

A new and improved quantitative recovery analysis for iterative hard thresholding algorithms in compressed sensing

Coralia Cartis* and Andrew Thompson†

September 20, 2013

Abstract

We present a new recovery analysis for a standard compressed sensing algorithm, Iterative Hard Thresholding (IHT) (Blumensath and Davies, 2008), which considers the fixed points of the algorithm. In the context of arbitrary measurement matrices, we derive a sufficient condition for convergence of IHT to a fixed point and a necessary condition for the existence of fixed points. These conditions allow us to perform a sparse signal recovery analysis in the deterministic noiseless case by implying that the original sparse signal is the unique fixed point and limit point of IHT, and in the case of Gaussian measurement matrices and noise by generating a bound on the approximation error of the IHT limit as a multiple of the noise level. By generalizing the notion of fixed points, we extend our analysis to the variable stepsize Normalised IHT (N-IHT) (Blumensath and Davies, 2010). For both stepsize schemes, we obtain asymptotic phase transitions in a proportional-dimensional framework, quantifying the sparsity/undersampling trade-off for which recovery is guaranteed. Exploiting the reasonable average-case assumption that the underlying signal and measurement matrix are independent, comparison with previous results within this framework shows a substantial quantitative improvement.

1 Introduction

Compressed Sensing (CS) seeks to recover sparse or compressible signals from undersampled linear measurements [11, 12, 15]; it asserts that the number of measurements should be proportional to the information content of the signal, rather than its dimension. More specifically, one seeks to recover a sparse signal from noisy linear measurements. We refer to a signal which has at most k nonzero entries as being k -sparse, and the problem can be stated as follows.

Sparse recovery from noisy measurements: *Recover a k -sparse signal $x^* \in \mathbb{R}^N$ from the linear measurements*

$$b = Ax^* + e \in \mathbb{R}^n, \quad (1.1)$$

where $e \in \mathbb{R}^n$ is an unknown noise vector and where $0 < 2k \leq n \leq N$.

Since the introduction of CS in 2004, many algorithms have been proposed to solve this seemingly (and generally) intractable problem; see [31] for a recent survey. A common approach is to solve, by means of classical or recently-proposed optimization methods, a (convex or nonconvex) optimization relaxation that penalizes the lack of sparsity of x by means of l_p -norms with $0 < p \leq 1$. Alternatively, greedy methods — such as (Orthogonal) Matching Pursuit [14, 24, 28], SP [13], CoSAMP [25], amongst others — can be used to tackle the so-called l_0 -problem directly, namely,

$$\min_{x \in \mathbb{R}^N} \Psi(x) \stackrel{\text{def}}{=} \frac{1}{2} \|Ax - b\|^2 \quad \text{subject to } \|x\|_0 \leq k, \quad (1.2)$$

where $\|\cdot\|_0$ counts the number of nonzero entries of the argument. Problem (1.2) is nonconvex, with many local minimizers and in the perfect case of zero noise, with a (unique) global minimizer at the k -sparse

*Mathematical Institute, University of Oxford, Radcliffe Observatory Quarter, Oxford, OX2 6GG, UK.
email: coralia.cartis@maths.ox.ac.uk.

†Department of Mathematics, Duke University, Box 90320, Durham, NC 27708-0320, USA.
email: thompson@math.duke.edu.

vector x^* [26] that we are aiming to recover. It is now well known that, under certain conditions, many of these algorithms have stable recovery properties, namely that the error in approximating the original signal is some (usually small) multiple of the noise level, which further implies exact recovery of the original signal x^* in the absence of noise [6, 31].

Here, we focus on a simple and widely-used greedy technique – Iterative Hard Thresholding (IHT) [9, 10] – that generates feasible steepest descent steps for problem (1.2), obtained by projecting steps along the negative gradient direction of Ψ onto the l_0 -norm constraint by means of the *hard threshold* operator which simply sets all but the k largest in magnitude coefficients of a vector to zero. IHT [9] performs gradient projection with constant stepsize, while Normalised Iterative Hard Thresholding (N-IHT) [10] employs a variable stepsize scheme.

Early CS theory focussed upon algorithms which solve convex relaxations of (1.2) [11, 15], and perhaps for this reason IHT algorithms were slow to gain acceptance in the CS community. However, more recently, empirical studies in [7] have shown that, surprisingly, these nonconvex approaches are in fact competitive in practice in terms of sparse recovery properties with more established CS algorithms based on l_1 -minimization. In common with many gradient methods proposed for l_1 -based CS recovery, IHT algorithms also have low computational cost, with the most costly operations being matrix-vector products and hard threshold operations. As we detail below, existing theoretical recovery analyses are unduly pessimistic and fail to account for this excellent practical behaviour of IHT variants, and it is the aim of this paper to improve quantitative recovery guarantees of IHT algorithms by means of a new probabilistic analysis.

Regarding state-of-the-art theoretical properties, Blumensath and Davies [9] obtained the first convergence result for IHT, proving convergence to a fixed point of IHT/local minimizer of (1.2) provided the spectral norm of the measurement matrix A is less than one, a somewhat restrictive condition. The same authors [8] then proved that stable recovery is guaranteed provided A satisfies a *restricted isometry property* (RIP) [12], which requires the matrix to act as a near isometry on all k -sparse vectors, a now ubiquitous tool in CS recovery analysis. Other RIP-based recovery conditions were subsequently obtained for IHT in [19–21], and for N-IHT in [10].

Determining whether a given measurement matrix satisfies a restricted isometry property is in itself, however, NP-hard. It has been shown [3] that certain random matrices, such as Gaussian matrices in which each entry is i.i.d. Gaussian, satisfy the RIP provided

$$n \geq C \cdot k \ln \left(\frac{N}{k} \right).$$

Quantifying the constant C is, however, vital to practitioners who wish to know how aggressively a signal may be undersampled given its dimension and sparsity. Based on the RIP analysis in [8, 10, 19–21], quantitative results were obtained for IHT in [6, 30] for the case of Gaussian matrices in an asymptotic framework in which the problem dimensions are assumed to grow proportionally. We will refer to such a framework as the *proportional-growth asymptotic*, defined as follows.

Definition 1.1 (Proportional-growth asymptotic [5]). *We say that a sequence of problem sizes (k, n, N) , where $0 < k \leq n \leq N$, grows proportionally if, for some $\delta \in (0, 1]$ and $\rho \in (0, 1]$,*

$$\frac{n}{N} \rightarrow \delta \quad \text{and} \quad \frac{k}{n} \rightarrow \rho \quad \text{as } (k, n, N) \rightarrow \infty.$$

This framework, advocated by Donoho and others [16, 17], defines a two-dimensional phase space for asymptotic analysis in which the variables δ and ρ have a simple practical interpretation. The parameter δ is the ratio by which the signal is undersampled (an *undersampling ratio*), while the ratio ρ indicates how many measurements need to be taken as a multiple of the sparsity (an *oversampling ratio*).

By making use of RIP analysis for Gaussian matrices, first performed in [5] and subsequently improved upon in [2], it was shown in [6, 30] that all of the RIP conditions proved to date for IHT algorithms are pessimistic compared to these algorithms’ numerically-observed average-case behaviour. This is not

altogether surprising, since the RIP gives worst-case guarantees. There is, therefore, a need for improved quantitative recovery guarantees for IHT algorithms which narrow the gap between theoretical guarantees and observed performance. This is in contrast to l_1 -minimization, for which average-case phase transitions in the proportional-growth asymptotic have been precisely determined for Gaussian matrices [16].¹

The main contributions of this paper are as follows:

1) We present an entirely new recovery analysis of IHT algorithms. In the context of constant stepsize IHT, whereas previous recovery analyses [8, 10, 19–21] take the direct approach of bounding the approximation error from iteration to iteration, we take a two-part approach in which we analyse the fixed points of the algorithm. First, we prove a *stable point condition*, namely a necessary condition for there to be a fixed point on a given support. Second, we give a *convergence condition* which guarantees the convergence of IHT to one of its fixed points. In the case of no noise, this analysis allows us to establish conditions under which, surprisingly, IHT converges to its unique fixed point, namely, the original signal x^* ; noise-dependent recovery results are also given. By extending the notion of a fixed point to the (new) concept of an α -stable point, we obtain similar recovery results for the variable-stepsizes N-IHT.

2) We use average-case assumptions to obtain improved recovery phase transitions for IHT algorithms with Gaussian matrices and Gaussian noise. While it is possible to analyse the stable point condition using the RIP [30], we take a different approach. Because the stable point condition has no dependence upon the iterates of the algorithm, it is amenable to analysis for Gaussian matrices under the average-case assumption that the measurement matrix is independent of the signal – a realistic assumption in CS. We derive precise distributions of this condition’s constituent terms, and obtain large deviations bounds on these terms over all possible support sets in the proportional-dimensional asymptotic; in this context, we deduce bounds on some *independent RIP* constants that occur naturally in our results. For the convergence condition, we still make use of the RIP, and upper bounds thereon in the proportional-growth asymptotic for Gaussian matrices. However, the RIP condition involved is substantially weaker than any others that have appeared in the literature to date for IHT algorithms. Combining these results, we obtain recovery phase transitions for IHT and N-IHT, namely regions of the phase plane in which stable recovery is guaranteed. In the case of zero noise, we have exact recovery of the original signal, and our analysis also implies a radically new insight: that, within some region of the phase plane, IHT algorithms have a single fixed point, namely the original signal. In the case of noise, we derive *stability factors* which bound the approximation error as a multiple of the expectation of the noise. Comparison with state-of-the-art results that have been quantified in the phase transition framework in [6, 30], shows a substantial quantitative improvement, both in terms of recovery guarantees as expressed by the height of the phase transitions and of robustness to noise as expressed by the size of the stability factors; thus narrowing the gap to observed average-case behaviour. In particular, for the variable-stepsizes N-IHT, we obtain about a factor 10 improvement in the height of the phase transition over best-known results.

Outline of the paper. We begin in Section 2 by describing in more detail the generic IHT algorithm and two stepsize scheme variants, IHT and N-IHT. In Section 3, we introduce our new recovery analysis, proving our stable point condition, and convergence conditions for both stepsize schemes. Then we focus our attention for the remainder of the paper upon Gaussian matrices: in Section 4, we prove various distributional and large deviations results, and we use these in Section 5 to obtain recovery phase transitions in the proportional-growth asymptotic. We illustrate and discuss our results in Section 6, after which we conclude.

Notation. We let $\|\cdot\|$ denote the Euclidean norm. The support set of the k -sparse signal x^* we aim to recover will be denoted by $\text{supp}(x^*) = \Lambda$ with cardinality $|\Lambda| = k$. Given some index set $\Gamma \subseteq \{1, 2, \dots, N\}$, we define the complement of Γ to be $\Gamma^C = \{1, 2, \dots, N\} \setminus \Gamma$. We write x_Γ for the restriction of the vector x to the coefficients indexed by the elements of Γ , and we write A_Γ for the restriction of the matrix A to those columns indexed by the elements of Γ .

¹To the best of our knowledge, the average-case analysis techniques of approximate message passing (Donoho, Maleki and Montanari, 2009) cannot be applied to IHT methods because the hard thresholding operator is not Lipschitz-continuous.

2 Iterative hard thresholding algorithms

Let us describe in detail the algorithms that are the focus of the analysis in this paper. Generically, on each hard thresholding iteration m , a steepest descent step, possibly with linesearch, is calculated for the objective Ψ in (1.2), namely, a move is performed from the current iterate x^m along the negative gradient of Ψ ,

$$-\nabla\Psi(x^m) = -A^T(Ax^m - b). \quad (2.1)$$

The resulting step is then projected onto the (nonconvex) l_0 -constraint in (1.2) using the so-called *hard threshold* operator $\mathcal{H}_k(\cdot)$ defined as

$$\mathcal{H}_k(x) \stackrel{\text{def}}{=} \arg \min_{\|z\|_0 \leq k} \|z - x\|.$$

As the name suggests, $\mathcal{H}_k(\cdot)$ is indeed a thresholding operator, keeping the k largest entries in magnitude of its argument and setting the rest to zero, namely,

$$\mathcal{H}_k(x) = \begin{cases} x_i & \text{for } i \in \Gamma, \\ 0 & \text{for } i \notin \Gamma, \end{cases} \quad \text{where } \Gamma \stackrel{\text{def}}{=} \{ \text{indices of the } k \text{ largest in magnitude entries of } x \} \quad (2.2)$$

(See [30, Lemma 1.10] for a proof of (2.2) given its definition.) To avoid a situation in which the support set Γ is not uniquely defined, if for instance some of the coefficients are equal in magnitude, then a support set for the identical components can be selected either randomly or according to some predefined ordering.

The generic IHT algorithm, that includes variants allowing constant or variable linesearch choices, can be summarized as follows².

Algorithm 2.1: Generic Iterative Hard Thresholding (G-IHT) algorithm [9, 10].

Given A , b and k for problem (1.1), do:

Step 0: Set $x^0 = 0$ and $m = 0$.

While some termination criterion is not satisfied, do:

Step 1: Compute

$$x^{m+1} := \mathcal{H}_k \{ x^m - \alpha^m A^T(Ax^m - b) \}, \quad (2.3)$$

with \mathcal{H}_k defined in (2.2) and $\alpha^m > 0$ some (pre-defined or computed) stepsize.

Step 2: Set $m = m + 1$ and return to Step 1.

In our analysis, we will consider the possibly infinite sequence of iterates generated by G-IHT, though in practice a useful termination criterion such as requiring the residual to be sufficiently small, would need to be employed. Two popular stepsize choices will be addressed: *constant stepsize* $\alpha^m = \alpha \in (0, 1)$ for all m , with the resulting G-IHT variant being denoted simply as IHT [9], and *variable stepsize* as prescribed in the Normalised IHT (N-IHT) variant proposed in [10].

The IHT variant of G-IHT can be summarized as follows.

Algorithm 2.2: Iterative Hard Thresholding (IHT) algorithm [9].

Given some $\alpha > 0$, on each iteration $m \geq 0$ of G-IHT, do:

In **Step 1**, set α^m in (2.3) as follows:

$$\alpha^m := \alpha. \quad (2.4)$$

²The reason we introduce the G-IHT framework is to allow a more concise presentation of results for the (fully specified) IHT variants that are of interest.

The N-IHT variant defined below follows [10], having the stepsize α^m chosen according to an *exact linesearch* [27] when the support set of consecutive iterates stays the same, and using a *shrinking* strategy when the support set changes so as to ensure sufficient decrease in the objective of (1.2).

Algorithm 2.3: Normalised Iterative Hard Thresholding (N-IHT) algorithm [10].

Given some $c \in (0, 1)$ and $\kappa > 1/(1 - c)$, on each iteration $m \geq 0$ of G-IHT, do:

In **Step 1**, compute α^m in (2.3) as follows:

Step 1.0: Set $\Gamma^m := \text{supp}(x^m)$ and

$$\alpha^m := \frac{\|A_{\Gamma^m}^T(b - Ax^m)\|^2}{\|A_{\Gamma^m} A_{\Gamma^m}^T(b - Ax^m)\|^2}. \quad (2.5)$$

Compute $\tilde{x}^{m+1} := H_k \{x^m + \alpha^m A^T(b - Ax^m)\}$. If $\text{supp}(\tilde{x}^{m+1}) = \Gamma^m$, terminate with α^m given in (2.5).

While $\alpha^m \geq (1 - c) \frac{\|\tilde{x}^{m+1} - x^m\|^2}{\|A(\tilde{x}^{m+1} - x^m)\|^2}$, do:

Step 1.1: $\alpha^m := \alpha^m / [\kappa(1 - c)]$;

Step 1.2: $\tilde{x}^{m+1} := H_k \{x^m + \alpha^m A^T(b - Ax^m)\}$;

End.

Under the (weakest) assumptions of this paper, we can ensure that both the exact linesearch and the shrinkage stepsizes in N-IHT are well-defined, until termination; see Section 3.2.2. The shrinkage iteration between Steps 1.1–1.2 of N-IHT can be shown to terminate in finitely many steps [10].

3 Deterministic conditions for a recovery analysis

In this section we derive conditions for IHT algorithms when applied to general measurement matrices A . Hence we make the following common assumption for compressed sensing algorithms.

A.1 The matrix A is in $2k$ -general position, namely any $2k$ of its columns are linearly independent.

The (weak) assumption **A.1** is equivalent to the condition that, for any Γ such that $|\Gamma| = 2k$, the matrix $A_\Gamma^T A_\Gamma$ is nonsingular. Thus, whenever **A.1** holds and there is no noise in the system (i.e., $e = 0$ in (1.1)), we have $\|A(x^* - x)\| > 0$ for any k -sparse $x \neq x^*$, and so x^* is the unique k -sparse exact solution to the linear system $b = Ax^*$. Note also that **A.1** holds if A is in general position. It is also satisfied with probability 1 if A is a Gaussian matrix and $2k \leq n$ [22].

The results derived in this section come in two parts: a necessary condition for the existence of (generalized) fixed points of G-IHT and a sufficient condition guaranteeing convergence for particular stepsize schemes. In the deterministic noiseless case, the former condition can be used to guarantee the existence of at most one fixed point, namely, the original signal x^* ; thus, provided we also have convergence of the algorithm to *some* such fixed/stable point, signal recovery is ensured. The below deterministic results also yield similar recovery properties (based on proximity/closeness of fixed points to the original signal) in the presence of noise and Gaussian measurement matrices, as we show in later sections.

3.1 A stable-point condition

We introduce the concept of an $\underline{\alpha}$ -stable point of G-IHT, a generalization of fixed points.

