

Does exercise and advice for subacute low back pain work via changes in depressive symptoms? A causal mediation analysis

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Introduction

Acute low back pain is common, but patients often experience rapid improvements in pain, disability and return to work within the first six weeks. During the subacute phase (six weeks – three months since onset), further improvement is apparent, resulting in 52% of patients being pain-free, and 71% being disability-free at three months after onset. However, if symptoms persist beyond the subacute period, patients typically do not experience much improvement. At one year after onset, only 57% are pain-free (Pengel, Herbert, Maher & Refshauge, 2003) and 75% are disability-free (Henschke et al., 2008). Back pain that progresses to chronicity can lead to profound individual and population burden (Hoy et al., 2014; Hartvigsen, Hancock, Kongsted et al., 2018).

According to recent UK and US clinical practice guidelines for low back pain treatment (NICE guidelines 2016; Qaseem, Wilt, McLean, et al. 2017), a shift in thinking about the primary care management of low back pain is encouraged, with the focus on nondrug treatment options and a stratified management approach according to scores from prognostic screening questionnaires (NICE guidelines 2016). For example, all patients with nonspecific low back pain should be offered information on the nature of low back pain, reassurance about the likely low risk of serious underlying disease and advice on evidence-based self-management (Traeger, Buchbinder, Harris & Maher, 2017). Following, patients with a high risk of poor outcome in the acute stage could benefit from specific nonpharmacologic treatments such as the advice to stay active, exercise, or psychologically informed physiotherapy (Traeger, Buchbinder, Harris & Maher, 2017; Foster, Anema, Cherkin et al., 2018). However, recommendations are based on low to moderate quality of evidence with small effect sizes and working mechanisms of these treatments are unknown.

Pengel et al. (2007) conducted a factorial randomized, placebo-controlled trial to estimate the effect of physiotherapist-directed exercise and/or advice intervention to reduce pain and improve function in patients with subacute low back pain. The trial concluded that a clinically worthwhile improvement of pain at 6 weeks occurred in some participants that received the combined advice and exercise intervention.

Although the effectiveness of this intervention is clear, there is limited evidence for the underlying mechanisms of this intervention. That is, we do not understand *how* exercise and advice reduces pain and disability in patients with subacute low back pain. It is important to understand treatment mechanisms to guide the adaptation and implementation of this effective intervention (Lee et al., 2016; Lee & Lamb, 2017).

The presence of psychological factors, such as depressive symptoms, in people with subacute low back pain is associated with an increased risk of poor outcome (Hartvigsen et al., 2018; Pinheiro et al., 2016). This suggests that depressive symptoms could be an important treatment target for the prevention of chronic back pain. In other words, reducing depressive symptoms could lead to reduced pain and increased function in patients with subacute low back pain. The intervention tested by Pengel et al. (2007) included two strategies that could have an effect on depressive symptoms (exercise and advice). That is, exercise could reduce depressive symptoms via the modulation of neurobiological and physiological mechanisms (Cooney et al., 2013); and advice could also reduce depressive symptoms via reassurance and the correction of unhelpful beliefs regarding low back pain. The aim of the study was to estimate the extent to which depressive symptoms mediate the effect of physiotherapy treatment (i.e. exercise and advice) on pain and disability in a population with subacute low back pain.

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85 **Methods**

86 This study is a secondary analysis of a factorial randomized controlled trial. The methods of
87 the trial are described elsewhere (Pengel et al., 2007). The original study protocol was
88 approved by the Institutional Review Board of the University of Sydney and relevant area
89 health services, and was registered in the Australian Clinical Trials Registry (number
90 12605000039684). We registered and locked a pre-specified analysis protocol for the current
91 study on the “Open Science Framework” prior to data analysis (XXX, 2017). Our a priori
92 protocol specified disability as the primary outcome that we later changed to pain as primary
93 outcome and disability as secondary outcome to bring this analysis in line with the original
94 trial.

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96 ***Participants***

97 Participants were recruited via advertisements in newspapers, invitations to hospital waiting
98 lists for physiotherapy treatment of low back pain, or by direct referral by a healthcare
99 professional. Persons between 18 and 80 years of age with nonspecific low back pain lasting
100 for at least six weeks, but no longer than twelve weeks were included. Sixteen
101 physiotherapists from seven clinics in Australia and New Zealand provided the intervention
102 and control sessions.

