

1 **ABSTRACT**

2 *Purpose:*

3 The HPS2-THRIVE trial among patients at high risk of vascular disease found that adding ER niacin-  
4 laropiprant to intensive statin-based LDL-lowering therapy had no benefit on cardiovascular outcomes.  
5 However, it also identified previously unrecognized serious adverse effects (including new-onset  
6 diabetes, bleeding and infection). Our objective was to explore the safety profile of niacin-laropiprant  
7 and determine whether any types of participant were at lower (or higher) risk of its adverse effects.

8 *Methods:*

9 HPS2-THRIVE was a randomized, double-blind trial of niacin-laropiprant (2000/40 mg daily) versus  
10 placebo among 25,673 patients at high risk of vascular disease. Information on all serious adverse  
11 events was collected during a median of 3.9 years of study treatment. Effects of niacin-laropiprant on  
12 new-onset diabetes, disturbances of diabetes control, bleeding, infection and gastrointestinal upset  
13 were estimated by (i) time after randomization; (ii) severity; (iii) baseline characteristics; (iv) baseline  
14 risk of the adverse event of interest; and (v) risk of major vascular event.

15 *Findings:*

16 The hazard ratio for new onset diabetes with niacin/laropiprant was 1.32 (95% CI 1.16-1.51;  $p < 0.001$ ),  
17 which corresponded to an absolute excess of 4 people (95% CI 2-6) developing diabetes per thousand  
18 person-years in the study population as a whole. Among the 8299 participants with diabetes at  
19 baseline, the hazard ratio for serious disturbances in diabetes control was 1.56 (95% CI 1.35-1.80)  
20 corresponding to an absolute excess of 12 (95% CI 8-16) per 1000 person-years. The hazard ratio for  
21 serious bleeding was 1.38 (95% CI 1.17-1.63;  $p < 0.001$ ) corresponding to an absolute excess of 2 (95%  
22 CI 1-3) per 1000 person-years; and for serious infection, it was 1.22 (95% CI 1.11-1.34;  $p < 0.001$ )  
23 corresponding to an absolute excess of 4 (95% CI 2-6) per thousand person-years. The excess risks of  
24 these serious adverse events was larger in the first year after starting niacin-laropiprant than in later  
25 years (except for the excess of infection which did not appear to attenuate with time), and the risks of

26 non-fatal and fatal events were similarly increased. The absolute excesses of each of these adverse  
27 effects were similar regardless of the baseline risk of the outcome.

28 *Implications:*

29 Clinicians or patients considering the use of niacin (in addition to, or instead of, a statin) despite the  
30 lack of evidence of cardiovascular benefits (at least when added to effective statin therapy) should  
31 take account of the significant risks of these serious adverse effects when making such decisions.

32 HPS2-THRIVE is registered on [clinicaltrials.gov](https://clinicaltrials.gov) (NCT00461630).

33

34 Keywords: niacin, adverse effects, cardiovascular disease, infection, bleedingse

35 **CONDENSED ABSTRACT**

36 Niacin-laropiprant causes a wide-range of adverse effects including new-onset diabetes, bleeding and  
37 infection. The absolute excess these is similar in nearly all of the different types of participant studied,  
38 suggesting that there is no type of patient who can take niacin-laropiprant without being at increased  
39 risk of such outcomes.

40

41 **Clinical perspectives**

42 Competency in medical knowledge: Niacin-laropiprant causes a number of adverse effects, including  
43 new-onset diabetes mellitus, bleeding and infection.

44 Competency in patient care: Consideration should be given to these adverse effects before niacin is  
45 initiated and during any consultation with a patient currently receiving niacin.

