f_t \kappa) (w_t^* - z_t) + \delta_t E_t \left\{ (1 - s_{t+1} - f_t \right. \\
&\quad \left. + \kappa f_t s_{t+1}) \left[\lambda (J_{t+1,t+1} - V_{t+1}) + \frac{\lambda}{1 - \lambda} \gamma_{t+1} + \psi (J_{t+1,t+1} - J_{t,t+1}) \right] \right\} \\
&= \lambda (J_{t,t} - V_t) + \frac{\lambda}{1 - \lambda} \gamma_t
\end{aligned}$$

or

$$\begin{aligned}
& (1 - f_t \kappa) (w_t^* - z_t) \\
&= \lambda (J_{t,t} - V_t) \\
&\quad - \delta_t E_t \left\{ (1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \left[\lambda (J_{t+1,t+1} - V_{t+1}) + \psi (J_{t+1,t+1} - J_{t,t+1}) \right] \right\} \\
&\quad + \frac{\lambda}{1 - \lambda} \left\{ \gamma_t - \delta_t E_t \left[(1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \gamma_{t+1} \right] \right\}. \tag{A.39}
\end{aligned}$$

From (A.27) (subtracting the second line below from both sides),

$$\begin{aligned}
& J_{t,t} - V_t \\
&\quad - \delta_t E_t \left\{ (1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \left[(J_{t+1,t+1} - V_{t+1}) + \frac{\psi}{\lambda} (J_{t+1,t+1} - J_{t,t+1}) \right] \right\} \\
&= (1 - q_t \kappa) (p_t - w_t^*) + c_t \\
&\quad + \delta_t E_t \left[(1 - s_{t+1} - q_t + q_t \kappa s_{t+1}) (J_{t+1,t+1} - V_{t+1} - \psi (J_{t+1,t+1} - J_{t,t+1})) \right] \\
&\quad - \delta_t E_t \left\{ (1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \left[(J_{t+1,t+1} - V_{t+1}) + \frac{\psi}{\lambda} (J_{t+1,t+1} - J_{t,t+1}) \right] \right\}
\end{aligned}$$

or

$$\begin{aligned}
& J_{t,t} - V_t \\
&\quad - \delta_t E_t \left\{ (1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \left[(J_{t+1,t+1} - V_{t+1}) + \frac{\psi}{\lambda} (J_{t+1,t+1} - J_{t,t+1}) \right] \right\} \\
&= (1 - q_t \kappa) (p_t - w_t^*) + c_t + \delta_t E_t \left[\left[(1 - s_{t+1} - q_t + q_t \kappa s_{t+1}) \right. \right. \\
&\quad \left. \left. - (1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \right] (J_{t+1,t+1} - V_{t+1}) \right] \\
&\quad - \psi \delta_t E_t \left[(1 - s_{t+1} - q_t + q_t \kappa s_{t+1}) (J_{t+1,t+1} - J_{t,t+1}) \right. \\
&\quad \left. + (1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \frac{1}{\lambda} (J_{t+1,t+1} - J_{t,t+1}) \right]
\end{aligned}$$

or

$$\begin{aligned}
& J_{t,t} - V_t \\
& - \delta_t E_t \left\{ (1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \left[(J_{t+1,t+1} - V_{t+1}) + \frac{\psi}{\lambda} (J_{t+1,t+1} - J_{t,t+1}) \right] \right\} \\
& = (1 - f_t \kappa) (p_t - w_t^*) + \kappa (f_t - q_t) (p_t - w_t^*) + c_t \frac{f_t}{q_t} - c_t \frac{f_t - q_t}{q_t} \\
& + (f_t - q_t) \delta_t E_t [(1 - \kappa s_{t+1}) (J_{t+1,t+1} - V_{t+1})] \\
& - \psi \delta_t E_t \left\{ \left[\left(1 + \frac{1}{\lambda}\right) (1 - s_{t+1}) - \left(\frac{f_t}{\lambda} + q_t\right) (1 - \kappa s_{t+1}) \right] (J_{t+1,t+1} - J_{t,t+1}) \right\}
\end{aligned}$$

or

$$\begin{aligned}
& J_{t,t} - V_t \\
& - \delta_t E_t \left\{ (1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \left[(J_{t+1,t+1} - V_{t+1}) + \frac{\psi}{\lambda} (J_{t+1,t+1} - J_{t,t+1}) \right] \right\} \\
& = (1 - f_t \kappa) (p_t - w_t^*) + c_t \frac{f_t}{q_t} + (f_t - q_t) \left\{ -\frac{c_t}{q_t} + \kappa (p_t - w_t^*) \right. \\
& \quad \left. + \delta_t E_t [(1 - \kappa s_{t+1}) (J_{t+1,t+1} - V_{t+1})] \right\} \\
& - \psi \delta_t E_t \left\{ \left[\left(1 + \frac{1}{\lambda}\right) (1 - s_{t+1}) - \left(\frac{f_t}{\lambda} + q_t\right) (1 - \kappa s_{t+1}) \right] (J_{t+1,t+1} - J_{t,t+1}) \right\}.
\end{aligned}$$

