

Assessing insomnia after stroke: A diagnostic validation of the Sleep Condition Indicator in self-reported stroke survivors

Background: Insomnia is common after stroke and is associated with poorer recovery and greater risk of subsequent strokes. Yet, no insomnia measures have been validated in English-speaking individuals affected by stroke.

Aims: This prospective diagnostic validation study investigated the discriminatory validity and optimal diagnostic cut-off of the Sleep Condition Indicator when screening for DSM-5 insomnia disorder post-stroke.

Methods: A convenience sample of 180 (60.0% women, mean age = 49.61 ± 12.41 years) community-based, adult (≥ 18 years) self-reported stroke survivors completed an online questionnaire. Diagnosis of DSM-5 insomnia disorder was based on analysis of a detailed sleep history questionnaire. Statistical analyses explored discriminant validity, convergent validity, relationships with demographic and mood variables, and internal consistency. Receiver operating characteristic (ROC) curves were plotted to assess diagnostic accuracy.

Results: Data from the sleep history questionnaire suggested that 75 participants (41.67%) met criteria for DSM-5 insomnia disorder, 33 (18.33%) exhibited symptoms of insomnia but did not meet diagnostic criteria, and 72 (40.0%) had no insomnia symptoms at the time of assessment. The SCI demonstrated 'excellent' diagnostic accuracy in the detection of insomnia post-stroke, with an AUC of 0.86 (95% CI [0.81, 0.91]). The optimal cut-off was determined as being ≤ 13 , yielding a sensitivity of 88.0%, and a specificity of 71.43%.

Conclusions: The findings of this study demonstrate the SCI to be a valid and reliable method with which to diagnose DSM-5 insomnia disorder and symptoms post-stroke. However, a lower threshold than is used in the general population may be necessary after stroke.

What is already known on this topic: Several meta-analyses have demonstrated the prevalence of insomnia after stroke to be higher than that found in the general population. However, currently no insomnia measure has been validated for use in English-speaking individuals who have had a stroke.

What this study adds: This is the first diagnostic validation of a diagnostic measure for insomnia in a sample of English-speaking stroke survivors.

How this study might affect research, practice or policy: The findings confirm the validity of the SCI post-stroke indicating a threshold of ≤ 13 to be most optimal. Use of the conventional threshold of ≤ 16 would lead to an inflated false-positive rate.

Contributorship Statement

D.M., M.G., J.E., and S.B., contributed to the conceptualisation, methodology, resources, and project administration. M.G., J.E., S.B., and M.H., additionally contributed to supervision of the project, and funding acquisition. D.M., led the visualisation, writing of the original draft, data curation, investigation, software, and formal analysis. M.M. assisted with the data curation and formal analysis. All authors contributed to the reviewing and editing of the manuscript, and M.G. is guarantor.

INTRODUCTION

A growing body of evidence indicates a bi-directional relationship between sleep and stroke [1]. Insomnia, broadly defined as difficulty with initiating, maintaining, or returning to sleep, which impacts daily functioning, is estimated to affect 32.21% to 50.4% of people affected by stroke; although estimates vary depending on diagnostic criteria and phase of recovery [2,3]. Prior to an individual's first stroke, insomnia symptoms are associated with a 30% increased risk of stroke [4]; and post-stroke, are associated with poorer functional and cognitive outcomes [5], and psychological distress [6]. Correspondingly, clinical guidelines and stroke organisations encourage post-stroke sleep assessment [7,8], to facilitate stroke management and reduce the risk of recurrence.

With respect to the assessment and diagnosis of insomnia disorder in the general population the use of clinical interviews investigating an individual's sleep history is recommended [9]. Psychometric tools commonly comprise part of this process, yet diagnostic validations of insomnia measures post-stroke have only recently been conducted. Chinese translations of the Insomnia Severity Index, Pittsburgh Sleep Quality Index, and Athens Insomnia Scale have demonstrated 'outstanding' diagnostic accuracy post-stroke [10]. However, none of these measures sufficiently capture the diagnostic criteria for insomnia disorder detailed in the Diagnostic and Statistical Manual of Mental Disorders – 5th edition (DSM-5) [11].

