



Excess deaths associated with covid-19 pandemic in 2020: age and sex disaggregated time series analysis in 29 high income countries

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ABSTRACT OBJECTIVE

To estimate the direct and indirect effects of the covid-19 pandemic on mortality in 2020 in 29 high income countries with reliable and complete age and sex disaggregated mortality data.

DESIGN

Time series study of high income countries.

SETTING

Austria, Belgium, Czech Republic, Denmark, England and Wales, Estonia, Finland, France, Germany, Greece, Hungary, Israel, Italy, Latvia, Lithuania, the Netherlands, New Zealand, Northern Ireland, Norway, Poland, Portugal, Scotland, Slovakia, Slovenia, South Korea, Spain, Sweden, Switzerland, and United States.

PARTICIPANTS

Mortality data from the Short-term Mortality Fluctuations data series of the Human Mortality Database for 2016-20, harmonised and disaggregated by age and sex.

INTERVENTIONS

Covid-19 pandemic and associated policy measures.

MAIN OUTCOME MEASURES

Weekly excess deaths (observed deaths versus expected deaths predicted by model) in 2020, by sex and age (0-14, 15-64, 65-74, 75-84, and ≥85 years), estimated using an over-dispersed Poisson regression

model that accounts for temporal trends and seasonal variability in mortality.

RESULTS

An estimated 979 000 (95% confidence interval 954 000 to 1 001 000) excess deaths occurred in 2020 in the 29 high income countries analysed. All countries had excess deaths in 2020, except New Zealand, Norway, and Denmark. The five countries with the highest absolute number of excess deaths were the US (458 000, 454 000 to 461 000), Italy (89 100, 87 500 to 90 700), England and Wales (85 400, 83 900 to 86 800), Spain (84 100, 82 800 to 85 300), and Poland (60 100, 58 800 to 61 300). New Zealand had lower overall mortality than expected (−2500, −2900 to −2100). In many countries, the estimated number of excess deaths substantially exceeded the number of reported deaths from covid-19. The highest excess death rates (per 100 000) in men were in Lithuania (285, 259 to 311), Poland (191, 184 to 197), Spain (179, 174 to 184), Hungary (174, 161 to 188), and Italy (168, 163 to 173); the highest rates in women were in Lithuania (210, 185 to 234), Spain (180, 175 to 185), Hungary (169, 156 to 182), Slovenia (158, 132 to 184), and Belgium (151, 141 to 162). Little evidence was found of subsequent compensatory reductions following excess mortality.

CONCLUSION

Approximately one million excess deaths occurred in 2020 in these 29 high income countries. Age standardised excess death rates were higher in men than women in almost all countries. Excess deaths substantially exceeded reported deaths from covid-19 in many countries, indicating that determining the full impact of the pandemic on mortality requires assessment of excess deaths. Many countries had lower deaths than expected in children <15 years. Sex inequality in mortality widened further in most countries in 2020.

Introduction

During the SARS-CoV-2 pandemic, national governments have reported the number of deaths from covid-19, often on a daily basis. However, widespread heterogeneity exists in the accuracy and completeness of reported deaths from covid-19 across countries and jurisdictions.^{1,2} Measures taken to handle the covid-19 pandemic have also varied substantially across countries.^{1,3-8} Therefore, assessment of the full impact of the pandemic on mortality should include both the

WHAT IS ALREADY KNOWN ON THIS TOPIC

Assessment of the full impact of the pandemic on mortality in different populations should include both its direct effect on deaths from covid-19 and its indirect effect on deaths from other diseases

This requires estimation of “excess deaths”—the difference between the number of deaths from all causes during the pandemic and the expected number of deaths given a historical baseline from recent years

Studies reporting on excess mortality have not accounted for temporal and seasonal trends, or differences in the age and sex composition, across the countries when estimating excess mortality

WHAT THIS STUDY ADDS

The five countries with the highest absolute number of excess deaths were the US, Italy, England and Wales, Spain, and Poland; New Zealand had lower overall mortality than expected

The highest excess death rates (per 100 000) in men were in Lithuania, Poland, Spain, Hungary, and Italy; the highest rates in women were in Lithuania, Spain, Hungary, Slovenia, and Belgium

Accounting for difference in age, the excess death rates were much higher in men than women in almost all countries

direct effect of the pandemic on deaths from covid-19 and the indirect effect of the pandemic on deaths from other causes, as might be expected from the disruption to health services or from wider economic and social changes.⁹⁻¹⁷

Assessing the overall impact of the pandemic on mortality requires measurement of “excess deaths,” calculated as the difference between the number of deaths from all causes that occurred during the course of the pandemic and the expected number of deaths based on a historical baseline from recent years.¹⁸⁻²¹ Mortality below the expected levels is called “avoided mortality” or “mortality deficit,” whereas “mortality displacement” (or “harvesting”) is characterised by a period of excess deaths followed by a period of mortality deficit (see glossary in supplementary materials for more information).¹³⁻²² Mortality displacement indicates that people who have died during the course of an event (in this case, the covid-19 pandemic) would have died soon after in the absence of the event (that is, their death was brought forward by a short period). Previous studies have varied in their use of historical baselines to calculate excess deaths. Some have compared deaths during the pandemic with those reported in 2019 or with a simple average of the preceding few years.²⁰ However, estimates of expected deaths should ideally account for recent temporal trends in mortality that are likely to have occurred in the absence of the pandemic.²¹⁻²³⁻²⁴

Focusing on excess deaths from all causes also reduces the potential for misclassification of deaths from covid-19 and other causes, which facilitates comparisons between countries.¹⁸⁻²¹ The reliable ascertainment of deaths from covid-19 depends on widespread testing for covid-19 and on the methods used to assign cause of death, both of which vary between countries.¹⁻² In addition, as the case fatality from covid-19 is strongly related to age, comparisons between countries should account for this by stratifying or standardising excess deaths for age.²⁵⁻²⁶

We report the findings on excess deaths from the covid-19 pandemic in 29 countries after accounting for temporal trends and seasonal variations in mortality within countries, and for demographic differences between countries, to give age and sex specific excess mortality in each country during 2020.

Methods

Study design

We did a time series analysis of weekly mortality data collected from 29 member countries of the Organisation for Economic Cooperation and Development (OECD). Data for this study came from the Human Mortality Database. The database collates mortality and population data from authoritative national agencies and is maintained by the Department of Demography at the University of California, Berkeley, USA, and the Max Planck Institute for Demographic Research in Rostock, Germany.²⁷⁻²⁸ Data are made available in a standardised format for each country, disaggregated by age and sex.

Since the start of the covid-19 pandemic, weekly national mortality data have been collected from many of the countries represented in the database, to support objective and comparable assessments of the scale of short term variations in mortality (Short-term Mortality Fluctuations data series). We used age and sex specific mortality data from this series. Further details on the sources of data and on the methods for collection and standardisation of data for the series are available in the supplementary table S1.²⁹⁻³⁰

Weekly mortality data were available for most OECD countries in the database for the period 2016-19, but delays in registration of vital events meant that weekly mortality data for some countries were available for only part of 2020. We restricted our analyses to 29 countries with complete weekly data in 2020 (52 weeks). Weekly mortality data were available by sex and by 5 year age groups (0-4, 5-9, 10-14, 15-19, 20-24, 25-29, 30-34, 35-39, 40-44, 45-49, 50-54, 55-59, 60-64, 65-69, 70-74, 75-79, 80-84, and ≥85 years) in 23 countries and by sex and broad age groups (0-14, 15-64, 65-74, 75-84, and ≥85 years) for the remaining six countries (England and Wales (combined), Germany, Israel, South Korea, New Zealand, and the US).

Statistical analysis

For each country, we compared the age and sex specific number of deaths in each week in 2020 (“observed weekly deaths”) with the age and sex specific number of expected deaths in each week given historical trends, to estimate the weekly number of excess deaths. We estimated weekly expected deaths by using an over-dispersed Poisson model that accounts for temporal trends and seasonal and natural variability in mortality.²³ The model has been validated in simulation and when using historical data.²³

For each age and sex specific group in a given country, the model used mortality data for 2016-19 to estimate the expected number of deaths in each group for each week of 2020. To ameliorate the influence of past major events on these estimates, past periods (2016-19) with atypical mortality (for example, heat waves, influenza outbreak) were excluded from the model (supplementary table S2).

Specifically, let $Y_{s,c,a}(t)$ be the number of deaths at week (t) for individuals of sex (s), in country (c), and age group (a). Note that (s) represents either female or male, (c) can be any of the 29 OECD countries, and (a) takes on any of the aforementioned age groups. Assuming that $Y_{s,c,a}(t) \sim \text{Poisson}(\mu_{s,c,a}(t))$, our mean model is:

$$\mu_{s,c,a}(t) = N_{s,c,a}(t) \exp\{\beta t + g_{s,c,a}(w_t)\} \text{ for } t \in I_c(1)$$

The expected number of deaths at week (t) for individuals of sex (s), in country (c), and age group (a) is represented with $\mu_{s,c,a}(t)$; $N_{s,c,a}(t)$ is an offset that accounts for the population size; β represents a linear effect of time that accounts for slow moving changes in mortality; $g_{s,c,a}(w_t)$ is a function that accounts for seasonal trends, where $w_t \in \{1, 52\}$ represents a week of the year; and I_c is a country specific training interval that is used to fit the model. For the smooth estimates of percentage change from average, let (t') be a week in 2020; then:

$$\lambda_{s,c,a}(t') = \mu_{s,c,a}^{\wedge}(t') \exp\{f_{s,c,a}(t')\} \quad (2)$$

In model (2), $\lambda_{s,c,a}(t')$ represents the average number of observed deaths at week (t') for individuals of sex (s), in country (c), and age group (a). The number of deaths at week (t') in the counterfactual scenario of no pandemic is represented with $\mu_{s,c,a}^{\wedge}(t')$ and it is used as an offset. Lastly, the function $f_{s,c,a}(t')$ is a natural cubic spline with 12 internal knots. It follows that $\gamma_{s,c,a}^{\wedge}(t') = \lambda_{s,c,a}^{\wedge}(t') / \mu_{s,c,a}^{\wedge}(t') - 1$ represents a smooth estimate of percentage change from average at t' . We used the *excessmort* R package to fit our modelling scheme and provide a detailed description of the model in the supplementary methods.

We report the excess deaths and excess death rates in each country overall, by sex, and by age. To avoid a false sense of precision,²¹ numbers <1000 are rounded to the nearest ten, numbers between 1000 and <100 000 are rounded to the nearest hundred, and those $\geq 100 000$ are rounded to the nearest thousand. To facilitate the comparison of excess deaths between countries, the age and sex specific excess death rates in each country for 2020 were directly standardised to the 2013 European standard population to calculate the age standardised excess death rate in each country in males and females.³¹ We produced age standardised rates by using 5 year age groups for 19 countries but broad age groups (0-14, 15-64, 65-74, 75-84, and ≥ 85 years) for the remaining countries for which either data by 5 year age groups were not available or the weekly age and sex specific death counts were too small (such as in Estonia, Lithuania, Northern Ireland, and Norway). Rates in the broad age groups were adjusted, using methods developed by the Max Planck Institute for Demographic Research on data from the core Human Mortality Database, to those that would be expected given the age distribution of the standard population; they were thus comparable to rates calculated using 5 year age groups.³²

We also compared the overall number of estimated excess deaths with the reported number of deaths from covid-19 in each country (source: covid-19 time series data collated by Johns Hopkins University Center for Systems Science and Engineering, available at <https://github.com/datasets/covid-19/tree/main/data>). We used RStudio (version 1.4.1103) and Stata SE (v.15.1) for all statistical analyses. All data used in this analysis are fully anonymised and aggregated without any identifiable information.

Patient and public involvement

The rapid nature of this research project in the context of the ongoing covid-19 pandemic meant that participants were not involved in the development of the research question or the outcome measures, or in the design or implementation of the study; neither were participants asked to advise on the interpretation or writing of the manuscript. However, the findings will be widely disseminated to the public through official channels (press release, blogs, institutional websites, and repositories), personal communications, and social communication tools.

Results

Excess deaths

Table 1 shows the estimated number of excess deaths in 2020 for the 29 countries included in the analysis, overall and separately for men and women. All countries had excess deaths for 2020, except Norway and Denmark, for which we found no evidence of a difference between the observed and expected number of deaths, and New Zealand, for which we found evidence of fewer observed than expected deaths. Overall, we estimated that about one million (979 000, 95% confidence interval 954 000 to 1 001 000) excess deaths occurred in the 29 countries analysed. The five countries with the highest absolute number of excess deaths were the US (458 000, 454 000 to 461 000), Italy (89 100, 87 500 to 90 700), England and Wales (85 400, 83 900 to 86 800), Spain (84 100, 82 800 to 85 300), and Poland (60 100, 58 800 to 61 300), whereas New Zealand had lower overall mortality than expected (−2500, −2900 to −2100). In most countries, the number of excess deaths was somewhat higher in men than in women.

The estimated number of excess deaths by age groups in men and women are available in supplementary table S3. Observed deaths in children <15 years were similar to the expected levels in most countries and lower than expected in some countries (for example, England and Wales, France, Israel, Italy, New Zealand, South Korea, and the US). The total number of excess deaths was largely concentrated among people aged ≥ 75 years, followed by people aged 65-74 years. Excess deaths in people aged 15-64 years were generally lower than in the older age groups with the exception of the US, where the total number of excess deaths was approximately 131 000 in people aged 15-64 years (91 500 (90 200 to 92 800) in men and 39 500 (38 500 to 40 500) in women), higher than those estimated among people aged 65-74 and 75-84 years. New Zealand had lower than expected deaths across all the age groups.

The distribution of excess deaths over time varied by country. Figure 1 shows the estimated monthly (aggregated from the weekly estimates) excess deaths (per 100 000 population) in 2020 for men and women, and figure 2 shows the percentage deviation from the expected deaths across weeks of 2020. Some countries did not see major increases in excess deaths at any point during 2020, including Denmark, Norway, South Korea, and New Zealand. Other countries reported one or more major waves of excess deaths, after which excess mortality fell to lower levels.