46

## 47 INTRODUCTION

48 Despite current best practice (including intensive statin-based [low density lipoprotein LDL](#)-cholesterol  
49 [\[LDL-C\]](#) reduction), patients with atherosclerotic disease remain at significant risk of future occlusive  
50 vascular events <sup>1,2</sup>. Niacin has been recommended to reduce this residual risk because of its  
51 apparently beneficial effects on the lipid profile (including lowering LDL-C, triglycerides and  
52 lipoprotein (a) and increasing [high density lipoprotein cholesterol \[HDL-C\]](#)) <sup>3</sup>. However, two  
53 randomized trials have found that despite these effects, niacin does not reduce the risk of vascular  
54 events when added to intensive statin-based LDL lowering therapy <sup>4,5</sup>. Furthermore, the large HPS2-  
55 THRIVE (Heart Protection Study 2–Treatment of HDL to Reduce the Incidence of Vascular Events)  
56 trial not only confirmed previously recognised adverse effects of niacin (in particular, cutaneous  
57 reactions, gastrointestinal upset and disturbance of diabetes control) but it also identified previously  
58 unrecognized serious adverse effects (including new-onset diabetes mellitus, bleeding and infection).  
59 As a consequence, the US Food and Drug Administration has withdrawn the indication for niacin to  
60 be used in combination with a statin<sup>6</sup>.

61 HPS2-THRIVE tested extended release (ER) niacin with laropiprant (niacin-laropiprant), a  
62 prostaglandin D2 receptor (DP<sub>1</sub>) antagonist. It is not possible to determine from the HPS2-THRIVE  
63 trial alone whether the harmful effects observed were due to niacin or laropiprant. However, similar  
64 effects with ER niacin given on its own were observed in the AIM-HIGH (Atherothrombosis  
65 Intervention in Metabolic Syndrome with Low HDL/High Triglycerides: Impact on Global Health  
66 Outcomes) trial <sup>7</sup>. In combination, therefore, these trials strongly suggest that niacin was responsible  
67 for the adverse effects that were observed.

68 It is uncertain whether certain types of patient are at higher or lower risks of the harmful effects of  
69 niacin. Furthermore, the predictors of outcomes such as bleeding (in the absence of a recent acute  
70 coronary syndrome or atrial fibrillation) and infection among patients with atherosclerotic disease are  
71 not well known. As some physicians and patients may still wish to continue to use niacin, knowledge  
72 of the risk factors for the harmful effects of niacin and whether the risks of these adverse effects vary

73 by those risk factors may be clinically useful. We have used the HPS2-THRIVE trial database to  
74 investigate whether there are subgroups of patients – in particular those at high risk of vascular  
75 disease or with a dyslipidaemia (low HDL-C/high triglycerides) who may be more likely to be  
76 considered for niacin therapy – in whom the harmful effects of niacin are less significant.

77 **METHODS**

78 *Study participants and procedures*

79 The design and methods of the HPS2-THRIVE trial have been described previously<sup>5,8</sup>. In brief, men  
80 and women aged 50-80 years were eligible if they had a history of myocardial infarction,  
81 cerebrovascular or peripheral arterial disease, or diabetes mellitus with evidence of symptomatic  
82 coronary disease. Patients were excluded if they had clinically significant hepatic, renal, muscle-  
83 related or other disease. Willing and eligible patients entered a two phase pre-randomization run-in  
84 period. During the first phase, participants' background LDL lowering therapy was standardized to  
85 either simvastatin 40 mg daily or ezetimibe/simvastatin 10/40 mg daily; during the second phase they  
86 took active niacin-laropiprant (1000/20 mg daily for 4 weeks then increased to 2000/40 mg daily).  
87 Participants who did not report adverse events and who remained eligible were randomly assigned to  
88 receive niacin-laropiprant (2000/40 mg daily) or matching placebo. Following randomization, follow-  
89 up assessments were scheduled at 3 and 6 months and then every 6 months for an average of about 4  
90 years at which all serious adverse events (SAE) were recorded. Non-serious adverse events were not  
91 recorded unless they were considered to be the reason for stopping the study treatment.

92 *Study outcomes*

93 New-onset diabetes was defined based on self-report by the participant or new use of hypoglycaemic  
94 medication. The definition of baseline diabetes also included HbA1c  $\geq 48$  mmol/mol, fasting glucose  
95  $\geq 7.0$  mmol/L or random (ie, <8 hours fasting) glucose  $> 11.1$  mmol/L. Disturbances in diabetic control  
96 included minor disturbances (eg, hyperglycaemia requiring hospitalisation for optimisation of blood  
97 glucose) and major disturbances (eg, ketoacidosis or hyperosmolar hyperglycaemic state). Infections  
98 and gastrointestinal events were defined as any SAE included in Medical Dictionary for Regulatory  
99 Activities (MedDRA version 14.0) system organ classes "infection and infestation" and  
100 "gastrointestinal disorder" (excluding any bleeding events) respectively. Bleeding was defined as any  
101 SAE included in the MedDRA "haemorrhage" standard medical query. If the SAE resulted in death,  
102 further information was requested from the local study site for central adjudication according to pre-