A valid diagnosis is dependent on specific diagnostic criteria [9], with which a scale intended for clinical use should be aligned. The Sleep Condition Indicator [12] is a short self-report questionnaire, designed to screen for insomnia disorder in accordance with DSM-5 criteria [11]. A recent validation of the Indonesian translation of the SCI [13] suggested the tool was valid for use post-stroke but recommended using a higher diagnostic threshold of ≤ 23 rather than ≤ 16 used in the general population [12,14]. However, existing validations of SCI translations do demonstrate heterogeneity in estimates of optimal cut-offs, [15–18], suggesting possible confounds from translation and/or nuanced cultural differences.

Accurate diagnoses are inextricably linked to appropriate treatment. Thus, a scale intended for clinical, or research use must be able to discriminate accurately between individuals with and without the condition of interest. Therefore, this study investigated the discriminant validity and optimal threshold of the English language SCI when detecting insomnia disorder, and symptoms, in a sample of self-reported stroke survivors. Secondary objectives involved assessing convergent validity, examining relationships between the SCI and variables relating to mood and general demographics, and assessing the internal consistency of the SCI after stroke. Additionally, we explored the discriminant validity, optimal threshold, and internal validity of the SCI-2 [19], a two-item version of the SCI comprised of items 3 and 7 from the full scale.

MATERIALS AND METHODS

The pre-registered protocol for this cross sectional prospective diagnostic validation study can be found at: <https://doi.org/10.17605/OSF.IO/4DGXW>.

Participants and Recruitment

Using voluntary response sampling via social media (Twitter/X, Instagram, Facebook, Reddit), community-based adult (≥ 18 years) stroke survivors were recruited across the UK from June 2022 to October 2023. Participants experiencing jetlag (crossed ≥ 2 time zones in the week preceding participation), undergoing treatment(s) that may affect their sleep (e.g. cancer treatment), or working nightshifts were excluded. Participants were not asked to describe whether they had experienced a haemorrhagic or ischaemic stroke, to avoid the risk of individuals self-excluding due to not knowing or remembering this information. Consent was obtained electronically, and no compensation was provided for participation. Responses were anonymous, and participants had the option of sharing their email addresses for future contact by the research team. Email addresses could not be linked to survey responses.

Ethical Favourable Opinion

A favourable ethical opinion was obtained from the University of Glasgow College of Science and Engineering Research Ethics Committee (#300210126). Optional aphasia friendly materials were available to all participants via a link on the participant information page (see supplementary material item 1).

Sample Size Estimation

An a priori power analysis was conducted using the 'MKPower' package [20]. Initial calculations were performed using parameters obtained from the existing literature. Expected prevalence was set to 32%, as estimated in Baylan et al. [2], and expected sensitivity was set to 89%, estimated by averaging the results of four existing validations of the Sleep Condition Indicator which used the conventional cut-off of ≤ 16 [12,15,21,22]. To achieve a statistical power of 0.8, with $\alpha = 0.05$, and δ (delta) = 0.1, the required sample size was 316. Following data collection from the first 109 participants, sample size estimations were re-evaluated due to stroke specific parameters being made available in Hasan et al. [13]. Expected sensitivity was set to 94% [13] and expected prevalence to 40.37% (taken from the actual prevalence in the preliminary data). With all other parameters unchanged, the estimated sample size required to achieve 80% statistical power was 176.

Measures and Materials

Participants completed an online survey comprising questionnaires pertaining to sleep, mental health, and stroke. Demographic information was also collected and comprised age, gender, time since most recent stroke, side of body affected by stroke, and total number of strokes. The index test being evaluated in this study was the SCI [12], total scores of which range from 0-32, with scores ≤ 16 considered indicative of insomnia disorder in the general population [12,14]. The validity of a two-item variant (SCI-2) [19] was also examined. Reference standard classifications of DSM-5 'Insomnia disorder', 'Insomnia symptoms', or 'no insomnia' were made following review of a

comprehensive sleep history questionnaire [23], for use in people following Acquired Brain Injury, adapted for use online by people after stroke. Index and reference tests were completed consecutively. Participants were classified as having 'insomnia disorder' if they reported difficulties with initiating or maintaining sleep, or early morning awakening, for at least three nights per week, which caused distress or impairment throughout the day, and symptoms had persisted for at least three months. Participants were classified as having 'Insomnia symptoms' if they displayed symptoms of insomnia but did not meet DSM-5 thresholds for severity, duration or frequency, or their symptoms were likely to be better explained by another sleep disorder [11].