Consistent with the timing of known peaks of covid-19 infection, many countries had major waves of excess deaths in the northern hemisphere spring (March-May) and in autumn-winter (October-December). In spring, the level of excess deaths was particularly high (with some weeks of >50% excess deaths) in Italy, Spain, England and Wales, Scotland, Northern Ireland, Belgium, and the Netherlands. In autumn-winter, levels of excess deaths were particularly high in Austria, Belgium, Czech Republic,

Table 1 | Estimated number of excess deaths in 2020 in 29 high income countries, by sex

Country	Total No (95% CI)	No (95% CI) male	No (95% CI) female
United States	458 000 (454 000 to 461 000)	245 000 (243 000 to 247 000)	213 000 (210 000 to 215 000)
Italy	89 100 (87 500 to 90 700)	49 000 (47 900 to 50 000)	40 200 (39 000 to 41 300)
England and Wales	85 400 (83 900 to 86 800)	45 000 (44 000 to 46 000)	40 400 (39 400 to 41 400)
Spain	84 100 (82 800 to 85 300)	41 100 (40 200 to 42 000)	43 000 (42 100 to 43 800)
Poland	60 100 (58 800 to 61 300)	35 400 (34 500 to 36 300)	24 700 (23 800 to 25 600)
France	43 500 (42 000 to 45 000)	24 600 (23 500 to 25 700)	18 900 (17 800 to 20 000)
Germany	25 900 (24 000 to 27 800)	17 700 (16 400 to 19 000)	8200 (6900 to 9600)
Belgium	17 900 (17 200 to 18 500)	9000 (8600 to 9500)	8800 (8400 to 9300)
Hungary	16 600 (15 900 to 17 300)	8100 (7600 to 8500)	8500 (8000 to 9000)
Netherlands	15 300 (14 600 to 16 100)	8800 (8200 to 9300)	6500 (6000 to 7100)
Czech Republic	14 400 (13 700 to 15 000)	8000 (7500 to 8500)	6400 (5900 to 6900)
Sweden	9300 (8700 to 9800)	5100 (4700 to 5500)	4100 (3700 to 4500)
Portugal	8500 (7800 to 9100)	4000 (3500 to 4500)	4500 (4000 to 5000)
Greece	7700 (7000 to 8300)	3500 (3100 to 4000)	4100 (3600 to 4600)
Austria	6800 (6200 to 7300)	3800 (3400 to 4200)	2900 (2500 to 3300)
Scotland	6800 (6400 to 7300)	3900 (3600 to 4200)	2900 (2600 to 3300)
Switzerland	6800 (6300 to 7300)	4000 (3700 to 4400)	2800 (2400 to 3200)
Lithuania	6800 (6500 to 7200)	3700 (3500 to 4000)	3100 (2800 to 3400)
Slovakia	4400 (4000 to 4900)	2300 (2000 to 2700)	2100 (1800 to 2400)
South Korea	4000 (2900 to 5000)	1800 (1000 to 2600)	2200 (1400 to 2900)
Slovenia	3200 (3000 to 3500)	1600 (1400 to 1800)	1600 (1400 to 1800)
Northern Ireland	2200 (1900 to 2400)	1000 (860 to 1200)	1100 (980 to 1300)
Israel	2000 (1600 to 2400)	1300 (1000 to 1600)	690 (400 to 990)
Finland	1000 (550 to 1500)	690 (360 to 1000)	320 (0 to 640)
Latvia	820 (490 to 1100)	270 (50 to 490)	550 (310 to 790)
Estonia	670 (430 to 910)	330 (160 to 490)	340 (170 to 520)
Norway	-70 (-470 to 320)	100 (-180 to 370)	-170 (-450 to 110)
Denmark	-160 (-610 to 300)	-40 (-360 to 290)	-120 (-440 to 200)
New Zealand	-2500 (-2900 to -2100)	-1200 (-1500 to -970)	-1300 (-1500 to -1000)

Excess deaths in 2020 calculated as difference in observed deaths and expected deaths predicted using over-dispersed Poisson model that accounts for temporal trends and seasonal and natural variability. Numbers <1000 are rounded to nearest ten; numbers between 1000 and <100 000 are rounded to nearest hundred, and numbers ≥100 000 are rounded to nearest thousand. Data: Short-term Mortality Fluctuations data series of Human Mortality Database

Hungary, Italy, Lithuania, Poland, Slovakia, Slovenia, and Switzerland. Within countries, we found little evidence of a difference between men and women in trends over time in the proportional increases in excess deaths in 2020.

In some countries, during some periods of the year, the observed number of deaths fell below the expected number of deaths, most notably in New Zealand in the middle of the year. Detailed age and sex specific estimated number of excess deaths and percentage deviation from expected deaths over time are provided in supplementary figure S1 (A-E) and supplementary figure S2 (A-E), respectively. Although excess mortality was higher in the older age groups in most of the countries, the observed mortality was lower than expected in children <15 years in many countries.

Crude and age standardised excess death rates

Figure 3 shows the age specific excess death rates (per 100 000) in each country, separately in men and women. The excess death rate increased exponentially with age in both sexes, except in those countries with no evidence of excess deaths, for which we observed a reduction in excess death rates at older ages; a particularly marked reduction in excess death rates occurred in New Zealand at ages ≥65 years. In most countries, age specific excess death rates were higher in men than in women, and the absolute difference in rates between the sexes tended to increase with age. However, in the

US, in contrast to other countries, the excess death rate (per 100 000) was higher among women (1899, 1855 to 1943) than men (1719, 1653 to 1784) at ages ≥85 years.

Figure 4 shows the crude and age standardised excess death rates in men and women. The highest crude excess death rates (per 100 000) in men were in Lithuania (285, 259 to 311), Poland (191, 184 to 197), Spain (179, 174 to 184), Hungary (174, 161 to 188), and Italy (168, 163 to 173); the highest rates in women were in Lithuania (210, 185 to 234), Spain (180, 175 to 185), Hungary (169, 156 to 182), Slovenia (158, 132 to 184), and Belgium (151, 141 to 162). Although we found little evidence of differences between men and women in the crude excess death rates in 2020, almost all countries with excess deaths had higher age standardised rates for men than women (fig 4). The age standardised rates also varied substantially within men and women across the countries analysed. The highest age standardised excess death rates (per 100 000) in men were in Lithuania (360, 324 to 396), Poland (298, 289 to 307), Hungary (235, 217 to 254), Slovenia (222, 189 to 256), and the Czech Republic (211, 194 to 228); the highest age standardised rates in women were in Lithuania (152, 134 to 171), Hungary (144, 133 to 156), the US (131, 129 to 133), Spain (125, 121 to 129), and Northern Ireland (121, 97 to 146). The lowest age standardised excess mortality rates for both men and women were in New Zealand, Denmark, Norway, South Korea, and Finland.

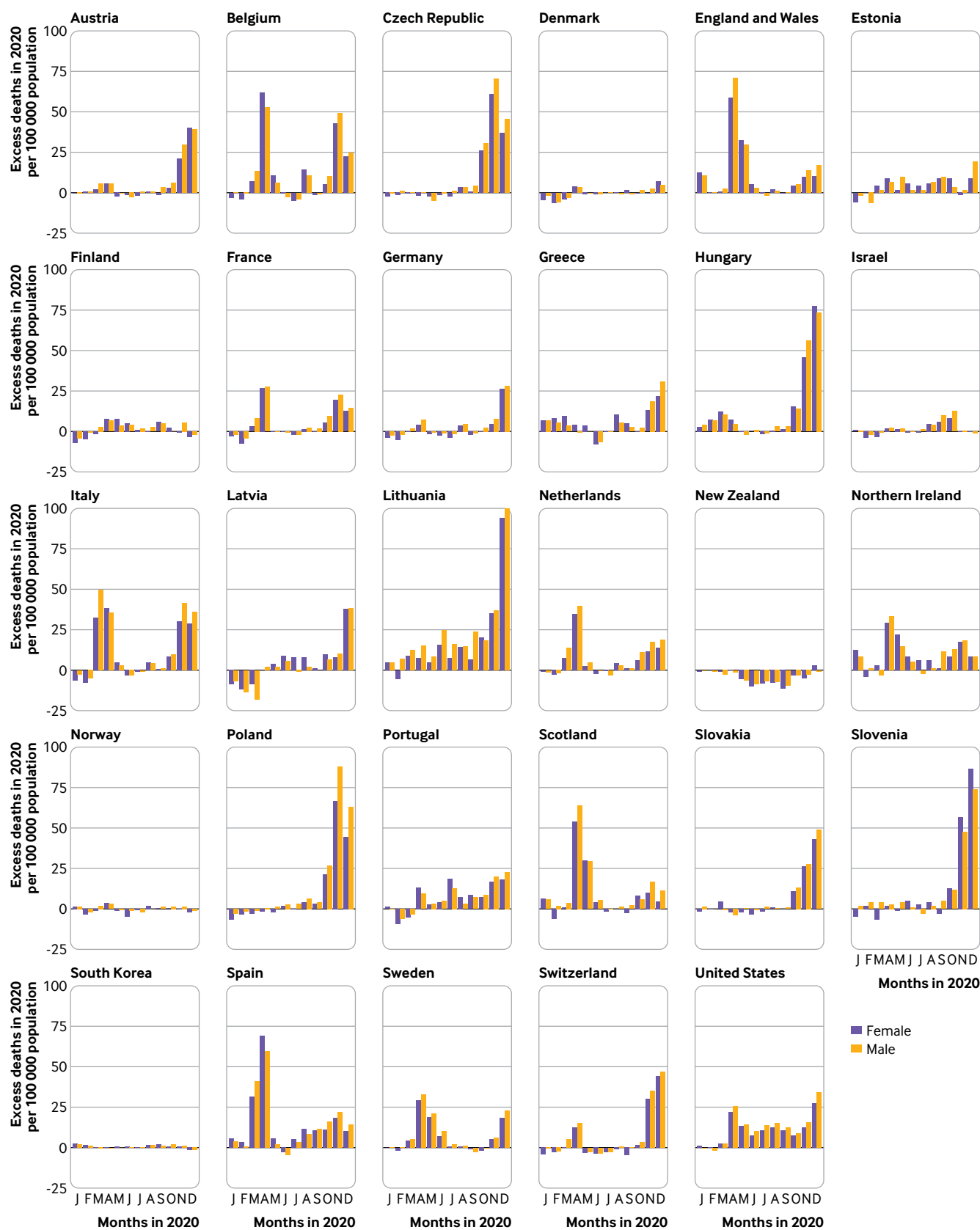


Fig 1 | Monthly excess deaths (per 100 000) in 29 high income countries in 2020, all ages, by sex. Excess deaths in 2020 were calculated as difference in observed deaths and expected deaths predicted using over-dispersed Poisson model that accounts for temporal trends and seasonal and natural variability. Estimated excess deaths for each week of 2020 were aggregated into months. Data: Short-term Mortality Fluctuations data series of Human Mortality Database

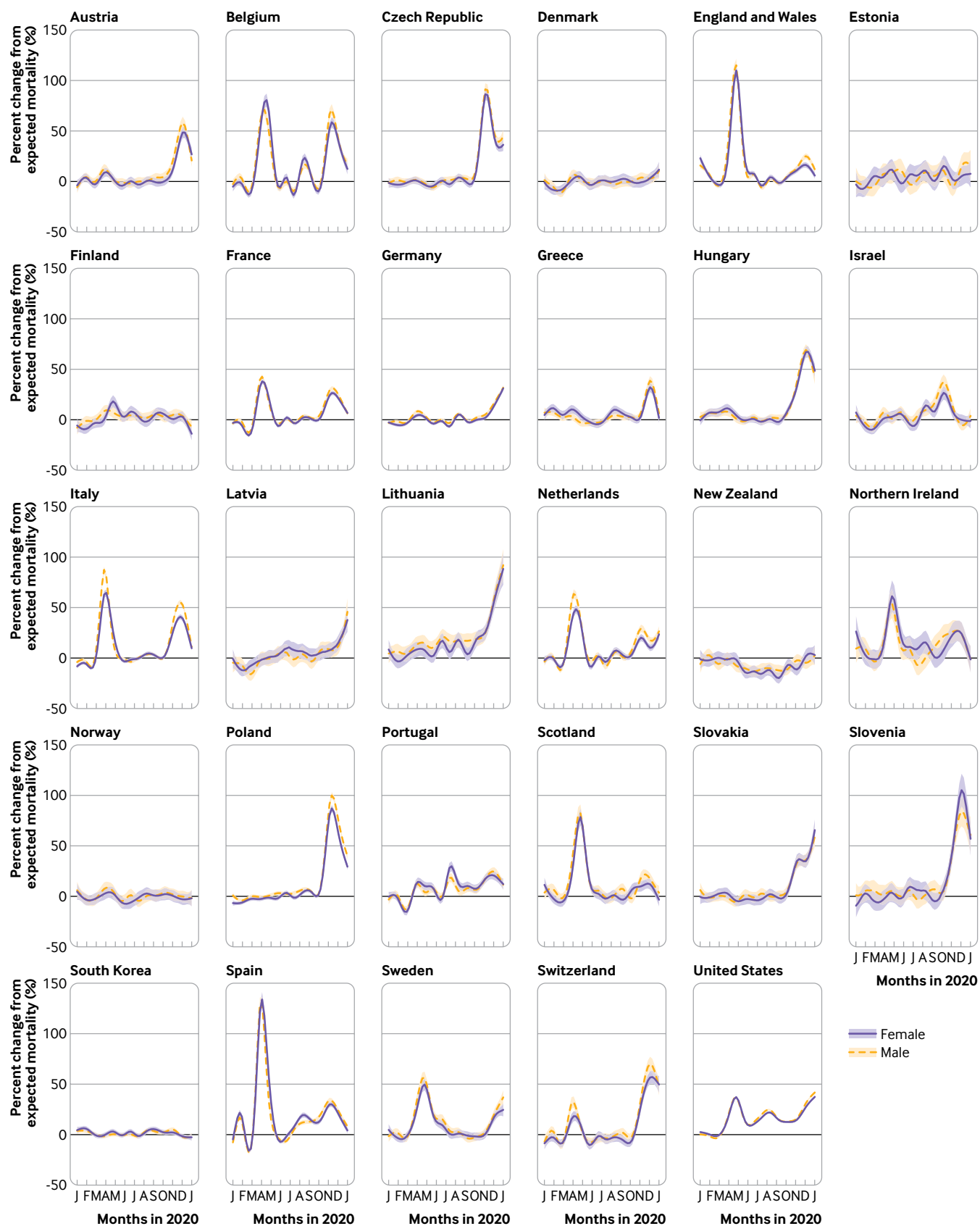


Fig 2 | Weekly percentage deviation from expected deaths in 29 countries in 2020, all ages, by sex. Weekly percentage deviation of observed deaths versus expected deaths predicted using over-dispersed Poisson model that accounts for temporal trends and seasonal and natural variability. Percentage changes from mean during study period were modelled using smooth function of time. Data: Short-term Mortality Fluctuations data series of Human Mortality Database