103 specified criteria by physicians who were unaware of the study treatment assignment. Non-fatal SAEs,  
104 other than those included in the composite primary outcome major vascular event (MVE: non-fatal  
105 MI, coronary death, stroke or arterial revascularization) or pre-specified safety outcomes (myopathy  
106 and hepatitis), were not adjudicated.

### 107 *Statistical analysis*

108 The effects of niacin-laropiprant on new-onset diabetes, bleeding and infection (separately) were  
109 examined by (i) time after randomization (i.e. in the first and later years after randomization); (ii)  
110 severity (defined by requirement for hospitalisation and whether the SAE resulted in death or not); (iii)  
111 subgroups of participant derived from baseline characteristics (as defined in the main trial data  
112 analysis plan <sup>8</sup> with the addition of individuals with the combination of HDL-C <0.9 mmol/L and  
113 triglycerides >1.7 mmol/L); (iv) baseline risk of the adverse event; and (v) baseline risk of MVE. All  
114 analyses were based on comparisons between all participants randomly assigned niacin-laropiprant  
115 versus all those assigned placebo (i.e., properly intention-to-treat analyses).

116 In order to develop baseline risk scores for harmful effects and for MVE, Cox proportional hazards  
117 models with stepwise model selection were used to identify significant baseline predictors of each  
118 relevant outcome (using a p-value threshold of 0.01), with randomized treatment allocation always  
119 included in the model. These independent risk factors were then used to construct a joint model which  
120 estimated each individual's risk of the outcome. The baseline characteristics considered and the  
121 variables selected for inclusion in each joint risk factor model are shown in webtable 1. Tertiles of the  
122 risk score (based on numbers of events rather than participants) were used to determine higher, mid,  
123 and lower risk groups shown in the figures.

124 The absolute excess of some particular adverse event due to niacin-laropiprant was defined as the  
125 difference between the two treatment arms in the number of events as a proportion of person years at  
126 risk, expressed as events per 1000 person years. Hazard ratios describing proportional effects of  
127 niacin-laropiprant (and interactions between individual baseline characteristics and randomized  
128 treatment) were estimated by Cox proportional hazards models based on the first occurrence of each

129 particular event after randomization. Tests for trend and heterogeneity were used to explore whether  
130 the effect of treatment varied by the various groups/categories examined. No formal correction for  
131 multiple testing was made, but p values were not to be considered significant unless  $p < 0.01$  due to the  
132 *post hoc* nature of these analyses and the number of analyses conducted.

133

134 **RESULTS**

135 Between January 2007 and May 2010, 51,698 patients were screened and 42,424 entered the pre-  
136 randomization run-in (of whom 6365 entered the active niacin-laropiprant phase directly). 4055 of the  
137 36,059 participants (11%) who entered the LDL standardization phase withdrew compared to 12,696  
138 of the 38,369 (33%) who entered the active niacin-laropiprant phase. 25,673 participants were  
139 randomized and followed for a mean of 3.6 years (median 3.9 years).<sup>5</sup> Study average adherence to  
140 niacin-laropiprant was 77.7% (i.e. the proportion reporting taking most of their study treatment at  
141 each study visit). 25.4% of participants assigned niacin-laropiprant discontinued the study drug  
142 compared to 16.6% of those assigned placebo (p<0.001; [see webtable 2](#)). Further details of the run-in  
143 period and baseline characteristics have been published elsewhere <sup>8</sup>.

144 *New-onset diabetes mellitus*

145 Among the 17,374 participants without diabetes mellitus at baseline, 494 (1.64% per year [py]) of  
146 those assigned niacin-laropiprant developed new-onset diabetes compared to 376 (1.23% py) of those  
147 assigned placebo (table 1); this corresponds to a one-third proportional increase (hazard ratio [HR]  
148 1.33; 95% CI 1.16-1.52; p<0.001) and an absolute excess of 4 people (95% CI 2-6) per thousand  
149 person-years developing diabetes in the study population as a whole. The proportional effect was  
150 greater in the first year of treatment (HR 2.02; 95% CI 1.58-2.58) than in later years (HR 1.09; 95%  
151 CI 0.93-1.28; p for heterogeneity <0.001).