Hence (A.39) can be written

$$\begin{aligned}
& (1 - f_t \kappa) (w_t^* - z_t) \\
& = \lambda \left[(1 - f_t \kappa) (p_t - w_t^*) + c_t \frac{f_t}{q_t} \right] \\
& + \lambda (f_t - q_t) \left\{ -\frac{c_t}{q_t} + \kappa (p_t - w_t^*) + \delta_t E_t [(1 - \kappa s_{t+1}) (J_{t+1,t+1} - V_{t+1})] \right\} \\
& - \lambda \psi \delta_t E_t \left\{ \left[\left(1 + \frac{1}{\lambda}\right) (1 - s_{t+1}) - \left(\frac{f_t}{\lambda} + q_t\right) (1 - \kappa s_{t+1}) \right] (J_{t+1,t+1} - J_{t,t+1}) \right\} \\
& + \frac{\lambda}{1 - \lambda} \{ \gamma_t - \delta_t E_t [(1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \gamma_{t+1}] \}. \tag{A.40}
\end{aligned}$$

With new vacancies created for which the wage is negotiated at every t , free entry (A.3) implies $V_t = V_{t+1} = 0$. From (A.2) with $V_t = V_{t+1} = 0$,

$$\kappa J_{t,t} + (1 - \kappa) \delta_t E_t \tilde{J}_{t,t+1} - \frac{c_t}{q_t} = 0, \text{ for all } t.$$

Substitution for $J_{t,t}$ from (A.25) in this, with use of the free entry condition (A.3) so $V_{t+1} = 0$, yields

$$\kappa (p_t - w_t^*) + \kappa \delta_t E_t \left[(1 - s_{t+1}) \tilde{J}_{t,t+1} \right] + (1 - \kappa) \delta_t E_t \tilde{J}_{t,t+1} - \frac{c_t}{q_t} = 0, \text{ for all } t,$$

or

$$\kappa (p_t - w_t^*) + \delta_t E_t \left[(1 - \kappa s_{t+1}) \tilde{J}_{t+1,t+1} \right] - \frac{c_t}{q_t} = 0, \text{ for all } t.$$

With the use of (A.21), this can be written

$$\kappa (p_t - w_t^*) + \delta_t E_t \left\{ (1 - \kappa s_{t+1}) \left[J_{t+1,t+1} - \psi (J_{t+1,t+1} - J_{t,t+1}) \right] \right\} - \frac{c_t}{q_t} = 0, \text{ for all } t,$$

or

$$\kappa (p_t - w_t^*) + \delta_t E_t \left[(1 - \kappa s_{t+1}) J_{t+1,t+1} \right] - \frac{c_t}{q_t} = \psi \delta_t E_t \left[(1 - \kappa s_{t+1}) (J_{t+1,t+1} - J_{t,t+1}) \right],$$

for all t . (A.41)

With the use of (A.41) and $V_{t+1} = 0$, (A.40) can be written

$$\begin{aligned}
& (1 - f_t \kappa) (w_t^* - z_t) \\
&= \lambda \left[(1 - f_t \kappa) (p_t - w_t^*) + c_t \frac{f_t}{q_t} \right] \\
&\quad + \lambda (f_t - q_t) \psi \delta_t E_t [(1 - \kappa s_{t+1}) (J_{t+1,t+1} - J_{t,t+1})] \\
&\quad - \lambda \psi \delta_t E_t \left\{ \left[\left(1 + \frac{1}{\lambda}\right) (1 - s_{t+1}) - \left(\frac{f_t}{\lambda} + q_t\right) (1 - \kappa s_{t+1}) \right] (J_{t+1,t+1} - J_{t,t+1}) \right\} \\
&\quad + \frac{\lambda}{1 - \lambda} \{ \gamma_t - \delta_t E_t [(1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \gamma_{t+1}] \}
\end{aligned}$$