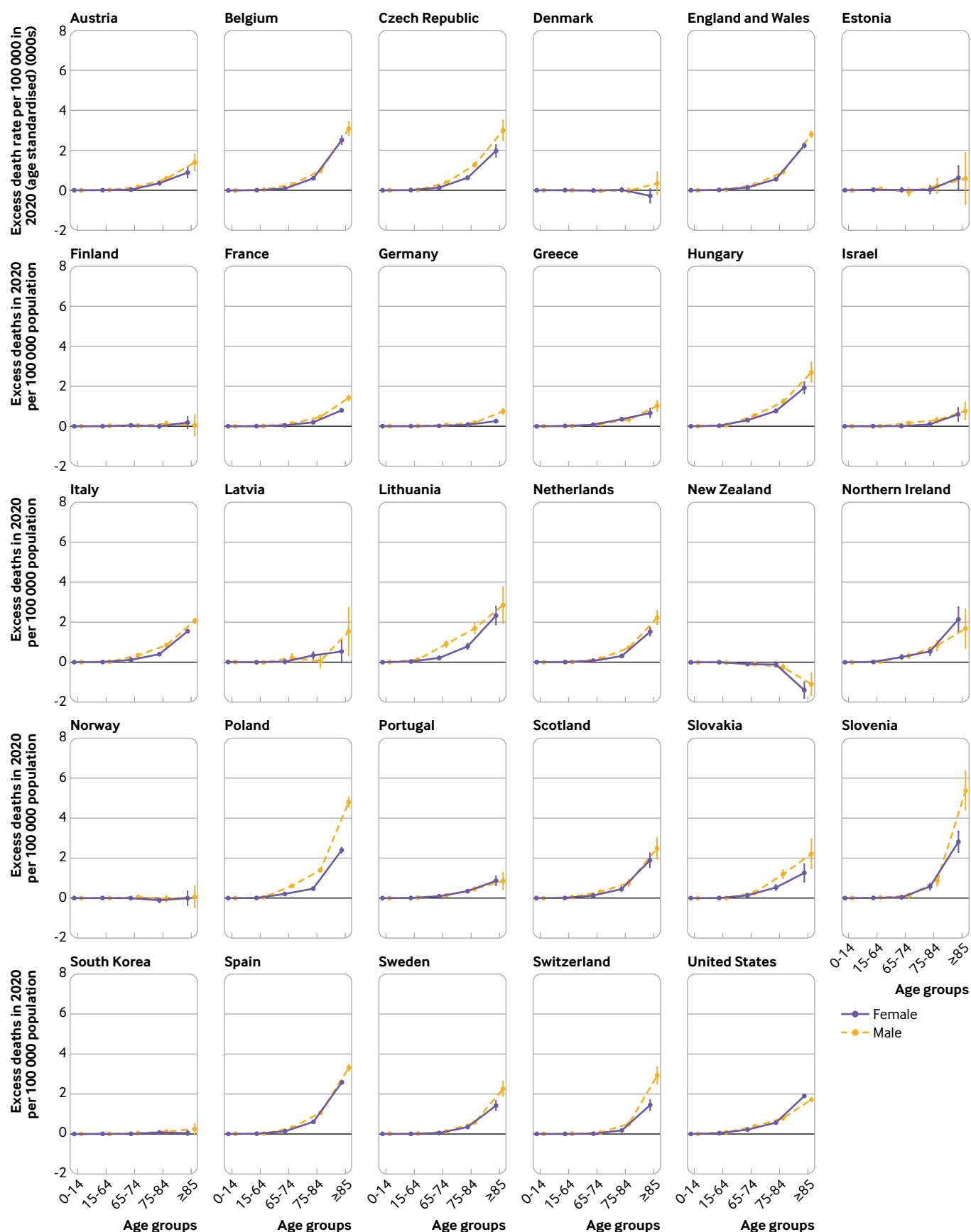


Fig 3 | Excess death rates in 29 high income countries in 2020, by sex and age. Excess death rate (per 100 000) in 2020 was calculated as difference in observed deaths and expected deaths predicted using over-dispersed Poisson model that accounts for temporal trends and seasonal and natural variability. Age standardised within each age group. Data: Short-term Mortality Fluctuations data series of Human Mortality Database

Comparison of estimated excess deaths and reported covid-19 deaths

In most countries, the estimated number of excess deaths exceeded the number of reported deaths from covid-19 (table 2). For example, in both the US and the UK (that is, England, Wales, Northern Ireland, and Scotland combined), the estimated excess deaths were more than 30% higher than the number of reported covid-19 deaths, and they were more than 50% higher in some other countries, including Spain, Poland, Hungary, Greece, Lithuania, Slovakia, Estonia, and South Korea. However, New Zealand, Norway, Denmark, Israel, France, Germany, Belgium, and Switzerland had a higher number of reported covid-19 deaths than estimated excess deaths.

Annual mortality 2016-20

The impact of the pandemic on trends in annual mortality rates for men and women in 2016-20 for each country is shown in supplementary figure S3. Men had consistently higher age standardised mortality rates than women across the period in all the countries. In general, age standardised mortality rates declined from 2016 to 2019, in both men and women. Whereas Denmark, Finland, South Korea, Norway, and New Zealand continued to show reduced mortality rates in 2020, most other countries experienced a marked increase in rates in both men and women, together with a widening of the gaps in mortality rates between the sexes. Sex specific annual mortality rates are shown for age groups 0-14, 15-64, 65-74, 75-84, and ≥85 years in supplementary figure S4 (A-E).

Model fit

Supplementary figure S5 (A-B) shows the weekly percentage deviation from model predicted excess deaths in 2016-20, along with the 95% confidence intervals, separately for males and females. It shows that the relative excess deaths fluctuated around zero between 2016 and 2019 in most countries indicating goodness of fit with a median of the median absolute deviation of 0.059 (interquartile range 0.042-0.068) (detailed country specific estimates, separately in males and females, are provided in supplementary table S4). However, some periods of moderate increase in excess mortality occurred during some winter months in some countries, consistent with known periods of excess mortality from seasonal influenza. We excluded these periods from the time periods used to estimate the counterfactual expected mortality (list of country specific dates excluded is available in supplementary table S2).

Discussion

In this international comparative study of excess deaths in relation to the covid-19 pandemic in 29 high income countries, all countries had excess deaths in 2020, except New Zealand, Norway, and Denmark. The highest excess deaths were in the US, Italy, England and Wales, Spain, and Poland. The highest excess death rates in men were in Lithuania, Poland, Spain,

Hungary, and Italy; the highest rates in women were in Lithuania, Spain, Hungary, Slovenia, and Belgium. Total deaths were below the expected level in New Zealand. Even though little sex difference existed in terms of total number or crude excess death rates, the age standardised excess death rate was noticeably higher in men than in women in most countries except New Zealand, Denmark, South Korea, Greece, Norway, Finland, Estonia, Northern Ireland, and Latvia. Most of the countries had underreporting of deaths from covid-19, an increase in non-covid-19 deaths, or a combination of both. On the other hand, the estimated number of excess deaths was lower than the reported covid-19 deaths in some countries (for example, Israel, France, Germany) owing to a reduction in non-covid-19 deaths in specific subgroups of the population (see supplementary methods for more details). Despite a recent trend of declining annual mortality rate between 2016 and 2019, the annual mortality rate increased substantially in 2020 in most of the countries, in both men and women. A similar trend was observed in sex differences in annual age standardised mortality rates.

Comparison with previous literature

Several previous studies have reported country specific excess deaths for the first few months of 2020.^{18 19 21 33-36} Mortality in younger age groups was reported to be at or lower than expected levels, which is consistent with our findings.^{19 20} Estimated excess deaths in the US are consistent with the projections of more than 400 000 excess deaths and recent estimates produced by the US Centers for Disease Control and Prevention (CDC).^{18 37} Drawing on data up to May 2020, previous research reported that Italy, Spain, and England and Wales had the highest number of excess deaths, which is consistent with our findings.^{21 38} An estimated 47 243 excess deaths were reported in England and Wales up to May 2020,³³ and 2400-4000 excess deaths were estimated to have occurred in Portugal up to April 2020.³⁴ The estimated number of excess deaths in England (excluding Wales) was 69 925 between 21 March and 25 December according to Public Health England (PHE).¹⁹ Although not directly comparable (our estimates include aggregate data from England and Wales), our estimate of excess deaths is likely higher than the estimates from PHE. This could largely be due to different methods of estimation of the counterfactual expected counts and the varying time periods used. The PHE model, like most of the models estimating excess deaths, used the average of the past few years' data without considering the natural and seasonal variabilities in mortality pattern over the control period. Our estimate of excess deaths in Poland is noticeably higher than previously reported, which could also be attributed to a marked increase in mortality in October-December 2020.²¹ Recent reports, however, noted that in 2020 Poland recorded the highest number of deaths since the second world war.³⁹ New Zealand was reported to have deaths at or below expected levels, which is consistent with our findings.^{21 40} Our estimates show that Sweden had the

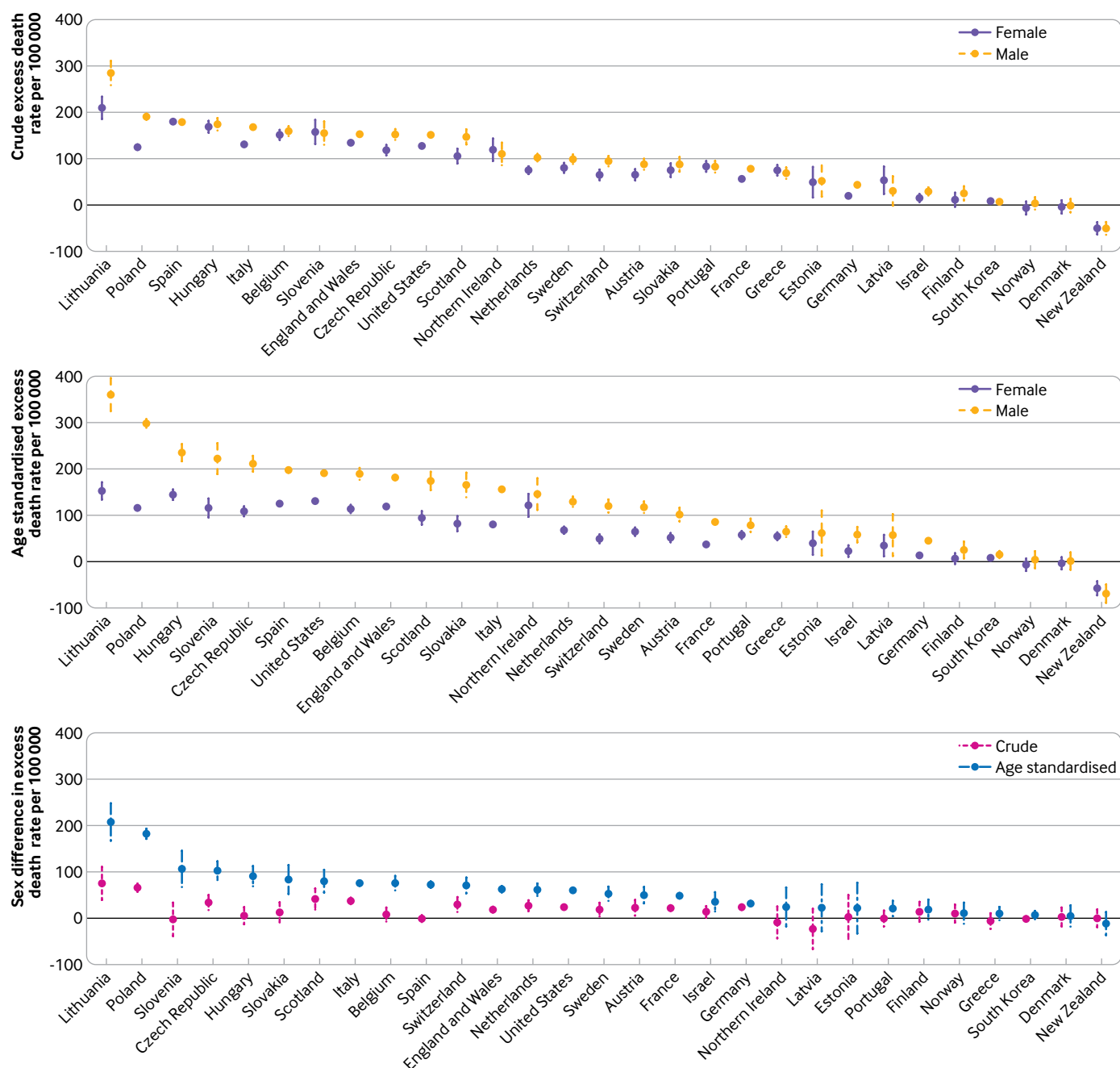


Fig 4 | Crude and age standardised excess death rates in 29 high income countries in 2020, by sex. Excess death rate (per 100 000) in 2020 was calculated as difference in observed deaths and expected deaths predicted using over-dispersed Poisson model that accounts for temporal trends and seasonal and natural variability. Age standardised, where indicated, to 2013 European standard population. Bottom panel: estimates above horizontal line at zero indicate higher excess death rate in men. Data: Short-term Mortality Fluctuations data series of Human Mortality Database

highest excess deaths among the Nordic countries, which is also consistent with a previous report.⁴¹ Our estimated number of excess deaths in South Korea is lower than those reported by Statistics Korea.^{42–43} This is because these reports compared the mortality in 2020 with that in 2019, whereas our model examined the difference in 2020 compared with the expected mortality based on 2016–19 data.

Previous studies reported a disproportionately higher toll of covid-19 mortality in men,^{44–46} but our detailed analysis showed that this could vary from

no difference to a substantial difference across the 29 countries. A previous study concluded that no sex difference existed in terms of excess deaths.²¹ By contrast, our analysis shows that many countries with no sex difference in terms of crude excess death rate did show a remarkable sex difference after standardisation for age. However, we observed no sex difference in New Zealand, Denmark, South Korea, Greece, Norway, Finland, Estonia, Northern Ireland, or Latvia, even after age standardisation.²¹ Among the Nordic countries, Sweden showed the largest gaps

Table 2 | Difference between reported number of covid-19 deaths and estimated number of excess deaths associated with covid-19 pandemic in 2020 in 29 high income countries

Country	Total covid-19 deaths reported* (A)	Total excess deaths estimated† (B)	Difference between excess deaths and reported covid-19 deaths (B–A)	Ratio of excess deaths to reported covid-19 deaths, % (B/A×100)
United States	339 014	458 000	118 986	135
United Kingdom‡	70 860	94 400	23 540	133
Italy	71 925	89 100	17 175	124
Spain	49 824	84 100	34 276	169
Poland	27 118	60 100	32 982	222
France	62 867	43 500	–19 367	69
Germany	30 297	25 900	–4 397	85
Belgium	19 200	17 900	–1300	93
Hungary	9047	16 600	7553	183
Netherlands	11 090	15 300	4210	138
Czech Republic	11 044	14 400	3356	130
Sweden	8279	9300	1021	112
Portugal	6619	8500	1881	128
Greece	4606	7700	3094	167
Switzerland	7210	6800	–410	94
Austria	5881	6800	919	116
Lithuania	1613	6800	5187	422
Slovakia	1773	4400	2627	248
South Korea	819	4000	3181	488
Slovenia	2565	3200	635	125
Israel	3226	2000	–1226	62
Finland	524	1000	476	191
Latvia	559	820	261	147
Estonia	204	670	466	328
Norway	421	–70	NA	NA
Denmark	1174	–160	NA	NA
New Zealand	25	–2500	NA	NA

NA=not applicable.

*Reported covid-19 deaths in 2020 (up to 27 December 2020, end of 52nd week). Source: Unified covid-19 dataset maintained by Johns Hopkins University Center for Systems Science and Engineering (<https://raw.githubusercontent.com/datasets/covid-19/main/data/countries-aggregated.csv>).

†Total excess deaths estimated from model. Numbers <1000 are rounded to nearest ten; numbers between 1000 and <100 000 are rounded to nearest hundred, and numbers ≥100 000 are rounded to nearest thousand.

‡Total number of excess deaths in England and Wales, Scotland, and Northern Ireland. See supplementary methods for simpler explanation of discrepancies between reported covid-19 deaths and estimated excess deaths.

in terms of sex inequality in age standardised excess death rate, with a much higher mortality rate in men than in women. This may be a result of a multitude of factors, including occupational and lifestyle factors between the sexes that may increase the likelihood of exposure to SARS-CoV-2 among men than women or differences in underlying comorbidities between men and women that might have increased the likelihood of death following infection.^{45 47-49} Many of these countries rank very high on the gender equality index, but no clear relation existed between gender equality and differences in age standardised excess death rate between men and women, with some countries that rank highly on gender equality indices having large differences between sexes and vice versa.^{50 51} New Zealand managed to keep deaths below the expected level in all the age groups in both men and women, with a corresponding narrowing of the gender inequality. The success in managing the covid-19 pandemic in New Zealand and South Korea to date may be attributed to their early viral elimination policies.⁵²⁻⁵⁴

Previous studies in the US reported that about two thirds of the estimated excess deaths were reported as deaths from covid-19,⁵⁵ which is lower than our estimate of 74% in the US. This may be indicative of a better reporting of, or an increase in, covid-19 deaths, or a combination of both, later in the pandemic.