152 After allowance for multiple comparisons, the proportional effect of niacin-laropiprant was broadly  
153 similar in the different types of participant studied (including those with low HDL-C/high  
154 triglycerides), with the exception of subgroups defined by adiposity (webfigure 1). In univariate  
155 analyses, there was evidence of an interaction (p<0.001) with body mass index (BMI) and, more  
156 weakly, with waist circumference (p=0.01): for each 5 kg/m<sup>2</sup> greater BMI, there was a 14% [95% CI  
157 10-18%] proportional reduction in the relative risk of new onset diabetes associated with assignment  
158 to niacin-laropiprant. Among participants with BMI <25 kg/m<sup>2</sup> (mean 23.1 kg/m<sup>2</sup>), 98 (4.2%) of those  
159 assigned niacin-laropiprant developed new-onset diabetes compared to 48 (2.0%) of those assigned

160 placebo (HR 2.09; 95% CI 1.48-2.96). By contrast, among participants with BMI  $\geq 30$  kg/m<sup>2</sup> (mean  
161 33.1 kg/m<sup>2</sup>), there were 165 (8.6%) versus 148 (8.1%) cases (HR 1.06; 95% CI 0.85-1.32).

162 Various independent predictors (identified by stepwise selection) were associated with the risk of  
163 new-onset diabetes (webtable 1). A risk score based on these variables was strongly predictive of  
164 new-onset diabetes and differentiated participant groups well: rate of new-onset diabetes among  
165 placebo-allocated participants in the lowest risk group was 1.7% compared to 25.9% in highest risk  
166 group (figure 1A). However, after allowance for multiple comparisons, the absolute effects of niacin-  
167 laropiprant on new-onset diabetes did not differ by baseline risk of new-onset diabetes or of MVE  
168 (figure 1A), or among participants with low baseline HDL-C/high triglycerides (webfigure 2A).

#### 169 *Disturbances of diabetes control*

170 Among the 8299 participants known to have diabetes at baseline, niacin-laropiprant was associated  
171 with a 56% proportional increase (HR 1.56; 95% CI 1.35-1.80) in individuals having serious  
172 disturbances of diabetes control, corresponding to an absolute excess of 12 people (95% CI 8-16;  
173 table 1) per 1000 person years in the overall study population. Most of the absolute excess was in  
174 minor disturbances, although the proportional effect on major disturbances appeared larger. The  
175 proportional effects were also larger in the first year after treatment started than in later years (p for  
176 heterogeneity <0.001).

#### 177 *Bleeding*

178 Overall, 326 (0.71% py) participants assigned niacin-laropiprant had at least one bleeding SAE during  
179 follow-up versus 238 (0.52% py) assigned placebo, corresponding to a 38% proportional increase (HR  
180 1.38; 95% CI 1.17-1.63; p<0.001) and an absolute excess of 2 people (95% CI 1-3; table 2) per 1000  
181 person-years. Most bleeds were intracranial (including haemorrhagic stroke) or gastrointestinal, but  
182 there was little evidence that the proportional effects of niacin-laropiprant varied at different sites (p  
183 for heterogeneity = 0.18).

184 About one-fifth of these bleeds were fatal, and the proportional effect of niacin-laropiprant did not  
185 appear to differ by severity of bleeding (p for heterogeneity = 0.64). The proportional effect appeared  
186 to be greater during the first year after treatment started (HR 2.00; 95% CI 1.42-2.83) than in later  
187 years (HR 1.22; 95% CI 1.01-1.48; p for heterogeneity = 0.01), and the absolute excess was also  
188 larger (albeit the difference was not statistically definite) in the first year than in subsequent years (4  
189 [95% CI 2-6] versus 1 [95% CI 0-3] per 1000 per year (p for heterogeneity = 0.02).

190 The proportional effect of niacin-laropiprant on bleeding was similar in different types of participant  
191 (webfigure 3). Although the risk score discriminated participants based on their risk of bleeding, the  
192 absolute effect of niacin-laropiprant did not vary by baseline risk of bleeding or MVE (figure 1B) or  
193 baseline low HDL-C/high triglycerides (webfigure 2B).