or

$$\begin{aligned}
& (1 - f_t \kappa) (w_t^* - z_t) \\
&= \lambda \left[(1 - f_t \kappa) (p_t - w_t^*) + c_t \frac{f_t}{q_t} \right] \\
&\quad + \lambda \psi \delta_t E_t \left\{ \left[(f_t - q_t) (1 - \kappa s_{t+1}) \right. \right. \\
&\quad \left. \left. - \left(1 + \frac{1}{\lambda}\right) (1 - s_{t+1}) + \left(\frac{f_t}{\lambda} + q_t\right) (1 - \kappa s_{t+1}) \right] (J_{t+1,t+1} - J_{t,t+1}) \right\} \\
&\quad + \frac{\lambda}{1 - \lambda} \{ \gamma_t - \delta_t E_t [(1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \gamma_{t+1}] \}
\end{aligned}$$

or

$$\begin{aligned}
& (1 - f_t \kappa) (w_t^* - z_t) \\
&= \lambda \left[(1 - f_t \kappa) (p_t - w_t^*) + c_t \frac{f_t}{q_t} \right] \\
&\quad + \lambda \psi \delta_t E_t \left\{ \left[f_t \left(1 + \frac{1}{\lambda}\right) (1 - \kappa s_{t+1}) - \left(1 + \frac{1}{\lambda}\right) (1 - s_{t+1}) \right] (J_{t+1,t+1} - J_{t,t+1}) \right\} \\
&\quad + \frac{\lambda}{1 - \lambda} \{ \gamma_t - \delta_t E_t [(1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \gamma_{t+1}] \}
\end{aligned}$$

or

$$\begin{aligned}
& (1 - f_t \kappa) (w_t^* - z_t) \\
&= \lambda \left[(1 - f_t \kappa) (p_t - w_t^*) + c_t \frac{f_t}{q_t} \right] \\
&\quad + \psi (1 + \lambda) \delta_t E_t \left\{ [f_t (1 - \kappa s_{t+1}) - (1 - s_{t+1})] (J_{t+1,t+1} - J_{t,t+1}) \right\} \\
&\quad + \frac{\lambda}{1 - \lambda} \left\{ \gamma_t - \delta_t E_t [(1 - s_{t+1} - f_t + \kappa f_t s_{t+1}) \gamma_{t+1}] \right\}.
\end{aligned}$$

With the definition $\delta_t^\kappa = \delta_{t-1} (1 - s_t - f_{t-1} + \kappa s_t f_{t-1})$, this can be written

$$\begin{aligned}
& (1 - f_t \kappa) (w_t^* - z_t) \\
&= \lambda \left[(1 - f_t \kappa) (p_t - w_t^*) + c_t \frac{f_t}{q_t} \right] + \frac{\lambda}{1 - \lambda} [\gamma_t - E_t (\delta_{t+1}^\kappa \gamma_{t+1})] \\
&\quad - \psi (1 + \lambda) E_t [\delta_{t+1}^\kappa (J_{t+1,t+1} - J_{t,t+1})].
\end{aligned} \tag{A.42}$$

With the use of (A.24) averaged over all jobs, equation (A.42) can be written as (3).

Appendix C Data

Series	Description	Source	Series ID	Sample
Labor share	nfb, sa, level	BLS	PRS85006173	1948q1-2011q3
Output per person	nfb, sa, index	BLS	PRS85006163	1948q1-2011q4
Unemployment	total, sa, end of quarter	BLS	LNS13000000	1948q1-2011q4
Unemployment	total, nsa, end of quarter	BLS	LNU03000000	1948q1-2011q4
Unemployment rate	sa, quarterly average	BLS	LNS14000000Q	1948q1-2011q4
Unempl. < 5 weeks	total, sa, monthly	BLS	LNS13008396	1948m1-2011m12
Unempl. < 5 weeks	total, nsa, monthly	BLS	LNU03008396	1948m1-2011m12
Unempl. 5 – 14 weeks	total, sa, monthly	BLS	LNS13008396	1948m1-2011m12
Unempl. 5 – 14 weeks	total, nsa, monthly	BLS	LNU03008756	1948m1-2011m12
Unempl. < 14 weeks	total, last month of quarter	derived	–	1948q1-2011q4
Employment	nf, total, sa, end of quarter	BLS	CES0000000001	1948q1-2011q4
Employment	nf, total, nsa, end of quarter	BLS	CEU0000000001	1948q1-2011q4
Help Wanted Index (HWI)	index, sa	MEI		1951q1-2008q1
Job openings rate	$\frac{\text{openings}}{\text{openings} + \text{employment}}$	Barnichon (2010)	–	1951q1-2011q3
Job openings	nf, total, sa, total over quarter	BLS/JOLTS	JTS00000000JOL	2001q1-2011q4
Separations	nf, total, sa, total over quarter	BLS/JOLTS	JTS00000000TSL	2001q1-2011q4
3mTbill	sec. market rate, quart. av.	FRED	TB3MS_20120305	1948q1-2011q4
Implicit price deflator	nfb, index	BLS	PRS85006143	1948q1-2011q4