Our study found that the existing sex inequality in annual standardised mortality rate, with higher rates in men than in women, was further widened in most of the countries during the first calendar year of the covid-19 pandemic. A large body of research reported that past pandemics have also widened existing disparities.⁵⁶⁻⁵⁹ A reduction in mortality (avoided mortality) was seen in children <15 years in some countries in 2020, which could be associated with a reduction in unintentional injuries (including drowning and car accidents) and a lower burden of cardiovascular and respiratory diseases potentially associated with a change in lifestyle or environmental factors during the periods of lockdown.⁶⁰⁻⁶³ However, long term data on the cause of deaths will be needed to fully examine the effects, because lockdown may have had other unintended consequences (for example, worsening mental health), which may take a longer time to affect overall mortality.¹³

In contrast to other countries, the excess death rates in women aged ≥85 years in the US were higher than in their male counterparts. The exact reasons for such a discrepancy will require detailed exploration of the cause of deaths in this age group, but we propose some potential reasons. Firstly, according to the CDC's recent estimates of cause of deaths by age and sex, this is the only age group in which women had higher

death rates than men across the causes of death in 2020—namely, covid-19 deaths (70 985 v 50 801), total deaths (616 124 v 393 995), pneumonia deaths (54 521 v 46 828), pneumonia and covid-19 deaths (25 989 v 23 086), influenza deaths (1065 v 734), and pneumonia, influenza, or covid-19 deaths (100 473 v 75 178).⁶⁴ Secondly, covid-19 deaths in US nursing homes have been described as catastrophic with >40% share of covid-19 deaths among this population that is only 0.6% of the total US population.^{65–67} Moreover, all cause mortality in US nursing homes had been decreasing until 2017.⁶⁸ With the decreasing trend, the expected all cause mortality in 2020 would be lower than the 2017 estimates of 534 714 deaths.⁶⁸ However, contrary to the expectations, the all cause mortality increased to 585 429 in 2020, with the highest number of deaths occurring in the ≥85 age group (n=328 387).⁶⁹ Thirdly, 59% of nursing home residents are women,⁷⁰ and elderly female nursing home residents have a higher burden of obesity (36.1% v 31.8%) and hypertension (86% v 81.5%) than their male counterparts.⁷¹ Fourthly, as depicted in supplementary figure S4 (E), although all cause mortality rates in men ≥85 remained very similar to those in the past few years, the death rates in women ≥85 years decreased sharply in 2019, but the increase in the death rate in 2020 was much steeper in women than in men. All these factors may have contributed to a higher estimated excess death rate in women than men in the ≥85 year age category in the US.

Strengths and limitations of study

In this large international comparative study, we examined excess deaths associated with the covid-19 pandemic in 2020 by using detailed age and sex specific mortality data with robust and validated analytical methods. Methods for estimating excess deaths have varied substantially. The most commonly used method is to compare deviations in mortality with the average of the preceding few years.²⁰ Comparing the deaths in the timeline of interest with the simple average of the previous few years ignores critical factors such as recent general health improvement and decline in mortality with time, seasonal and natural variability, and demographic changes.^{23–24} More nuanced methods of outbreak detection modelling have been proposed to detect aberrations in mortality on the basis of historical trends.^{72–74} These methods provide a point estimate and an upper bound for expected deaths and produce a conservative estimate of excess deaths by considering counts above these thresholds.¹⁸ More recently, Rizzi and Vaupel proposed a method to “shortcast” (neologism for short term forecasting) the average number of deaths in the absence of an event that relies exclusively on death counts.²⁴ Their approach is appealing owing to its simplicity but was applied to only four countries, and, as the authors specifically noted, this method is designed for short term forecasting. Therefore, this method may not be suitable for the covid-19 pandemic, which continues to exist more than a year after its

emergence. Moreover, this method does not account for changing demographic composition. Karlinksky and Kobak take a much more traditional route to assessing excess mortality.⁷⁵ For each country, they fit a linear regression with a slope for year and week specific coefficients without taking into account the possible changes in the population. To calculate counterfactual counts, they simply predict fitted values for 2020. Their model relies on marginal death counts without stratification by age or sex, and hence is limited to exploring the differences by two of the most important predictors of mortality. Our dataset is richer than that used by the groups described above because it provides more granular data by age and sex in a large number of countries. Furthermore, our statistical model does not have the aforementioned limitations. We accounted for changes in the demographic profile of all countries by using age and sex specific population estimates.

A unique strength of the database is the use of the authoritative national agencies for data collection, extensive checks and data validation procedures, comparability across countries and time, highly detailed and uniform data formats, and availability of data on survival to the highest ages. However, we were limited to an analysis of countries that reported weekly deaths by age and sex for the entire study period of 2016–20. We acknowledge that no consensus exists in the literature on the historical time period for the estimation of the counterfactual expected deaths. We were also limited by the completeness of the data (for example, Canada and Australia have not yet reported complete data in 2020).

As highlighted previously, the variability around the estimated excess death in children <15 years is larger than in other age groups owing to sparse data.²⁴ To make the analysis comparable, we presented our findings by five broader age groups. The 15–64 years age group represents the working age group, but it also consists of heterogeneous age groups in terms of risk of mortality. In an exploratory analysis, the rate of excess deaths was smaller in age groups <45 years but varied markedly in age groups 45–54 and 55–64 years (supplementary figure S6). Excess risk was also relatively low in females across all age groups (except in Hungary, where the rate was higher in women aged 55–64 years). Belgium, the Czech Republic, Hungary, Poland, and Scotland had the highest excess death rates in men aged 55–64 years (exceeding 100 excess deaths per 100 000). The aggregated analysis essentially masks these nuanced trends in mortality rates.

We did not have the data to examine the variability in excess deaths by other potentially critical factors such as ethnicity and socioeconomic status.^{45–59–76–81} We also did not examine within country regional differences in mortality owing to unavailability of such granular data. Even though we assumed a fairly stable population during 2020 within each country, some in-country migration or displacement could occur as a result of changes in job or financial situations. Although this assumption will more likely hold true for

most countries, an underestimation of such population displacement may have resulted in an underestimation of excess deaths.²³ Our analysis is restricted to the end of 2020 to make our estimates by calendar year comparable to the conventional mortality statistics reported by international agencies including the World Health Organization and the United Nations. Therefore, our study does not reflect the excess deaths experienced by some countries in early 2021. The indirect effects of a pandemic and its associated policy measures are multifactorial.¹³ Many of these may need a longer timeframe to have a measurable effect on mortality. Our analysis was unable to measure these effects. Also, we did not have access to individual patient level data to explore the cause of deaths, which would provide important insights about the sources and potential mechanisms of the direct and indirect effects of the pandemic and its associated policy measures.¹³ Although no visible mortality displacement occurred in our study, mortality displacement in one group may be masked by the ongoing excess mortality in other population groups. More studies with longer follow-up data are needed to disentangle this phenomenon.

Policy implications and future directions

Our study adds important insights on the direct and indirect effects of the covid-19 pandemic on total mortality. It underscores the importance of availability of age and sex disaggregated data for more nuanced analysis and estimation of the direct and indirect effect of the pandemic. A lack of detailed data from lower and middle income countries, especially those in Asia and Africa, calls for a globally coordinated effort to improve the local capacity in collecting and reporting critical vital statistics data promptly to aid evidence based healthcare policy decisions. New Zealand stood out as the only country that had a lower than expected mortality across all the age groups, in both men and women, with no sex difference in excess death rates, which could potentially be attributed to the country's elimination strategy early in the pandemic.^{52 53} Our findings also suggest that many countries had an underestimation or underreporting of covid-19 deaths, a substantial increase in non-covid-19 deaths, or both. Reliable and timely monitoring of excess deaths would help to inform public health policy in investigating the sources of excess mortality in populations and would help to detect important social inequalities in the impact of the pandemic to inform more targeted interventions.⁸²

Our study also highlights the need for nuanced analysis taking into account the other potential sources of social inequalities including ethnicity and socioeconomic status. This study provides a detailed and robust assessment of the impact of the covid-19 pandemic on total mortality up to the point that mass vaccination programmes started to become widely available throughout these countries. Future work will be needed to understand the impact of national vaccination programmes on mortality in 2021.

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Data sharing: All the data used in this study are publicly available and properly cited. However, more guided instruction to get access to the data for transparency and reproducibility will be provided on request.

The lead author affirms that the manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned (and, if relevant, registered) have been explained.

Dissemination to participants and related patient and public communities: We will widely disseminate the main findings to members of the public through official (press release, institutional websites, and repositories), personal, and social media. We also plan to write a BMJ Opinion article to describe it in more general terms for the members of the public.

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Web appendix: Supplementary materials

Excess deaths associated with covid-19 pandemic in 2020: age and sex disaggregated time series analysis in 29 high income countries

Supplementary Appendix

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Glossary

Observed deaths: Deaths reported by the individual countries.

Expected deaths: Deaths that would be expected in 2020 based on historical data (2016-2019), separately calculated by country, calendar year, week of the year, age, and sex. These numbers are calculated from our model after taking into account the recent temporal trend and seasonal variation in mortality.

Excess deaths: Observed minus expected deaths.

Avoided mortality: When the observed deaths are below the expected deaths. Also known as “mortality deficit”.

Mortality displacement: When a period of excess deaths is followed by a mortality deficit.

This indicates that people who have died would have died even in the absence of the event (Covid-19 pandemic, in this case), but their death was brought forward by the event.

Supplementary Methods

Statistical model:

Denote $Y_{s,c,a}(t)$ as the number of deaths at week t for individuals of sex s , in country c , and age group a . Assume that $Y_{s,c,a}(t) \sim \text{Poisson}(\mu_{s,c,a}(t))$, then our mean model is:

$$\mu_{s,c,a}(t) = N_{s,c,a}(t) \exp\{\beta t + g_{s,c,a}(w_t)\} \text{ for } t \in I_c \quad (1)$$

The expected number of deaths at week t for individuals of sex s , in country c , and age group a is represented with $\mu_{s,c,a}(t)$, $N_{s,c,a}(t)$ is an offset that accounts for the population size, β represents a linear effect of time that accounts for slow-moving changes in mortality, $g_{s,c,a}(w_t)$ is a function that accounts for seasonal trends, where $w_t \in \{1, \dots, 52\}$ represents a week of the year, and I_c is a country-specific training interval that is used to fit the model (Supplemental Table S2). For the function $g_{s,c,a}(t)$, we use a Fourier transform with two harmonics:

$$g_{s,c,a}(w_t) = \sum_{k=1}^2 A_{s,c,a,k} \times \cos\left(\frac{2\pi k w_t}{52}\right) + B_{s,c,a,k} \times \sin\left(\frac{2\pi k w_t}{52}\right) \quad (2)$$

, where the $\mathbf{A} = [A_{s,c,a,1}, A_{s,c,a,2}]$ and $\mathbf{B} = [B_{s,c,a,1}, B_{s,c,a,2}]$ parameters are estimated from data. We fit model (1) via Maximum Likelihood and estimated the counterfactual death counts in 2020 for each sex-age group in each country.

Now, we turn the focus to obtaining smooth estimates of percent change from average. Let t' be a week in 2020, then:

$$\lambda_{s,c,a}(t') = \hat{\mu}_{s,c,a}(t') \exp\{f_{s,c,a}(t')\} \quad (3)$$

In model (3), $\lambda_{s,c,a}(t')$ represents the average number of observed deaths at week t' for individuals of sex s , in country c , and age group a . The number of deaths at week t' in the counterfactual scenario of no pandemic is represented with $\hat{\mu}_{s,c,a}(t')$ and it is used as an offset. Lastly, the function $f_{s,c,a}(t')$ is a natural cubic spline with 12 internal knots. It follows that $\hat{\gamma}_{s,c,a}(t') = \hat{\lambda}_{s,c,a}(t') / \hat{\mu}_{s,c,a}(t') - 1$ represents a smooth estimate of percent change from average at t' . Using a first-order Taylor approximation and the assumptions from above we have that:

$$\text{Var}(\hat{\gamma}_{s,c,a}(t')) \approx \frac{\hat{\lambda}_{s,c,a}^2(t')}{\hat{\mu}_{s,c,a}^2(t')} \left(\text{Var}(\log \hat{\lambda}_{s,c,a}(t')) + \hat{\mu}_{s,c,a}^{-2}(t') \text{Var}(\log \hat{\mu}_{s,c,a}(t')) \right) \quad (4)$$

Furthermore, we computed excess deaths at t' with:

$$\hat{\Delta}_{s,c,a}(t') = Y_{s,c,a}(t') - \hat{\mu}_{s,c,a}(t') \quad (5)$$

Using a first-order Taylor approximation and the assumptions from above we have that:

$$Var(\hat{\Delta}_{s,c,a}(t')) \approx (\hat{\phi}_{s,c,a} \hat{\lambda}_{s,c,a}(t') + \hat{\mu}_{s,c,a}(t') Var(\log \hat{\mu}_{s,c,a}(t'))) \quad (6)$$

, where $\hat{\phi}_{s,c,a}$ is the estimated dispersion parameter from model (3).

Finally, to get marginal effects for males and females in each country, we took a weighted average of the age-specific estimates. Specifically, the percent change from average for individuals of sex s in country c is:

$$\hat{\gamma}_{s,c}(t') = \frac{\sum_{a \in Age\ groups} \hat{\mu}_{s,c,a}(t') \times \hat{\gamma}_{s,c,a}(t')}{\sum_{a \in Age\ groups} \hat{\mu}_{s,c,a}(t')} = \sum_{a \in Age\ groups} \hat{\pi}_{s,c,a}(t') \times \hat{\gamma}_{s,c,a}(t') \quad (7)$$

, where $\hat{\pi}_{s,c,a}(t') = 1 / \sum_{a \in Age\ groups} \hat{\mu}_{s,c,a}(t')$. Note that we can obtain variability estimates for $\hat{\gamma}_{s,c}(t')$ in equation (3) using a first-order Taylor approximation.

Discrepancies between the reported Covid-19 deaths and the estimated excess deaths:

Many countries reported their official statistics on Covid-19 deaths. Previous research highlighted that there could be under-reporting of Covid-19 due to a host of reasons including a difference in the definition of Covid-19 deaths, inadequate testing facility to precisely diagnose the cause of death, some degrees of diagnostic inaccuracy in the diagnostic tests used to detect the SARS-CoV-2 virus, delay in reporting etc. Therefore, when we see that the reported number of Covid-19 deaths is lower than the estimated excess deaths, it is understood more intuitively. However, a higher excess death does not necessarily mean that there was an under-reporting or underestimation of the Covid-19 deaths. It could also be due to an increase in mortality due to causes other than Covid-19 (for example, more deaths due to delayed medical care, suicide, domestic violence etc. potentially associated with the disruption of healthcare services due to prolonged lockdown). In reality the difference is more likely due to a mix of these factors.

In contrast, it is less intuitive to understand why Covid-19 death counts could be higher than the estimated excess deaths. For example, if we conclude that the estimated number of excess deaths in a specific country was 1,000, but the country already has reported, let's say, 1,500 deaths due to Covid-19. So, how is it possible that the excess death is lower than the

reported Covid-19 deaths? Aren't all deaths due to Covid-19 already part of the 'excess' deaths? The answer lies in the estimation process.