Convergent validity was assessed by exploring correlations between the SCI and the Insomnia Severity Index (ISI) [24], a 7-item insomnia measure, scored from 0-28 with higher scores denoting greater insomnia severity. A Chinese translation of the ISI was recently validated for use after stroke [10].

Depression was measured via the Patient Health Questionnaire (PHQ-9) [25], a 9-item self-report scale scored from 0-27, where greater scores denote greater severity of depressive symptoms. Anxiety was measured via the Generalised Anxiety Disorder Scale (GAD-7) [26], a 7-item self-report measure scored from 0-21, where greater scores denote greater severity of anxious symptoms. Stroke severity was assessed via the Stroke Impact Scale Short Form (SIS-SF) [27], an 8-item self-report score scored from 8 to 40, where lower scores denote greater severity of impact from stroke.

Data were collected via an online survey platform (<https://www.qualtrics.com/>). Classifications were determined by two researchers (DM, MM) independently, blind to SCI scores, following calibration of scoring by MG on 5 responses. Percent agreement between assessor classifications was computed for the first 25 participants, with 100% agreement attained.

Statistical Analyses

Statistical analyses were conducted using R (version 4.3.1). Forced responses prevented missing values on the sleep history questionnaire and the SCI. Missing data from other scales (GAD-7, PHQ-9, SIS-SF, ISI) or demographic information were omitted in their respective analyses. As surveys could be completed across multiple sessions, mean durations greater than 1.5x the interquartile range ($n = 18$) were omitted when calculating summary statistics to curtail the impact of outlier effects.

Normality was assessed visually and with Anderson-Darling tests. Descriptive analysis of demographic and psychometric variables was conducted across classifications. Kruskal-Wallis tests were used for group comparisons of continuous variables, and Dunn's tests with Bonferroni corrections used for pairwise comparisons. Group differences in categorical data were explored using Fisher's exact test. Statistical significance was set at $p < 0.05$. Spearman's correlations were used to quantify relationships between the SCI and continuous variables, and confidence intervals were computed via bootstrapping ($n = 1000$). The internal consistency of all psychometric scales was assessed via Cronbach's α [28], and the commonly assumed threshold of ≥ 0.70 adopted to indicate acceptable reliability [29].

Receiver operating characteristic (ROC) curves were plotted using the 'pROC' package [30] to assess the SCI's diagnostic accuracy. Youden's J statistic [31] was calculated to determine the optimal diagnostic cut-off. The area under the curve (AUC) provided an estimate of overall accuracy. AUC values of 0.7 to 0.8 were considered 'acceptable', 0.8 to 0.9 'excellent', and > 0.9 'outstanding' diagnostic accuracy [32]. Confidence intervals were computed via bootstrapping ($n = 2000$).

RESULTS

Participant Characteristics

A total of 544 potential participants accessed the study platform (see Figure 1), of which 246 exited before providing consent, and 118 did not complete the SCI and/or sleep history. The remaining 180 provided consent and all met inclusion criteria. The mean survey duration was 24.48 minutes (SD = 10.56). Participant characteristics are summarised in Table 1.

After classification based on the sleep history, 75 (41.67%) participants met criteria for insomnia disorder, 33 (18.33%) exhibited insomnia symptoms without meeting diagnostic criteria, and 72 (40.0%) had no insomnia symptoms. A breakdown of participants by classification is available in Figure 1 (a confusion matrix of index versus reference standard classifications can be found in the supplementary material 2).

[Figure 1: Participant Flowchart].