Table 8: Data description and sources.

Abbreviations: nfb is nonfarm business sector, nf is nonfarm sector, sa is seasonally adjusted, nsa is non-seasonally adjusted, end of quarter is last month of quarter.

Sources: BLS is Bureau of Labor Statistics, JOLTS is Job Openings and Labor Turnover Survey, MEI is OECD Main Economic Indicators, FRED is St. Louis Fed’s Economic Database.

Table 8 specifies the raw data we use, with sources. Figure 1 uses the original data of Hagedorn and Manovskii (2008), downloaded from the *AER* website. This data contains series for the labour share, output per person, the unemployment rate and HWI from 1951q1 to 2004q4. All other empirical results in the paper use the most recent revisions of these series. A comparison of the series used by Hagedorn and Manovskii (2008) with the most recent revisions is in Table 9. For the labour share and productivity series, the revisions are minor. The unemployment and HWI series are identical.

Vacancies are measured using the approach described in Shimer (2005, p. 29), specifically

Series	st. dev.	first autocorrelation	correlation
Labor share, ours	0.020	0.899	0.989
Labor share, HM	0.020	0.907	-
log(prod)-HP trend, ours	0.0133	0.758	0.990
log(prod)-HP trend, HM	0.0132	0.765	-
Unemployment rate, ours	0.015	0.966	1
Unemployment rate, HM	0.015	0.966	-
HWI (unfiltered), ours	0.373	0.978	1
HWI (unfiltered), HM	0.373	0.978	-

Table 9: Comparison of Hagedorn and Manovskii (2008) data series (HM) with the most recent revisions (ours). HP trend is computed with smoothing parameter 1600.

removing variations at very low frequency (in logs), computed using the HP filter with smoothing parameter 10^5 . Barnichon (2010) provides a different vacancy series, available from his website. This is derived by merging newspaper advertising data (which was discontinued after 2008q1) with online advertising (available since 1995) to create a composite HWI index. Barnichon (2010) reports a vacancy rate series v^r that conforms to the JOLTS definition, that is, $v^r = \frac{V}{V+J}$, where V is total vacancies and J is employment. Using our data on employment, we obtain a measure of total vacancies as $V = \frac{v^r}{1-v^r} J$. (We compute this at the monthly frequency and then aggregate to quarterly).

Vacancy data are used primarily for a measure of the probability of filling a vacancy, q_t . Table 10 compares series for this probability using the different vacancy series. Our baseline measure (based on filtered HWI), constructed as in Shimer (2005), is very similar to the measure based on JOLTS data (correlation 0.88). It is also very similar to the one based on the vacancy data constructed by Barnichon (2010) (correlation 0.97). The measure that uses the raw (unfiltered) HWI data is very different from the others.

Summary statistics for the series that appear in the model of equation (3) are in Table 11.

Appendix D Additional empirical results

This appendix gives details of various robustness checks.

Correlations over JOLTS sample 2001q1-2008q1

Vacancy series	Unfiltered HWI	Barnichon (2010)	JOLTS
Filtered HWI	0.161	0.965	0.876
Unfiltered HWI	1	0.104	0.103
Barnichon (2010)		1	0.956
JOLTS			1

Summary statistics

	JOLTS sample 2001q1-2008q1		full sample 1951q1-2008q1	
Vacancy series	st. dev.	first autocorrelation	st. dev.	first autocorrelation
Filtered HWI	0.040	0.776	0.067	0.944
Unfiltered HWI	0.064	0.691	0.097	0.943
Barnichon (2010)	0.042	0.789	0.073	0.945
JOLTS	0.044	0.834	-	-

Table 10: Series for probability of filling a vacancy with different vacancy series.