Death and death rates are always non-negative (zero or more). However, because 'excess death' is defined as a difference between the observed and the expected number of deaths (observed minus expected, to be precise), it can be (and indeed was) negative in many countries (when the observed death was lower than the expected death), especially in specific subgroups (mostly in younger age groups). When we add negative numbers, the total would be lower than the sum of all the positive numbers. For example, if we have 300 excess deaths each in the four older age groups, we already have $4 \times 300 = 1,200$ excess deaths. But if there were 200 fewer deaths in the youngest age groups, then altogether we will have 1,000 deaths ($1,200 - 200 = 1,000$), which is lower than the total of the other four age groups!

Lastly, overall excess death is a combination of Covid-19 and non-Covid deaths. For example, let's say, we have 1,500 deaths from Covid-19, but 500 fewer deaths than expected for other non-Covid deaths (maybe due to lower road traffic accidents, or lower occupational hazards, or lower levels of pollution potentially associated with lockdown measures). The overall excess deaths would be 1,000 ($1,500 - 500 = 1,000$), which is lower than the reported Covid-19 deaths.

Population estimates

For each stratum, we assumed a constant rate of change in the population size between the mid-year estimates. Hence, from week 26, 2016, to week 26, 2020, we calculated population sizes via linear interpolation of the mid-year population estimates. For the first six months of 2016, we assumed the same rate of change as in mid-2016 to mid-2017, thus we extrapolated the population back in time using a linear model. Similarly, for the last six months of 2020, we assumed the same rate of change as in mid-2019 to mid-2020 and extrapolated using a linear model.

Crude and age-standardised annual mortality rates

We examined the annual mortality rate (per 100,000 population) in each of the aggregate age groups and for all ages, separately in men and women. To address the differential demographic structure across all the countries included in our analysis, we used age-standardised mortality rate (per 100,000 population) calculated by the direct method of age-

standardisation with the 2013 European Standard Population.[1] Since the death counts, especially in the youngest age groups (<15 years) could be very low in some countries, we applied the methods proposed by Fay and Feuer (modified by Tiwari et al.) to calculate the confidence intervals for age-standardised rates between 2016 and 2020, separately in men and women.[2,3]

Table S1: Sources of the Short-Term Mortality Fluctuation data for 29 high-income countries

Country code	Country name	Data source
AUT	Austria	Statistik Austria
BEL	Belgium	STATBEL and Eurostat
CHE	Switzerland	Eurostat and Swiss Federal Statistical Office
CZE	Czech Republic	Statistical Office and Eurostat
DEU	Germany	Statistisches Bundesamt
DNK	Denmark	Statistics Denmark
ESP	Spain	National Statistical Office
EST	Estonia	Eurostat
FIN	Finland	Statistics Finland
FRA	France	INSERM
ENW	England and Wales	Office for National Statistics
NIR	Northern Ireland	Eurostat
SCO	Scotland	National Statistical Office
GRC	Greece	Eurostat
HUN	Hungary	Eurostat
ISR	Israel	Stat Office
ITA	Italy	ISTAT
KOR	South Korea	Statistics Korea
LTU	Lithuania	Statistics Lithuania
LVA	Latvia	Eurostat
NLD	Netherlands	Statistics Netherlands
NOR	Norway	Statistics Norway
NZL	New Zealand	Statistics New Zealand
POL	Poland	Eurostat
PRT	Portugal	Direção-Geral da Saúde (General Directorate of Health)
SVK	Slovakia	Eurostat
SVN	Slovenia	Eurostat
SWE	Sweden	Statistics Sweden
USA	USA	Centers for Disease Control and Prevention

For details please see https://www.mortality.org/Public/STMF_DOC/STMFNote.pdf. The data was further harmonised for the purpose of this analysis (described in the main paper).

Table S2: Dates excluded from the counterfactual estimation for 2020 in 29 high-income countries

Country	Excluded dates
Austria, Belgium, Czech Republic, England & Wales, Estonia, Finland, Germany, Northern Ireland, Slovakia, Norway, Poland, Sweden, Switzerland, and Netherlands	2016-12-01 to 2017-04-30, & 2017-12-01 to 2018-04-30
Greece, Portugal	2016-12-01 to 2017-04-30, & 2018-12-01 to 2019-04-30
Hungary, Lithuania, Slovenia, Spain	2016-12-01 to 2017-04-30, 2017-12-01 to 2018-04-30, & 2018-12-01 to 2019-04-30
Denmark, Latvia, United States, Scotland, South Korea	2017-12-01 to 2018-04-30
France, Israel, Italy	2016-12-01 to 2017-04-30
New Zealand	2017-06-01 to 2017-08-31

Table S3: Estimated number (95% confidence interval) of excess deaths by age and sex in 2020 in 29 high-income countries										
	Estimated number of excess deaths in 2020 (95% confidence interval)									
	0 to 14 years		15 to 64 years		65 to 74 years		75 to 84 years		85 or above years	
Country	Female	Male	Female	Male	Female	Male	Female	Male	Female	Male
Austria	0 (-20 to 30)	-10 (-30 to 20)	170 (40 to 290)	570 (400 to 740)	140 (0 to 280)	650 (480 to 830)	1,300 (1,100 to 1,500)	1,600 (1,400 to 1,800)	1,300 (1,100 to 1,600)	1,000 (790 to 1,200)
Belgium	-90 (-120 to -50)	-140 (-180 to -110)	250 (100 to 400)	1,200 (1,000 to 1,400)	530 (360 to 690)	1,600 (1,400 to 1,800)	2,600 (2,400 to 2,800)	3,100 (2,900 to 3,300)	5,500 (5,200 to 5,900)	3,200 (3,000 to 3,500)
Czech Republic	-30 (-60 to -10)	-10 (-40 to 20)	270 (120 to 420)	760 (540 to 980)	960 (760 to 1,200)	2,200 (1,900 to 2,400)	2,500 (2,200 to 2,700)	3,200 (3,000 to 3,500)	2,700 (2,400 to 3,000)	1,800 (1,600 to 2,000)
Denmark	0 (-20 to 20)	20 (-10 to 40)	100 (0 to 200)	-100 (-240 to 30)	-40 (-170 to 80)	-100 (-250 to 60)	50 (-120 to 220)	-30 (-210 to 160)	-220 (-440 to -10)	170 (0 to 340)
England & Wales	-130 (-200 to -50)	-90 (-180 to -10)	3,600 (3,300 to 4,000)	6,600 (6,200 to 7,000)	4,400 (4,000 to 4,800)	8,000 (7,600 to 8,500)	11,400 (10,900 to 12,000)	15,500 (15,000 to 16,100)	21,100 (20,400 to 21,700)	14,900 (14,300 to 15,400)
Estonia	0 (-10 to 10)	-10 (-20 to 0)	140 (90 to 190)	270 (180 to 360)	10 (-50 to 80)	-40 (-130 to 40)	20 (-80 to 120)	60 (-20 to 150)	170 (50 to 290)	40 (-30 to 110)
Finland	-20 (-40 to 0)	-10 (-30 to 10)	-40 (-130 to 60)	380 (250 to 520)	170 (50 to 290)	80 (-70 to 240)	20 (-140 to 180)	200 (30 to 380)	190 (-40 to 420)	30 (-140 to 200)
France	-120 (-200 to -50)	-270 (-360 to -180)	180 (-170 to 530)	2,000 (1,500 to 2,500)	1,600 (1,300 to 2,000)	4,600 (4,100 to 5,100)	4,800 (4,300 to 5,200)	8,400 (7,800 to 8,900)	12,500 (11,600 to 13,300)	9,900 (9,300 to 10,500)
Germany	90 (20 to 170)	40 (-40 to 120)	340 (-80 to 760)	3,600 (3,000 to 4,100)	590 (140 to 1,000)	2,700 (2,200 to 3,300)	3,200 (2,400 to 3,900)	5,700 (4,900 to 6,500)	4,000 (3,100 to 5,000)	5,600 (4,900 to 6,300)
Greece	-40 (-70 to -10)	30 (0 to 60)	480 (340 to 610)	270 (70 to 470)	480 (330 to 640)	610 (400 to 820)	1,800 (1,500 to 2,000)	1,200 (890 to 1,400)	1,400 (1,100 to 1,800)	1,500 (1,200 to 1,800)
Hungary	-10 (-40 to 20)	-10 (-40 to 30)	820 (640 to 1,000)	1,500 (1,300 to 1,800)	1,900 (1,700 to 2,100)	2,500 (2,300 to 2,800)	3,100 (2,800 to 3,400)	2,600 (2,300 to 2,800)	2,700 (2,400 to 3,000)	1,400 (1,200 to 1,600)
Israel	-50 (-80 to -10)	-90 (-130 to -40)	60 (-40 to 160)	120 (-10 to 260)	20 (-100 to 130)	480 (350 to 620)	180 (30 to 330)	420 (270 to 580)	480 (280 to 690)	380 (210 to 550)
Italy	-40 (-100 to 10)	-70 (-130 to -10)	1,400 (1,100 to 1,700)	4,400 (4,000 to 4,800)	4,200 (3,900 to 4,600)	11,300 (10,800 to 11,700)	11,800 (11,200 to 12,300)	18,700 (18,100 to 19,300)	22,800 (21,900 to 23,700)	14,600 (14,000 to 15,300)
Latvia	10 (-10 to 20)	10 (0 to 30)	-30 (-120 to 50)	-90 (-210 to 40)	30 (-60 to 120)	160 (50 to 270)	340 (210 to 480)	20 (-100 to 140)	200 (50 to 350)	160 (70 to 240)
Lithuania	-10 (-20 to 10)	10 (-10 to 30)	350 (250 to 440)	1,200 (1,000 to 1,300)	370 (270 to 470)	970 (850 to 1,100)	1,200 (1,000 to 1,300)	1,100 (930 to 1,200)	1,300 (1,100 to 1,400)	480 (370 to 590)
Netherlands	20 (-20 to 50)	0 (-40 to 40)	210 (30 to 390)	730 (520 to 940)	670 (460 to 880)	1,700 (1,500 to 2,000)	1,900 (1,600 to 2,100)	3,600 (3,300 to 3,900)	3,800 (3,400 to 4,200)	2,700 (2,400 to 3,000)
New Zealand	-40 (-70 to -20)	-50 (-70 to -20)	-70 (-170 to 30)	-330 (-450 to -200)	-230 (-330 to -130)	-200 (-320 to -80)	-160 (-290 to -40)	-290 (-430 to -150)	-760 (-930 to -590)	-380 (-520 to -230)
Northern Ireland	-10 (-30 to 0)	-20 (-40 to 0)	70 (0 to 130)	170 (90 to 250)	230 (170 to 290)	260 (190 to 330)	330 (240 to 420)	400 (310 to 500)	540 (430 to 650)	220 (130 to 310)
Norway	-10 (-20 to 10)	-30 (-50 to 0)	30 (-60 to 110)	0 (-110 to 110)	-10 (-120 to 90)	150 (20 to 270)	-180 (-320 to -40)	-50 (-200 to 100)	-10 (-210 to 190)	20 (-140 to 180)
Poland	-140 (-200 to -80)	-30 (-90 to 30)	860 (540 to 1,200)	4,500 (4,000 to 5,000)	4,800 (4,400 to 5,200)	10,900 (10,400 to 11,400)	6,100 (5,600 to 6,500)	10,000 (9,500 to 10,400)	13,100 (12,600 to 13,700)	10,000 (9,600 to 10,400)
Portugal	-20 (-50 to 10)	-40 (-70 to -10)	260 (120 to 390)	740 (530 to 940)	600 (450 to 740)	710 (510 to 910)	1,700 (1,500 to 1,900)	1,500 (1,200 to 1,700)	2,000 (1,600 to 2,300)	1,100 (830 to 1,400)
Scotland	0 (-20 to 20)	-10 (-40 to 10)	120 (0 to 250)	870 (720 to 1,000)	400 (270 to 530)	870 (730 to 1,000)	890 (710 to 1,100)	1,100 (950 to 1,300)	1,500 (1,300 to 1,700)	1,000 (870 to 1,200)
Slovakia	-20 (-40 to 10)	-10 (-40 to 20)	60 (-60 to 180)	130 (-50 to 310)	440 (300 to 580)	650 (470 to 820)	900 (720 to 1,100)	1,100 (930 to 1,200)	710 (520 to 900)	480 (350 to 600)
Slovenia	0 (-10 to 10)	0 (-10 to 10)	30 (-30 to 90)	210 (120 to 300)	40 (-30 to 110)	130 (30 to 230)	530 (430 to 630)	490 (380 to 600)	1,000 (900 to 1,200)	760 (670 to 850)
South Korea	-120 (-170 to -70)	-130 (-190 to -80)	730 (460 to 990)	-650 (-1,100 to -230)	130 (-100 to 370)	800 (440 to 1,100)	1,200 (810 to 1,600)	1,300 (860 to 1,800)	220 (-270 to 710)	480 (140 to 820)
Spain	20 (-30 to 70)	60 (10 to 120)	1,800 (1,500 to 2,000)	3,100 (2,700 to 3,500)	3,400 (3,100 to 3,600)	6,800 (6,400 to 7,200)	11,200 (10,800 to 11,600)	14,200 (13,700 to 14,600)	26,600 (26,000 to 27,300)	17,000 (16,500 to 17,500)
Sweden	10 (-10 to 40)	50 (30 to 80)	70 (-50 to 180)	490 (350 to 630)	280 (130 to 420)	640 (470 to 820)	1,300 (1,100 to 1,500)	1,900 (1,700 to 2,200)	2,400 (2,100 to 2,700)	2,000 (1,800 to 2,300)
Switzerland	10 (-20 to 40)	30 (0 to 60)	10 (-100 to 120)	260 (130 to 400)	50 (-70 to 170)	290 (140 to 440)	570 (390 to 740)	1,300 (1,100 to 1,400)	2,200 (1,900 to 2,500)	2,200 (2,000 to 2,400)
United States	-360 (-580 to -140)	-750 (-1,000 to -500)	39,500 (38,500 to 40,500)	91,500 (90,200 to 92,800)	37,200 (36,200 to 38,100)	57,100 (55,900 to 58,200)	52,700 (51,600 to 53,900)	56,700 (55,500 to 57,900)	83,500 (82,100 to 85,000)	40,500 (39,400 to 41,700)

Table S4: Median absolute difference of the excess mortality model (2016-2019)

Country	Median absolute difference	
	Female	Male
Austria	0.056	0.055
Belgium	0.060	0.049
Switzerland	0.065	0.062
Czech Republic	0.050	0.054
Germany	0.029	0.031
Denmark	0.059	0.064
England & Wales	0.038	0.041
Spain	0.041	0.035
Estonia	0.123	0.109
Finland	0.066	0.073
France	0.036	0.029
Greece	0.069	0.059
Hungary	0.055	0.062
Israel	0.064	0.049
Italy	0.045	0.042
South Korea	0.038	0.034
Lithuania	0.090	0.086
Latvia	0.075	0.081
Northern Ireland	0.112	0.119
Netherlands	0.050	0.041
Norway	0.074	0.068
New Zealand	0.061	0.053
Poland	0.038	0.040
Portugal	0.068	0.060
Scotland	0.067	0.062
Slovakia	0.078	0.068
Slovenia	0.109	0.106
Sweden	0.056	0.061
United States	0.013	0.013

Figure S1: Monthly excess death rates (per 100,000) by age and sex in 2020 in 29 high-income countries

Fig. S1 (A): <15 years

Note: empty (zero length) bars indicate no excess deaths in those months.