Table 1: Participant Characteristics

	Total (n = 180)	Insomnia Disorder (n = 75)	Insomnia Symptoms (n = 33)	No Insomnia (n = 72)	
Variable	Mean (SD)	Mean (SD)	Mean (SD)	Mean (SD)	p value
Age	49.61 (12.41)	49.24 (12.04)	51.82 (11.56)	48.99 (13.18)	.689
Depression (PHQ-9)	10.80 (7.31)	12.97 (7.36) *	11.48 (7.53)	8.20 (6.35) *	<.001
Anxiety (GAD-7)	7.92 (6.19)	9.83 (6.55) *	7.06 (5.94)	6.31 (5.40) *	.004
Stroke Severity (SIS-SF)	30.02 (6.40)	29.00 (5.99) *	28.97 (6.70)	31.60 (6.46) *	.018
Time Since Stroke (Years)	4.16 (4.26)	3.81 (3.55)	3.02 (3.72)	5.09 (5.02)	.069
SCI	13.74 (7.51)	8.47 (4.60) **	13.52 (5.33) +°	19.35 (6.79) **	<.001
ISI	13.40 (6.59)	17.68 (5.59) **°	13.76 (4.31) °+	8.78 (5.25) **	<.001
Variable	n (%)	n (%)	n (%)	n (%)	p value
Gender					.087
Female	108 (60.00)	52 (69.33)	15 (45.45)	41 (56.94)	
Male	70 (38.89)	22 (29.33)	18 (54.55)	30 (41.67)	
Non-binary	2 (1.11)	1 (1.33)	0 (0)	1 (1.39)	
Number of Strokes					.336
1	152 (84.44)	62 (82.67)	32 (96.97)	58 (80.56)	
2	22 (12.22)	9 (12.00)	1 (3.03)	12 (16.67)	
3	2 (1.11)	1 (1.33)	0 (0)	1 (1.39)	
> 3	4 (2.22)	3 (4.00)	0 (0)	1 (1.39)	
Affected Side					.994
Don't remember	20 (11.11)	8 (10.67)	3 (9.09)	9 (12.50)	
Left	84 (46.67)	35 (46.67)	16 (48.48)	33 (45.83)	
Right	76 (42.22)	32 (42.67)	14 (42.42)	30 (41.67)	

n Number of participants. SCI Sleep Condition Indicator. ISI Insomnia Severity Index. SD Standard Deviation. Statistically significant differences exist between groups that are marked by the same symbol on one row. One participant was excluded from SIS-SF, PHQ-9 and GAD-7 analyses, and another from SIS-SF analyses, due to incomplete data. Both were classified as 'No Insomnia'.

The distribution of total SCI scores by classification is displayed in Figure 2. Skewness and kurtosis were 0.40 and 2.46, respectively. Visual inspection and Anderson-Darling tests ($p = .004$) indicated non-normal distribution of total SCI scores.

[Figure 2: Violin and boxplot graphs displaying the distribution of total SCI scores by classification]

Sensitivity and Specificity of the SCI

When detecting insomnia disorder post-stroke using the SCI, the AUC was 0.86 (95% CI [0.81, 0.91]; Figure 3), with an optimal cut-off of ≤ 13 , yielding sensitivity, specificity, and Youden's J values of 88.0%, 71.43%, and 0.59, respectively (see Table 2). When detecting insomnia symptoms, the AUC was 0.85 (95% CI [0.80, 0.91]; Figure 3), with an optimal cut-off of ≤ 14 , yielding sensitivity, specificity, and Youden's J values of 81.48%, 76.39% and 0.58, respectively (see Table 2). Both AUC curves are indicative of 'excellent' diagnostic accuracy [32]. Delong's test for two paired ROC curves found no statistical difference ($p = 0.91$) in the SCI's ability to detect insomnia disorder vs symptoms. Using the sensitivity observed when detecting insomnia disorder, with all other parameters unchanged, returns a power of 0.71.

[Figure 3: ROC curves displaying the accuracy of the SCI when detecting DSM-5 insomnia disorder and symptoms, including AUC values and 95% confidence intervals.]

Table 2: Sensitivity, specificity, and Youden's J statistic for select thresholds on the SCI when detecting DSM-5 insomnia disorder and insomnia symptoms.

Insomnia Disorder				Insomnia Disorder and Symptoms			
Threshold	Specificity (%)	Sensitivity (%)	Youden's J	Threshold	Specificity (%)	Sensitivity (%)	Youden's J
8	89.52	52.00	0.42	8	91.67	40.74	0.32
9	85.71	57.33	0.43	9	88.89	46.30	0.35
10	81.90	70.67	0.53	10	86.11	57.41	0.44
11	80.95	73.33	0.54	11	86.11	60.19	0.46
12	78.10	78.67	0.57	12	83.33	64.81	0.48
13	71.43	88.00	0.59	13	80.56	75.93	0.56
14	64.76	90.67	0.55	14	76.39	81.48	0.58
15	60.00	90.67	0.51	15	73.61	84.26	0.58
16	54.29	96.00	0.50	16	66.67	88.89	0.56
17	48.57	98.67	0.47	17	59.72	91.67	0.51