#### 194 *Infection*

195 Overall 1031 (2.32% py) participants assigned niacin-laropiprant versus 853 (1.90% py) assigned  
196 placebo had at least one infection reported as a SAE during follow-up, corresponding to a 22%  
197 proportional increase (HR 1.22; 95% CI 1.11-1.34; p<0.001) and an absolute excess of 4 people (95%  
198 CI 2-6; table 3) per thousand person-years. The most common sites of infection were lower  
199 respiratory and urinary, with little evidence that the proportional effect varied by site of infection (p  
200 for heterogeneity = 0.20; table 3). Only 5% of the infections resulted in death, and again there was  
201 little evidence that the proportional effect of niacin-laropiprant varied by severity of infection (p for  
202 heterogeneity = 0.51). The proportional effects appeared to be similar in the first year after treatment  
203 started (HR 1.34; 95% CI 1.13-1.59) and in later years (1.18; 95% CI 1.06-1.31; p for heterogeneity =  
204 0.21), as were the absolute effects (absolute excess 6 [95% CI 3-10] versus 3 [95% CI 1-6] per  
205 thousand per year).

206 The proportional effect of niacin-laropiprant on infection was similar in different types of participant  
207 (webfigure 4). However, the absolute effect did not appear to vary by baseline risk of infection or  
208 MVE (figure 1C) or baseline low HDL-C/high triglycerides (webfigure 2C).

209 **DISCUSSION**

210 Niacin-laropiprant causes a wide range of serious adverse effects, including new-onset diabetes,  
211 disturbances in diabetic control, bleeding, infection and gastrointestinal events. These are in addition  
212 to the larger number of non-serious adverse effects and effects on statin-associated myopathy<sup>5,8</sup>. The  
213 harmful effects of niacin-laropiprant described here appear to be consistent in different types of  
214 participant (with the exception of an interaction between adiposity and the effect on new-onset  
215 diabetes). Specifically, after identifying risk factors for each outcome of interest and combining these  
216 into a risk score, the absolute effect of niacin-laropiprant was similar regardless of baseline risk of the  
217 harmful effect in question. Therefore, whereas it was possible to identify a subgroup of participants at  
218 lower risk of the particular adverse event, it was not possible to identify a subgroup among whom the  
219 harmful effects of niacin-laropiprant were smaller. Furthermore, there was no evidence that  
220 participants either at high risk of vascular events or with a low HDL-C/high triglyceride  
221 dyslipidaemia (either of which might be considered an indication for niacin therapy) were at lower  
222 absolute risk of these harmful effects of niacin.

223

224 Niacin's effects on gastrointestinal events and disturbances of diabetes control are well-established<sup>9</sup>.  
225 However, it is not possible to determine from the HPS2-THRIVE data alone whether the effects of  
226 niacin-laropiprant observed on the previously unknown harmful effects (new-onset diabetes, bleeding  
227 and infection) were due to the ER niacin or the laropiprant. However, previous trials of niacin alone  
228 suggest that it was the ER niacin component that was responsible for the effects observed. For  
229 example, a meta-analysis investigating the effects of niacin alone on the development of new-onset  
230 diabetes found a 38% (95% CI 16-65%) increased risk of new-onset diabetes<sup>10</sup>, which is similar to  
231 the 33% (95% CI 16-52%) proportional increase that was observed in HPS2-THRIVE. Although  
232 insulin resistance (and consequently the risk of diabetes) increases with adiposity<sup>11</sup>, the HPS2-  
233 THRIVE results suggest that the adverse effect of niacin on new-onset diabetes reduces with  
234 increasing adiposity, which was not expected. The effects of niacin on insulin resistance and glucose

235 homeostasis are not fully understood <sup>12, 13</sup>. Genetic studies have suggested that genetic variants  
236 associated with lower LDL-C are associated with an increased risk of diabetes mellitus. <sup>14</sup> However,  
237 treatments which reduce LDL-C have discordant effects on the risk of diabetes (statins increase the  
238 risk modestly, <sup>15</sup> whereas cholesterol ester transfer protein inhibitors reduce the risk <sup>16</sup>) and the effect  
239 size observed in HPS2-THRIVE is disproportionately large compared to the reduction in LDL-C. The  
240 effect in HPS2-THRIVE was largely confined to the first year after the start of treatment, perhaps  
241 because the effects on glucose metabolism attenuate with time. Adherence to niacin-laropiprant also  
242 fell with time so this may also explain the apparent attenuation of effect during follow-up.