103

104 ***Intervention***

105 After providing informed consent, participants were allocated to one of four groups: exercise
106 and advice; exercise and placebo advice; placebo exercise and advice; or placebo exercise and
107 placebo advice. Participants received twelve sessions over six weeks. During three of the
108 twelve sessions, participants received advice or placebo advice. The exercise intervention

consisted of an individualized program including supervised physical therapy incorporating principles of cognitive behavioral therapy as well as home exercises, tailored to the specific functional problems of each patient. Placebo exercise involved deactivated pulsed ultrasonography and deactivated short-wave diathermy. Advice sessions aimed to encourage a graded return to normal activities in which therapists explained the benign nature of low back pain, addressed unhelpful beliefs, and emphasized the importance of staying active. In the placebo advice session, the physiotherapist listened empathically to the participant, but did not give advice for their back pain (Pengel et al., 2007).

Measures

Baseline assessment captured demographic, social, and pain characteristics of the sample, as well as baseline values of the outcome measures, mediator and confounders. Assessment of the mediator and outcomes was repeated at six weeks and three months.

Mediator: Depressive symptoms were measured with the depression subscale of the Depression, Anxiety, Stress Scale-21 (DASS-21) (Lovibond & Lovibond, 1995). The DASS-21 measures the negative emotional states of depression, anxiety, and stress. The depression subscale includes 7 questions about dysphoria, hopelessness, devaluation of life, self-deprecation, lack of interest/involvement, anhedonia, and inertia. Subjects report the extent to which they have experienced these states in the past week on a 4-point scale (0-3). Score is multiplied by 2, so that total score range is 0-42. We used the baseline and six weeks score in the analysis. The scale is stable over time (Brown, Chorpita, Korotitscw & Barlow, 1997) and responsive to treatments directed at mood problems (Ng et al., 2007).

Outcome measures: Outcomes were pain and disability. Average pain over the past week was measured on a numerical rating scale from 0 (no pain), to 10 (worst pain possible). Disability was measured using the Roland-Morris Disability Questionnaire (RMDQ) (Roland

& Morris, 1983). The RMDQ consists of 24 items describing difficulties people encounter when experiencing back pain. Subjects indicated which of these difficulties they currently encountered (scale range 0-24). We used the baseline and three month scores in the analysis.

Potential confounders: We used a structural approach to identify potential confounders for the mediation-outcome effect (Shrier & Platt, 2008). The confounders were selected based on theoretical and empirical evidence including the fear-avoidance model, which specifies that catastrophizing thoughts, fear-avoidance beliefs, and depression play a role in the development of its associated disability and persistence of chronic pain (Vlaeyen & Crombez, 1999; Vlaeyen & Linton, 2012). Selected confounders were: age, gender, and baseline assessments of kinesiophobia, pain self-efficacy, and pain catastrophizing. Baseline assessments of the mediator and outcomes (depression, pain and disability) were also included in the models. Kinesiophobia was measured using the Tampa Scale of Kinesiophobia (TSK) (Miller, Kori & Todd, 1991). Scores range from 17 (low fear of movement) to 68 (high fear of movement). Pain self-efficacy was measured with the Pain Self-Efficacy Questionnaire (PSEQ) (Nicholas, 2007). Scores range 0 (low self-efficacy) to 60 (high self-efficacy). Pain catastrophizing was measured with the Pain-Related Self-Statements Scale, catastrophizing subscale (Flor, Behle & Birbaumer, 1993). Scores range 0 (low) to 45 (high). Psychometric properties of these measures are within acceptable standards (Weermeijer & Meulders, 2018; Asghari & Nicholas 2009; Flor, Behle & Birbaumer, 1993)

Statistical analysis

Descriptive statistics

We present descriptive statistics for subjects who completed the baseline and one or both of the follow-up questionnaires. Baseline characteristics and outcome values with a normal distribution were presented as mean with their standard deviation (SD), non-normally

distributed data were presented as medians with 25th-75th percentiles (interquartile range). We excluded subjects who were lost to follow-up at both six weeks and three months.