Series	st. dev.	first autocorrelation
w_t	0.017	0.940
p_t	0.013	0.762
f_t	0.066	0.941
q_t	0.032	0.741
s_t	0.015	0.914
δ_t	0.006	0.681
π_t	0.007	0.818

Table 11: Summary statistics for the variables used in the econometric model. w_t and p_t are detrended using HP filter-based productivity trend.

Truncation length (quarters):	28	24	20	16	12	8
ψ	0.631	0.631	0.631	0.631	0.631	0.631
μ	0.000	0.000	0.000	0.000	0.000	0.000
z	1.000	1.000	1.000	1.000	1.000	1.000
c^K	0.000	0.000	0.000	0.000	0.000	0.000
c^W	0.083	0.083	0.083	0.083	0.083	0.083
ζ	0.000	0.000	0.000	0.000	0.000	0.000
κ	1.000	1.000	1.000	1.000	1.000	1.000
H^K	0.011	0.011	0.011	0.011	0.011	0.011
H^W	1.145	1.145	1.145	1.145	1.145	1.145
γ	-1.376	-1.376	-1.376	-1.376	-1.376	-1.376
λ	0.131	0.131	0.131	0.131	0.131	0.131
GMM objective	8.923	8.923	8.923	8.923	8.920	8.909

Table 12: Estimates of eq. (3) with truncation of the infinite sum at various quarters.

Table 12 gives estimates corresponding to column 4 of Table 4 for different truncation lengths of the infinite sum on the right-hand side of equation (3). The point estimates are identical to 3 decimal places for truncations of 8 or more quarters.

Table 13 gives estimates corresponding to Table 4 with wages and productivity separately detrended (using separate HP filters). The fit is slightly worse than with the baseline specification in Table 4. The point estimates of the parameters are similar, with somewhat higher β and lower λ (though still well within the confidence intervals given in Table 5 for the baseline specification), but with point estimates of zero for both the vacancy posting cost parameters. The point estimates of ψ remain significantly greater than zero, so our main conclusions are unaffected.

Table 14 gives estimates corresponding to Table 4 using the following instrument set: a constant, four lags of w_t and three lags of p_t , f_t , q_t , and s_t . The conclusions to be drawn remain unchanged.

Table 15 gives estimates corresponding to Table 4 with non-seasonally adjusted data for employment and unemployment. The fit is slightly worse than in the baseline specification in Table 4 but the conclusions we draw remain unchanged.

Table 16 gives estimates corresponding to Table 4 with vacancy data as computed by Barnichon (2010) and the sample extended by 3 years to 2007q4. These give somewhat higher point estimates

Parameter	1	2	3	4
ψ	0.682 (0.024)	0.682 (0.054)	0.682 (0.064)	0.516 (0.078)
μ	0.040 (0.179)	0.040 (0.190)	0.040 (0.292)	0.000 (0.271)
β	0.280 (1.450)	0.280 (1.476)	0.280 (5.208)	— (—)
z	0.930 (0.143)	0.930 (0.147)	0.930 (0.528)	1.000 (0.034)
c^K	0.000 (8.265)	0.000 (8.437)	0.000 (18.856)	0.000 (23.25)
c^W	0.000 (8.267)	0.000 (8.440)	0.000 (18.858)	0.000 (23.27)
ζ	0.097 (3×10^{13})	0.097 (3×10^{13})	0.097 (3×10^{13})	0.164 (3518)
κ	— (—)	0.000 (0.588)	0.000 (0.838)	1.000 (0.467)
H^K	— (—)	— (—)	0.000 (64.660)	1.470 (1217)
H^W	— (—)	— (—)	0.000 (64.812)	0.337 (1333)
γ	— (—)	— (—)	— (—)	−1.877 (150.3)
λ	— (—)	— (—)	— (—)	0.123 (9.841)
GMM objective	13.153	13.153	13.153	11.526
Hansen test p value	0.358	0.358	0.358	0.490
Hansen test proj. p value	0.726	0.726	0.726	0.832

Table 13: Estimates of eq. (3) as in Table 4 with wages detrended by an HP filter with smoothing parameter 1600, instead of by the productivity trend.