243

244 The effects on bleeding observed in HPS2-THRIVE are also consistent with those observed in  
245 previous trials of niacin alone. For example, in the Coronary Drug Project (CDP), which compared  
246 immediate release niacin 3 g daily (n=1119) versus placebo (n=2789) in male survivors of a  
247 myocardial infarction <sup>9</sup>, there was a non-significant excess of “black tarry stools” (presumably  
248 melaena due to gastrointestinal bleeding) among participants assigned niacin <sup>9</sup>. The AIM-HIGH trial  
249 was much smaller than HPS2-THRIVE and did not have the statistical power to detect the effect on  
250 serious bleeds that was observed in HPS2-THRIVE. So, although the proportional effect on SAEs due  
251 to bleeding (crude risk ratio 1.19; 95% CI 0.82-1.73) in the AIM-HIGH trial is consistent with that in  
252 HPS2-THRIVE (HR 1.38; 95% CI 1.17-1.63) there were too few such SAEs in AIM-HIGH for this  
253 difference (59 [3.4%] in the niacin group versus 49 [2.9%] in the placebo group; p=0.36) to be  
254 statistically significant <sup>7</sup>. However, when both serious and non-serious bleeding adverse events  
255 recorded in AIM-HIGH are considered (174 [10.1%] assigned niacin 137 [8.1%] assigned placebo),  
256 there was found to be a significant excess with niacin (crude risk ratio 1.25; 95% CI 1.01-1.55;  
257 p=0.04). The wide variety of sites at which bleeding was in excess suggests that niacin induces a  
258 generalized bleeding diathesis. It is known to reduce the platelet count by about 10% <sup>17</sup> and to inhibit  
259 platelet aggregation <sup>18</sup>, as well as reducing the concentrations of circulating coagulation factors and  
260 inhibitings thrombus formation *in vitro* . In addition, niacin has been found to exacerbate peptic ulcer  
261 disease <sup>19</sup> and, although patients with a recent diagnosis of a peptic ulcer were excluded from HPS2-

262 THRIVE, some of the excess in gastrointestinal bleeding may have been among people with  
263 undiagnosed peptic ulcers.

264

265 Infection was another previously unrecognized harmful effect of niacin. Information on infections was  
266 not collected during the CDP, but dysuria was significantly more common among niacin-allocated  
267 participants (31/1073 [2.9%] versus 32/2695 [1.2%];  $p < 0.001$ ), perhaps reflecting an excess of urinary  
268 tract infections<sup>9</sup>. AIM-HIGH provides clearer supporting evidence for the effect on infection  
269 observed in HPS2-THRIVE: with significant excesses in both serious infections (139 [8.1%] in the  
270 niacin group versus 98 [5.8%] in the placebo group; crude risk ratio 1.40; 95% CI 1.09-1.80) and in  
271 the combination of serious and non-serious infections (-674 [39.2%] versus 593 [35.0%]; crude risk  
272 ratio 1.12; 95% CI 1.03-1.23)<sup>7</sup>. Niacin produces only a modest reduction in the white cell count<sup>9</sup>, but  
273 it suppresses monocyte activation *in vivo* via the GPR109A receptor, resulting in reduced cytokine  
274 production after lipopolysaccharide stimulation<sup>20</sup>. In animal models, niacin induces  
275 immunosuppressive T regulatory cells<sup>21</sup> and its anti-inflammatory effects may be undesirably  
276 immunosuppressive, impairing the response to infection. The effect of niacin-laropiprant on infections  
277 did not appear to attenuate with time which is relevant for patients already receiving niacin who may  
278 be considering whether to continue or not<sup>22</sup>; a previous lack of side-effects does not guarantee they  
279 will not occur in the future.