Causal mediation analysis

Mediation analyses were performed for complete cases. We conducted a causal mediation analysis under the counterfactual framework (Imai, Keele & Tingley, 2010). We estimated the average causal mediation effect (ACME), average direct effect (ADE), and the total effect. The ACME is the effect of the exposure (group allocation) on the outcome (pain or disability) that is mediated by the hypothesized mediator (depressive symptoms). The effect of the exposure on the outcome that is not mediated through the selected mediator (depression) is represented by the ADE. Thus, the sum of the ACME and ADE equals the total effect. The proportion of the total effect that is channeled through the mediator (ACME) is expressed as the proportion mediated (Imai, Keele & Tingley, 2010). We estimated unstandardized point estimates of the ACME, ADE and total effect and their 95% confidence intervals using 1000 bootstrapped stimulations.

We constructed six single mediator models to quantify the ACME, ADE and total effect for each model. These models are outlined in Appendix 1. For each outcome (pain and disability), we specified three models with the exposure defined as: 1: exercise and advice versus placebo; 2: exercise versus placebo; 3: advice versus placebo. If the confidence intervals surrounding the point estimate for ACME included 0, we aimed to identify where the causal path broke down. We did this to understand why the intervention did not work through depressive symptoms. To do so, we fit two linear regression models: the mediator model (path a) and the outcome model (path b). We constructed the mediator model with exposure as the independent variable and the mediator as the dependent variable. The outcome model was specified with the exposure, mediator, exposure-mediator interaction term, and selected

confounders as independent variables, and the outcome as the dependent variable (VanderWeele 2015). Figure 1 provides an overview of all causal models.

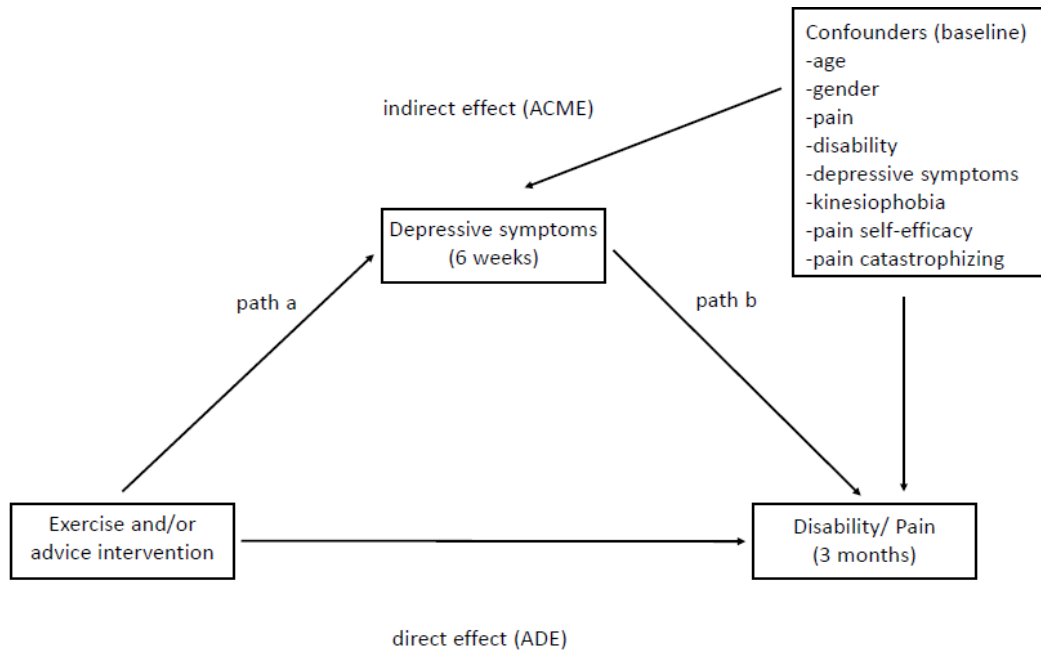


Figure 1: Mediator model. Hypothesized causal mechanisms with outcome pain and disability. The indirect effect, or ACME, is the effect of the exposure (group allocation) on pain or disability, mediated by depressive symptoms. The direct effect, or ADE, is the remaining effect of the exposure on pain or disability that is not mediated through depressive symptoms. Path a is the effect of the exposure on depressive symptoms. Path b is the effect of depressive symptoms on pain or disability.