Definition 3.1 ($\underline{\alpha}$ -stable points of G-IHT). Given $\underline{\alpha} > 0$ and an index set Γ with $|\Gamma| = k$, we say $\bar{x} \in \mathbb{R}^N$ is an $\underline{\alpha}$ -stable point of G-IHT on Γ if $\text{supp}(\bar{x}) \subseteq \Gamma$ and

$$\{A^T(b - A\bar{x})\}_\Gamma = 0 \quad \text{and} \quad (3.1)$$

$$\min_{i \in \Gamma} |\bar{x}_i| \geq \underline{\alpha} \max_{j \in \Gamma^C} |\{A^T(b - A\bar{x})\}_j|. \quad (3.2)$$

Note that in the noiseless case ($e = 0$ in (1.1)), the original signal x^* is clearly an $\underline{\alpha}$ -stable point on $\text{supp}(x) = \Lambda$, for any value of $\underline{\alpha} > 0$.

In the case of the constant-stepsize IHT algorithm, an α -stable point is nothing other than a *fixed point* of IHT (see Blumensath & Davies [9, Lemma 6]) or an L -stationary point of (1.2) as defined in [4, §2.3]. (Indeed, if a further IHT iteration is applied at a fixed point \bar{x} , there is no change in the support set; thus the gradient term on the complement of the support of \bar{x} must be suitably small, which is (3.2). Also, the coefficients on the support of \bar{x} must remain unchanged, and so we require the gradient on the support of \bar{x} to be zero, namely (3.1).) A generalization of the notion of a fixed point and L -stationary point to stable points is, however, required to allow for variable stepsize schemes in G-IHT³; we will be interested in values of $\underline{\alpha}$ that lower bound the stepsize α^m of G-IHT.

Next we show that any $\underline{\alpha}$ -stable point is a minimum-norm solution on some k -subspace.

Lemma 3.1. Let **A.1** hold and \bar{x} be an $\underline{\alpha}$ -stable point of G-IHT on Γ for some $\underline{\alpha} > 0$. Then

$$\bar{x}_\Gamma = A_\Gamma^\dagger b, \quad (3.3)$$

where A_Γ^\dagger is the Moore-Penrose pseudo-inverse, namely,

$$A_\Gamma^\dagger \stackrel{\text{def}}{=} (A_\Gamma^T A_\Gamma)^{-1} A_\Gamma^T. \quad (3.4)$$

Proof. It follows from (3.1) that $A_\Gamma^T(b - A_\Gamma \bar{x}_\Gamma) = 0$. By **A.1**, the pseudoinverse A_Γ^\dagger is well-defined and we may rearrange to give (3.3). \square

While the previous lemma tells us that any stable point is necessarily a minimum-norm solution on some k -subspace, the converse may not hold. Next, we give a more useful necessary condition for there to exist a stable point on a given support set. We will use the latter condition in a sufficient sense later on, to guarantee that under certain conditions, all G-IHT stable points are *close* to the underlying signal, which in the noiseless case reduces to G-IHT having *at most one* stable point, namely, the original signal.

Theorem 3.2 (Stable point condition; noise case). Consider problem (1.1) and let $\Lambda = \text{supp}(x^*)$. Suppose **A.1** holds and suppose there exists an $\underline{\alpha}$ -stable point of G-IHT on some Γ such that $\Gamma \neq \Lambda$. Then

$$\|A_\Gamma^\dagger A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^*\| + \|A_\Gamma^\dagger e\| \geq \underline{\alpha} \left\{ \|A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^*\| - \|A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) e\| \right\}, \quad (3.5)$$

where A_Γ^\dagger is defined in (3.4).

Proof. Assume \bar{x} is an $\underline{\alpha}$ -stable point on Γ . Since $\Gamma \setminus \Lambda \subseteq \Gamma$ and $\Lambda \setminus \Gamma \subseteq \Gamma^C$, where $\Lambda = \text{supp}(x^*)$, (3.2) implies that

$$\min_{i \in \Gamma \setminus \Lambda} |\bar{x}_i| \geq \underline{\alpha} \max_{j \in \Lambda \setminus \Gamma} |\{A^T(b - A\bar{x})\}_j|. \quad (3.6)$$

Definition 3.1 implies that $|\Gamma| = |\Lambda|$, and so $|\Gamma \setminus \Lambda| = |\Lambda \setminus \Gamma|$. This, properties of the Euclidean norm and (3.6) provide

$$\|\bar{x}_{\Gamma \setminus \Lambda}\|^2 \geq |\Gamma \setminus \Lambda| \left\{ \min_{i \in \Gamma \setminus \Lambda} |\bar{x}_i| \right\}^2 \geq |\Lambda \setminus \Gamma| \left\{ \underline{\alpha} \max_{j \in \Lambda \setminus \Gamma} |\{A^T(b - A\bar{x})\}_j| \right\}^2 \geq \underline{\alpha}^2 \|A_{\Lambda \setminus \Gamma}^T (b - A\bar{x})\|^2. \quad (3.7)$$

³When (3.1) holds for N-IHT, the exact linesearch stepsize (2.5) is not well-defined with its numerator and denominator both being zero.

Problem (1.1) and $x_{\Lambda^c}^* = 0$ imply

$$b = Ax^* + e = A_\Gamma x_\Gamma^* + A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* + e. \quad (3.8)$$

This and Lemma 3.1 now provide, under **A.1**,

$$\bar{x}_\Gamma = A_\Gamma^\dagger b = x_\Gamma^* + A_\Gamma^\dagger A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* + A_\Gamma^\dagger e,$$

where in the last equality, we used $A_\Gamma^\dagger A_\Gamma = I$. Therefore, since $x_{\Gamma \setminus \Lambda}^* = 0$, we deduce

$$\bar{x}_{\Gamma \setminus \Lambda} = \left(A_\Gamma^\dagger A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* + A_\Gamma^\dagger e \right)_{\Gamma \setminus \Lambda}$$

and so,

$$\|\bar{x}_{\Gamma \setminus \Lambda}\| \leq \left\| \left(A_\Gamma^\dagger A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* \right)_{\Gamma \setminus \Lambda} \right\| + \left\| \left(A_\Gamma^\dagger e \right)_{\Gamma \setminus \Lambda} \right\| \leq \|A_\Gamma^\dagger A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^*\| + \|A_\Gamma^\dagger e\|, \quad (3.9)$$

which upper bounds the left-hand side of (3.7). Under **A.1**, we may next use Lemma 3.1 and (3.8) to express the right-hand side of (3.7) independently of \bar{x} , as

$$A_{\Lambda \setminus \Gamma}^T (b - A\bar{x}) = A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) b = A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) (A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* + e),$$

where in the last equality, we used $A_\Gamma^\dagger A_\Gamma = I$. We therefore may deduce

$$\left\| A_{\Lambda \setminus \Gamma}^T (b - A\bar{x}) \right\| \geq \left\| A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* \right\| - \left\| A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) e \right\|. \quad (3.10)$$

Substituting (3.9) and (3.10) into (3.7), we arrive at (3.5). \square

Theorem 3.2 simplifies further in the noiseless case.

Corollary 3.3 (Stable point condition; noiseless case). *Consider problem (1.1) with $e \stackrel{\text{def}}{=} 0$ and let $\Lambda = \text{supp}(x^*)$. Suppose **A.1** holds and suppose there exists an $\underline{\alpha}$ -stable point of G-IHT on some Γ such that $\Gamma \neq \Lambda$. Then*

$$\left\| A_\Gamma^\dagger A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* \right\| \geq \underline{\alpha} \left\| A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* \right\|, \quad (3.11)$$

where A_Γ^\dagger is defined in (3.4).

Proof. The result follows immediately by setting $e \stackrel{\text{def}}{=} 0$ in (3.5). \square

Clearly, Corollary 3.3 implies that if the reverse inequality in (3.41) holds for all support sets $\Gamma \neq \Lambda$, then x^* is the only $\underline{\alpha}$ -stable point of G-IHT.

3.2 A convergence condition

This section gives conditions for IHT algorithms to convergence to stable points. Recalling (1.2) and (2.3), we introduce the notation

$$g^m \stackrel{\text{def}}{=} \nabla \Psi(x^m) \text{ and } \Gamma^m \stackrel{\text{def}}{=} \text{supp}(x^m), \text{ for all } m. \quad (3.12)$$

Some useful properties of the G-IHT iterates are given in the next lemma.

Lemma 3.4. *Apply the G-IHT algorithm to solve (1.2). Then the G-IHT iterates satisfy for all $m \geq 0$,*

$$\|x^{m+1} - x^m\|^2 + 2\alpha^m (g^m)^T (x^{m+1} - x^m) \leq 0 \quad (3.13)$$

and

$$\Psi(x^{m+1}) - \Psi(x^m) = (g^m)^T (x^{m+1} - x^m) + \frac{1}{2} \|A(x^{m+1} - x^m)\|^2. \quad (3.14)$$

Proof. Since the hard thresholding operation in (2.3) can be viewed as a projection onto the constraint of (1.2), we may rewrite the G-IHT iteration (2.3) as

$$x^{m+1} = \arg \min_{\|z\|_0 \leq k} \|z - \{x^m - \alpha^m g^m\}\|^2.$$

This further gives that

$$\|x^{m+1} - (x^m - \alpha^m g^m)\|^2 \leq \|x^m - (x^m - \alpha^m g^m)\|^2 = (\alpha^m)^2 \|g^m\|^2,$$

which expands to give $\|x^{m+1} - x^m\|^2 + 2\alpha^m (g^m)^T (x^{m+1} - x^m) + (\alpha^m)^2 \|g^m\|^2 \leq (\alpha^m)^2 \|g^m\|^2$, and so (3.13) holds. Since Ψ in (1.2) is quadratic, we have no remainder in the following second-order Taylor expansion

$$\begin{aligned} \Psi(x^{m+1}) - \Psi(x^m) &= [\nabla \Psi(x^m)]^T (x^{m+1} - x^m) + \frac{1}{2} (x^{m+1} - x^m)^T [\nabla^2 \Psi] (x^{m+1} - x^m) \\ &= (g^m)^T (x^{m+1} - x^m) + \frac{1}{2} (x^{m+1} - x^m)^T A^T A (x^{m+1} - x^m), \end{aligned}$$

and so (3.14) follows. \square

A sufficient condition for G-IHT convergence is given next.

Lemma 3.5. *Let A.1 hold and the G-IHT iterates satisfy*

$$\|x^{m+1} - x^m\|^2 \leq d [\Psi(x^m) - \Psi(x^{m+1})], \text{ for all } m \geq 0, \quad (3.15)$$

for some $d > 0$ independent of m and Ψ defined in (1.2). Assume that there exist $\bar{\alpha} \geq \underline{\alpha} > 0$ such

$$\bar{\alpha} \geq \alpha^m \geq \underline{\alpha} \text{ for all } m \geq 0. \quad (3.16)$$

Then $x^m \rightarrow \bar{x}$ as $m \rightarrow \infty$, where \bar{x} is an $\underline{\alpha}$ -stable point of G-IHT.

Proof. We deduce from (3.15) that

$$\sum_{m=0}^{\infty} \|x^{m+1} - x^m\|^2 \leq d \sum_{m=0}^{\infty} [\Psi(x^m) - \Psi(x^{m+1})] \leq d\Psi(x^0),$$

where to obtain the last inequality, we used $\Psi(x^m) \geq 0$. Thus convergent series properties provide

$$\|x^{m+1} - x^m\| \rightarrow 0 \text{ as } m \rightarrow \infty. \quad (3.17)$$

From (2.3) and (3.12), we deduce

$$x_{\Gamma^{m+1}}^{m+1} = x_{\Gamma^{m+1}}^m - \alpha^m g_{\Gamma^{m+1}}^m \text{ and } x_{(\Gamma^{m+1})^C}^{m+1} = 0.$$

Thus restricting (3.17) to Γ^{m+1} and using (3.16) provide

$$\|g_{\Gamma^{m+1}}^m\| \rightarrow 0 \text{ as } m \rightarrow \infty, \quad (3.18)$$

while restricting (3.17) to $\Gamma^m \setminus \Gamma^{m+1}$ yields

$$\|x_{\Gamma^m \setminus \Gamma^{m+1}}^m\| \rightarrow 0. \quad (3.19)$$

For $m \geq 0$, let y^m denote the minimum-norm solution on Γ^m , namely,

$$y_{\Gamma^m}^m \stackrel{\text{def}}{=} A_{\Gamma^m}^\dagger b \text{ and } y_{(\Gamma^m)^C}^m \stackrel{\text{def}}{=} 0, \quad (3.20)$$

which is well-defined due to A.1. Then (3.20) and $x_{(\Gamma^m)^C}^m = 0$ provide

$$\begin{aligned} \|y^{m+1} - x^m\| &\leq \|y_{\Gamma^{m+1}}^{m+1} - x_{\Gamma^{m+1}}^m\| + \|x_{(\Gamma^{m+1})^C}^m\| = \|A_{\Gamma^{m+1}}^\dagger b - x_{\Gamma^{m+1}}^m\| + \|x_{\Gamma^m \setminus \Gamma^{m+1}}^m\| \\ &= \|(A_{\Gamma^{m+1}}^T A_{\Gamma^{m+1}})^{-1} A_{\Gamma^{m+1}}^T (b - A_{\Gamma^{m+1}} x_{\Gamma^{m+1}}^m)\| + \|x_{\Gamma^m \setminus \Gamma^{m+1}}^m\| \\ &= \|(A_{\Gamma^{m+1}}^T A_{\Gamma^{m+1}})^{-1} g_{\Gamma^{m+1}}^m\| + \|x_{\Gamma^m \setminus \Gamma^{m+1}}^m\| \rightarrow 0, \text{ as } m \rightarrow \infty, \end{aligned}$$

where the limit follows from (3.18), (3.19), **A.1** and the fact that there are finitely many distinct support sets Γ^m , $m \geq 0$. This and (3.17) further give

$$\|y^m - x^m\| \longrightarrow 0 \text{ as } m \longrightarrow \infty, \quad (3.21)$$

and so for any $\epsilon > 0$, there exists $m_0 \geq 0$ such that

$$\|y^m - x^m\| \leq \epsilon, \text{ for all } m \geq m_0. \quad (3.22)$$

We denote the index set of changing minimal-norm solutions by

$$\mathcal{S} \stackrel{\text{def}}{=} \{m \geq m_0 : y^{m+1} \neq y^m\},$$

and we will show that \mathcal{S} is finite. Define

$$\epsilon \stackrel{\text{def}}{=} \frac{1}{4} \min_{m \in \mathcal{S}} \|y^{m+1} - y^m\|. \quad (3.23)$$

Note that $\epsilon > 0$ since there are finitely many distinct support sets Γ^m , $m \geq 0$. Then, the triangle inequality, (3.22) and (3.23) yield

$$\|x^{m+1} - x^m\| \geq \|y^{m+1} - y^m\| - \|y^{m+1} - x^{m+1}\| - \|y^m - x^m\| \geq 4\epsilon - \epsilon - \epsilon > \epsilon, \text{ for all } m \in \mathcal{S}.$$

This and (3.17) imply that \mathcal{S} must be finite and so there exists $m_1 \geq m_0$ such that $y^{m+1} = y^m = \bar{x}$ for all $m \geq m_1$, where $\bar{x}_\Gamma = A_\Gamma^\dagger b$ and $\bar{x}_{\Gamma^c} = 0$, for some Γ with $|\Gamma| = k$. This and (3.21) give

$$x^m \longrightarrow \bar{x}, \text{ as } m \longrightarrow \infty. \quad (3.24)$$

Clearly, (3.1) holds for the limit point \bar{x} of the iterates $\{x^m\}$. To complete the proof, it remains to establish (3.2). The thresholding operation that defines x^{m+1} in G-IHT gives that

$$\min_{i \in \Gamma^{m+1}} |x_i^{m+1}| \geq \max_{j \in (\Gamma^{m+1})^c} |\{x^m - \alpha^m g^m\}_j|, \text{ for all } m \geq 0, \quad (3.25)$$

and (3.16) implies that there exists a convergent subsequence of stepsizes,

$$\alpha^{m_r} \longrightarrow \tilde{\alpha} \geq \underline{\alpha} \text{ as } r \longrightarrow \infty. \quad (3.26)$$

Letting $\epsilon \stackrel{\text{def}}{=} \frac{1}{2} \min_{i \in \text{supp}(\bar{x})} \bar{x}_i$, (3.24) implies that $\|x^m - \bar{x}\| \leq \epsilon$, and so

$$\text{supp}(\bar{x}) \subseteq \Gamma^m, \text{ for all } m \text{ sufficiently large.} \quad (3.27)$$

Firstly, assume that $\text{supp}(\bar{x}) = \Gamma$. Then, since $|\Gamma| = |\Gamma^m| = k$, (3.27) implies that $\Gamma^m = \Gamma$ for all m sufficiently large, which together with (3.25), provides

$$\min_{i \in \Gamma} |x_i^{m+1}| \geq \max_{j \in \Gamma^c} |\{x^m - \alpha^m g^m\}_j|, \text{ for all } m \text{ sufficiently large.} \quad (3.28)$$

Passing to the limit in (3.28) on the subsequence m_r for which (3.26) holds, using (3.24), $\bar{x}_{\Gamma^c} = 0$ and the right-hand side of (3.16) imply (3.2) holds in this case. It remains to consider the case when $\text{supp}(\bar{x}) \subset \Gamma$. Then $\min_{i \in \Gamma} |\bar{x}_i| = 0$ and so (3.24) further provides

$$\min_{i \in \Gamma^{m+1}} |x_i^{m+1}| \longrightarrow 0 \text{ as } m \longrightarrow \infty. \quad (3.29)$$

Now (3.27) and again (3.24) provide

$$x_{\Gamma^{m+1}}^m \longrightarrow 0 \text{ as } m \longrightarrow \infty. \quad (3.30)$$

Passing to the limit in (3.25) on the subsequence m_r for which (3.26) holds, and using (3.29) and (3.30), we obtain that $g_{(\Gamma^{m+1})^c}^m \longrightarrow 0$ as $m \longrightarrow \infty$. This and (3.18) now give that $g^m = A^T(Ax^m - b) \longrightarrow 0$, which due to (3.24), implies that $A^T(b - A\bar{x}) = 0$ and so (3.2) trivially holds in this case. \square

In order to ensure (3.15) and (3.16), we make use of the well-known Restricted Isometry Constants (RIC) of the matrix A , defined as follows.