280

281 It is of note that the harmful effects detected by the randomised comparisons within HPS2-THRIVE  
282 were not detected by the expedited reporting required by regulatory authorities of any suspected  
283 unexpected serious adverse events (SUSARs) that were recorded in the trial<sup>23</sup>. This observation  
284 reinforces the point that such non-randomised pharmacovigilance data (whether from trials or from  
285 post-marketing surveillance) are typically only of value for detecting large increases in rare outcomes  
286 (such as the excess of statin-related myopathy that was seen when statin therapy was used in  
287 combination with ER niacin in China<sup>8,24</sup>) and not for detecting reliably more moderate effects on

288 relatively common serious adverse events (such as the 20-40% proportional increases in bleeding and  
289 infection in HPS2-THRIVE), despite attempts to use non-randomised observational data for such  
290 purposes <sup>25</sup>.

291

292 Niacin has been specifically recommended for patients who have the combination of low HDL-C  
293 levels and high triglyceride levels because it is partially effective at correcting this dyslipidaemia.  
294 However, niacin was not found to be effective at reducing the risk of vascular events (at least in the  
295 presence of effective statin therapy) either in the subgroup of over 4000 patients in HPS2-THRIVE  
296 with low HDL-C levels and high triglyceride, or in the AIM-HIGH population which selected  
297 specifically for such patients. Nor do the results in these trials indicate that niacin is less hazardous  
298 among such patients, or among other patients at high vascular risk who might also be encouraged to  
299 take niacin. Although niacin is no longer licensed for use in combination with statins in the US, it may  
300 still be considered for patients who do not wish to take statins. In such circumstances, the absolute  
301 LDL-C reduction with niacin is likely to be larger and this might be expected to yield a larger benefit.  
302 However, the present analyses indicate that the absolute excess of adverse effects would still be  
303 clinically significant. It may, therefore, be more prudent to consider using some other LDL-C  
304 lowering therapies (such as ezetimibe or, for patients at particularly high risk, PCSK9 inhibitors) <sup>26</sup>.

305 A limitation of these analyses is the pre-randomization run-in phase with niacin-laropiprant treatment  
306 that was intended to exclude individuals who could not tolerate it for at least two months. One-third of  
307 the people who started this run-in phase withdrew prior to randomization, with the commonest  
308 reasons being related to skin, gastrointestinal and diabetes-related adverse events <sup>8</sup>. The analyses  
309 presented here are, therefore, restricted to people able to tolerate niacin short-term and, as a  
310 consequence, the absolute excesses that were observed may underestimate the adverse effects of  
311 starting niacin in routine care. In addition, despite the addition of laropiprant it is possible that some  
312 participants still experienced flushing and were therefore aware of their treatment assignment which  
313 could have made them more likely to report adverse effects. However, as the major adverse effects

314 described here were not previously known to be associated with niacin therapy, this is not likely to  
315 bias these assessments.

316 In conclusion, a number of previously unrecognized harmful effects – namely new-onset diabetes,  
317 bleeding and infection – were observed in HPS2-THRIVE which appear to have been caused largely  
318 (if not wholly) by the niacin component of the study treatment. The absolute excesses of these and  
319 other adverse effects appeared to be similar in a wide variety of different patients. Consequently,  
320 clinicians or patients wishing to continue using niacin should consider the potential impact of these  
321 harmful effects in determining whether using niacin is worth the risk.

322

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447 **FIGURE LEGEND**

448 **Figure 1** Effects of niacin-laropiprant on new-onset diabetes, bleeding and infection by risk.  
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450 **TABLES**

451

452 Table 1: Effects of niacin-laropiprant on (A) new-onset diabetes by time; (B) Disturbances in diabetic  
453 control by description, severity and time