Sensitivity analyses for sequential ignorability

Because the exposure was randomly allocated, one can assume that the exposure-mediator and exposure-outcome effects are unconfounded. However, it is possible that the mediator-outcome effect could be confounded by unknown/unmeasured confounders (Imai, Keele & Tingley, 2010). Therefore, we conducted sensitivity analyses to assess the robustness of the ACME to the effects of unknown/unmeasured confounders. The level of confounding due to unknown/unmeasured confounders is represented by the correlation between the residuals from the mediator and outcome regression models (ρ). We explored how varying levels of ρ between -1 and 1 influenced the ACME by plotting the results. Analyses were performed

using the “mediation” package in R (The R Foundation for Statistical Computing) (Tingley, Hirose, Keele & Imai, 2014).

Results

Participant characteristics

After excluding cases lost to follow-up at both six weeks and three months, 240 participants remained from 259 participants in the original trial. In this sample, 60 received exercise and advice, 60 received exercise and placebo advice, 59 received placebo exercise and advice, and 61 received placebo exercise and placebo advice. In total, 120 received exercise, and 119 received advice (figure 2). Average age was 50.5 (SD 15.6) years, and 52% were male. Co-occurrence of leg pain was reported by 31.7%. Previous episode of back pain was experienced by 66.1%. Current regular exercise was practiced by 54.4%. Mean pain score at baseline was 5.4 (SD 2.0). Mean disability score at baseline was 8.4 (SD 4.9). Median depression score at baseline was 4 (IQR 2-12) (table 1).

According to the DASS cut-off scores by Lovibond & Lovibond (1995), 69.6% scored no depressive symptoms (score < 10), 9.6% had mild (score 10-13), 12.5% had moderate (14-20), and 8.3% had severe or extremely severe (21+) depressive symptoms. At 6 weeks, median depression score was 0 (0-6); 83.6 had no depressive symptoms, 4.7% had mild, 6.5% had moderate, and 5.2% had severe or extremely severe depressive symptoms. (table 2)

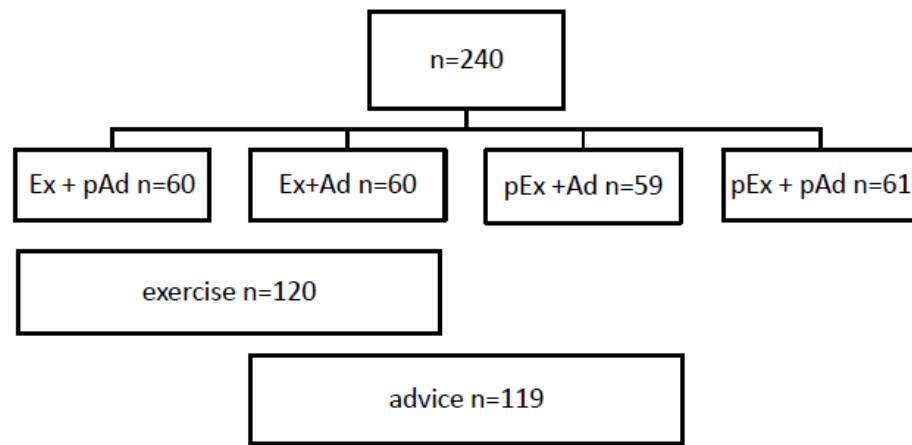


Figure 2. Participant flow chart. Ex: exercise, Ad: advice, pEx: placebo exercise, pAd: placebo advice