Definition 3.2. [5, 12] Define L_s and U_s , the lower and upper RIC constants of A of order s , to be, respectively,

$$L_s = 1 - \min_{1 \leq \|y\|_0 \leq s} \frac{\|Ay\|^2}{\|y\|^2} \quad \text{and} \quad U_s = \max_{1 \leq \|y\|_0 \leq s} \frac{\|Ay\|^2}{\|y\|^2} - 1. \quad (3.31)$$

Note that **A.1** is equivalent to the requirement that $L_{2k} < 1$.

In order to ensure (3.15) and (3.16) – using RICs or otherwise – we must specify the choice of stepsize α^m in G-IHT. (This is by contrast to the stable point condition for which only lower bounds on the stepsizes α^m matter.) Hence we now return to the constant-stepsize IHT and variable-stepsize N-IHT variants defined in Section 2.

3.2.1 A convergence condition for the IHT algorithm

In [9], Blumensath and Davies prove convergence of IHT iterates to a fixed point that is also a local minimizer of (1.2) (that may or may not be the original signal x^*) under the assumption that $\alpha\|A\|_2 < 1$. Similarly, Beck and Eldar [4, Theorem 3.2] show IHT iterates converge to an L -stationary point, an equivalent notion to that of a fixed point, under a commensurate condition on the stepsize, namely, $\alpha\|A^T A\| < 1$. Largely following the method of proof in [9], we now show that the requirement on the IHT stepsize in both these analyses can be weakened to a condition involving the RIC constant U_{2k} of A .

Theorem 3.6. Suppose that **A.1** holds, and that the IHT stepsize is chosen to satisfy

$$\alpha < \frac{1}{1 + U_{2k}}. \quad (3.32)$$

Then the IHT iterates $\{x^m\}$ converge to an α -stable point \bar{x} of IHT.

Proof. Let $m \geq 0$. Since the support size of the change to the iterates $x^{m+1} - x^m$ is at most $2k$, the upper RIC of A in (3.31) with $s = 2k$ provides $\|A(x^{m+1} - x^m)\|^2 \leq (1 + U_{2k})\|x^{m+1} - x^m\|^2$. Using this bound, and (3.13) with the choice (2.4), in (3.14), we obtain

$$\Psi(x^{m+1}) - \Psi(x^m) \leq -\frac{1}{2\alpha}\|x^{m+1} - x^m\|^2 + \frac{1}{2}(1 + U_{2k})\|x^{m+1} - x^m\|^2 = \frac{\alpha(1 + U_{2k}) - 1}{2\alpha}\|x^{m+1} - x^m\|^2,$$

which due to (3.32), implies that (3.15) holds with $d \stackrel{\text{def}}{=} 2\alpha/[1 - \alpha(1 + U_{2k})]$. Due to (2.4), (3.16) trivially holds with $\bar{\alpha} = \underline{\alpha} = \alpha$. Thus Lemma 3.5 applies, and so the IHT iterates x^m converge to an α -stable point of IHT. \square

3.2.2 A convergence condition for the N-IHT algorithm

Using the notation (3.31) and **A.1**, we obtain that the N-IHT stepsize α^m satisfies

$$\frac{1}{1 + U_k} \leq \alpha^m \leq \frac{1}{1 - L_k}, \quad \text{whenever } \alpha^m \text{ satisfies (2.5),} \quad (3.33)$$

and using also (2.3), that

$$\frac{1}{\kappa(1 + U_{2k})} \leq \alpha^m \leq \frac{1 - c}{1 - L_{2k}}, \quad \text{otherwise (i.e., whenever } \alpha^m \text{ is shrunk according to Steps 1.1–1.2).} \quad (3.34)$$

As the RICs of A are monotonically increasing with k and $\kappa, c \in (0, 1)$, (3.33) and (3.34) imply

$$\frac{1}{\kappa(1 + U_{2k})} \leq \alpha^m \leq \frac{1 - c}{1 - L_{2k}}, \quad \text{for all } m \geq 0. \quad (3.35)$$

Theorem 3.7. *Suppose **A.1** holds. Then the N-IHT iterates $\{x^m\}$ converge to a $[\kappa(1 + U_{2k})]^{-1}$ -stable point \bar{x} of N-IHT.*

Proof. Firstly, we consider the case when α^m satisfies (2.5). Then (3.12) implies $\Gamma^{m+1} = \Gamma^m$, and (2.3) implies

$$x_{\Gamma^m}^{m+1} = x_{\Gamma^m}^m - \alpha^m g_{\Gamma^m}^m. \quad (3.36)$$

Using (3.36), (2.5) becomes

$$\alpha^m = \frac{\|g_{\Gamma^m}^m\|^2}{\|A_{\Gamma^m} g_{\Gamma^m}^m\|^2} = \frac{\|x^{m+1} - x^m\|^2}{\|A(x^{m+1} - x^m)\|^2}. \quad (3.37)$$

Using that $x^{m+1} - x^m$ is supported on Γ^m , expressing $g_{\Gamma^m}^m$ from (3.36) and substituting into (3.14), we deduce that

$$\begin{aligned} \Psi(x^{m+1}) - \Psi(x^m) &= -\frac{1}{\alpha^m} (x_{\Gamma^m}^{m+1} - x_{\Gamma^m}^m)^T (x_{\Gamma^m}^{m+1} - x_{\Gamma^m}^m) + \frac{1}{2} \|A(x^{m+1} - x^m)\|^2 \\ &= -\frac{1}{\alpha^m} \|x^{m+1} - x^m\|^2 + \frac{1}{2\alpha^m} \|x^{m+1} - x^m\|^2 = -\frac{1}{2\alpha^m} \|x^{m+1} - x^m\|^2, \end{aligned} \quad (3.38)$$

where to obtain the second equality, we also used (3.37). Alternatively, when α^m is computed by shrinkage, we deduce that

$$\|A(x^{m+1} - x^m)\|^2 \leq \frac{1-c}{2\alpha^m} \|x^{m+1} - x^m\|^2.$$

Substituting this and (3.13) into (3.14), we obtain

$$\Psi(x^{m+1}) - \Psi(x^m) \leq -\frac{1}{2\alpha^m} \|x^{m+1} - x^m\|^2 + \frac{1-c}{2\alpha^m} \|x^{m+1} - x^m\|^2 = -\frac{c}{2\alpha^m} \|x^{m+1} - x^m\|^2. \quad (3.39)$$

Thus (3.38), (3.39) and $c \in (0, 1)$ imply that for all $m \geq 0$,

$$\|x^{m+1} - x^m\|^2 \leq \frac{2\alpha^m}{c} [\Psi(x^m) - \Psi(x^{m+1})] \leq \frac{2(1-c)}{c(1-L_{2k})} [\Psi(x^m) - \Psi(x^{m+1})],$$

due to (3.35). Hence (3.15) holds with $d \stackrel{\text{def}}{=} 2(1-c)/[c(1-L_{2k})]$, and so does (3.16) due to (3.35). Lemma 3.5 applies and together with (3.35) provides the required conclusion. \square

Note that due to (3.35), the shrinkage strategy, rather than the exact linesearch, determines the value of $\underline{\alpha}$ in Theorem 3.7, which is crucial for our phase transitions. However, we cannot guarantee that the less-conservative exact linesearch strategy is taken asymptotically.

3.3 Deterministic recovery conditions

In the case of zero measurement noise, combining the two parts of our analysis in Sections 3.1 and 3.2 respectively leads immediately to recovery conditions for both IHT and N-IHT.

Theorem 3.8. *Consider problem (1.1) with $e \stackrel{\text{def}}{=} 0$ and let $\Lambda = \text{supp}(x^*)$. Suppose that **A.1** holds, that the stepsize α of IHT satisfies (3.32), and that*

$$\left\| A_{\Gamma}^{\dagger} A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* \right\| < \alpha \left\| A_{\Lambda \setminus \Gamma}^T (I - A_{\Gamma} A_{\Gamma}^{\dagger}) A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* \right\| \quad (3.40)$$

for all $\Gamma \neq \Lambda$ such that $|\Gamma| = k$, where A_{Γ}^{\dagger} is defined in (3.4). Then the IHT iterates $\{x^m\}$ converge to its only fixed point, namely, the original signal x^* .

Proof. Under Assumption **A.1**, Corollary 3.3 and (3.40) imply that there exists no α -stable point on any $\Gamma \neq \Lambda$ such that $|\Gamma| = k$. It follows that any α -stable point is supported on Λ , and therefore by Lemma 3.1, it must coincide with x^* . Also under Assumption **A.1**, it follows from (3.32) and Theorem 3.6 that IHT converges to an α -stable point, and hence to x^* . Since a fixed point of IHT with stepsize α is the same as an α -stable point, we conclude the proof. \square

Theorem 3.9. Consider problem (1.1) with $e \stackrel{\text{def}}{=} 0$ and let $\Lambda = \text{supp}(x^*)$. Suppose that **A.1** holds and that

$$\left\| A_{\Gamma}^{\dagger} A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* \right\| < [\kappa(1 + U_{2k})]^{-1} \left\| A_{\Lambda \setminus \Gamma}^T (I - A_{\Gamma} A_{\Gamma}^{\dagger}) A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* \right\| \quad (3.41)$$

for all $\Gamma \neq \Lambda$ such that $|\Gamma| = k$, where A_{Γ}^{\dagger} is defined in (3.4). Then the N-IHT iterates $\{x^m\}$ converge to the original signal x^* .

Proof. Under Assumption **A.1**, Corollary 3.3 and (3.41) imply that there exists no $[\kappa(1 + U_{2k})]^{-1}$ -stable point on any $\Gamma \neq \Lambda$ such that $|\Gamma| = k$. It follows that any $[\kappa(1 + U_{2k})]^{-1}$ -stable point is supported on Λ , and therefore by Lemma 3.1 it must be x^* . Also under Assumption **A.1**, Theorem 3.7 implies that we also have convergence to an $[\kappa(1 + U_{2k})]^{-1}$ -stable point, which concludes the proof. \square

While Theorems 3.8 and 3.9 give conditions guaranteeing recovery, what is less clear is when one might expect these conditions to be satisfied. We provide answers to this question in the rest of the paper, quantifying when these conditions are satisfied in the case of Gaussian matrices. Furthermore, we also extend our analysis for Gaussian matrices to the case of measurements contaminated by Gaussian noise.

4 Probabilistic quantification of the deterministic analysis

4.1 Distribution results for the stable point condition

The aim of this section is to derive distribution results in the context of Gaussian measurement matrices for each of the terms in the stable point condition (3.5) of Theorem 3.2. We first give some definitions of Gaussian and Gaussian-related matrix variate distributions, along with some fundamental results concerning their Rayleigh quotients when applied to independent vectors.

We consider a particular kind of matrix variate Gaussian distribution in which all entries are i.i.d. Gaussian random variables, and a few other related distributions.

Definition 4.1 (Matrix variate Gaussian distribution [1]). We say that an $s \times t$ matrix B follows the matrix variate Gaussian distribution $B \sim \mathcal{N}_{s,t}(\mu, \sigma^2)$, if each entry of B independently follows the (univariate) Gaussian distribution $B_{ij} \sim \mathcal{N}(\mu, \sigma^2)$.

Definition 4.2 (Matrix variate Wishart distribution [1]). Let $B \sim \mathcal{N}_{s,t}(\mu, \sigma^2)$ such that $s \geq t$. Then we say that $B^T B$ follows a matrix variate Wishart distribution $\mathcal{W}_t(s; \mu, \sigma^2)$ with s degrees of freedom, mean μ and variance σ^2 .

Definition 4.3 (χ^2 and F distributions [1, pp.940,946]). Given a positive integer s , let $Z_i \sim \mathcal{N}(0, 1)$ be independent random variables for $1 \leq i \leq s$. Then we say $P = Z_1^2 + Z_2^2 + \dots + Z_s^2$ follows a chi-squared distribution with s degrees of freedom, and we write $P \sim \chi_s^2$. Furthermore, given positive integers s and t , if $P \sim \frac{1}{s}\chi_s^2$ and $Q \sim \frac{1}{t}\chi_t^2$ are independent random variables, we say that P/Q follows the F -distribution, and we write $P/Q \sim \mathcal{F}(s, t)$.

Useful results concerning the distributions of Rayleigh quotients related to Gaussian and Wishart matrices are given in the next lemma.

Lemma 4.1 (Distributions of Rayleigh quotients). Let $B \sim \mathcal{N}_{s,t}(0, \sigma^2)$ with $s \geq t$. Let $z \in \mathbb{R}^t$ be independent of B , and such that $\mathbb{P}(z \neq 0) = 1$. Then

$$\frac{z^T B^T B z}{z^T z} \sim \sigma^2 \chi_s^2 \text{ and is independent of } z; \quad (4.1)$$

$$\frac{z^T z}{z^T (B^T B)^{-1} z} \sim \sigma^2 \chi_{s-t+1}^2 \text{ and is independent of } z; \quad (4.2)$$

$$\frac{z^T (B^T B)^2 z}{z^T z} \text{ has the same distribution as } \{(B^T B)^2\}_{11}. \quad (4.3)$$

Proof. Let $B \sim \mathcal{N}_{s,t}(0, \sigma^2)$ with $s \geq t$. [22, Theorem 3.3.12] gives a more general result than (4.1) for when the entries of B are not necessarily independent. The present result follows by setting $\Sigma = \sigma^2 I$ for the covariance matrix. Similarly, (4.2) follows by setting $\Sigma = \sigma^2 I$ in [22, Corollary 3.3.14.1]. To prove (4.3), let $S = B^T B$ so that $S \sim \mathcal{W}_t(s; 0, \sigma^2)$ and let $Z \in \mathbb{R}^{t \times t}$ be any orthogonal matrix which is independent of B . Lemma 4.2 yields $BZ \sim \mathcal{N}_{s,t}(0, \sigma^2)$ independently of Z , and, writing $T := Z^T S Z$, we therefore have

$$T = Z^T S Z = Z^T B^T B Z = (BZ)^T BZ \sim \mathcal{W}_t(s; 0, \sigma^2), \quad (4.4)$$

independently of Z . In particular, let us fix the first column of Z as z normalized so that

$$Z = \begin{bmatrix} \frac{z}{\|z\|} & Z_2 \end{bmatrix},$$

which leads to

$$\begin{aligned} z^T S^2 z &= z^T (Z T Z^T)^2 z = z^T Z T Z^T Z T Z^T z = z^T Z T^2 Z^T z \\ &= z^T \begin{bmatrix} \frac{z}{\|z\|} & Z_2 \end{bmatrix} T^2 \begin{bmatrix} \frac{z^T}{\|z\|} \\ Z_2^T \end{bmatrix} z = \begin{bmatrix} \|z\| & 0 \end{bmatrix} T^2 \begin{bmatrix} \frac{\|z\|}{0} \end{bmatrix} = (T^2)_{11} \|z\|^2. \end{aligned}$$

Dividing by $\|z\|^2$ and using (4.4) then gives the desired result. \square

Crucial to our argument will be the well-known result that the central matrix variate Gaussian distribution defined in Definition 4.1 is invariant under transformation by an independent orthogonal matrix.

Lemma 4.2 (Orthogonal invariance [18]). *Let $B \sim \mathcal{N}_{s,t}(0, \sigma^2)$ and let $Z_1 \in \mathbb{R}^{s \times s}$ and $Z_2 \in \mathbb{R}^{t \times t}$ be orthogonal and independent of B . Then*

$$Z_1 B \sim \mathcal{N}_{s,t}(0, \sigma^2), \quad \text{independently of } Z_1, \quad (4.5)$$

and

$$B Z_2 \sim \mathcal{N}_{s,t}(0, \sigma^2), \quad \text{independently of } Z_2. \quad (4.6)$$

We now make the assumption that the measurement matrix in (1.1) is drawn from the (central) matrix variate Gaussian distribution with appropriate normalization.

A.2 The measurement matrix A has i.i.d. $\mathcal{N}(0, 1/n)$ entries, so that $A \sim \mathcal{N}_{n,N}(0, 1/n)$. Furthermore, A is independent of x^* .

Given Assumption **A.2** and the standard compressed sensing regime with $2k \leq n$, we can dispense with Assumption **A.1** [22, Theorem 3.2.1]⁴.

We also impose the additional assumption that measurement noise is itself Gaussian and independent of both the original signal and the measurement matrix.

A.3 The noise vector e has i.i.d. Gaussian entries $e_i \sim \mathcal{N}(0, \sigma^2/n)$, independently of A and x^* .

Note that, under Assumption **A.3**, $\mathbb{E}\|e\|^2 = \sigma^2$, so that $\|e\| \approx \sigma$.

We now give the main result of this section, in which we derive precise distributions for various expressions which make up the stable point condition (3.5) of Theorem 3.2, in terms of the χ^2 and F distributions.