Fig. S1 (B): 15-64 years

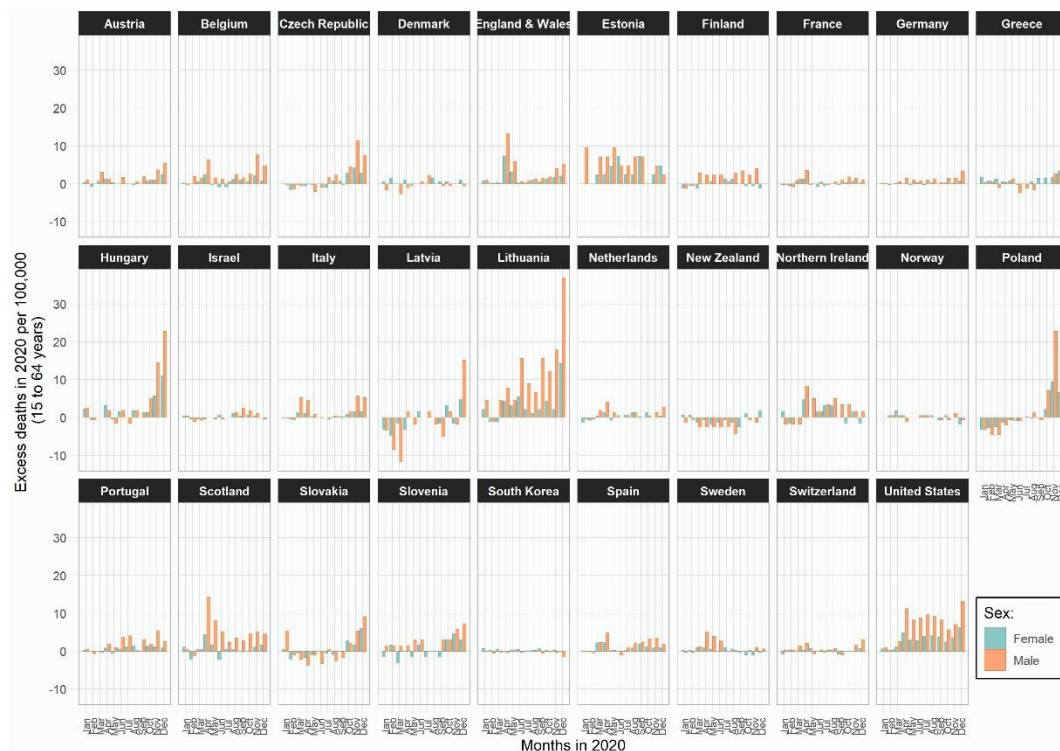


Fig. S1 (C): 65-74 years

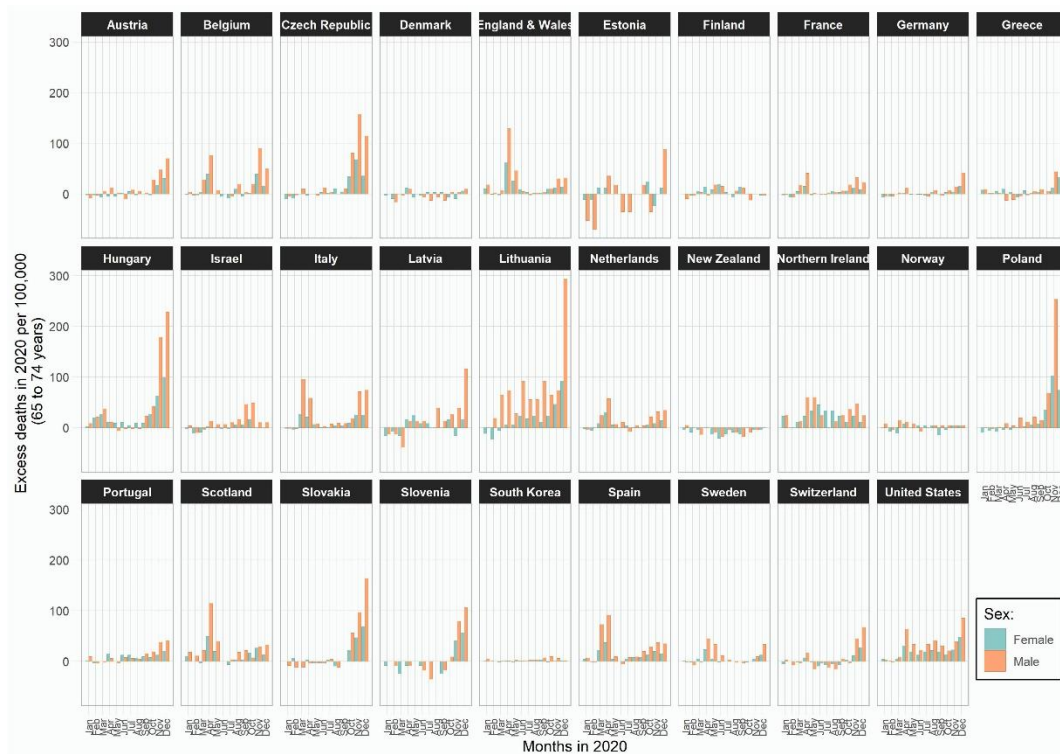


Fig. S1 (D): 75-84 years

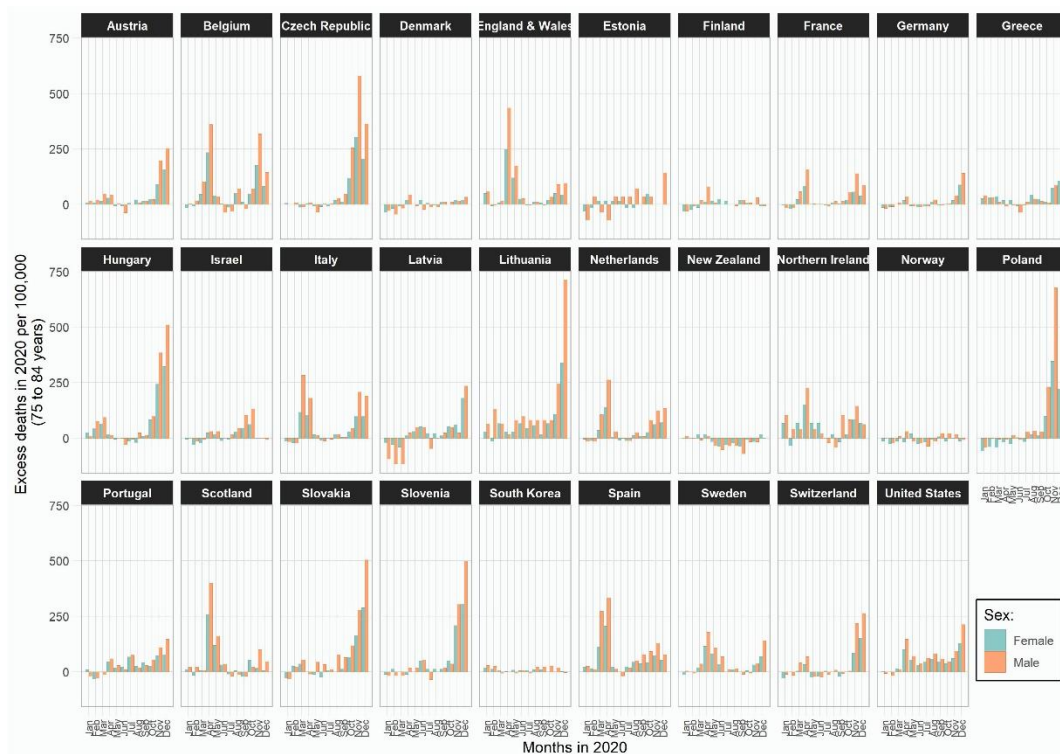


Fig. S1 (E): 85 years or above



Figure S2: Percent change in weekly excess deaths in 2020 by age and sex in 29 high-income countries