235 Table 1: Baseline characteristics

	Exercise only	Advice only	Exercise+Advice	Placebo exercise+ Placebo advice
No. of participants	120	119	60	61
Age (years)	49.7 (SD 15.9)	51.1 (SD 15.5)	50.5 (SD 15.7)	51.1 (SD 15.4)
Gender, male	63 (52.5%)	65 (54.6%)	32 (53.3%)	28 (45.9%)
Working before low back pain (yes)	70 (58.3%)	69 (58%)	37 (61.7%)	30 (49.2%)
Currently working	62 (51.7%)	62 (52.1%)	33 (55%)	26 (42.6%)
Currently smoking	18 (15%)	18 (15.1%)	11 (18.3%)	14 (23%)
Currently performs regular exercise	67 (55.8%)	66 (56.4%)	32 (53.3%)	28 (46.7%)
Previous episodes of low back pain	78 (65%)	83 (70.3%)	42 (70%)	39 (63.9%)
Pain referred to the leg	35 (29.2%)	40 (33.6%)	17 (28.3%)	18 (29.5%)
Pain in other areas (neck, shoulder)	27 (22.5%)	25 (21%)	11 (18.3%)	8 (13.1%)
Low back pain treatment (<6 wks)	65 (54.2%)	63 (52.9%)	35 (58.3%)	29 (47.5%)
Taking pain medication (<6 wks)	65 (54.2%)	70 (59.8%)	33 (56.9%)	39 (63%)
Pain severity last week	5.3 (SD 2.1)	5.4 (SD 2.1)	5.4 (SD 2.2)	5.4 (SD 1.8)
RMDQ score	8.6 (SD 4.9)	8.6 (SD 4.6)	8.9 (SD 4.8)	8.2 (SD 5.5)
DASS-Depression	4 (IQR 0-10)	4 (IQR 0-10)	4 (IQR 0-8)	4 (IQR 2-14)
DASS-Anxiety	2 (IQR 0-8)	2 (IQR 0-8)	2 (IQR 0-6)	2 (IQR 0-8)
DASS-Stress	10 (IQR 4-18)	10 (IQR 4-16)	8 (IQR 4-12)	10 (IQR 2-18)
TSK score	39.1 (SD 8.4)	39.1(SD 7.8)	38.9 (SD 8.0)	38.5 (SD 8.2)
PSE score	44.5 (SD 11.9)	45.3 (SD 11.7)	44.7 (SD 12.4)	43.3(SD 13.7)
PRSS catastrophizing	17.5 (SD 8.8)	17.6 (SD 9.8)	17.2 (SD 9.2)	18.5 (SD 7.7)

236 Variables are presented as n (%), mean(standard deviation or SD), or median with interquartile range, IQR).

237 RMDQ: Roland Morris Disability Questionnaire; DASS: Depression, Anxiety, Stress Scales-21; TSK: Tampa

238 Scale for Kinesiophobia (TSK); PSE: Pain Self-Efficacy questionnaire; PRSS: Pain-Related Self-Statements

239 Scale.

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Table 2: Depressive symptoms score (DASS)

Depressive symptoms	baseline	6 weeks
None	69.6%	83.6%
Mild	9.6%	4.7%
Moderate	12.5%	6.5%
Severe/ extremely	8.3%	5.2%

Mediation analyses

Pain outcome (models 1, 2 & 3)

From a sample of 225 complete cases for the pain outcome, we found that depressive symptoms did not mediate the effect of exercise plus advice on pain (ACME -0.05 [95%CI -0.24 to 0.15]). Most of the total effect (-1.23 [-1.94 to -0.49]) was mediated through unknown mechanisms (ADE -1.19 [-1.85 to -0.48]). For exercise alone and advice alone, there was also no significant mediating effect of depressive symptoms: exercise (ACME -0.09 [-0.29 to 0.05]), advice (ACME 0.005 [-0.17 to 0.19]). For exercise alone, the ADE (-0.50 [-1.09 to 0.12]) and total effect (-0.60 [-1.21 to 0.06]) were not significant. For advice alone, the ADE (-0.82 [-1.39 to -0.21]) and total effect (-0.81 [-1.42 to -0.20]) were not significant. See figure 3.

None of the exposures (combined exercise and advice, exercise alone, or advice alone) compared to placebo, had an effect on depressive symptoms (path a). Whereas the association between change in depressive symptoms and pain (path b) was significant for exercise combined with advice (regression coefficient = 0.06 [SE 0.03]) and advice alone (regression coefficient = 0.07 [SE 0.04]). For exercise alone, the association was of similar magnitude but not significant (regression coefficient = 0.06 [SE 0.04]).