Lemma 4.3 (Distribution results for the stable point condition). *Suppose Assumptions **A.2** and **A.3** hold, and let Γ and Λ be index sets of cardinality k , where $k < n$, such that $\Gamma \neq \Lambda$. Then*

$$\frac{\|A_{\Gamma}^\dagger A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^*\|}{\|x_{\Lambda \setminus \Gamma}^*\|} = \sqrt{F_\Gamma}, \quad \text{where } F_\Gamma \sim \frac{k}{n-k+1} F(k, n-k+1); \quad (4.7)$$

$$\frac{\|A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^*\|}{\|x_{\Lambda \setminus \Gamma}^*\|} \geq \left(\frac{n-k}{n} \right) \cdot R_\Gamma, \quad \text{where } R_\Gamma \sim \frac{1}{n-k} \chi_{n-k}^2; \quad (4.8)$$

⁴ [22, Theorem 3.2.1] states that $B^T B$ is positive definite with probability 1 when $B \sim \mathcal{N}_{s,t}(0, \sigma^2)$ with $s \geq t$.

$$\|A_\Gamma^\dagger e\| \leq \sigma \cdot \sqrt{G_\Gamma}, \quad \text{where } G_\Gamma \sim \frac{k}{n-k+1} F(k, n-k+1); \quad (4.9)$$

$$\|A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) e\| \leq \sigma \sqrt{\frac{k(n-k)}{n^2} \cdot (S_\Gamma)(T_\Gamma)}, \quad \text{where } S_\Gamma \sim \frac{1}{n-k} \chi_{n-k}^2, \quad T_\Gamma \sim \frac{1}{k} \chi_k^2. \quad (4.10)$$

Proof of (4.7): Let A_Γ have the singular value decomposition

$$A_\Gamma := U[D \mid 0]V^T = U_1 D V^T, \quad (4.11)$$

where $D \in \mathbb{R}^{k \times k}$ is diagonal, and where $V \in \mathbb{R}^{k \times k}$ and $U = [U_1 \mid U_2] \in \mathbb{R}^{n \times n}$ are orthogonal, with $U_1 \in \mathbb{R}^{n \times k}$. By Assumption **A.2**, A_Γ^\dagger is well-defined and we have the standard result

$$A_\Gamma^\dagger = V D^{-1} U_1^T, \quad (4.12)$$

and since $(A_\Gamma^T A_\Gamma)^{-1} = V D^{-2} V^T$, it follows by rearrangement that

$$D^{-2} = V^T (A_\Gamma^T A_\Gamma)^{-1} V. \quad (4.13)$$

Using (4.12) and (4.13), we have

$$\begin{aligned} \|A_\Gamma^\dagger A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^*\|^2 &= (x_{\Lambda \setminus \Gamma}^*)^T A_{\Lambda \setminus \Gamma}^T (A_\Gamma^\dagger)^T (A_\Gamma^\dagger) A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* \\ &= (x_{\Lambda \setminus \Gamma}^*)^T A_{\Lambda \setminus \Gamma}^T U_1 D^{-1} V^T V D^{-1} U_1^T A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* \\ &= (x_{\Lambda \setminus \Gamma}^*)^T A_{\Lambda \setminus \Gamma}^T U_1 D^{-2} U_1^T A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* \\ &= (x_{\Lambda \setminus \Gamma}^*)^T A_{\Lambda \setminus \Gamma}^T U_1 V^T (A_\Gamma^T A_\Gamma)^{-1} V U_1^T A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^*. \end{aligned} \quad (4.14)$$

By Lemma 4.2, we have $U^T A_{\Lambda \setminus \Gamma} \sim \mathcal{N}_{n,r}(0, 1/n)$, independently of U , where $r := |\Lambda \setminus \Gamma|$. Since $U_1^T A_{\Lambda \setminus \Gamma}$ is a submatrix of $U^T A_{\Lambda \setminus \Gamma}$, it follows that $U_1^T A_{\Lambda \setminus \Gamma} \sim \mathcal{N}_{k,r}(0, 1/n)$, independently of U . Writing $C := V U_1^T A_{\Lambda \setminus \Gamma} \in \mathbb{R}^{k \times r}$, we also have by Lemma 4.2 that $C \sim \mathcal{N}_{k,r}(0, 1/n)$, independently of both U and V , and therefore independently of A_Γ . Substituting for C in (4.14), we have

$$\begin{aligned} \frac{\|A_\Gamma^\dagger A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^*\|^2}{\|x_{\Lambda \setminus \Gamma}^*\|^2} &= \frac{(x_{\Lambda \setminus \Gamma}^*)^T C^T (A_\Gamma^T A_\Gamma)^{-1} C x_{\Lambda \setminus \Gamma}^*}{(x_{\Lambda \setminus \Gamma}^*)^T x_{\Lambda \setminus \Gamma}^*} \\ &= \frac{(x_{\Lambda \setminus \Gamma}^*)^T C^T (A_\Gamma^T A_\Gamma)^{-1} C x_{\Lambda \setminus \Gamma}^*}{(x_{\Lambda \setminus \Gamma}^*)^T C^T C x_{\Lambda \setminus \Gamma}^*} \cdot \frac{(x_{\Lambda \setminus \Gamma}^*)^T C^T C x_{\Lambda \setminus \Gamma}^*}{(x_{\Lambda \setminus \Gamma}^*)^T x_{\Lambda \setminus \Gamma}^*}, \end{aligned} \quad (4.15)$$

where x^* , C and A_Γ are all independent. Now it follows from Lemma 4.1 that

$$\frac{(x_{\Lambda \setminus \Gamma}^*)^T C^T C x_{\Lambda \setminus \Gamma}^*}{(x_{\Lambda \setminus \Gamma}^*)^T x_{\Lambda \setminus \Gamma}^*} \sim \frac{1}{n} \chi_k^2 \quad \text{and} \quad \frac{(x_{\Lambda \setminus \Gamma}^*)^T C^T C x_{\Lambda \setminus \Gamma}^*}{(x_{\Lambda \setminus \Gamma}^*)^T C^T (A_\Gamma^T A_\Gamma)^{-1} C x_{\Lambda \setminus \Gamma}^*} \sim \frac{1}{n} \chi_{n-k+1}^2, \quad (4.16)$$

where both distributions are independent of each other. Combining (4.15) and (4.16) leads us to conclude

$$\frac{\|A_\Gamma^\dagger A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^*\|^2}{\|x_{\Lambda \setminus \Gamma}^*\|^2} \sim \frac{\chi_k^2}{\chi_{n-k+1}^2} = \frac{k}{n-k+1} \mathcal{F}(k, n-k+1),$$

where in the last step we use the fact that the two distributions are independent, which proves (4.7).

Proof of (4.8): Using (4.11) and (4.12), we have

$$A_\Gamma A_\Gamma^\dagger = U_1 D V^T V D^{-1} U_1^T = U_1 U_1^T = U \begin{bmatrix} I & 0 \\ 0 & 0 \end{bmatrix} U^T,$$

and writing $I = U U^T$,

$$I - A_\Gamma A_\Gamma^\dagger = U \left\{ \begin{bmatrix} I & 0 \\ 0 & I \end{bmatrix} - \begin{bmatrix} I & 0 \\ 0 & 0 \end{bmatrix} \right\} U^T = U \begin{bmatrix} 0 & 0 \\ 0 & I \end{bmatrix} U^T = U_2 U_2^T, \quad (4.17)$$

which in turn gives

$$A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) A_{\Lambda \setminus \Gamma} = A_{\Lambda \setminus \Gamma}^T U_2 U_2^T A_{\Lambda \setminus \Gamma}. \quad (4.18)$$

Writing $F := U_2^T A_{\Lambda \setminus \Gamma}$, we have $U^T A_{\Lambda \setminus \Gamma} \sim \mathcal{N}_{n,r}(0, 1/n)$ by Lemma 4.2, and since $U_2^T A_{\Lambda \setminus \Gamma} \in \mathbb{R}^{(n-k) \times r}$ is a submatrix of $U^T A_{\Lambda \setminus \Gamma}$, it follows that

$$F \sim \mathcal{N}_{(n-k),r}(0, 1/n). \quad (4.19)$$

Substituting for F in (4.18) gives

$$A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) A_{\Lambda \setminus \Gamma} = F^T F. \quad (4.20)$$

Now, writing $M := F^T F$, and using (4.20) and (4.3) of Lemma 4.1, we deduce

$$\frac{\|A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^*\|^2}{\|x_{\Lambda \setminus \Gamma}^*\|^2} = \frac{\|F^T F x_{\Lambda \setminus \Gamma}\|^2}{\|x_{\Lambda \setminus \Gamma}^*\|^2} = \frac{(x_{\Lambda \setminus \Gamma}^*)^T (F^T F)^2 x_{\Lambda \setminus \Gamma}^*}{(x_{\Lambda \setminus \Gamma}^*)^T x_{\Lambda \setminus \Gamma}^*} \sim (M^2)_{11}. \quad (4.21)$$

To obtain a lower bound in terms of the chi-squared distribution, note that

$$(M^2)_{11} = \sum_{i=1}^r M_{i1}^2 = M_{11}^2 + \sum_{i=2}^r M_{i1}^2 \geq M_{11}^2. \quad (4.22)$$

Meanwhile it follows from (4.19) and (4.3) that

$$M_{11} = \sum_{i=1}^{n-k} F_{i1}^2 \sim \frac{1}{n} \chi_{n-k}^2,$$

which combines with (4.21) and (4.22) to give (4.8).

Proof of (4.9): By (4.12), we have

$$A_\Gamma^\dagger e = V D^{-1} U_1^T e = V D^{-1} p, \quad (4.23)$$

where $p := U_1^T e \in \mathbb{R}^{n-k}$. Using Assumption **A.3**, we may view e as a one-column Gaussian matrix, such that $e \sim \mathcal{N}_{n,1}(0, \sigma^2/n)$, it follows from Lemma 4.2 that

$$p \sim \mathcal{N}_{k,1}(0, \sigma^2/n), \quad (4.24)$$

independently of U and therefore independently of A_Γ . Substituting (4.13) into (4.23) then gives

$$\|A_\Gamma^\dagger e\|^2 = \|V D^{-1} p\|^2 = \|D^{-1} p\|^2 = p^T D^{-2} p = p^T V^T (A_\Gamma^T A_\Gamma)^{-1} V p = q^T (A_\Gamma^T A_\Gamma)^{-1} q, \quad (4.25)$$

where $q := V p \in \mathbb{R}^k$. It now follows from (4.24) and Lemma 4.2 that $q \sim \mathcal{N}_{k,1}(0, \sigma^2/n)$, independently of V and therefore independently of A_Γ , and consequently that

$$q^T q \sim \sigma^2 \chi_k^2. \quad (4.26)$$

By (4.2) of Lemma 4.1,

$$\frac{q^T q}{q^T (A_\Gamma^T A_\Gamma)^{-1} q} \sim \frac{1}{n} \chi_{n-k+1}^2. \quad (4.27)$$

Since q and A_Γ are independent, we may combine (4.25), (4.26) and (4.27) to give

$$\|A_\Gamma^\dagger e\| \sim \sigma \sqrt{G_\Gamma}, \quad \text{where } G_\Gamma \sim \frac{k}{n-k+1} F(k, n-k+1), \quad (4.28)$$

and (3.9) now follows.

Proof of (4.10): Using (4.17), we have

$$A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) e = A_{\Lambda \setminus \Gamma}^T U_2 U_2^T e = A_{\Lambda \setminus \Gamma}^T U_2 f = B^T f, \quad (4.29)$$

where $B := U_2^T A_{\Lambda \setminus \Gamma} \sim \mathcal{N}_{n-k, r}(0, 1/n)$ by Lemma 4.2, and where

$$f := U_2^T e \sim \mathcal{N}_{n-k, 1}(0, \sigma^2/n) \quad (4.30)$$

by Lemma 4.2. Now let B have singular value decomposition

$$W[F \mid 0]Y^T = W_1 F Y^T, \quad (4.31)$$

where $F \in \mathbb{R}^{r \times r}$ is diagonal, and where $Y \in \mathbb{R}^{r \times r}$ and $W = [W_1 \mid W_2] \in \mathbb{R}^{(n-k) \times (n-k)}$ are orthogonal, noting that $W_1 \in \mathbb{R}^{(n-k) \times r}$. We have

$$g := W_1^T f \sim \mathcal{N}_{r, 1}(0, \sigma^2/n) \quad (4.32)$$

by (4.30) and Lemma 4.2, and we may apply (4.29) to give

$$\begin{aligned} \|A_{\Lambda \setminus \Gamma}^T (I - A_\Gamma A_\Gamma^\dagger) e\|^2 &\leq \|B^T f\|^2 = \|Y F W_1^T f\|^2 \\ &= \|F g\|^2 = g^T F^2 g = g^T Y^T (B^T B) Y g = h^T (B^T B) h, \end{aligned} \quad (4.33)$$

where $h := Y g \in \mathbb{R}^k$. Since $h \sim \mathcal{N}_{r, 1}(0, \sigma^2/n)$ by (4.32) and Lemma 4.2, it follows that

$$h^T h \sim \sigma^2 \chi_r^2 \leq \sigma^2 \chi_k^2, \quad (4.34)$$

since a χ_r^2 random variate may be viewed as a truncation of its extension to a χ_k^2 random variate. By (4.1) of Lemma 4.1,

$$\frac{h^T B^T B h}{h^T h} \sim \frac{1}{n} \chi_{n-k}^2. \quad (4.35)$$

Combining (4.33), (4.34) and (4.35) then proves (4.10). \square

In order for the (converse of the) stable point condition (3.5) to provide a recovery result regarding the proximity of all stable points to the underlying signal, we need to quantify the quantities in Lemma 4.3 on all possible fixed points on support sets Γ of cardinality k . Similarly, the convergence conditions in Section 3.2 involve RIP constants which again involve looking over combinatorially many supports. Thus we need to derive union bounds for the relevant distributions involved in the stable point and convergence conditions.

4.2 Large deviation results involving Gaussian matrices

In this section, we derive large deviations results for quantities relating to Gaussian matrices within the proportional-growth asymptotic that is defined on page 2. We define three tail bound functions.

Definition 4.4 (χ^2 tail bounds). Let $\delta \in (0, 1]$, $\rho \in (0, 1)$ and $\lambda \in (0, 1]$. Let $\mathcal{IU}(\delta, \rho, \lambda)$ be the unique solution to

$$\nu - \ln(1 + \nu) = \frac{2H(\delta\rho)}{\lambda} \quad \text{for } \nu > 0, \quad (4.36)$$

and let $\mathcal{IL}(\delta, \rho, \lambda)$ be the unique solution to

$$-\nu - \ln(1 - \nu) = \frac{2H(\delta\rho)}{\lambda} \quad \text{for } \nu \in (0, 1), \quad (4.37)$$

where $H(\cdot)$ is the Shannon entropy with base e logarithms [5], namely,

$$H(p) := -p \ln(p) - (1 - p) \ln(1 - p). \quad (4.38)$$

That \mathcal{IU} is well-defined follows since the left-hand side of (4.36) is zero at $\nu = 0$, tends to infinity as $\nu \rightarrow \infty$, and is strictly increasing on $\nu > 0$. Similarly, \mathcal{IL} is well-defined since the left-hand side of (4.37) is zero at $\nu = 0$, tends to infinity as $\nu \rightarrow 1$, and is strictly increasing on $\nu \in (0, 1)$.

Definition 4.5 (F tail bound). Let $\delta \in (0, 1]$ and $\rho \in (0, 1/2]$. Let $\mathcal{IF}(\delta, \rho)$ be the unique solution in f to

$$\ln(1 + f) - \rho \ln f = 2H(\delta\rho) + H(\rho) \quad \text{for } f > \frac{\rho}{1 - \rho}, \quad (4.39)$$

where $H(\cdot)$ is defined in (4.38).

That \mathcal{IF} is well-defined follows since the left-hand side of (4.39) is equal to $H(\rho)$ at $f = \rho/(1 - \rho)$, tends to infinity as $f \rightarrow \infty$, and is strictly increasing on $f > \rho/(1 - \rho)$.

Defining S_n as

$$S_n \stackrel{\text{def}}{=} \left\{ 1, \dots, \binom{N}{k} \right\}, \quad (4.40)$$

we have the following large deviation bound for a combinatorial number of χ^2 distributions.

Lemma 4.4 (Large deviations result for χ^2). Let $l \in \{1, \dots, n\}$ and let the random variables $X_l^i \sim \frac{1}{l} \chi_l^2$ for all $i \in S_n$, and let $\epsilon > 0$. In the proportional-growth asymptotic, let $l/n \rightarrow \lambda \in (0, 1]$. Then

$$\mathbb{P}\{\cup_{i \in S_n} [X_l^i \geq 1 + \mathcal{IU}(\delta, \rho, \lambda) + \epsilon]\} \rightarrow 0 \quad (4.41)$$

and

$$\mathbb{P}\{\cup_{i \in S_n} [X_l^i \leq 1 - \mathcal{IL}(\delta, \rho, \lambda) - \epsilon]\} \rightarrow 0, \quad (4.42)$$

exponentially in n , where $\mathcal{IU}(\delta, \rho, \lambda)$ and $\mathcal{IL}(\delta, \rho, \lambda)$ are defined in (4.36) and (4.37) respectively.

The proof of Lemma 4.4 is delegated to Appendix A. It employs asymptotic results derived by Temme [29] for the incomplete gamma function which is related to the χ^2 distribution.

Lemma 4.5 (Large deviations result for F). Let the random variables $X_n^i \sim \frac{k}{n-k+1} \mathcal{F}(k, n-k+1)$ for all $i \in S_n$, and let $\epsilon > 0$. In the proportional-growth asymptotic,

$$\mathbb{P}\{\cup_{i \in S_n} [X_n^i \geq \mathcal{IF}(\delta, \rho) + \epsilon]\} \rightarrow 0, \quad (4.43)$$

exponentially in n , where $\mathcal{IF}(\delta, \rho)$ is defined in (4.39).