Values for all thresholds can be found in supplementary material 3. Optimal thresholds are displayed in bold. Youden's J - ($J = \text{sensitivity} + \text{specificity} - 1$)

Sensitivity and Specificity of the SCI-2

ROC analysis was performed on the summed totals for items 3 and 7 of the SCI (Figure 4), forming the SCI-2 [19]. When detecting DSM-5 insomnia in stroke survivors, the AUC for the SCI-2 was 0.79 (95% CI [0.73, 0.85]), with an optimal cut-off of ≤ 2 . Corresponding with a specificity of 68.57%, sensitivity of 76.0%, and a Youden's J of 0.45. Delong's test provided statistical evidence ($p < .001$) that the SCI outperformed the SCI-2. Sensitivity and specificity values at all thresholds can be found in supplementary material 4. The positive and negative predictive values of the SCI in this sample, using the optimal threshold of ≤ 13 , were 0.69 and 0.89, respectively. Predictive values at a range of hypothetical prevalence rates can be found in supplementary material 5.

[Figure 4: ROC curves displaying the accuracy of the SCI-2 when detecting DSM-5 insomnia disorder and symptoms, including AUC values and 95% confidence intervals.]

Convergent Validity

Spearman's rank correlation was computed to assess the relationship between total scores on the ISI and SCI, with ties handled by average ranks. Results revealed a large negative correlation between the two scales ($r(178) = -0.86$, 95% CI [-0.91, -0.80]; Figure 5).

[Figure 5: Scatter plot visualising the relationship between total SCI and ISI scores.]

Associations between the SCI, demographic, stroke, and mood variables.

Spearman's rho was computed to determine the strength of the relationship between total SCI scores and all other continuous variables. Results are presented in Table 3.

Table 3: *Spearman's Correlation Coefficients, Confidence Intervals, and P values for Relationships Between the SCI and Select Demographic and Psychometric Variables*

Variable	<i>r</i>	95% CI	<i>p</i> value
Age	-0.07	(-0.20, 0.08)	.357
PHQ-9	-0.63	(-0.72, -0.52)	<.001
GAD-7	-0.47	(-0.57, -0.33)	<.001
SIS-SF	0.47	(0.34, 0.57)	<.001
Time Since Stroke	0.03	(-0.12, 0.18)	.713

CI – Confidence Interval

Internal Validity

The reliability of all included psychometric scales was assessed via Cronbach's α , and all deemed acceptable. For the SCI (n = 180), Cronbach's α was 0.84 (95% CI [0.80, 0.87]), for the SCI-2 (n = 180) α was 0.76 (95% CI [0.69, 0.82]), for the GAD-7 (n = 179) α was 0.91 (95% CI [0.89, 0.93]), for the PHQ-9 (n = 179), α was 0.89 (95% CI [0.87, 0.92]), and for the SIS-SF (n = 178) α was 0.83 (95% CI [0.78, 0.86]).

DISCUSSION

This is the first study examining the SCI's validity and reliability, and that of any insomnia screening tool, in a sample of English-speaking stroke survivors. Results from 180 participants indicated that 41.7% met DSM-5 criteria for insomnia disorder, and 60.0% exhibited insomnia symptoms (including those classified as insomnia disorder), consistent with meta-analytic findings on post-stroke insomnia prevalence (2,3), and supporting the representativeness of the sample. The remaining 40% had no insomnia symptoms.

The SCI demonstrated ‘excellent’ diagnostic accuracy when screening for DSM-5 insomnia disorder (AUC = 0.86) and/or symptoms (AUC = 0.85) post-stroke. Moreover, examination of internal consistency indicated acceptable reliability (Cronbach’s α = 0.84). The SCI-2 demonstrated ‘acceptable’ accuracy, with an optimal cut-off of ≤ 2 ; paralleling validations in the general population [19]. However, the full SCI demonstrated significantly higher accuracy, and should be the preferred assessment method where possible.