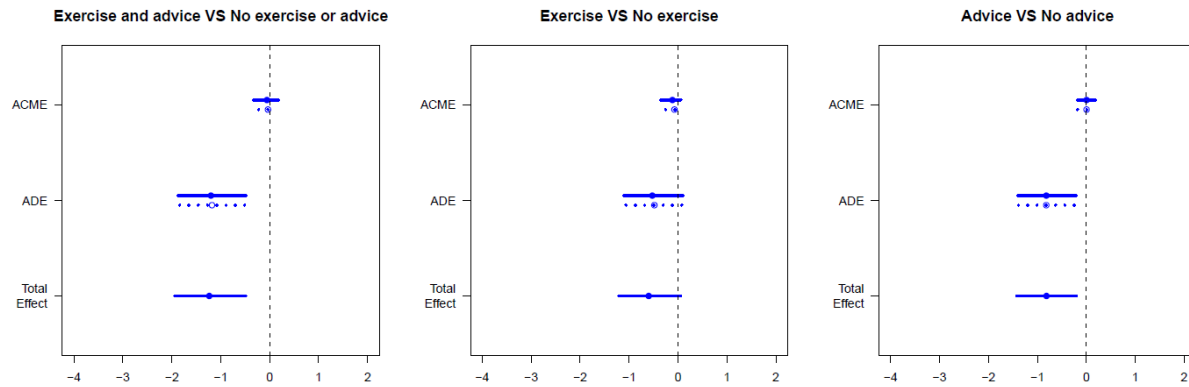


Figure 3. Effect size including 95% confidence interval of average causal mediation effect (ACME), average direct effect (ADE) and total effect for the model with the outcome pain and exposure advice and exercise, exercise alone, and advice alone. Solid line: effects for intervention group, striped line: effects for placebo group. Significant effects are visualized by a 95% confidence interval not including 0.

Disability outcome (models 4, 5 & 6)

From a sample of 225 cases for the disability outcome, we found that depressive symptoms did not mediate the effect of exercise and advice on disability (ACME -0.10 [95%CI -0.59 to 0.38]). The ADE (-1.70 [-2.89 to -0.57]) and the total effect (-1.80 [-3.09 to -0.47]) were significant. For exercise alone and advice alone, there was also no significant mediating effect of depressive symptoms: exercise (ACME -0.23 [-0.61 to 0.11]), advice (ACME 0.005 [-0.39 to 0.36]). For exercise alone, there was no significant ADE or total effects (ADE -0.46 [-1.41 to 0.52], total effect -0.69 [-1.70 to 0.39]). For advice alone, there were significant ADE (-1.30 [-2.28 to -0.28]) and total effects (-1.30 [-2.31 to -0.28]). See figure 4. None of the exposures (combined exercise and advice, exercise alone, or advice alone) had an effect on depressive symptoms (path a). Whereas the association between change in depressive symptoms and disability (path b) was significant for all three exposures: exercise and advice (regression coefficient = 0.17 [SE 0.05]), exercise (regression coefficient = 0.16 [SE 0.05]), advice (regression coefficient = 0.16 [SE 0.06]).

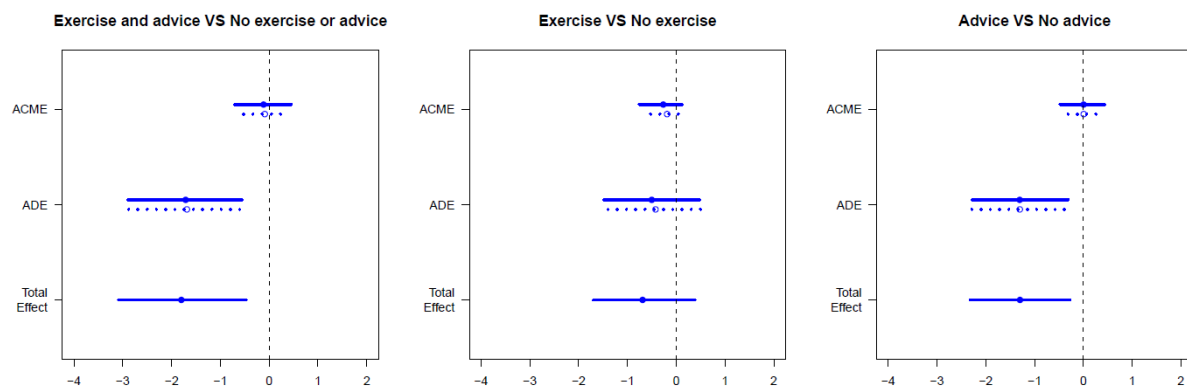


Figure 4. Effect size including 95% confidence interval of average causal mediation effect (ACME), average direct effect (ADE) and total effect for the model with the outcome disability and exposure advice and exercise, exercise alone, and advice alone. Solid line: effects for intervention group, striped line: effects for placebo group. Significant effects are visualized by a 95% confidence interval not including 0.