The proof of Lemma 4.5 is delegated to Appendix A. It employs asymptotic results derived by Temme [29] for the incomplete beta function which is related to the F distribution.

We will also make use of upper bounds on RIP constants of Gaussian matrices in the proportional-dimensional asymptotic.

Lemma 4.6 (Gaussian RIP bounds [2, Theorem 2.3]). Suppose $A \sim \mathcal{N}_{n,N}(0, 1/n)$ has RIP constants L_k and U_k as defined in Definition 3.2, and let the implicit but computable expressions $\mathcal{L}(\delta, \rho)$ and $\mathcal{U}(\delta, \rho)$ be defined as in [2, Definition 2.2]. Then, for any fixed ϵ , in the proportional-growth asymptotic,

$$\mathbb{P}[L_k < \mathcal{L}(\delta, \rho) + \epsilon] \rightarrow 1 \quad \text{and} \quad \mathbb{P}[U_k < \mathcal{U}(\delta, \rho) + \epsilon] \rightarrow 1,$$

exponentially in n .

Comparison of Lemmas 4.4 and 4.6; RIP versus Independent RIP constants. Suppose $A \sim \mathcal{N}_{n,N}(0, 1/n)$, let Γ be an index set of cardinality k and fix $\epsilon > 0$. Then in the conditions of Lemma 4.6, in the proportional-growth asymptotic, for any $y \in \mathbb{R}^k$,

$$1 - \mathcal{L}(\delta, \rho) - \epsilon < \frac{\|A_\Gamma y\|^2}{\|y\|^2} < 1 + \mathcal{U}(\delta, \rho) + \epsilon. \quad (4.44)$$

However, if y is independent of A , we may set $\lambda = 1$ in Lemma 4.4, giving in the proportional-growth asymptotic,

$$1 - \mathcal{IL}(\delta, \rho, 1) - \epsilon < \frac{\|A_{\Gamma}y\|^2}{\|y\|^2} < 1 + \mathcal{IU}(\delta, \rho, 1) + \epsilon. \quad (4.45)$$

Comparing (4.44) and (4.45), we see that $\mathcal{IU}(\delta, \rho, 1)$ and $\mathcal{IL}(\delta, \rho, 1)$ may be viewed as upper bounds on ‘independent RIP’ constants for Gaussian matrices.

Figure 4.1 gives plots of the ‘independent RIP’ bounds for Gaussian matrices $\mathcal{IU}(\delta, \rho, 1)$ and $\mathcal{IL}(\delta, \rho, 1)$ derived in this paper, along with plots of the RIP bounds for Gaussian matrices $\mathcal{U}(\delta, \rho)$ and $\mathcal{L}(\delta, \rho)$ in [2]. One observes empirically the inequalities

$$\mathcal{IU}(\delta, \rho, 1) < \mathcal{U}(\delta, \rho) \quad \text{and} \quad \mathcal{IL}(\delta, \rho, 1) < \mathcal{L}(\delta, \rho).$$

A simple interpretation is that the additional information that the matrix and vector are independent allows us to tighten the bounds in (4.44) to obtain (4.45). This consideration accounts for a large part of the quantitative improvement that is obtained in this paper over existing recovery results for IHT algorithms which rely solely upon the RIP. Of course, our improved analysis is only possible because our proposed stable point condition (3.5) can exploit the assumption of matrix-vector independence.

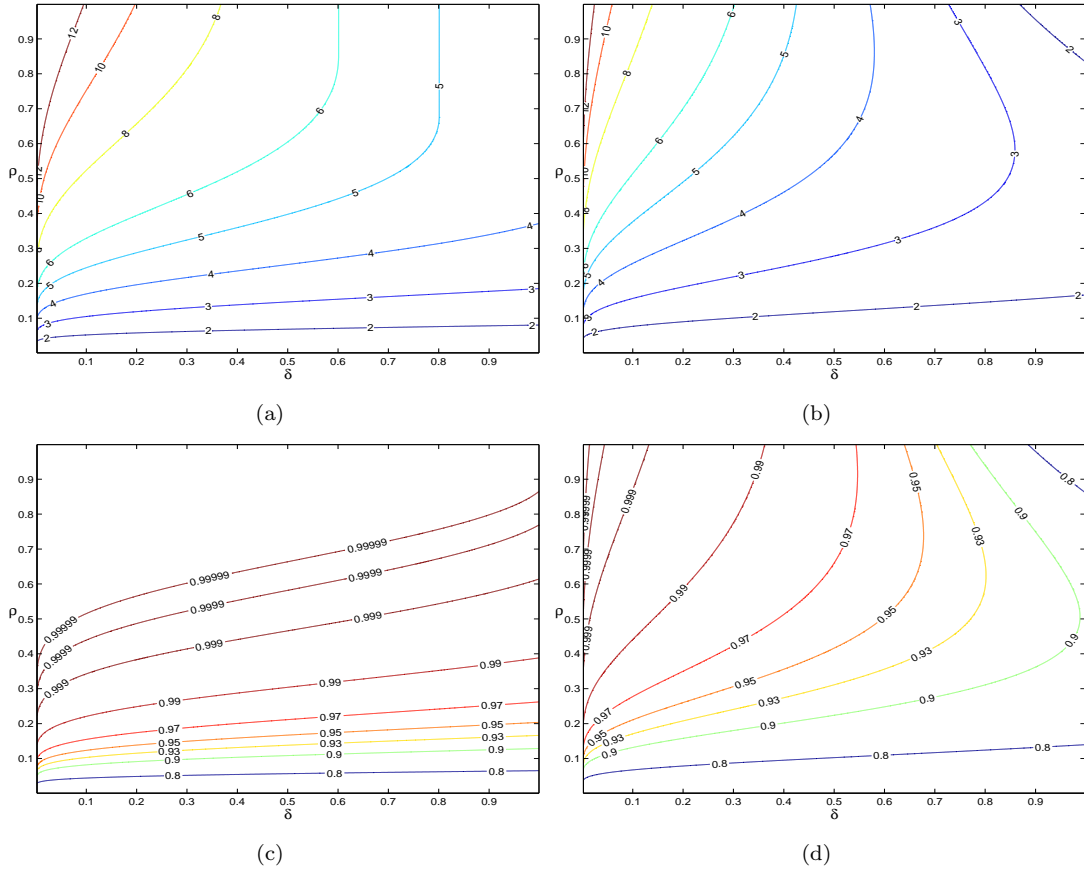


Figure 4.1: A comparison of standard RIP bounds [2] and ‘independent RIP’ bounds for Gaussian matrices: (a) $\mathcal{U}(\delta, \rho)$ (b) $\mathcal{IU}(\delta, \rho, 1)$ (c) $\mathcal{L}(\delta, \rho)$ (d) $\mathcal{IL}(\delta, \rho, 1)$.

5 Novel recovery analysis for IHT algorithms

5.1 Definitions of recovery phase transitions

In this section, we define asymptotic recovery phase transitions for G-IHT variants.

Definition 5.1 (Phase transition functions for IHT and N-IHT). *Given $\delta \in (0, 1]$, define $\hat{\rho}^{IHT}(\delta)$ to be the unique solution to*

$$\frac{\sqrt{\mathcal{IF}(\delta, \rho)}}{(1 - \rho)[1 - \mathcal{IL}(\delta, \rho, 1 - \rho)]} = \frac{1}{1 + \mathcal{U}(\delta, 2\rho)} \quad \text{for } \rho \in (0, 1/2], \quad (5.1)$$

and define $\hat{\rho}^{N-IHT}(\delta)$ to be the unique solution to

$$\frac{\sqrt{\mathcal{IF}(\delta, \rho)}}{(1 - \rho)[1 - \mathcal{IL}(\delta, \rho, 1 - \rho)]} = \frac{1}{\kappa[1 + \mathcal{U}(\delta, 2\rho)]} \quad \text{for } \rho \in (0, 1/2], \quad (5.2)$$

where \mathcal{IF} is defined in (4.39), \mathcal{IL} is defined in (4.37), \mathcal{U} is defined in [2, Definition 2.2], and κ is an N-IHT algorithm parameter.

A proof that $\hat{\rho}^{IHT}(\delta)$ and $\hat{\rho}^{N-IHT}(\delta)$ are well-defined can be found in [30, Section 5.2], and plots of these functions are given in Section 6.

We proceed to our main recovery results for IHT algorithms, beginning with constant stepsize IHT.

5.2 Recovery results for IHT

We begin by introducing several definitions.

Definition 5.2 (Stability factor for IHT). *Given $\delta \in (0, 1]$, $\rho \in (0, 1/2]$ and $\alpha > 0$, provided*

$$\alpha > \frac{\sqrt{\mathcal{IF}(\delta, \rho)}}{(1 - \rho)[1 - \mathcal{IL}(\delta, \rho, 1 - \rho)]}, \quad (5.3)$$

define

$$a(\delta, \rho) \stackrel{\text{def}}{=} \frac{\sqrt{\mathcal{IF}(\delta, \rho)} + \alpha \sqrt{\rho(1 - \rho)[1 + \mathcal{IU}(\delta, \rho, 1 - \rho)][1 + \mathcal{IU}(\delta, \rho, \rho)]}}{\alpha(1 - \rho)[1 - \mathcal{IL}(\delta, \rho, 1 - \rho)] - \sqrt{\mathcal{IF}(\delta, \rho)}}, \quad (5.4)$$

and

$$\xi(\delta, \rho) \stackrel{\text{def}}{=} \sqrt{\mathcal{IF}(\delta, \rho)[1 + a(\delta, \rho)]^2 + [a(\delta, \rho)]^2}, \quad (5.5)$$

where \mathcal{IF} is defined in (4.39), \mathcal{IU} is defined in (4.36), and where \mathcal{IL} is defined in (4.37).

Note that (5.3) ensures that the denominator in (5.4) is strictly positive and that $a(\delta, \rho)$ is therefore well-defined. The function $\xi(\delta, \rho)$ will represent a stability factor in our results, bounding the approximation error of the output of IHT as a multiple of the noise level σ .

Definition 5.3 (Support set partition for IHT). *Suppose $\delta \in (0, 1]$, $\rho \in (0, 1/2]$ and $\alpha > 0$. Given $\zeta > 0$, let us write*

$$a^*(\delta, \rho; \zeta) \stackrel{\text{def}}{=} a(\delta, \rho) + \zeta, \quad (5.6)$$

where $a(\delta, \rho)$ is defined in (5.2), let us write $\{\Gamma_i : i \in S_n\}$ for the set of all possible support sets of cardinality k , and let us disjointly partition $S_n \stackrel{\text{def}}{=} \Theta_n^1 \cup \Theta_n^2$ such that

$$\Theta_n^1 \stackrel{\text{def}}{=} \left\{ i \in S_n : \|x_{\Lambda \setminus \Gamma_i}^*\| > \sigma \cdot a^*(\delta, \rho; \zeta) \right\}; \quad \Theta_n^2 \stackrel{\text{def}}{=} \left\{ i \in S_n : \|x_{\Lambda \setminus \Gamma_i}^*\| \leq \sigma \cdot a^*(\delta, \rho; \zeta) \right\}. \quad (5.7)$$

We recall that Λ is defined to be the support of the original signal x^* . Note that the partition $S_n := \Theta_n^1 \cup \Theta_n^2$ defined in (5.7) also depends on ζ , though we omit this dependency from our notation for the sake of brevity. Note also that if $\Gamma_i = \Lambda$, then $\|x_{\Lambda \setminus \Gamma_i}^*\| = 0$ and $i \in \Theta_n^2$. In other words, the index corresponding to Λ is contained in Θ_n^2 .

Let us outline how our argument will proceed. The partition in (5.7) has been defined in such a way that, provided (5.3) holds, an analysis of the stable point condition (3.5) shows that there are asymptotically no α -stable points on any Γ_i such that $i \in \Theta_n^1$, and this is proved in Lemma 5.1. On the other hand, it is also possible to use the large deviations results of Section 4.2 to bound the error in approximating x^* by any α -stable point on Γ_i such that $i \in \Theta_n^2$, which is achieved by Lemma 5.3. It follows that, for any $\alpha > 0$, all α -stable points have bounded approximation error. Finally, Lemma 5.4 builds on the convergence result in Theorem 3.6 and gives a condition on the stepsize α which asymptotically guarantees convergence of IHT to some α -stable point. Combining all three results, we have convergence to some α -stable point with guaranteed approximation error, provided the conditions in each lemma hold; combining the conditions leads to the phase transition defined in (5.1).

We first show that, asymptotically, there are no α -stable points on any Γ_i with $i \in \Theta_n^1$, and we write NSP_α for this event.

Lemma 5.1. *Choose $\zeta > 0$. Suppose Assumptions A.2 and A.3 hold, as well as (5.3). Then, in the proportional-growth asymptotic, there are no α -stable points on any Γ_i such that $i \in \Theta_n^1$, with probability tending to 1 exponentially in n .*

Proof. For any Γ_i such that $i \in \Theta_n^1$, we have $\Gamma_i \neq \Lambda$, and we may therefore use Theorem 3.2 and Lemma 4.3 with $\Gamma := \Gamma_i$ to deduce that a necessary condition for there to be an α -stable point on Γ_i is

$$\|x_{\Lambda \setminus \Gamma_i}^*\| \cdot \sqrt{F_{\Gamma_i}} + \sigma \cdot \sqrt{G_{\Gamma_i}} \geq \alpha \left[\left(\frac{n-k}{n} \right) \|x_{\Lambda \setminus \Gamma_i}^*\| \cdot R_{\Gamma_i} - \sigma \cdot \sqrt{\frac{k(n-k)}{n^2} \cdot S_{\Gamma_i} \cdot T_{\Gamma_i}} \right], \quad (5.8)$$

where

$$\begin{aligned} F_{\Gamma_i} &\sim \frac{k}{n-k+1} F(k, n-k+1); & G_{\Gamma_i} &\sim \frac{k}{n-k+1} F(k, n-k+1); \\ R_{\Gamma_i} &\sim \frac{1}{n-k} \chi_{n-k}^2; & S_{\Gamma_i} &\sim \frac{1}{n-k} \chi_{n-k}^2; & T_{\Gamma_i} &\sim \frac{1}{k} \chi_k^2. \end{aligned}$$

We also have, by (5.7),

$$\sigma \leq \frac{\|x_{\Lambda \setminus \Gamma_i}^*\|}{a^*(\delta, \rho; \zeta)} \quad (5.9)$$

for any Γ_i such that $i \in \Theta_n^1$. Since $\Gamma_i \neq \Lambda$, $\|x_{\Lambda \setminus \Gamma_i}^*\| > 0$, and substitution of (5.9) into (5.8), rearrangement and division by $\|x_{\Lambda \setminus \Gamma_i}^*\|$ yields

$$a^*(\delta, \rho; \zeta) \left[\alpha \left(\frac{n-k}{n} \right) \cdot R_{\Gamma_i} - \sqrt{F_{\Gamma_i}} \right] \leq \sqrt{G_{\Gamma_i}} + \alpha \sqrt{\frac{k(n-k)}{n^2} \cdot S_{\Gamma_i} \cdot T_{\Gamma_i}}.$$

Consequently,

$$\begin{aligned} &\mathbb{P}(\overline{NSP_\alpha}) \\ &= \mathbb{P}\{\cup_{i \in \Theta_n^1} (\exists \text{ an } \alpha\text{-stable point supported on } \Gamma_i)\} \\ &\leq \mathbb{P}\left\{ \bigcup_{i \in \Theta_n^1} \left[a^*(\delta, \rho; \zeta) \left[\alpha (1 - \rho_n) \cdot R_{\Gamma_i} - \sqrt{F_{\Gamma_i}} \right] \leq \sqrt{G_{\Gamma_i}} + \alpha \sqrt{\rho_n(1 - \rho_n) \cdot S_{\Gamma_i} \cdot T_{\Gamma_i}} \right] \right\}, \end{aligned} \quad (5.10)$$

where we write ρ_n for the sequence of values of the ratio k/n . For brevity's sake, let us define

$$\Phi[\rho, F, G, R, S, T] \stackrel{\text{def}}{=} \sqrt{G} + \alpha \sqrt{\rho(1 - \rho)(S)(T)} - a^*(\delta, \rho; \zeta) \cdot \left[\alpha(1 - \rho) \cdot R - \sqrt{F} \right], \quad (5.11)$$

so that (5.10) may be equivalently written as

$$\mathbb{P}(\overline{NSP_\alpha}) \leq \mathbb{P}\{\cup_{i \in \Theta_n^1} (\Phi[\rho_n, F_{\Gamma_i}, G_{\Gamma_i}, R_{\Gamma_i}, S_{\Gamma_i}, T_{\Gamma_i}] \geq 0)\}. \quad (5.12)$$

Given some $\epsilon > 0$, we now define

$$\begin{aligned} F^* &\stackrel{\text{def}}{=} G^* \stackrel{\text{def}}{=} \mathcal{IF}(\delta, \rho) + \epsilon; & R^* &\stackrel{\text{def}}{=} 1 - \mathcal{IL}(\delta, \rho, 1 - \rho) - \epsilon; \\ S^* &\stackrel{\text{def}}{=} 1 + \mathcal{IU}(\delta, \rho, 1 - \rho) + \epsilon; & T^* &\stackrel{\text{def}}{=} 1 + \mathcal{IU}(\delta, \rho, \rho) + \epsilon. \end{aligned} \quad (5.13)$$