Estimates of diagnostic accuracy in the current study are similar, albeit slightly lower, to recent validations in stroke populations of the SCI in Indonesian [13] and the Insomnia Severity Index, Pittsburgh Sleep Quality Index, and Athens Insomnia Index in Chinese [10]. Nonetheless, comparisons should be drawn cautiously. Existing validations of SCI translations demonstrate heterogeneity in estimates of optimal cut-offs [15–18], suggesting possible confounds from translation and/or nuanced cultural differences. Additionally, a recent meta-analysis demonstrated that condition prevalence may influence the sensitivity and specificity of binary diagnostic tests [33]. The prevalence of insomnia in the previous validations [10,13] was 15.6% and 31.0%, respectively, both being lower than the current study (41.7%) and previous estimates of post-stroke insomnia prevalence [2,3]. Thus, the variance in prevalence should be considered when comparing studies of diagnostic accuracy, and when considering a screening tool for clinical or research use.

The optimal cut-off for detecting insomnia disorder using the SCI in the current study was found to be ≤ 13 . This is lower than both the conventional cut-off for the general population of ≤ 16 [12,14], and the optimal cut-off of ≤ 23 observed in the validation of the Indonesian translation of the SCI post-stroke [13]. Were the conventional cut-off adopted in the current sample, 96% of positive insomnia disorder cases would be correctly identified. However, approximately 45.7% of participants without insomnia disorder would be incorrectly classified as a positive case (i.e., having insomnia disorder). Depending on the intended use for the scale, there may be instances where clinicians or researchers wish to maximise sensitivity at the expense of specificity (or vice versa) but should do so

with caution. Use of an erroneously high threshold on the SCI will lead to an increased false positive rate, potentially leading to inappropriate treatment and use of resources in clinical settings; and may cast doubt on the validity of findings in research.

Understanding the heterogeneity of optimal diagnostic thresholds is important for ensuring the SCI is appropriately used in different clinical samples, and future research should explore this further. One possibility is that poorer sleep post-stroke [34] may necessitate a lower threshold. Mean SCI scores in the current study ($\bar{x} = 13.74$, $SD = 7.51$) are slightly lower, and display greater variance than results from a retrospective exploration of scores in 200,000 adults who completed the SCI ($\bar{x} = 14.97$, $SD = 5.93$; [14]). However, the results of the latter [14] may be biased towards people with sleep impairment, and the true population mean SCI scores are likely higher. Variation in reference standards likely also contribute to variability in optimal cut-offs, and ideally studies seeking to validate measures of insomnia should select reference standards in accordance with guidelines on the assessment and diagnosis of insomnia disorder [9]. In studies of diagnostic accuracy, the reference standard assumes infallibility. Yet, the index test may not always be solely responsible for the discrepancy [35]. The conventional cut-off of ≤ 16 [12,14] was determined as the minimum score necessary for putative DSM-5 insomnia disorder, with convergent validity assessed against the ISI. The ISI boasts commendable accuracy but standalone, is not a recommended tool for diagnosing DSM-5 insomnia disorder [9], and is less specific than clinical interviews. Thus, when comparing the optimal diagnostic cut-offs between studies, one should be cognisant of the possibility of differences in reference standards explaining some or all the variation in optimal cut-offs.

Spectrum bias, a phenomenon whereby accuracy of a diagnostic test may be confounded by symptom severity [36], may further explain discrepancies in optimal diagnostic thresholds between studies. Scores closer to the cut-off are more prone to misclassification, therefore, studies where index test scores demonstrate a bimodal distribution have a greater risk of spectrum bias. Hasan et al. [13] excluded 164 participants with sleep or psychiatric disorders other than insomnia in their

validation. Many of these participants would likely have scored lower on the SCI regardless of insomnia classification, as suggested by the relationships between SCI score and measures of mental health in the current study and corroborated by meta-analytic results from Hertenstein et al. [37]. Likely as a result of their inclusion criteria, mean SCI scores in Hasan et al. [13] for participants without insomnia were 29.55 (SD = 1.96), with little variance. Compared to 19.33 (SD = 6.81) in the current study, which is more akin to previous research [34]. Such a skewed distribution of scores allows specificity to remain high at higher thresholds, contributing to potentially misleading estimates of accuracy and optimal cut-offs. Hasan and colleagues demonstrate that the Indonesian SCI can discriminate post-stroke insomnia disorder in the absence of potential confounds. The present study adds to this by validating the English SCI in self-reported stroke survivors and doing so whilst including participants with comorbid psychological or sleep disorders; thereby increasing generalisability of the findings. Results of the current study, when considered in the context of the existing literature, highlight the importance of considering spectrum bias in studies of diagnostic accuracy. When considering a screening tool, researchers and clinicians should consider whether a study's selection criteria are clinically meaningful to the population they intend to use the test in.