Sensitivity analysis for sequential ignorability

All six sensitivity analyses indicated that our estimates of the ACME were robust to the influence of potential unknown/unmeasured confounders. Even when we assume a large magnitude of unknown confounding effects (ρ at the extremes of -1 and 1), the ACME only varied by a small magnitude (Appendix 2 and 3).

Post-hoc sensitivity analysis

A visual inspection of the distribution of depression scores showed a large peak at 0 (zero-inflated). We conducted a post-hoc sensitivity analysis to check whether modelling the mediator using a negative binominal regression model would change the results of the effect of the exposure on the mediator (path a). The negative binominal model produced a regression coefficient of -0.04 (SE 0.31) for the exercise and advice intervention, -0.26 (SE 0.27) for the exercise alone intervention, and 0.14 (SE 0.27) for the advice alone intervention. These effect sizes approximated the coefficients of the linear regression models in the original analysis.

308 Table 3: Mediation effects

	Total effect	ADE	ACME	Path a	Path b
Model 1 (pain)					
Average effect	-1.23	-1.19	-0.05	-0.17 SE 1.16	0.06 SE 0.03*
Adex vs. placebo	(-1.94 to -0.49)*	(-1.85 to -0.48)*	(-0.24 to 0.15)		
Model 2 (pain)					
Average effect	-0.60	-0.50	-0.09	-1.12 SE 1.01	0.06 SE 0.04
Ex vs. placebo	(-1.21 to 0.06)	(-1.09 to 0.12)	(-0.29 to 0.05)		
Model 3 (pain)					
Average effect	-0.81	-0.82	0.005	0.60 SE 1.01	0.07 SE 0.04*
Ad vs. placebo	(-1.42 to -0.20)*	(-1.39 to -0.21)*	(-0.17 to 0.19)		
Model 4 (disability)					
Average effect	-1.80	-1.70	-0.10	-0.17 SE 1.16	0.17 SE 0.05*
Adex vs. placebo	(-3.09 to -0.47)*	(-2.89 to -0.57)*	(-0.59 to 0.38)		
Model 5 (disability)					
Average effect	-0.69	-0.46	-0.23	-1.12 SE 1.01	0.16 SE 0.05*
Ex vs. placebo	(-1.70 to 0.39)	(-1.41 to 0.52)	(-0.61 to 0.11)		
Model 6 (disability)					
Average effect	-1.30	-1.30	0.005	0.60 SE 1.01	0.16 SE 0.06
Ad vs. placebo	(-2.31 to -0.28)*	(-2.28 to -0.28)*	(-0.39 to 0.36)		

309 Results are presented as estimate (95% confidence interval) or estimate with standard error (SE). Adex:
310 combined advice and exercise intervention, Ex: exercise alone, Ad: advice alone. The indirect effect, or ACME,
311 is the effect of the exposure (group allocation) on pain or disability, mediated by depressive symptoms. The
312 direct effect, or ADE, is the remaining effect of the exposure on pain or disability that is not mediated through
313 depressive symptoms. The sum of the ACME and ADE equals the total effect. Path a is the effect of the
314 exposure on depressive symptoms. Path b is the effect of depressive symptoms on pain or disability. Models
315 were adjusted for confounders: age, gender, baseline values of pain, disability, depressive symptoms,
316 kinesiophobia, pain self-efficacy, pain catastrophizing. *= p<0.05

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Discussion

Depressive symptoms did not mediate the effect of exercise and/or advice on pain and disability in patients with subacute low back pain. A likely explanation for this is due to the lack of effect of exercise and advice on depressive symptoms. Second, we found a positive relationship between depressive symptoms at week six, and pain and disability at three months. This suggests that people with subacute low back pain and depressive symptoms could benefit from a targeted approach that specifically addresses depressive symptoms.