Using (5.13), we deduce from (5.12) that

$$\begin{aligned} &\mathbb{P}(\overline{NSP_\alpha}) \\ &\leq \mathbb{P}\{\cup_{i \in \Theta_n^1} (\Phi[\rho_n, F_{\Gamma_i}, G_{\Gamma_i}, R_{\Gamma_i}, S_{\Gamma_i}, T_{\Gamma_i}] \geq \Phi[\rho_n, F^*, G^*, R^*, S^*, T^*])\} \end{aligned} \quad (5.14)$$

$$+ \mathbb{P}\{\Phi[\rho_n, F^*, G^*, R^*, S^*, T^*] \geq \Phi[\rho, F^*, G^*, R^*, S^*, T^*] + \epsilon\} \quad (5.15)$$

$$+ \mathbb{P}\{\Phi[\rho, F^*, G^*, R^*, S^*, T^*] + \epsilon \geq 0\}, \quad (5.16)$$

since the event in the right-hand side of (5.12) lies in the union of the three events in (5.14), (5.15) and (5.16). Now (5.16) is a deterministic event, and $a^*(\delta, \rho; \zeta)$ has been defined in such a way that, for any $\zeta > 0$, provided ϵ is taken sufficiently small, the event has probability 0. This follows from (5.3), (5.4), (5.6), and by the continuity of Φ . The event (5.15) is also deterministic, and by continuity and since $\rho_n \rightarrow \rho$, it follows that there exists some \tilde{n} such that

$$\mathbb{P}\{\Phi[\rho_n, F^*, G^*, R^*, S^*, T^*] \geq \Phi[\rho, F^*, G^*, R^*, S^*, T^*] + \epsilon\} = 0 \quad \text{for all } n \geq \tilde{n}.$$

Taking limits as $n \rightarrow \infty$, the terms (5.15) and (5.16) are zero, leaving only (5.14), and we have

$$\begin{aligned} &\lim_{n \rightarrow \infty} \mathbb{P}(\overline{NSP_\alpha}) \\ &\leq \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^1} (\Phi[\rho_n, F_{\Gamma_i}, G_{\Gamma_i}, R_{\Gamma_i}, S_{\Gamma_i}, T_{\Gamma_i}] \geq \Phi[\rho_n, F^*, G^*, R^*, S^*, T^*])\} \\ &\leq \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^1} (F_{\Gamma_i} \geq F^*)\} + \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^1} (G_{\Gamma_i} \geq G^*)\} + \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^1} (R_{\Gamma_i} \leq R^*)\} \\ &+ \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^1} (S_{\Gamma_i} \geq S^*)\} + \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^1} (T_{\Gamma_i} \geq T^*)\}, \end{aligned} \quad (5.17)$$

where the last line follows from the monotonicity of Φ with respect to F, G, R, S and T . Since $\Theta_n^1 \subseteq S_n$, we may apply Lemmas 4.4 and 4.5 to (5.17), and we deduce $\mathbb{P}(\overline{NSP_\alpha}) \rightarrow 0$ as $n \rightarrow \infty$, exponentially in n , as required. \square

In the case of IHT applied to problems with zero noise, the above result has a surprising corollary: a condition can be given which guarantees that, with overwhelming probability, the underlying k -sparse signal x^* is the algorithm's only fixed point. In other words, within some portion of phase space, there is only one possible solution to which the IHT algorithm can converge, namely the underlying signal x^* . This is remarkable since IHT is a gradient projection algorithm for the nonconvex problem (1.2) which can be shown to have a combinatorially large number of local minimizers. The conclusion is that the properties of Gaussian matrices ensure that, within this region of phase space, the IHT algorithm will never 'get stuck' at an unwanted local minimizer, thus exhibiting a behaviour one would usually only expect if a convex problem was being solved. The result follows.

Corollary 5.2 (Single fixed point condition; noiseless case). *Suppose Assumption A.2 holds, as well as (5.3), and that $e = 0$. Then, in the proportional-growth asymptotic, x^* is the only fixed point of IHT with stepsize α , with probability tending to 1 exponentially in n .*

Proof. Lemma 5.1 establishes that, if (5.3) holds, there are asymptotically no α -stable points on any Γ_i such that $i \in \Theta_n^1$. Setting $\sigma \stackrel{\text{def}}{=} 0$ in (5.7), we have $i \in \Theta_n^2 \Rightarrow \Gamma_i = \Lambda$. Therefore any α -stable point is supported on Λ , and Lemma 3.1 implies that it must be x^* . However, any fixed point of IHT with

stepsize α is necessarily an α -stable point, and therefore x^* is also the only fixed point of IHT with stepsize α . \square

Next, we show that any α -stable points on Γ_i with $i \in \Theta_n^2$ are ‘close’ to x^* .

Lemma 5.3. *Suppose Assumptions A.2 and A.3 hold, as well as (5.3). Then there exists ζ sufficiently small such that, in the proportional-growth asymptotic, any α -stable point \bar{x} on Γ_i such that $i \in \Theta_n^2$ satisfies*

$$\|\bar{x} - x^*\| \leq \xi(\delta, \rho) \cdot \sigma, \quad (5.18)$$

with probability tending to 1 exponentially in n , where $\xi(\delta, \rho)$ is defined in (5.5).

Proof. If $\sigma = 0$, the result follows trivially from Lemma 5.2, so let us assume that $\sigma > 0$. Suppose \bar{x} is a minimum-norm solution on Γ , so that $\bar{x}_\Gamma = A_\Gamma^\dagger b$. Then, using $A_\Gamma^\dagger A_\Gamma = I$, we have

$$\begin{aligned} (\bar{x} - x^*)_\Gamma &= A_\Gamma^\dagger (A_\Gamma x_\Gamma^* + A_{\Gamma^C} x_{\Gamma^C}^* + e) - x_\Gamma^* \\ &= x_\Gamma^* + A_\Gamma^\dagger (A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* + A_{(\Lambda \cup \Gamma)^C} x_{(\Lambda \cup \Gamma)^C}^* + e) - x_\Gamma^* \\ &= A_\Gamma^\dagger (A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* + e) + x_\Gamma^* - x_\Gamma^* \\ &= A_\Gamma^\dagger (A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* + e), \end{aligned} \quad (5.19)$$

while

$$(\bar{x} - x^*)_{\Gamma^C} = -x_{\Gamma^C}^*. \quad (5.20)$$

Combining (5.19) and (5.20) using the triangle inequality, we may bound

$$\begin{aligned} \|\bar{x} - x^*\|^2 &= \|(\bar{x} - x^*)_\Gamma\|^2 + \|(\bar{x} - x^*)_{\Gamma^C}\|^2 \\ &= \|A_\Gamma^\dagger (A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^* + e)\|^2 + \|x_{\Gamma^C}^*\|^2 \\ &\leq \left[\|A_\Gamma^\dagger A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^*\| + \|A_\Gamma^\dagger e\| \right]^2 + \|x_{\Lambda \setminus \Gamma}^*\|^2 \end{aligned} \quad (5.21)$$

We may deduce, by (4.7) of Lemma 4.3,

$$\|A_\Gamma^\dagger A_{\Lambda \setminus \Gamma} x_{\Lambda \setminus \Gamma}^*\|^2 = \|x_{\Lambda \setminus \Gamma}^*\|^2 \cdot P_\Gamma, \quad \text{where } P_\Gamma \sim \frac{k}{n-k+1} F(k, n-k+1), \quad (5.22)$$

and by (4.9) of Lemma 4.3,

$$\|A_\Gamma^\dagger e\|^2 = \sigma^2 \cdot Q_\Gamma \leq \sigma^2 \cdot Q_\Gamma, \quad \text{where } Q_\Gamma \sim \frac{k}{n-k+1} F(k, n-k+1). \quad (5.23)$$

Substituting (5.22) and (5.23) into (5.21), we have

$$\|\bar{x} - x^*\|^2 \leq \left[\|x_{\Lambda \setminus \Gamma}^*\| \cdot \sqrt{P_\Gamma} + \sigma \cdot \sqrt{Q_\Gamma} \right]^2 + \|x_{\Lambda \setminus \Gamma}^*\|^2, \quad (5.24)$$

and we may use (5.7) to further deduce

$$\begin{aligned} \|\bar{x} - x^*\|^2 &\leq \sigma^2 \left[a^*(\delta, \rho; \zeta) \cdot \sqrt{P_\Gamma} + \sqrt{Q_\Gamma} \right]^2 + [a^*(\delta, \rho; \zeta)]^2 \cdot \sigma^2 \\ &= \sigma^2 \left\{ \left[a^*(\delta, \rho; \zeta) \cdot \sqrt{P_\Gamma} + \sqrt{Q_\Gamma} \right]^2 + [a^*(\delta, \rho; \zeta)]^2 \right\}. \end{aligned} \quad (5.25)$$

For the sake of brevity, let us define

$$\Psi(P, Q) := \sqrt{\left(a^*(\delta, \rho; \zeta) \cdot \sqrt{P} + \sqrt{Q} \right)^2 + a^*(\delta, \rho; \zeta)^2}, \quad (5.26)$$

so that (5.25) may equivalently be written as

$$\|\bar{x} - x^*\| \leq \sigma \cdot \Psi[P_\Gamma, Q_\Gamma]. \quad (5.27)$$

Given $\zeta > 0$, let us define

$$P^* = Q^* := \mathcal{IF}(\delta, \rho) + \zeta. \quad (5.28)$$

Now we use (5.27) to perform a union bound over all Γ_i such that $i \in \Theta_n^2$, writing \bar{x}_i for the minimum-norm solution on Γ_i , giving

$$\begin{aligned} & \mathbb{P}\{\exists \text{ some } \Gamma_i \text{ such that } i \in \Theta_n^2 \text{ and } \|\bar{x}_i - x^*\| > \sigma \cdot \Psi[P^*, Q^*]\} \\ &= \mathbb{P}\left\{\bigcup_{i \in \Theta_n^2} (\|\bar{x}_i - x^*\| > \sigma \cdot \Psi[P^*, Q^*])\right\} \end{aligned} \quad (5.29)$$

$$\leq \mathbb{P}\left\{\bigcup_{i \in \Theta_n^2} (\|\bar{x}_i - x^*\| > \sigma \cdot \Psi[P_{\Gamma_i}, Q_{\Gamma_i}])\right\} \quad (5.30)$$

$$+ \mathbb{P}\left\{\bigcup_{i \in \Theta_n^2} (\sigma \cdot \Psi[P_{\Gamma_i}, Q_{\Gamma_i}] \geq \sigma \cdot \Psi[P^*, Q^*])\right\}, \quad (5.31)$$

since the event in (5.29) lies in the union of the two events in (5.30) and (5.31). It is an immediate consequence of (5.27) that the event in (5.30) has probability 0. Taking limits of (5.31) as $n \rightarrow \infty$, and cancelling σ , we have

$$\begin{aligned} & \lim_{n \rightarrow \infty} \mathbb{P}\{\exists \text{ some } \Gamma_i \text{ such that } i \in \Theta_n^2 \text{ and } \|\bar{x}_i - x^*\| > \sigma \cdot \Psi[P^*, Q^*]\} \\ &\leq \lim_{n \rightarrow \infty} \mathbb{P}\left\{\bigcup_{i \in \Theta_n^2} (\Psi[P_{\Gamma_i}, Q_{\Gamma_i}] \geq \Psi[P^*, Q^*])\right\} \\ &\leq \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^2} (P_{\Gamma_i} \geq P^*)\} + \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^2} (Q_{\Gamma_i} \geq Q^*)\}, \end{aligned} \quad (5.32)$$

where we used the monotonicity of Ψ with respect to P and Q in the last line. Since $\Theta_n^2 \subseteq S_n$, and using (5.22) and (5.23), we may apply Lemma 4.5 to (5.32), yielding that each of the limits in the right-hand side of (5.32) converges to zero exponentially in n , and so finally

$$\lim_{n \rightarrow \infty} \mathbb{P}\{\exists \text{ some } \Gamma_i \text{ such that } i \in \Theta_n^2 \text{ and } \|\bar{x}_i - x^*\| > \sigma \cdot \Psi[a^*(\delta, \rho; \zeta), P^*, Q^*]\} = 0,$$

exponentially in n . Since by Lemma 3.1, any stable point is necessarily a minimum-norm solution, and recalling the definition of $a^*(\delta, \rho; \zeta)$ in (5.6), $\Psi(a, P, Q)$ in (5.26), and the definitions of P^* , Q^* in (5.28), we have

$$\lim_{n \rightarrow \infty} \mathbb{P}\left\{\exists \text{ some } \alpha\text{-stable point } \bar{x}_i \text{ on } \Gamma_i \text{ such that } i \in \Theta_n^2 \text{ and } \|\bar{x}_i - x^*\| > \sigma \sqrt{\mathcal{IF}(\delta, \rho) [1 + a(\delta, \rho) + \zeta]^2 + [a(\delta, \rho) + \zeta]^2}\right\} = 0, \quad (5.33)$$

with convergence exponential in n . Finally, by continuity,

$$\begin{aligned} & \|\bar{x}_i - x^*\| > \sigma \sqrt{\mathcal{IF}(\delta, \rho) [1 + a(\delta, \rho)]^2 + 1 + [a(\delta, \rho)]^2} \\ & \implies \|\bar{x}_i - x^*\| > \sigma \sqrt{\mathcal{IF}(\delta, \rho) [1 + a(\delta, \rho) + \zeta]^2 + [a(\delta, \rho) + \zeta]^2}, \end{aligned}$$

for some ζ suitably small, and the result now follows from the definition of $\xi(\delta, \rho)$ in (5.5). \square

In the context of IHT, we obtain the following convergence result in the proportional-dimensional asymptotic framework.

Lemma 5.4. *Suppose Assumption A.2 holds and that the stepsize α of IHT is chosen to satisfy*

$$\alpha < \frac{1}{1 + \mathcal{U}(\delta, 2\rho)}. \quad (5.34)$$

Then, in the proportional-growth asymptotic, IHT converges to an α -stable point with probability tending to 1 exponentially in n .

Proof. Given (5.34), we may apply Lemma 4.6 with ϵ sufficiently small to deduce $\alpha(1 + U_{2k}) < 1$, with probability tending to 1 exponentially in n . Under Assumption A.2, we may then apply Theorem 3.6 and deduce convergence of IHT to an α -stable point. \square

We now combine Lemmas 5.1, 5.3 and 5.4 and prove a recovery result for IHT, one of the main results of the paper.

Theorem 5.5 (Recovery result for IHT; noise case). *Suppose Assumptions A.2 and A.3 hold, suppose that*

$$\rho < \hat{\rho}^{IHT}(\delta), \quad (5.35)$$

where $\hat{\rho}^{IHT}(\delta)$ is defined in (5.1), and that the IHT stepsize α satisfies

$$\frac{\sqrt{\mathcal{IF}(\delta, \rho)}}{(1 - \rho)[1 - \mathcal{IL}(\delta, \rho, 1 - \rho)]} < \alpha < \frac{1}{1 + \mathcal{U}(\delta, 2\rho)}. \quad (5.36)$$

Then, in the proportional-growth asymptotic⁵, IHT converges to \bar{x} that is close to x^ in the sense that (5.18) holds with probability tending to 1 exponentially in n .*

Proof. First note that (5.35) implies that the interval in (5.36) is well-defined. Provided α is chosen to satisfy (5.36), (5.34) holds, and under Assumption A.2, we may apply Lemma 5.4 to deduce convergence of IHT to an α -stable point. On the other hand, Lemma 5.1 establishes that there are asymptotically no α -stable points on any Γ_i such that $i \in \Theta_n^1$, while we may apply Lemma 5.3 to deduce that any α -stable points on any Γ_i such that $i \in \Theta_n^2$ satisfy (5.18). \square

In the absence of noise, the same phase transition guarantees exact recovery of the original signal x^* .

Corollary 5.6 (Recovery result for IHT; noiseless case). *Suppose Assumption A.2 holds, as well as (5.35), and that α satisfies (5.36) and the noise $e \stackrel{\text{def}}{=} 0$. Then, in the proportional-growth asymptotic, IHT converges to x^* with probability tending to 1 exponentially in n .*

Proof. The result follows by setting $\sigma \stackrel{\text{def}}{=} 0$ in Theorem 5.5. \square

Comparing our result with previous RIP-based recovery results for IHT, Theorem 5.5 proves a phase transition that is equally valid over a continuous stepsize range. In contrast, the recovery results in [6, 8, 20, 21] either require a specific fixed stepsize or degrade with the choice of stepsize.

5.3 Recovery results for N-IHT

In the case of N-IHT, it is possible to prove convergence to an $\underline{\alpha}(\delta, \rho; \epsilon)$ -stable point, where

$$\underline{\alpha}(\delta, \rho; \epsilon) \stackrel{\text{def}}{=} \{\kappa[1 + \mathcal{U}(\delta, 2\rho) + \epsilon]\}^{-1}, \quad (5.37)$$

for some $\epsilon > 0$. Due to the dependence of $\underline{\alpha}(\delta, \rho; \epsilon)$ upon (δ, ρ) , we need to replace Definition 5.2 with the following definition.

⁵In other words, we consider instances of the Gaussian random variables A and e for a sequence of triples (k, n, N) where $n \rightarrow \infty$, where n is the number of measurements, N , the signal dimension and k , the sparsity of the underlying signal.