This study has several strengths. A priori power analyses, a rarity in studies of diagnostic accuracy [38], ensured the study was suitably powered to obtain reliable estimates of accuracy. Consecutive completion of index and reference tests reduced the risk of temporal changes in disease state. Finally, stratifying analyses by insomnia classification reduced the risk of spectrum bias and prevalence influencing estimates of accuracy [33].

Nonetheless, voluntary response sampling does carry a risk of response bias and may lead to over-representation of participants with an interest in sleep, or more severe sleep difficulties. However, prevalence in the current sample aligns with that of previous meta-analytic estimates [2,3]. Recruiting participants via social media may have led to a younger than expected stroke sample (mean age = 49.61 ± 12.41 years). Future research may wish to consult older stroke survivors to

explore methods of increasing the accessibility of online stroke research to a wider age range. Nevertheless, stroke can affect people of any age, and research representative of younger stroke survivors is necessary, particularly to understand the challenges of improving life after stroke for working age individuals - perhaps with young families - compared to older adults. Results of a linear regression from a large cross-sectional study ($n = 200,000$) by Espie and colleagues [14] did demonstrate that SCI score was predicted to decrease 0.057 points per year of life. Findings from the current study demonstrated no significant relationship between age, or gender, and total SCI score, however, it is likely that a much larger sample size would be needed to detect an effect as small as that described in Espie et al. [14]. Thus, although the modulating effect of demographic characteristics on the accuracy of the SCI was not explored, current findings provide no evidence to suggest that such an effect exists. Due to the online nature of recruitment, participant self-reports of stroke were not verified by another means. Various efforts were made however, to minimise the risk of individuals without a history of stroke volunteering to take part. Firstly, recruitment adverts were shared explicitly by stroke researchers and third sector organisations supporting stroke survivors, targeting stroke peer support groups. Moreover, offering no incentive to participants removed the risk of individuals taking part without a history of stroke for financial gain. Furthermore, while the results of a systematic review by Woodfield et al. [39] demonstrate that the sensitivity and positive predictive value of self-reported stroke is low in samples with low prevalence, specificity (96 to 99.1%) and negative predictive value (88.2 to 99.9%) were consistently high. As one would expect, authors report that positive predictive values increased in line with prevalence. Thus, in samples with high prevalence of stroke, such as in the current study, the risk of false positives is likely to be small. We therefore believe self-reports of stroke in the present sample to be reliable. Nevertheless, where time and resources allow, future researchers may wish to verify self-reported stroke status by recruiting participants from stroke outpatient clinics or obtaining clinical records. Moreover, the number of incomplete responses ($n = 118$) is worth noting. This may indicate systemic challenges related to conducting stroke research online, or difficulties faced by potential participants with

language difficulties. Relatedly, the duration of the survey was relatively long, and future research may consider limiting survey length further to increase accessibility. Finally, participants were not asked which type of stroke they had. This was a methodological decision to avoid excluding participants who did not know or remember. Existing evidence does indicate potential relationships between insomnia and stroke type/location [40], therefore future researchers may wish to consider purposive sampling methods that allow exploration of whether these phenomena may influence insomnia presentation and the accuracy of diagnostic measures.

In summary, this study confirms the validity and reliability of the English version of the SCI when detecting DSM-5 insomnia in self-reported stroke survivors and indicates that a lower threshold than is traditionally used should be considered when using the SCI in the stroke population. At the optimal diagnostic threshold of ≤ 13 , the SCI demonstrates 'excellent' diagnostic accuracy. The brevity of the tool makes it attractive for clinical and research use and the SCI-2 does offer an acceptable shorter alternative. Nevertheless, the accuracy of the SCI-2 is outperformed by the full SCI, and use of the latter should be preferred where possible. In conclusion, the SCI should be considered a valid and reliable screening tool for DSM-5 insomnia disorder and insomnia symptoms post-stroke.

DECLARATION OF CONFLICTING INTERESTS

The authors declare that there are no conflicts of interest.

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