Fig. S2 (A): <15 years

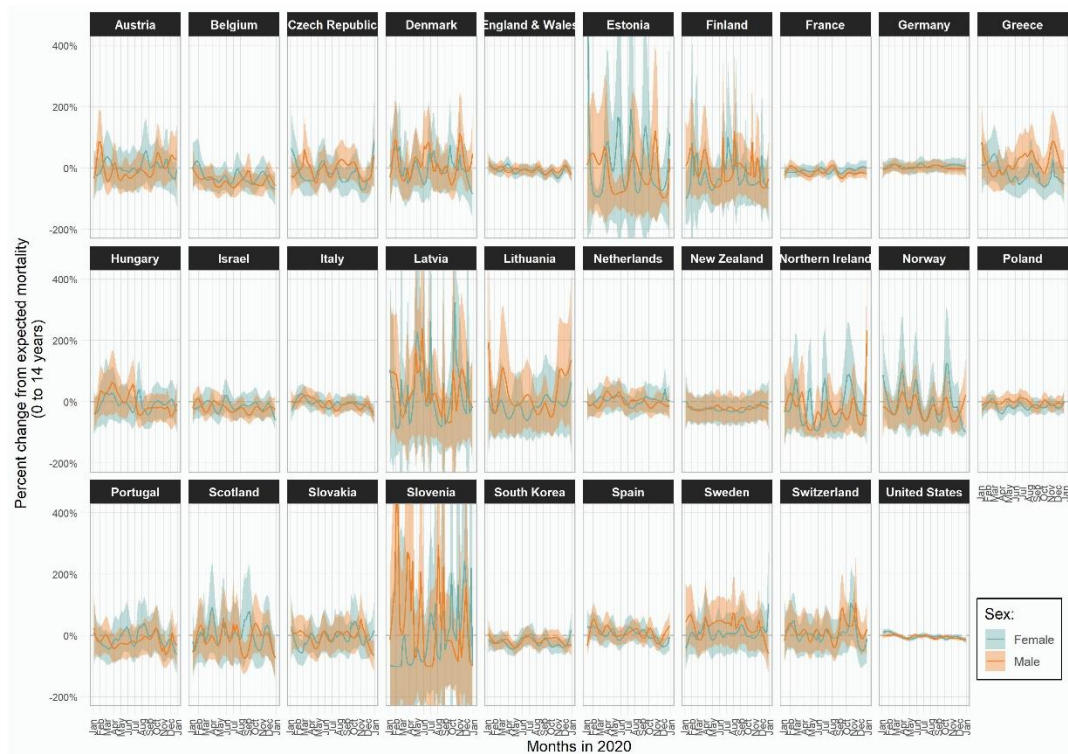


Fig. S2 (B): 15-64 years

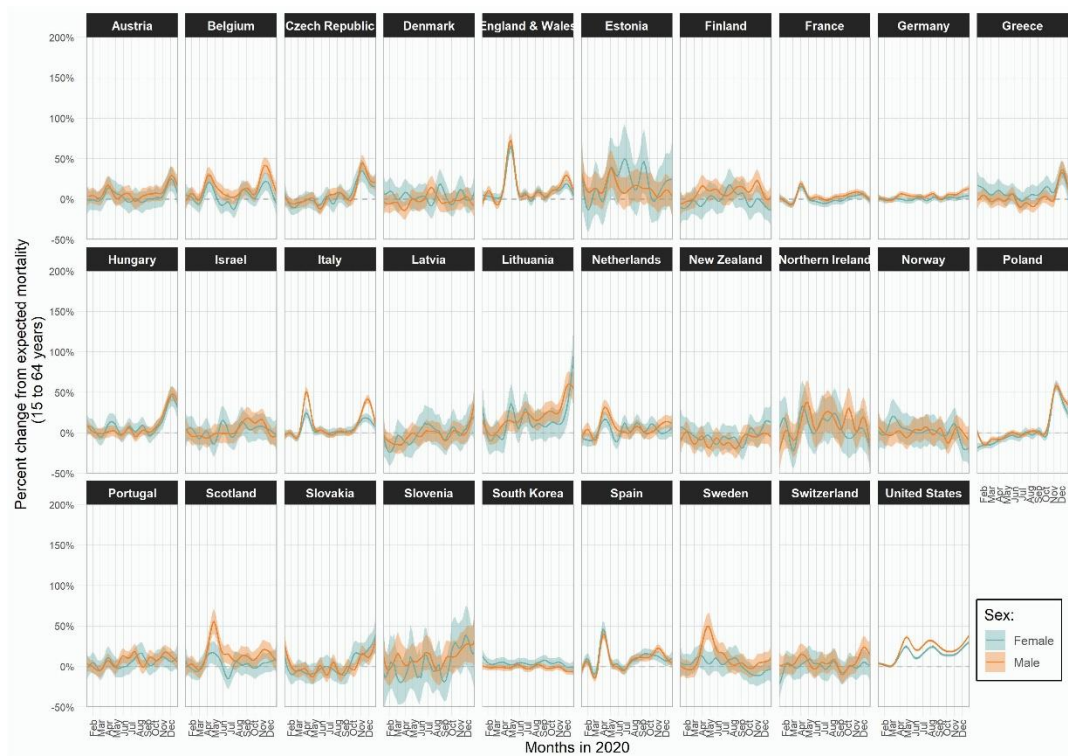


Fig. S2 (C): 65-74 years

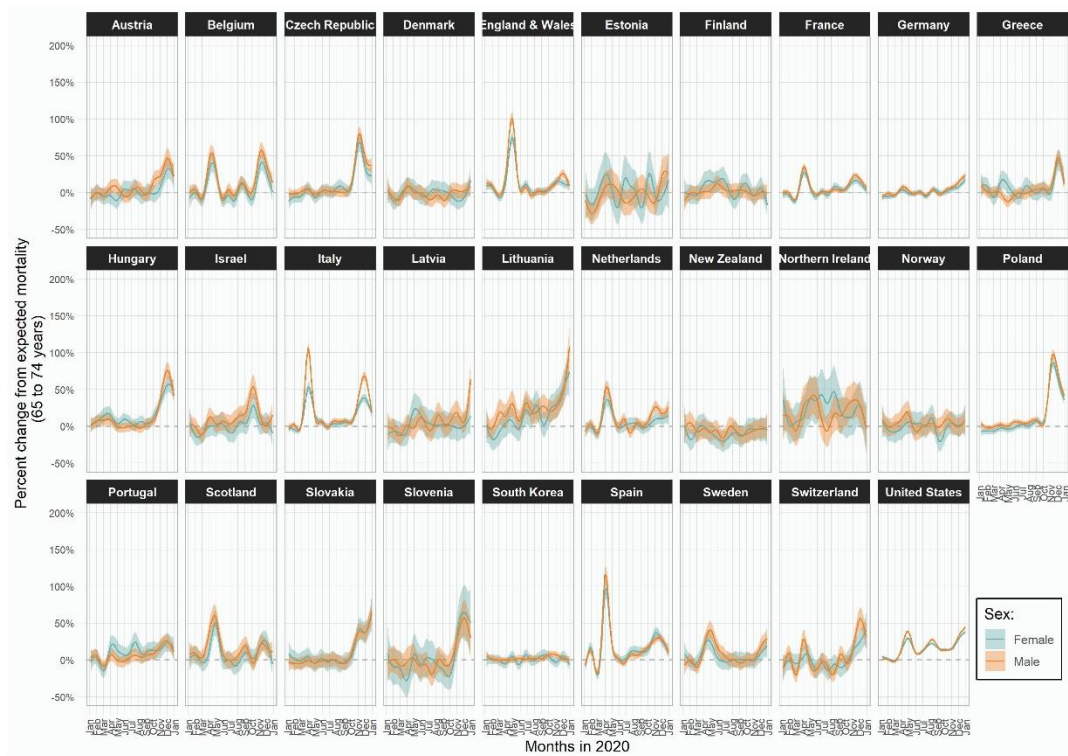


Fig. S2 (D): 75-84 years

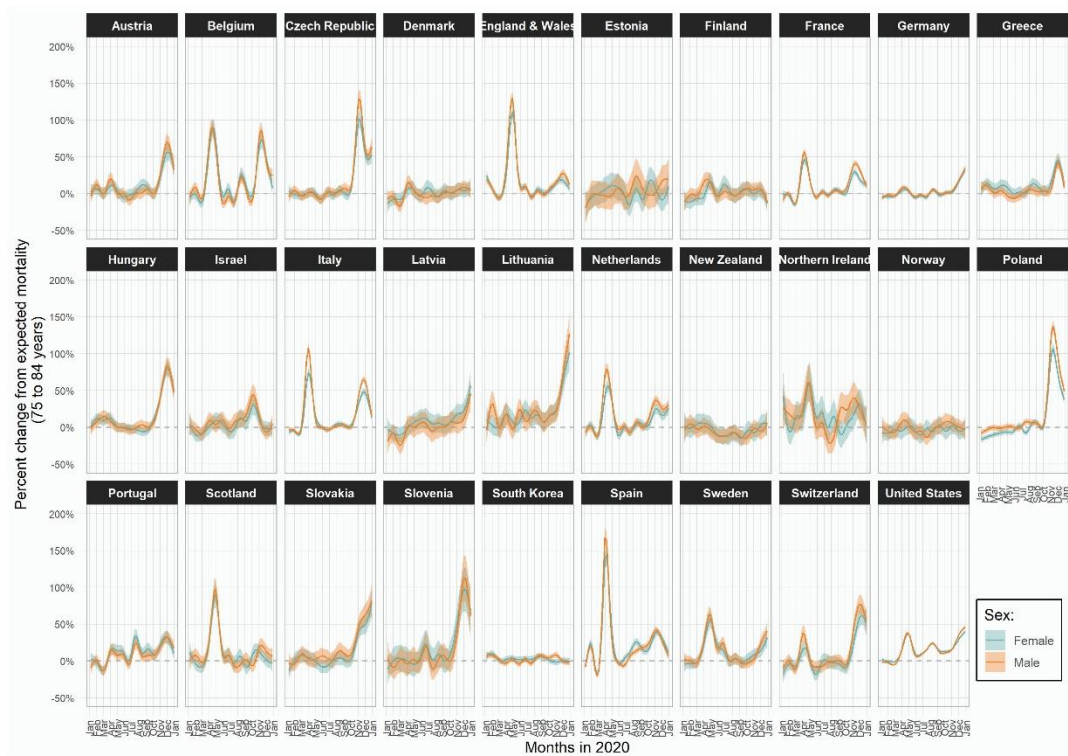


Fig. S2 (E): 85 years or above

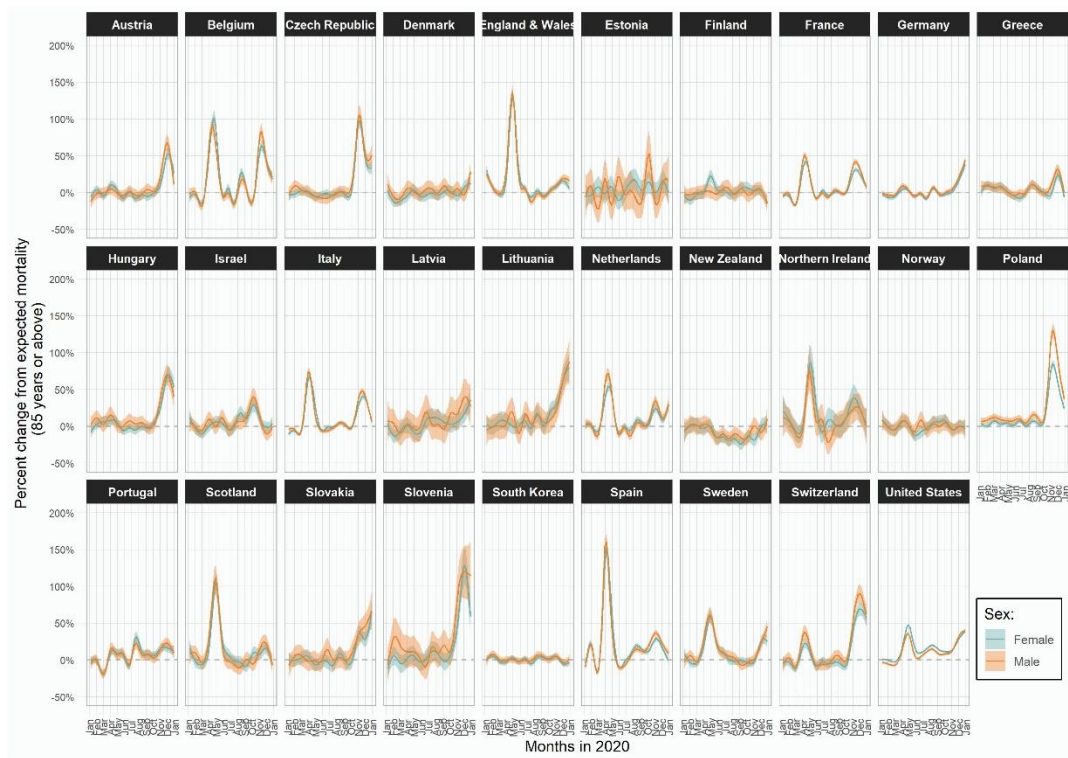


Figure S3: Age-standardised annual mortality rate in men and women in 29 high-income countries, 2016-2020

Fig. S3 (A): Age-standardised annual mortality rate in men and women

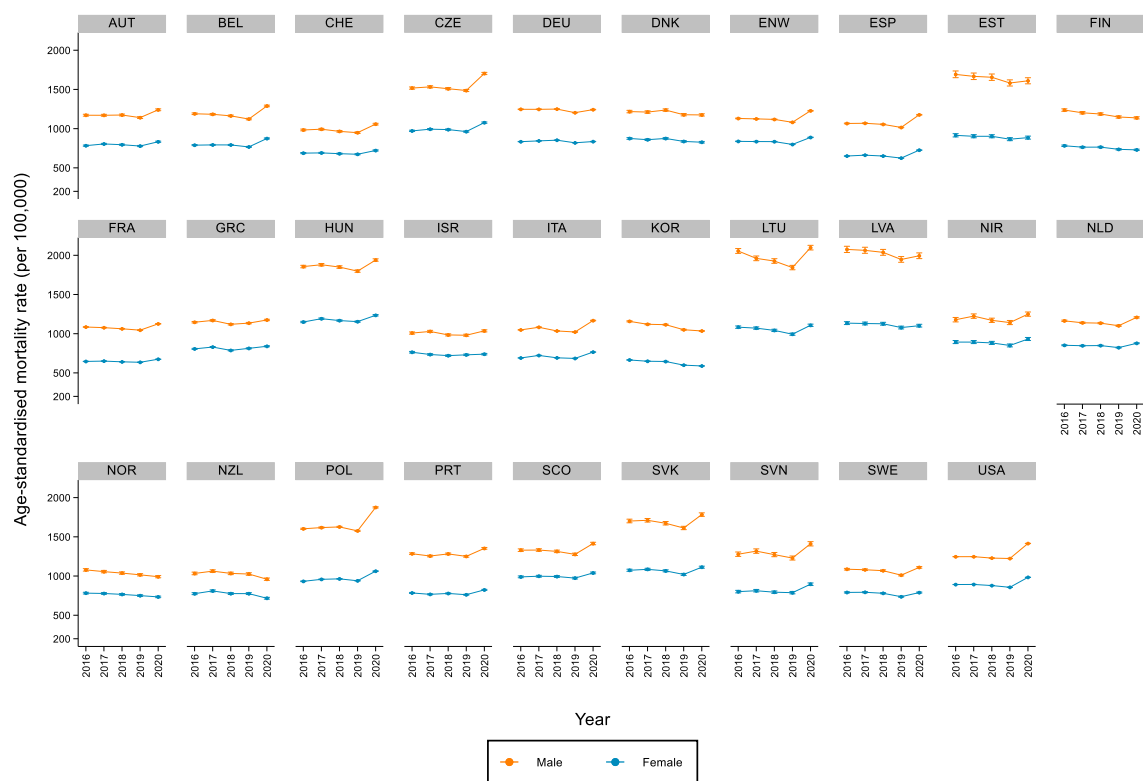
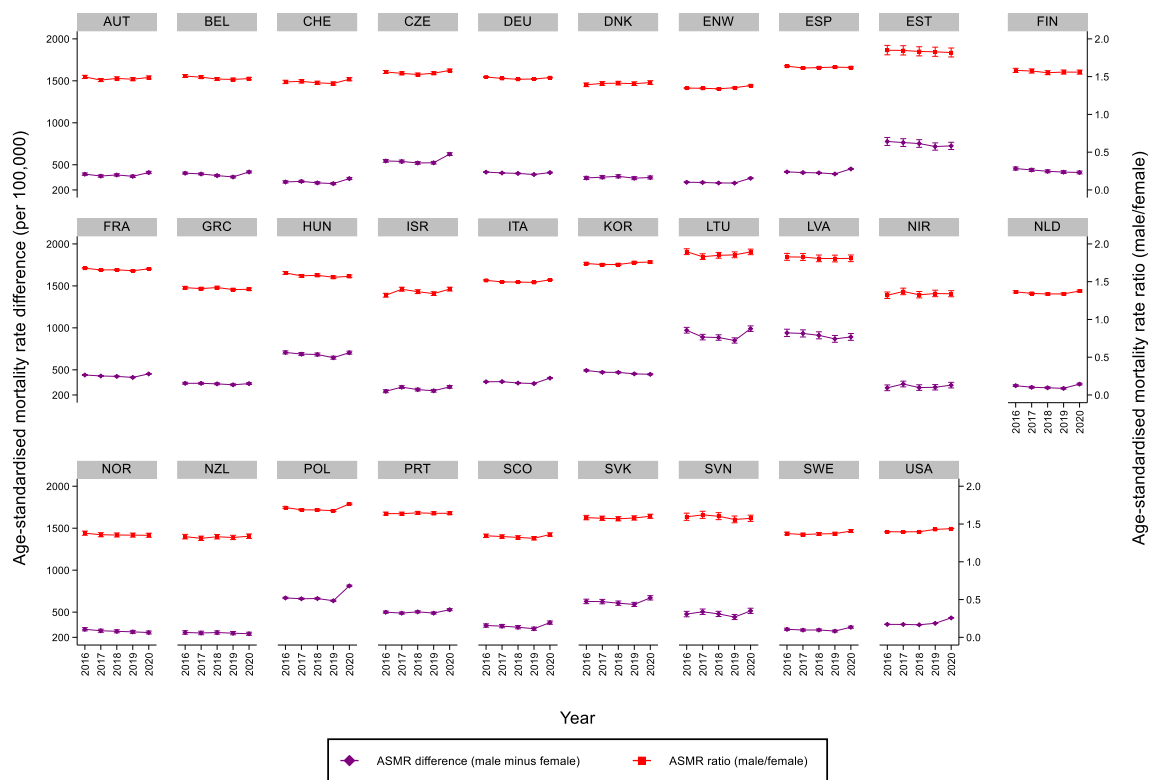


Fig. S3 (B): Age-standardised annual mortality rate difference and ratio between men and women



-ASMR: Age-standardised mortality rate

List of country code and the country names is available in the Supplementary Table S1. The death rates were standardised to the 2013 European Standard Population.

Figure S4: Crude annual mortality rate in 29 high-income countries, 2016-2020
Fig. S4 (A): <15 years