Definition 5.4 (Stability factor for N-IHT). Given $\delta \in (0, 1]$ and $\rho \in (0, 1/2]$, provided

$$\rho < \hat{\rho}^{N-IHT}(\delta), \quad (5.38)$$

where $\hat{\rho}^{N-IHT}(\delta)$ is defined in (5.2), let

$$a(\delta, \rho) \stackrel{\text{def}}{=} \frac{\sqrt{\mathcal{IF}(\delta, \rho)} + \{\kappa[1 + \mathcal{U}(\delta, 2\rho)]\}^{-1} \sqrt{\rho(1 - \rho)[1 + \mathcal{IU}(\delta, \rho, 1 - \rho)][1 + \mathcal{IU}(\delta, \rho, \rho)]}}{(1 - \rho)\{\kappa[1 + \mathcal{U}(\delta, 2\rho)]\}^{-1}[1 - \mathcal{IL}(\delta, \rho, 1 - \rho)] - \sqrt{\mathcal{IF}(\delta, \rho)}}, \quad (5.39)$$

and

$$\xi(\delta, \rho) \stackrel{\text{def}}{=} \sqrt{\mathcal{IF}(\delta, \rho) [a(\delta, \rho)]^2 + [a(\delta, \rho)]^2}, \quad (5.40)$$

where \mathcal{IF} is defined in (4.39), \mathcal{IU} is defined in (4.36), \mathcal{IL} is defined in (4.37), and where \mathcal{U} is defined in [2, Definition 2.2].

Note that (5.38) ensures that the denominator in (5.39) is strictly positive and that $a(\delta, \rho)$ is therefore well-defined.

The reader may verify by comparison with Definitions 5.2 that the α terms have been replaced by the term $\{\kappa[1 + \mathcal{U}(\delta, 2\rho)]\}^{-1}$. Since the proof follows broadly the same approach as for IHT, we proceed to the main recovery result, relegating the proof to Appendix B.

Theorem 5.7 (Recovery result for N-IHT; noise case). Suppose Assumptions A.2 and A.3 hold, as well as (5.38). Then, in the proportional-growth asymptotic, N-IHT converges to \bar{x} such that

$$\|\bar{x} - x^*\| \leq \xi(\delta, \rho) \cdot \sigma, \quad (5.41)$$

with probability tending to 1 exponentially in n , where $\xi(\delta, \rho)$ is defined in (5.40).

In the case of zero noise, Theorem 5.7 simplifies to an exact recovery result.

Corollary 5.8 (Recovery result for N-IHT; noiseless case). Suppose Assumption A.2 holds, as well as (5.38), and that the noise $e \stackrel{\text{def}}{=} 0$. Then, in the proportional-growth asymptotic, N-IHT converges to x^* with probability tending to 1 exponentially in n .

Proof. The result follows by setting $\sigma \stackrel{\text{def}}{=} 0$ in Theorem 5.7. □

6 Illustration and discussion of recovery phase transitions

Noiseless case. The recovery phase transitions given in Definition 5.1 for IHT and N-IHT (with $\kappa = 1.1$) respectively are displayed in Figure 6.1. Exact recovery in the case of zero noise is guaranteed asymptotically for (δ, ρ) pairs falling below the respective curves. The best-known exact recovery phase transitions obtained in [30] from previous RIP analysis are included for comparison: the IHT phase transition applies the RIP bounds in [2] to Foucart's analysis in [19], while an extension of the same approach leads to the phase transition for N-IHT. An RIP analysis of the stable point approach adopted in this paper was also carried out in [30], and the resulting phase transitions are also displayed in Figure 6.1. We see a considerable improvement over the phase transitions corresponding to previous RIP analysis, with recovery being guaranteed for IHT for values of ρ around 1.7 times higher than before, and for N-IHT around 10 times higher than before. Figure 6.2 displays the inverse of the phase transition for each stepsize scheme. Previous RIP analysis requires a lower bound of $n \geq 234k$ measurements to guarantee recovery using IHT, and $n \geq 1617k$ using N-IHT. By comparison, we reduce these lower bounds to $n \geq 138k$ for IHT and $n \geq 154k$ for N-IHT. It should also be added that our result for IHT holds for a continuous stepsize range, while the result based upon [19], in keeping with all other similar RIP-based results for IHT (see [30]), holds true only if the stepsize is optimized to a particular value.

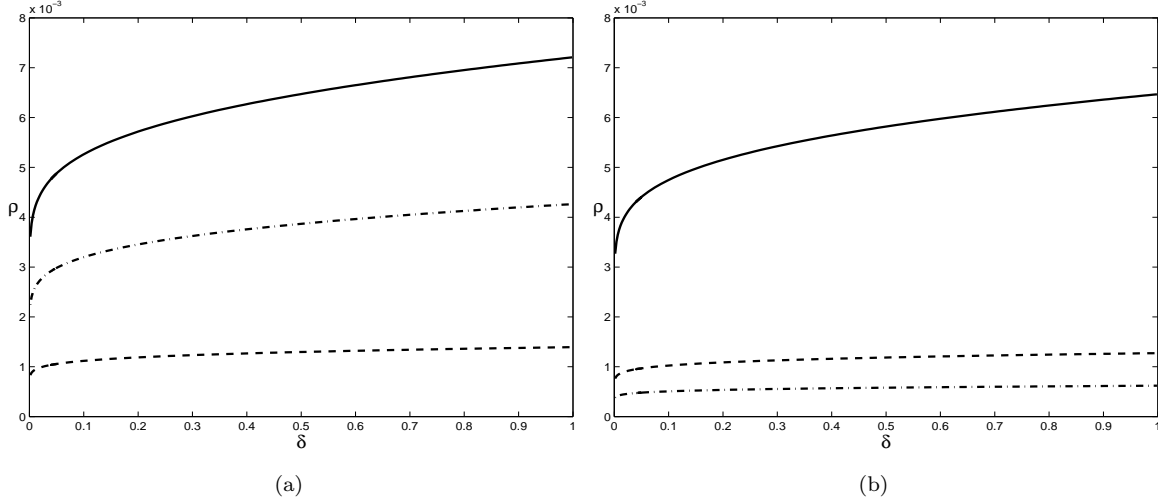


Figure 6.1: Our average-case phase transitions for IHT algorithms (unbroken) compared with the best-known RIP-based phase transitions based on our stable point analysis [30] (dashed) and the analysis in [19] (dash-dot): (a) IHT (b) N-IHT.

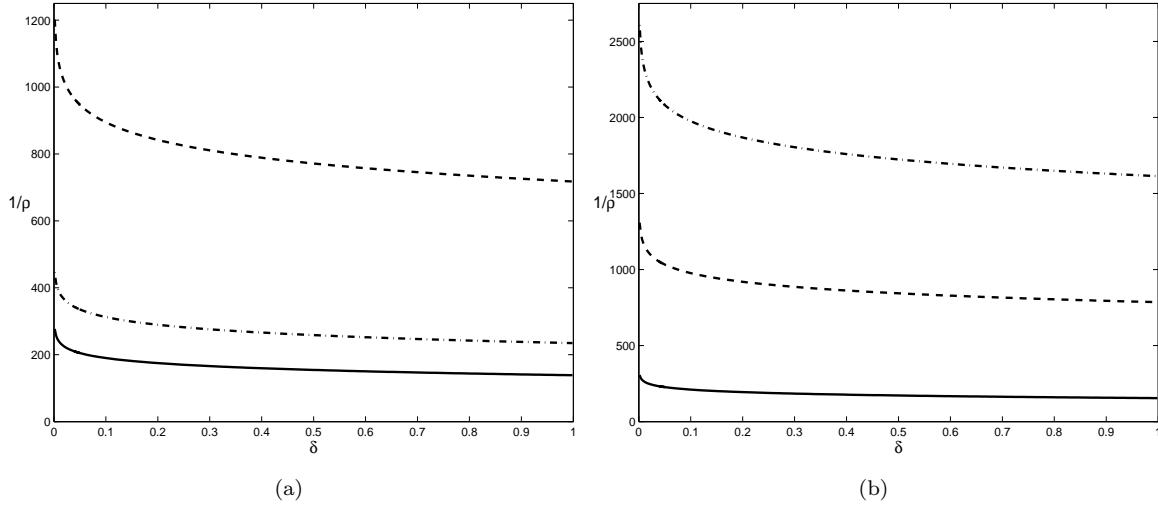


Figure 6.2: Inverse of the phase transitions in Figure 6.1: (a) IHT (b) N-IHT.

Interpretation of recovery results as lower bounds on a weak phase transition. We have obtained an improvement by switching to a new method of analysis which allows us to leverage the assumption that the measurements are statistically independent of the signal. The latter has allowed us to make the transition from worst-case to average-case analysis.

The distinction between worst-case and average-case phase transitions can also be found in the phase transitions of Donoho and Tanner for recovery using l_1 -minimization [17], where successful recovery by means of l_1 -minimization is shown to be equivalent to the neighbourliness of projected l_1 balls [16]. There are both strong and weak version of neighbourliness: strong neighbourliness guarantees recovery of *any* signal by means of l_1 -minimization, while weaker forms of neighbourliness assume either a randomly-chosen support and/or randomly-chosen sign pattern for the signal. It is appropriate then to see our derived phase transitions as lower bounds on a weak phase transition for IHT, in contrast to an RIP analysis which gives lower bounds on the strong phase transition. The notions of weakness are comparable but not identical: in the case of l_1 -minimization, some dependency between the signal and measurement matrix is permitted:

it is only required that the support set and sign pattern of the signal are chosen independently of the matrix. However, independence is the only assumption we place upon the signal, and beyond this there is no further restriction upon the signal's coefficients.

It is worth pointing out that it is the weak phase transition that is observed empirically for recovery by means of l_1 -minimization, and the same is also to be expected for IHT algorithms. While we obtain a significant improvement, our lower bound is still pessimistic compared to the weak phase transition observed empirically, though we have succeeded in narrowing the gap between the two. It is no surprise that our results do not give the precise weak phase transition, due to the continued (but limited) use of worst-case techniques, such as the RIP and large deviations analysis. However, the use of the average-case independence assumption to analyse the stable point condition has allowed us to break free in part from the restrictions of worst-case analysis.

Choice of stepsize for IHT. Corollary 5.6 guarantees exact recovery using IHT provided the stepsize α falls within the interval given in (5.36), provided this interval is well-defined. In fact, an inspection of the proof of these two results reveals that the lower bound in (5.36) arises from the stable point condition, while the upper bound in (5.36) arises from the convergence condition. Figure 6.3 illustrates these bounds for the case $\delta = 0.5$. We see that, as ρ is increased, the admissible stepsize range contracts, until a critical ρ -value is reached at which the interval is no longer well-defined.

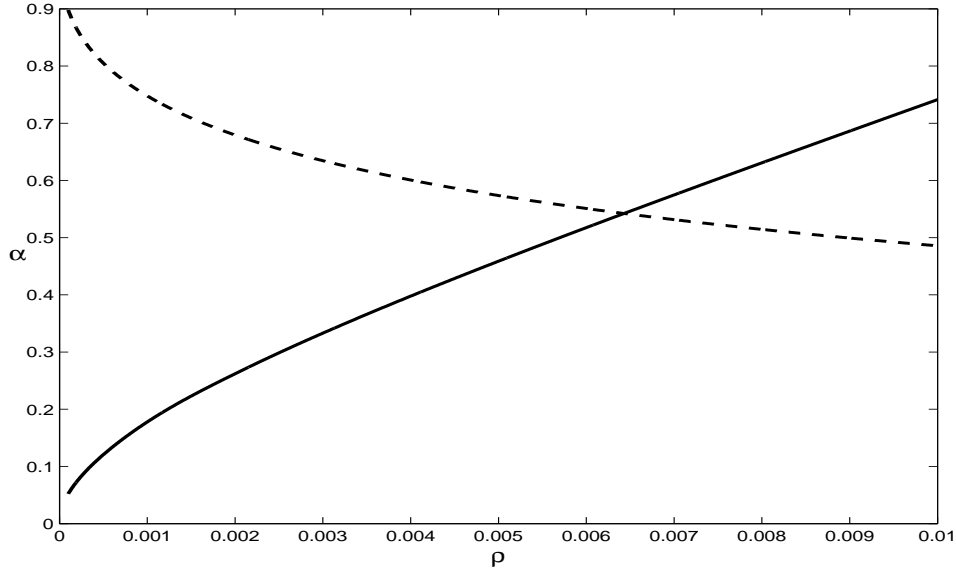


Figure 6.3: Lower bound (unbroken) and upper bound (dashed) on the IHT stepsize for $\delta = 0.5$.

It has been observed empirically [23] that care must be taken to ensure that the IHT stepsize is neither too small or too large. Our analysis gives theoretical insight into this observation: the stepsize must be small enough to ensure that the algorithm converges, but large enough to ensure that it does not converge to fixed points other than the underlying sparse signal.

Extension to noise. In the case where measurements are contaminated by noise, exact recovery of the original signal is impossible. However, Theorems 5.5 and 5.7 guarantee that, in the same region of phase space defined by the exact recovery phase transition, the approximation error of the output of IHT/N-IHT is asymptotically bounded by some known stability factor multiplied by the noise level σ . Figure 6.4 plots this noise stability factor $\xi(\delta, \rho)$ for each of the two stepsize schemes considered ($\kappa = 1.1$ for N-IHT). In keeping with the results in [5] and [6], we observe that the stability factor tends to infinity as the transition point is reached.

For both IHT and N-IHT, in the region for which the stability factors derived in this paper are defined, they are everywhere lower than the corresponding stability factors derived from the previous analysis

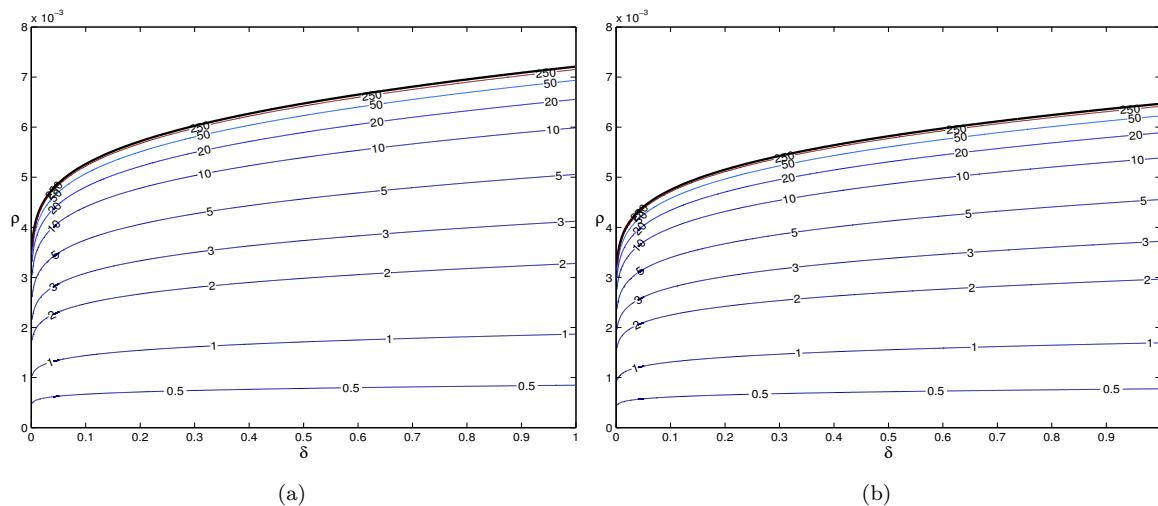


Figure 6.4: Plot of the stability factor $\xi(\delta, \rho)$ for (a) IHT (b) N-IHT.

in [19]; see [30] for a comparison. It should be pointed out that we have obtained improved stability results by imposing additional restrictions upon the noise, namely that the noise is Gaussian distributed and independent of the signal and measurement matrix. This assumption is in keeping with our aim of performing average-case analysis. Our analysis could, however, be altered to deal with the case of non-independent noise by making more use of the RIP, though this would lead to larger stability constants.

We have also extended our analysis in [30] to the case of signals which are only approximately k -sparse, for both IHT and N-IHT, though we omit this extension in the present work for the sake of brevity. In this extension, a stability factor is derived which multiplies the unrecoverable energy of the signal, due to both measurement noise and inaccuracy of the k -sparse model.

7 Conclusions and future directions

While CS was first developed within the framework of l_1 -minimization, there is growing evidence that recovery algorithms which do not rely on convex relaxation and the l_1 -norm can be equally effective in practice [7]. Two such examples are the gradient-based IHT [9] and N-IHT [10] algorithms, which also have favourable computational efficiency in comparison with other CS approaches. It is important that a CS recovery algorithm is supported by theory which quantitatively determines the degree of undersampling that the algorithm permits. Such results now exist for l_1 -minimization, where precise phase transitions have been determined within a proportional-growth asymptotic framework in the case of Gaussian matrices [17]. By contrast, worst-case recovery guarantees for IHT algorithms using the RIP are pessimistic in comparison with observed empirical behaviour [6].

To address this issue, we introduced a new method of recovery analysis for IHT algorithms in which we analysed the algorithms' *stable points*, a generalization of the notion of fixed points. By making the realistic assumption of independence between the signal and measurement matrix, we obtained the first average-case recovery guarantees for IHT algorithms and Gaussian measurement matrices in the phase transition framework. In contrast to RIP analysis, which leads to lower bounds on the strong phase transition, we obtained lower bounds on a weak phase transition for recovery using IHT algorithms, which is the notion of practical interest. By breaking free in part from the restrictions of worst-case analysis, we have obtained, to the best of our knowledge, the highest phase transitions yet guaranteeing exact recovery of sparse signals by means of IHT and N-IHT. Our results extend to the realistic model of noisy measurements, guaranteeing an improved robustness to these inaccuracies.

The ultimate remaining goal of the work is to fully close the gap between theoretical guarantees

and empirical performance for IHT algorithms. At present, the continued use of worst-case methods of analysis such as union bounds over combinatorially many support sets is a hindrance to significant further improvements in phase transitions. It is an open question whether such a strong requirement is necessary for ensuring signal recovery on average. Though we have obtained quantitative results only for Gaussian matrices here, many other families of random or randomized measurement matrices exhibit similar empirical behaviour and are important to practitioners. Obtaining quantitative guarantees for IHT algorithms applied to such CS measurement schemes is an open avenue of research.