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	<b>Niacin-laropiprant N (%)</b>	<b>Placebo N (%)</b>	<b>Hazard ratio (95% CI)</b>	<b>Absolute excess per 1000 participants per year (95% CI)</b>
Randomized participants	12838	12835		
<b>(A) New-onset diabetes mellitus</b>				
Non-diabetic at randomization	8704	8670		
<b>Time from randomization</b>				
<1 year	193 (2.26%)	96 (1.12%)	2.02 (1.58-2.58)	11 (8, 15)
≥ 1 year	301 (1.39%)	280 (1.28%)	1.09 (0.93-1.28)	1 (-1, 3)
<b>Any new-onset diabetes</b>	494 (1.64%)	376 (1.23%)	1.33 (1.16-1.52)	4 (2, 6)
<b>(B) Disturbances in diabetic control</b>				
Diabetic at randomization	4134	4165		
<b>Description of event</b>				
Minor hyperglycaemic disturbance	361 (2.59%)	242 (1.66%)	1.56 (1.32-1.83)	9 (6, 13)
Major hyperglycaemic disturbance	42 (0.28%)	12 (0.08%)	3.56 (1.88-6.77)	2 (1, 3)
Hypoglycaemia	46 (0.31%)	31 (0.21%)	1.51 (0.96-2.38)	1 (-0, 2)
Other diabetic disturbance	44 (0.30%)	48 (0.32%)	0.93 (0.62-1.40)	-0 (-1, 1)
<b>Time from randomization</b>				
<1 year	219 (5.51%)	83 (2.03%)	2.71 (2.11-3.49)	35 (27, 43)
≥1 year	241 (2.48%)	228 (2.19%)	1.13 (0.94-1.35)	3 (-1, 7)
<b>Any diabetic disturbance</b>	460 (3.36%)	311 (2.14%)	1.56 (1.35-1.80)	12 (8, 16)

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**Table 2: Effects of niacin-laropiprant on bleeding by site, severity and time**

	Niacin-laropiprant N (%)	Placebo N (%)	Hazard ratio (95% CI)	Absolute excess per 1000 participants per year (95% CI)
Randomized participants	12838	12835		
<b>Description of event</b>				
Gastrointestinal bleeding	109 (0.24%)	71 (0.15%)	1.54 (1.14-2.08)	1 (0, 1)
Intracranial	139 (0.30%)	119 (0.26%)	1.17 (0.92-1.50)	0 (-0, 1)
Other	82 (0.18%)	49 (0.11%)	1.68 (1.18-2.39)	1 (0, 1)
<b>Severity</b>				
Non-Fatal	286 (0.63%)	198 (0.43%)	1.45 (1.21-1.74)	2 (1, 3)
Fatal	60 (0.13%)	49 (0.11%)	1.23 (0.84-1.79)	0 (-0, 1)
<b>Time from randomization</b>				
<1 year	96 (0.75%)	48 (0.38%)	2.00 (1.42-2.83)	4 (2, 6)
≥ 1 year	230 (0.70%)	190 (0.57%)	1.22 (1.01-1.48)	1 (0, 2)
<b>Any bleeding SAE</b>	<b>326 (0.71%)</b>	<b>238 (0.52%)</b>	<b>1.38 (1.17-1.63)</b>	<b>2 (1, 3)</b>

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**Table 3: Effects of niacin-laropiprant on infection by site, severity and time**

	Niacin-laropiprant N (%)	Placebo N (%)	Hazard ratio (95% CI)	Absolute excess per 1000 participants per year (95% CI)
Randomized participants	12838	12835		
<b>Description of event</b>				
Lower respiratory	547 (1.21%)	471 (1.03%)	1.17 (1.03-1.32)	2 (0, 3)
Urinary tract	114 (0.25%)	107 (0.23%)	1.07 (0.82-1.39)	0 (-0, 1)
Abdominal GI	82 (0.18%)	65 (0.14%)	1.27 (0.91-1.75)	0 (-0, 1)
Skin	67 (0.15%)	40 (0.09%)	1.68 (1.14-2.49)	1 (0, 1)
Other	306 (0.67%)	223 (0.49%)	1.38 (1.16-1.64)	2 (1, 3)
<b>Severity</b>				
Non-Fatal	994 (2.24%)	825 (1.84%)	1.22 (1.11-1.33)	4 (2, 6)
Fatal	52 (0.11%)	43 (0.09%)	1.21 (0.81-1.82)	0 (-0, 1)
<b>Time from randomization</b>				
<1 year	311 (2.47%)	233 (1.84%)	1.34 (1.13-1.59)	6 (3, 10)
≥1 year	720 (2.27%)	620 (1.93%)	1.18 (1.06-1.31)	3 (1, 6)
<b>Any infection</b>	1031 (2.32%)	853 (1.90%)	1.22 (1.11-1.34)	4 (2, 6)

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