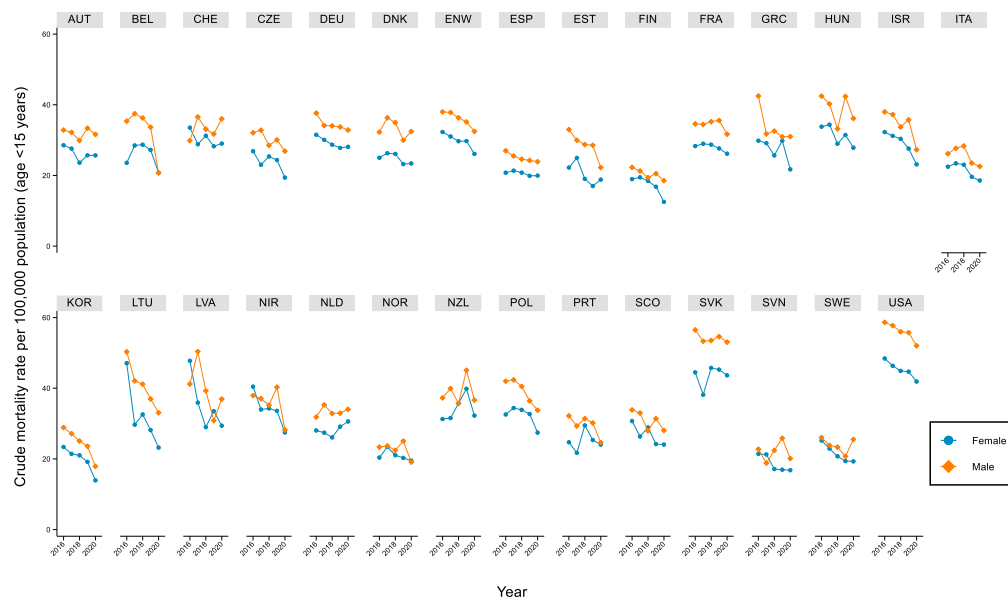


Fig. S4 (B): 15-64 years

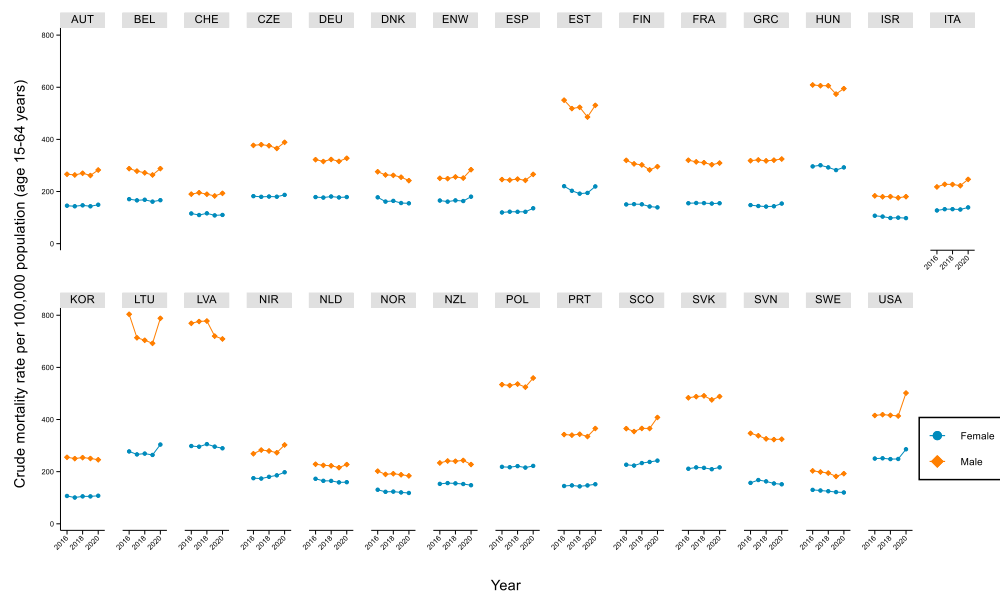


Fig. S4 (C): 65-74 years

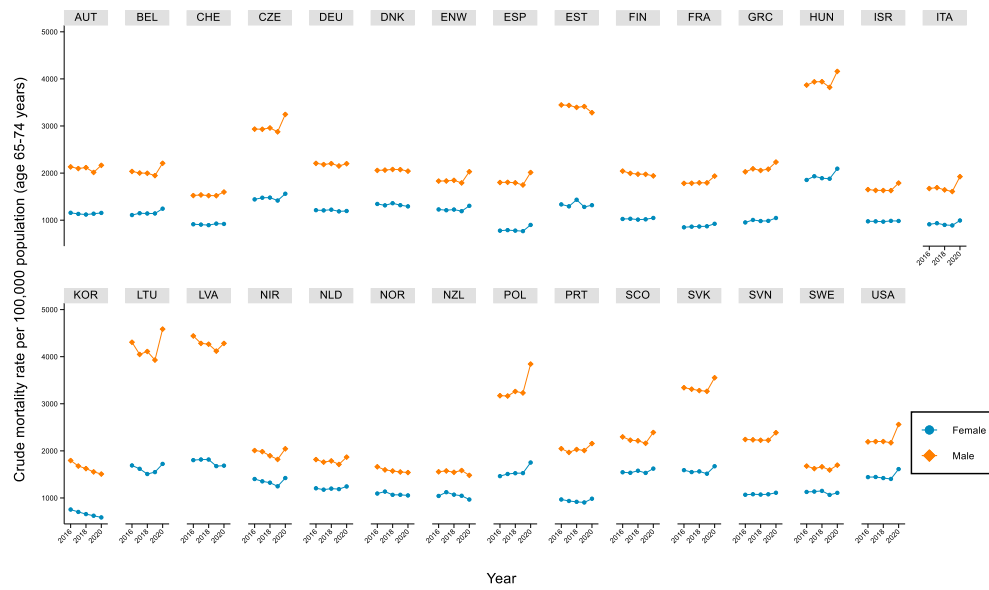


Fig. S4 (D): 75-84 years

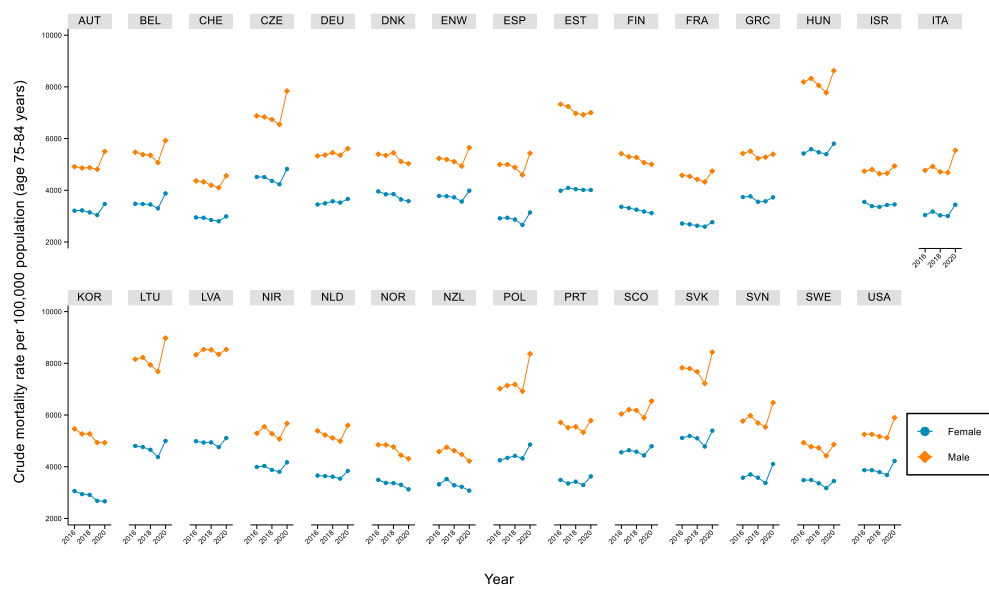


Fig. S4 (E): 85 years or above

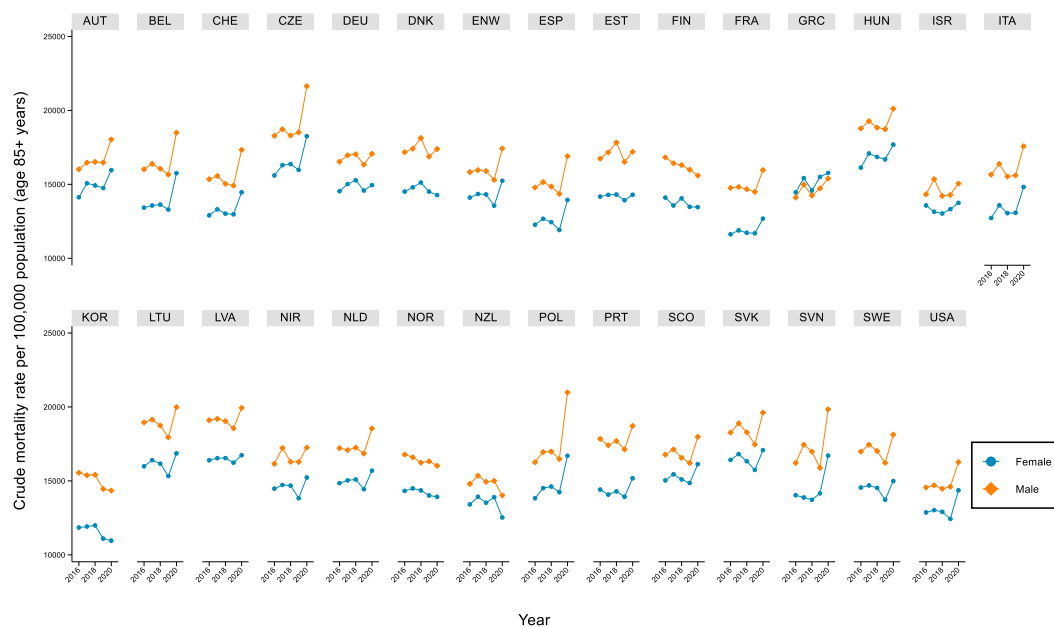


Figure S5: Percent change in weekly excess deaths in 29 high-income countries, 2016-2019
(Shaded regions indicate 95% confidence interval)

Fig. S5 (A): Female

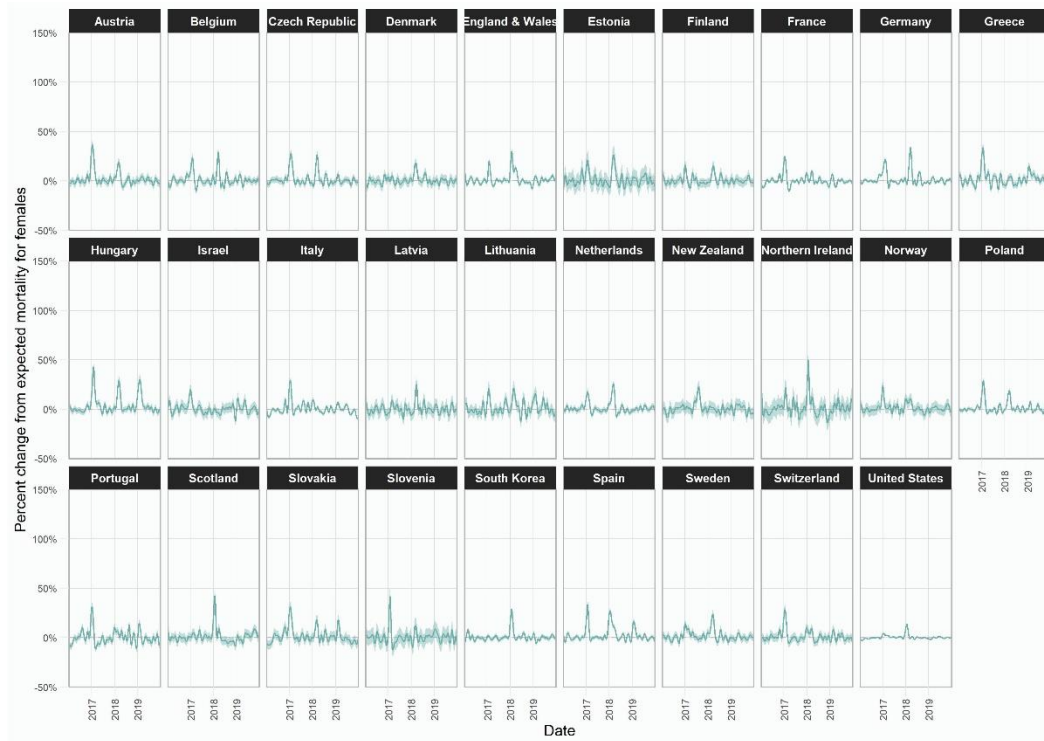


Fig. S5 (B): Male

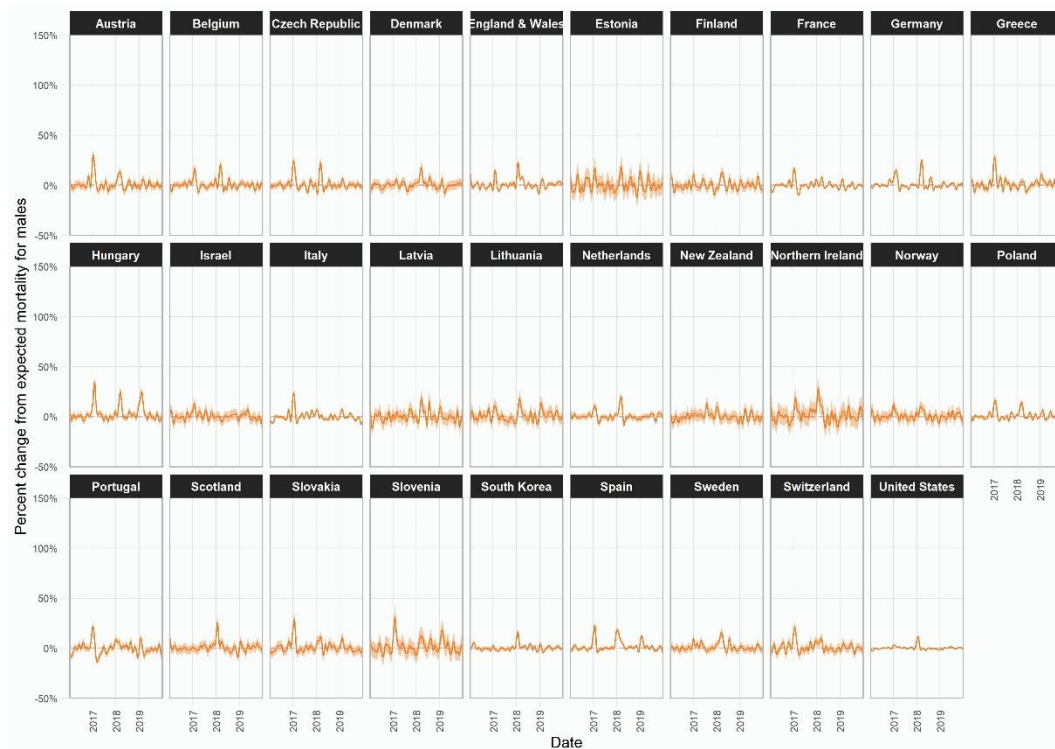
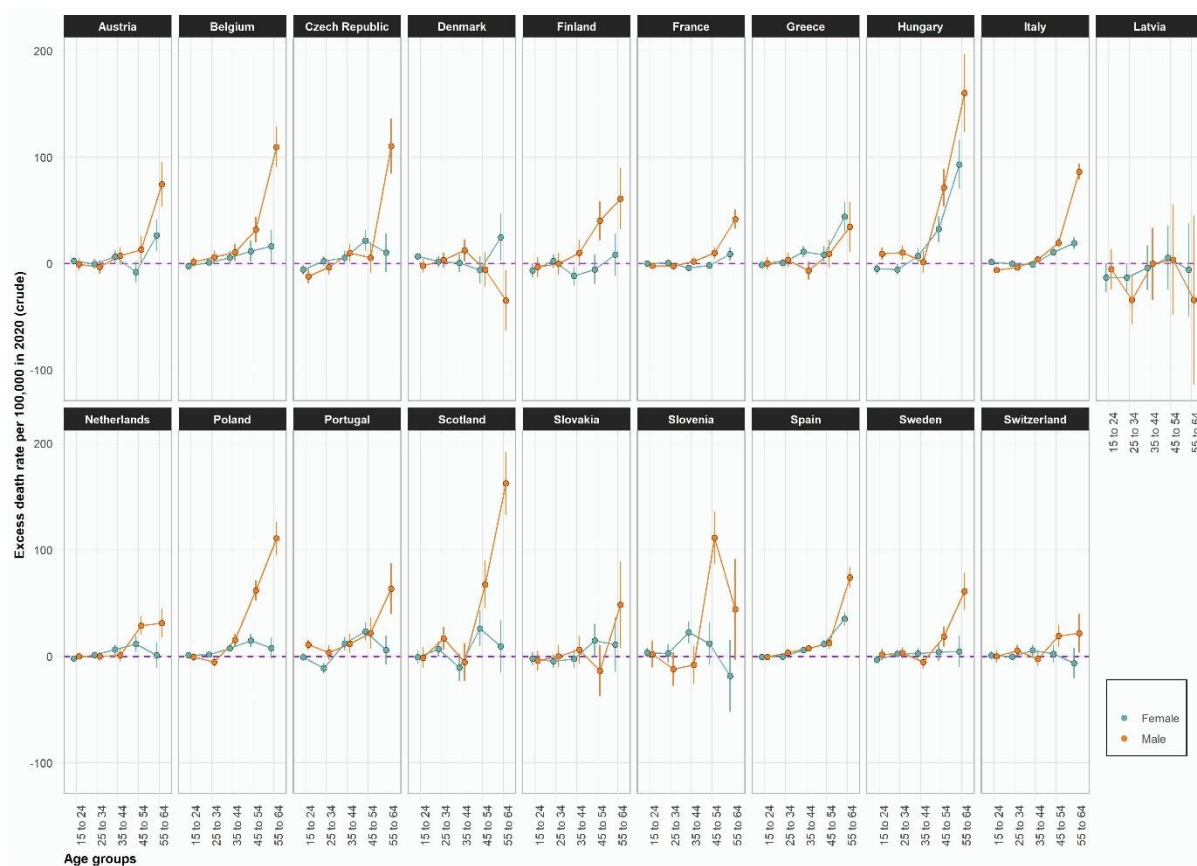


Figure S6: Excess death rates (per 100,000) in 19 high-income countries in 2020 among people 15-64 years, by sex and 10-year age groups



References:

- 1 European Commission. Eurostat. *Revision of the European Standard Population: report of Eurostat's task force : 2013 edition*. LU: : Publications Office 2013. <https://data.europa.eu/doi/10.2785/11470> (accessed 14 Jan 2021).
- 2 Fay MP, Feuer EJ. Confidence intervals for directly standardized rates: a method based on the gamma distribution. *Stat Med* 1997;**16**:791–801. doi:10.1002/(sici)1097-0258(19970415)16:7<791::aid-sim500>3.0.co;2-#
- 3 Tiwari RC, Clegg LX, Zou Z. Efficient interval estimation for age-adjusted cancer rates. *Stat Methods Med Res* 2006;**15**:547–69. doi:10.1177/0962280206070621