An interesting finding is that level of depressive symptoms in our sample were comparable to normative data for the Australian population (Lovibond & Lovibond, 1995). About seventy percent in our sample did not experience depressive symptoms, ten percent reported mild levels, and twenty percent reported moderate to severe levels of depressive symptoms. This is in contrast to the strong evidence that exists that depressive symptoms are elevated in acute low back pain (Shaw et al. 2016). This could have induced less power and may explain the absence of a mediating effect through depression because most participants did not have high depressive symptoms at baseline. Because our sample did not include a sufficient number of people with high depressive symptoms, we could not conduct subgroup analyses with current data. Mansell et al. (2016) evaluated the working mechanisms of a psychologically informed physiotherapy treatment for a subgroup of low back pain patients with high risk of poor outcome. Mansell et al. combined catastrophizing, fear-avoidance, anxiety and depression into a single mediator: pain-related distress. In this high risk group, the authors found that pain-related distress mediated the effect of the treatment (Mansell et al., 2016). The discrepancy between the findings of Mansell et al. (2016) and our current study might be an indication of the importance of stratified primary care based on the patient's risk of persistent

disabling pain as levels of depressive symptoms are a factor in currently widely used screening tools, such as the STarT Back (Hill, Dunn, Lewis et. al., 2008).

The significant association between depressive symptoms at week 6 and pain and disability scores at three months suggests that people with subacute low back pain could benefit from a targeted approach that specifically addresses depressive symptoms. Insights into the treatment of subacute low back pain has emerged over recent years since publication of the original trial some 10 years ago. For example, a systematic review provided strong evidence that patient education reduced psychological distress and use of health care related to low back pain. However, did not improve pain or function in the acute phase. (Traeger, Hubscher, Henschke et al. 2015). Also, the effects of a large variety of other interventions ranging from education and advice, brief interventions with light mobilization, diverse exercise programs, graded activity, psychological counselling or multidisciplinary biopsychosocial rehabilitation for subacute low back pain are still debated. (Marin et al. 2017)

Although guidelines recommend that low back pain should be managed in primary care, providing education and advice to stay active and at work, using a biopsychosocial framework, this is often not the reality in clinical practice (Foster, Anema, Cherkin et al., 2018). Improved and better integrated education of health-care professionals could support implementation of practice guidelines for low back pain (Foster, Anema, Cherkin et al., 2018). Results from this trial provide more insight in working mechanisms of exercise and advice in the treatment of subacute low back pain in primary care, therewith providing

researchers with new hypothesis for further research to improve treatment and providing clinicians with more understanding of working mechanisms of treatment in clinical practice. When clinicians can understand the rationale of a treatment, one is more likely to adopt to that treatment.

Strength and weaknesses

As it is not known what works for whom in the treatment of subacute low back pain, it is important that new knowledge is created by designing appropriate studies that incorporate mediation analysis a priori. Ideally, this means comparing two treatment arms, of which one should be devoted in changing the a priori hypothesized mediator, while the other treatment arm should provide another active treatment that is believed not to specifically target the mediating factor. However, when not available, it is feasible to reuse available data from older trials as we did to generate new knowledge. The strength of this study is the application of a causal mediation analysis on a well-conducted large randomized, placebo-controlled, clinical trial. This design allows for a rigorous evaluation of working mechanisms of an effective intervention. By application of the counterfactual framework for causal inference we have provided estimates that are based on causal definitions while explicitly outlining the assumptions required for causal interpretation. Primarily, we found our sensitivity analyses for sequential ignorability robust across all models, suggesting that high levels of unknown/unmeasured confounding for the mediator-outcome effect would not invalidate our results. A limitation is the relatively small percentage of patients who experienced moderate or severe depressive symptoms. Also, treatment content for subacute low back pain has evolved ever since designing the treatment of the original study. Further, we deviated from our a priori protocol by conducting additional analyses with pain intensity as the

outcome. Last, we were unable to account for the zero-inflated distribution of the mediator in our causal mediation analysis due to software limitations. However, our sensitivity analysis of the exposure-mediator effect to account for the zero-inflated distribution approximated the results from the original analysis.

Directions for future research

The significant direct effect indicates that there are other, unknown, mediators responsible for the effect of exercise and/ or advice on pain and disability in subacute low back pain treatment. There are other theories that could be tested, for example the mediating role of pain catastrophizing and self-efficacy. Further, stratification of the population into patients with high and low levels of depressive symptoms might provide a better insight in the mediating role of depressive symptoms. Future trials testing the efficacy of exercise and advice should plan and design a priori mediation analyses to identify mediators.

In conclusion, depressive symptoms did not mediate the effect of an exercise and/or advice intervention on pain or disability in patients with subacute low back pain with relatively low levels of depressive symptoms.

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