Parameter	1	2	3	4
ψ	0.746 (0.028)	0.746 (0.054)	0.746 (0.062)	0.632 (0.079)
μ	0.256 (0.188)	0.256 (0.222)	0.256 (0.259)	0.060 (0.542)
β	0.020 (3.493)	0.020 (4.041)	0.020 (11.940)	— (—)
z	0.950 (0.180)	0.950 (0.210)	0.950 (0.621)	1.000 (0.104)
c^K	0.330 (5375.722)	0.330 (7596.439)	0.330 (9385.721)	0.704 (3308.491)
c^W	0.504 (5463.599)	0.504 (7717.699)	0.504 (9784.890)	0.482 (3642.685)
ζ	0.162 (9281.905)	0.162 (13099.239)	0.162 (16085.326)	0.800 (1074.968)
κ	— (—)	0.000 (0.544)	0.000 (0.660)	0.931 (0.450)
H^K	— (—)	— (—)	0.000 (1431.706)	0.629 (31165.073)
H^W	— (—)	— (—)	0.000 (1435.087)	6.284 (32967.394)
γ	— (—)	— (—)	— (—)	−11.523 (4905.867)
λ	— (—)	— (—)	— (—)	0.013 (5.633)
GMM objective	11.511	11.511	11.511	8.562
Hansen test p value	0.319	0.319	0.319	0.286
Hansen test proj. p value	0.829	0.829	0.829	0.953

Table 14: Estimates of eq. (3) as in Table 4 with the instrument set consisting of a constant, four lags of w_t and three lags of p_t , q_t , f_t and s_t .

Parameter	1	2	3	4
ψ	0.736 (0.050)	0.736 (0.054)	0.736 (0.072)	0.616 (0.078)
μ	0.301 (0.272)	0.301 (0.283)	0.301 (0.273)	0.242 (0.304)
β	0.013 (3.022)	0.013 (3.071)	0.013 (4.625)	— (—)
z	0.951 (0.162)	0.951 (0.166)	0.951 (0.242)	0.952 (0.113)
c^K	0.294 (10293.212)	0.294 (10302.540)	0.294 (11957.181)	0.000 (15722.291)
c^W	1.334 (10626.352)	1.334 (10640.468)	1.334 (12461.001)	1.159 (16188.884)
ξ	0.000 (7971.002)	0.000 (7980.958)	0.000 (9111.905)	0.033 (12910.579)
κ	— (—)	0.000 (0.508)	0.000 (0.626)	1.000 (0.319)
H^K	— (—)	— (—)	0.000 (797.681)	0.002 (1386.153)
H^W	— (—)	— (—)	0.000 (797.131)	0.127 (1403.494)
γ	— (—)	— (—)	— (—)	−0.055 (33.678)
λ	— (—)	— (—)	— (—)	0.005 (2.496)
GMM objective	17.824	17.824	17.824	15.150
Hansen test p value	0.086	0.086	0.086	0.056
Hansen test proj. p value	0.400	0.400	0.400	0.585

Table 15: Estimates of eq. (3) as in Table 4 with non-seasonally adjusted data for employment and unemployment.

Parameter	1	2	3	4
ψ	0.674 (0.034)	0.670 (0.056)	0.670 (0.068)	0.670 (0.071)
μ	0.000 (0.366)	0.000 (0.370)	0.000 (0.388)	0.000 (0.453)
β	0.386 (1.937)	0.385 (2.040)	0.385 (17.793)	— (—)
z	0.909 (0.283)	0.908 (0.301)	0.908 (2.810)	0.911 (2.622)
c^K	0.000 (221.351)	0.000 (250.214)	0.000 (749.895)	0.000 (788.170)
c^W	0.028 (221.437)	0.030 (250.321)	0.030 (751.319)	0.031 (789.446)
ξ	0.248 (6147.263)	0.269 (6341.802)	0.269 (18459.787)	0.356 (16322.833)
κ	— (—)	0.000 (0.468)	0.000 (0.534)	0.000 (0.762)
H^K	— (—)	— (—)	0.000 (140.474)	0.000 (155.331)
H^W	— (—)	— (—)	0.000 (141.168)	0.000 (155.220)
γ	— (—)	— (—)	— (—)	−0.004 (1.244)
λ	— (—)	— (—)	— (—)	0.622 (42.277)
GMM objective	17.178	17.163	17.163	17.157
Hansen test p value	0.143	0.144	0.144	0.103
Hansen test proj. p value	0.442	0.443	0.443	0.444

Table 16: Estimates of eq. (3) as in Table 4 with vacancy data computed by Barnichon (2010) and the sample extended by 3 years to 2007q4.

for β and λ than the baseline specification in Table 4 (though still well within the confidence intervals given in Table 5 for the baseline specification). But the estimates of the wage rigidity parameter ψ are not very different, so our main conclusions remain unchanged.