References

- [1] M. Abramowitz and I. Stegun. A handbook of mathematical functions. *Dover, 9th Edition*, 1964.
- [2] B. Bah and J. Tanner. Improved bounds on restricted isometry constants for Gaussian matrices. *SIAM Journal on Matrix Analysis*, 31(5):2882–2898, 2010.
- [3] R. Baraniuk and M. Davenport and R. DeVore and M. Wakin. A Simple Proof of the Restricted Isometry Property for Random Matrices. *Constructive Approximation*, 28(3):253–263, 2008.
- [4] A. Beck and Y. Eldar. Sparsity constrained nonlinear optimization: optimality conditions and algorithms. Technical Report, EE Department, Technion; [arXiv:1203.4580\[cs.IT\]](#), 2012.
- [5] J. D. Blanchard, C. Cartis, and J. Tanner. Compressed sensing: how sharp is the restricted isometry property? *SIAM Review*, 53(1):105–125, 2011.
- [6] J. D. Blanchard, C. Cartis, J. Tanner, and A. Thompson. Phase transitions for greedy sparse approximation algorithms. *Applied and Computational Harmonic Analysis*, 30:188–203, 2011.
- [7] J. D. Blanchard and J. Tanner. Performance comparisons for greedy algorithms in compressed sensing. *submitted*, 2013.
- [8] T. Blumensath and M. E. Davies. Iterative hard thresholding for compressed sensing. *Applied and Computational Harmonic Analysis*, 27(3):265–274, 2009.
- [9] T. Blumensath and M. E. Davies. Iterative thresholding for sparse approximations. *Journal of Fourier Analysis and Applications*, 14(5):629–654, 2008.
- [10] T. Blumensath and M. E. Davies. Normalised iterative hard thresholding: guaranteed stability and performance. *IEEE Journal of Selected Topics in Signal Processing*, 4(2):298–309, 2010.
- [11] E. Candès and J. Romberg and T. Tao. Robust uncertainty principles: exact signal reconstruction from highly incomplete frequency information. *Communications on Pure and Applied Mathematics*, 59(8):1207–1223, 2006.
- [12] E. J. Candes and T. Tao. Decoding by linear programming. *IEEE Transactions on Information Theory*, 51(12): 4203–4215, 2005.
- [13] W. Dai and O. Milenkovic. Subspace pursuit for compressive sensing signal reconstruction. *IEEE Transactions on Information Theory*, 55(5):2230–2249, 2009.
- [14] G. Davis, S. Mallat, and Z. Zhang. Adaptive time-frequency decompositions. *Optical Engineering*, 33:2183–2191, 1994.
- [15] D. Donoho. Compressed Sensing. *IEEE Transactions on Information Theory*, 52(4):1289–1306, 2006.
- [16] D. Donoho. High-dimensional centrally symmetric polytopes with neighborliness proportional to dimension. *Technical report, Computational Geometry (online)*, 2005.

- [17] D. L. Donoho and J. Tanner. Precise undersampling theorems. *Proceedings of the IEEE*, 98(6): 913–924, 2010.
- [18] A. Edelman and N. Raj Rao. Random matrix theory. *Acta Numerica*, 14:233–297, 2005.
- [19] S. Foucart. Hard thresholding pursuit: an algorithm for compressed sensing. *SIAM Journal on Numerical Analysis*, 49(6):2543–2563, 2011.
- [20] S. Foucart. Sparse recovery algorithms: sufficient conditions in terms of restricted isometry constants. *Springer Proceedings in Mathematics, Approximation Theory XIII* 65–77, San Antonio, 2010.
- [21] R. Garg and R. Kandekhar. Gradient descent with sparsification: an iterative algorithm for sparse recovery with restricted isometry property. *International Conference on Machine Learning*, 2009.
- [22] A. K. Gupta and D. K. Nagar. *Matrix Variate Distributions*. Chapman & Hall, 1999.
- [23] A. Maleki and D. L. Donoho. Optimally tuned iterative reconstruction algorithms for compressed sensing. *Journal of Selected Topics in Signal processing* 4(2): 330–341, 2010.
- [24] S. Mallat and Z. Zhang. Matching pursuit in a time-frequency dictionary. *IEEE Transactions on Signal Processing* 41(12):3397–3415, 1993.
- [25] D. Needell and J. Tropp. Cosamp: iterative signal recovery from incomplete and inaccurate samples. *Applied and Computational Harmonic Analysis*, 26(3):301–321, 2008.
- [26] M. Nikolova. Description of the minimizers of least squares regularized with l_0 -norm. Uniqueness of the global minimizer. *SIAM Journal on Imaging Sciences* 6(2):904–937, 2013.
- [27] J. Nocedal and S. J. Wright. *Numerical Optimization*. Second edition, Springer-Verlag, New York, USA, 2006.
- [28] Y. Pati, R. Rezaifar, and P. Krishnaprasad. Orthogonal matching pursuit: recursive function approximation with applications to wavelet decomposition. In *Proceedings of the 27th Asilomar Conference on Signals, Systems and Computers*, Los Alamitos, CA, 1993.
- [29] N. Temme. The uniform asymptotic expansion of a class of integrals related to cumulative distribution functions. *SIAM Journal of Mathematical Analysis*, 13(2):239–253, 1982.
- [30] A. Thompson. Quantitative analysis of algorithms for compressed signal recovery. *PhD thesis, School of Mathematics, University of Edinburgh*, 2012.
- [31] J. A. Tropp and S. J. Wright. Computational methods for sparse solution of linear inverse problems. *Proceedings of the IEEE*, 98(5):948–958, 2010.

Appendix A

Proofs of results in Section 4.2

We make use of asymptotic results derived by Temme [29] for the incomplete gamma and beta functions, which are related to the χ^2 and F distributions respectively. We denote by $P(s, t)$ the lower regularized incomplete gamma function $P(s, t)$ [29], and we let $Q(s, t) = 1 - P(s, t)$ be the upper regularized incomplete gamma function. We also define the complementary error function $\operatorname{erfc}(\omega)$ in the usual way as

$$\operatorname{erfc}(\omega) \stackrel{\text{def}}{=} \frac{2}{\sqrt{\pi}} \int_{\omega}^{\infty} e^{-u^2} du.$$

The result for the gamma function follows.

Lemma A.1 (Gamma asymptotic [29, Section 3.4 and (2.20)]). For $0 < s < t$,

$$Q(s, t) = \frac{1}{2} \operatorname{erfc} \left(\eta_Q \sqrt{\frac{s}{2}} \right) - R_s(\eta_Q) \quad \text{where} \quad \eta_Q = \sqrt{2 \left[\frac{t}{s} - \ln \left(1 + \frac{t}{s} \right) \right]}, \quad (\text{A.1})$$

and for $s > t > 0$,

$$P(s, t) = \frac{1}{2} \operatorname{erfc} \left(-\eta_P \sqrt{\frac{s}{2}} \right) + R_s(\eta_P) \quad \text{where} \quad \eta_P = -\sqrt{2 \left[-\frac{t}{s} - \ln \left(1 - \frac{t}{s} \right) \right]}, \quad (\text{A.2})$$

where $R_s(\cdot)$ is a residual term. Furthermore, if t/s remains fixed so that η is held constant,

$$R_s(\eta) = \mathcal{O} \left(\frac{1}{\sqrt{s}} \right) e^{-\frac{1}{2} s \eta^2} \quad \text{for } s \text{ sufficiently large.} \quad (\text{A.3})$$

Lemma A.2. Let $0 < l \leq n$ and let the random variable $X_l \sim \frac{1}{l} \chi_l^2$. Let $l/n \rightarrow \gamma \in (0, 1]$ as $n \rightarrow \infty$. Then, for any $\nu > 0$,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \ln \mathbb{P}(X_l^i \geq 1 + \nu) = -\frac{\gamma}{2} [\nu - \ln(1 + \nu)] \quad (\text{A.4})$$

and, for any $\nu \in (0, 1)$,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \ln \mathbb{P}(X_l^i \leq 1 - \nu) = -\frac{\gamma}{2} [-\nu - \ln(1 - \nu)]. \quad (\text{A.5})$$

Proof. We first show (A.4). We have

$$\mathbb{P}(X_l \geq 1 + \nu) = \mathbb{P}[\chi_l^2 \geq l(1 + \nu)] = Q \left[\frac{l}{2}, \frac{l(1 + \nu)}{2} \right], \quad (\text{A.6})$$

where the first step follows from the definition of X_l , and the second step follows from the properties of the χ^2 distribution. We can further express the right-hand side of (A.6) by using (A.1) with $s = l/2$ and $t = l(1 + \nu)/2$, which then gives

$$\mathbb{P}(X_l \geq 1 + \nu) = \frac{1}{2} \operatorname{erfc} \left(\frac{\eta_Q}{2} \sqrt{l} \right) - R_l(\eta_Q), \quad (\text{A.7})$$

where

$$\eta_Q \stackrel{\text{def}}{=} \sqrt{2[\nu - \ln(1 + \nu)]}. \quad (\text{A.8})$$

Applying a standard exponential tail bound on the complementary error function erfc to (A.7) then gives

$$\mathbb{P}(X_l \geq 1 + \nu) \leq \frac{1}{2} e^{-\frac{1}{4} l \eta_Q^2} - R_l(\eta_Q), \quad (\text{A.9})$$

to which we can apply (A.3) to obtain

$$\mathbb{P}(X_l \geq 1 + \nu) = \mathcal{O}(1) e^{-\frac{1}{4} l \eta_Q^2} \quad \text{for all } l \text{ sufficiently large.}$$

Taking logarithms, letting $n \rightarrow \infty$ and recalling that $l/n \rightarrow \gamma$, we deduce

$$\lim_{n \rightarrow \infty} \frac{1}{n} \ln \mathbb{P}(X_l \geq 1 + \nu) = \lim_{n \rightarrow \infty} \frac{1}{n} \ln \mathcal{O}(1) + \lim_{n \rightarrow \infty} \frac{1}{n} \cdot \left(-\frac{1}{4} l \eta_Q^2 \right) = -\frac{\gamma}{4} \eta_Q^2,$$

which together with (A.8) yields (A.4). The proof for the lower tail is similar, since the distribution function of X_l is given by

$$\mathbb{P}(X_l \leq 1 - \nu) = \mathbb{P}[\chi_l^2 \leq l(1 - \nu)] = P \left[\frac{l}{2}, \frac{l(1 - \nu)}{2} \right],$$

which further becomes, due to (A.2) with $s = l/2$ and $t = l(1 - \nu)/2$,

$$\mathbb{P}(X_l \leq 1 - \nu) = \frac{1}{2} \operatorname{erfc} \left(-\frac{\eta_P}{2} \sqrt{l} \right) + R_l(\eta_P),$$

where $\eta_P = -\sqrt{2[-\nu - \ln(1 - \nu)]}$. The bound (A.5) now follows similarly to (A.4). \square

We will need the following lemma which gives the limit of a binomial coefficient in the proportional-growth asymptotic.

Lemma A.3 (Combinatorial limit). *In the proportional-dimensional asymptotic,*

$$\lim_{n \rightarrow \infty} \frac{1}{n} \ln \binom{N}{k} = \frac{H(\delta\rho)}{\delta}, \quad (\text{A.10})$$

where $H(\cdot)$ is defined in (4.38).

Proof. In the proportional-dimensional asymptotic,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \ln \binom{N}{k} = \lim_{n \rightarrow \infty} \frac{N}{n} \cdot \frac{1}{N} \ln \binom{N}{k} = \frac{1}{\delta} \cdot H(\delta\rho),$$

where the last step follows from Stirling's formula. \square

Proof of Lemma 4.4 (Large deviation result for χ^2). Union bounding $\mathbb{P}(X_l^i \geq 1 + \nu)$ over all $i \in S_n$ gives

$$\mathbb{P}\{\cup_{i \in S_n} (X_l^i \geq 1 + \nu)\} \leq \sum_{i \in S_n} \mathbb{P}(X_l^i \geq 1 + \nu) = |S_n| \cdot \mathbb{P}(X_l^1 \geq 1 + \nu). \quad (\text{A.11})$$

Taking logarithms and limits of the right-hand side of (A.11), using (A.4) and (A.10), we have

$$\lim_{n \rightarrow \infty} \frac{1}{n} \ln [|S_n| \cdot \mathbb{P}(X_l^1 \geq 1 + \nu)] = H(\delta\rho) - \frac{\lambda}{2} [\nu - \ln(1 + \nu)],$$

and so (A.11) implies that, for any $\eta > 0$,

$$\frac{1}{n} \ln \mathbb{P}\{\cup_{i \in S_n} (X_l^i \geq 1 + \nu)\} \leq H(\delta\rho) - \frac{\lambda}{2} [\nu - \ln(1 + \nu)] + \eta, \quad (\text{A.12})$$

for all n sufficiently large. By the definition of $\mathcal{IU}(\delta, \rho, \lambda)$ in (4.36), and since $[\nu - \ln(1 + \nu)]$ is strictly increasing on $\nu > 0$, then, for any $\epsilon > 0$, setting $\nu := \nu^* = \mathcal{IU}(\delta, \rho, \lambda) + \epsilon$ and choosing η sufficiently small in (A.12) ensures

$$\frac{1}{n} \ln \mathbb{P}\{\cup_{i \in S_n} (X_l^i \geq 1 + \nu^*)\} \leq -c_Q \quad \text{for all } n \text{ sufficiently large,}$$

where c_Q is some positive constant, from which it follows that

$$\mathbb{P}\{\cup_{i \in S_n} (X_l^i \geq 1 + \nu^*)\} \leq e^{-c_Q \cdot n} \quad \text{for all } n \text{ sufficiently large,}$$

and (4.41) follows. Combining the same union bound argument with the lower tail result of Lemma A.2 shows that, if we take $\nu^* = \mathcal{IL}(\delta, \rho, \lambda) + \epsilon$ for some $\epsilon > 0$, then

$$\frac{1}{n} \ln \mathbb{P}\{\cup_{i \in S_n} (X_l^i \leq 1 - \nu^*)\} \leq -c_P \quad \text{for all } n \text{ sufficiently large,}$$

where c_P is some positive constant, and (4.42) follows similarly to (4.41). \square

For the F -distribution, we need an asymptotic result concerning the regularized incomplete beta function [29], which we denote by $I_\beta(d_1, d_2)$.

Lemma A.4 (Beta asymptotic [29, Section 3.3.2 and (2.20)]). *For $d_1 > d_2 > 0$,*

$$I_\beta(d_1, d_2) = \frac{1}{2} \operatorname{erfc} \left(\eta_I \sqrt{\frac{d_1 + d_2}{2}} \right) + S_n(\eta_I) \quad (\text{A.13})$$

where

$$-\frac{1}{2} \eta_I^2 = \left(\frac{d_1}{d_1 + d_2} \right) \ln \left[\frac{\beta(d_1 + d_2)}{d_1} \right] + \left(\frac{d_2}{d_1 + d_2} \right) \ln \left[\frac{(1 - \beta)(d_1 + d_2)}{d_2} \right], \quad (\text{A.14})$$

where

$$\operatorname{sgn}(\eta_I) = \operatorname{sgn}\left(\beta - \frac{d_1}{d_1 + d_2}\right), \quad (\text{A.15})$$

and where $S_n(\cdot)$ is a residual term. Furthermore,

$$S_n(\eta_I) = \mathcal{O}\left(\frac{1}{\sqrt{s}}\right) e^{-\frac{1}{2}s\eta_I^2} \quad \text{for } l \text{ sufficiently large,} \quad (\text{A.16})$$

uniformly in η_I on compactly-supported subsets of \mathbb{R} .

Lemma A.5. Let the random variable $X_n \sim \frac{k}{n-k+1} \mathcal{F}(k, n-k+1)$. Provided

$$f > \frac{\rho}{1-\rho}, \quad (\text{A.17})$$

in the proportional-growth asymptotic,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \ln \mathbb{P}(X_n \geq f) = \frac{1}{2} [\ln(1+f) - \rho \ln f - H(\rho)]. \quad (\text{A.18})$$

Proof. We have

$$\mathbb{P}[\mathcal{F}(d_1, d_2) \geq \beta] = I_{\left(\frac{d_2}{d_1 + d_2}\right)}\left(\frac{d_2}{2}, \frac{d_1}{2}\right), \quad (\text{A.19})$$

where the first step follows from the definition of X_n , and the second step follows from the properties of the F -distribution. Now $n \geq 2k$, and therefore $\frac{n-k+1}{2} > \frac{k}{2}$, and so we may apply (A.13) with $d_1 = k$, $d_2 = n - k + 1$ and $\beta = \left(\frac{n-k+1}{k}\right) f$ to the right-hand side of (A.19) to obtain

$$\mathbb{P}[\mathcal{F}(d_1, d_2) \geq \beta] = \frac{1}{2} \operatorname{erfc}\left(-\frac{\eta_I}{2} \sqrt{n+1}\right) + S_n(\eta_I), \quad (\text{A.20})$$

where

$$-\frac{1}{2}\eta_I^2 = \left(\frac{n-k+1}{n+1}\right) \ln \left[\frac{n+1}{(n-k+1)(1+f)}\right] + \left(\frac{k}{n+1}\right) \ln \left[\frac{(n+1)f}{k(1+f)}\right], \quad (\text{A.21})$$

and where

$$\operatorname{sgn}(\eta_I) = \operatorname{sgn}\left(\frac{1}{1+f} - \frac{n-k+1}{n+1}\right). \quad (\text{A.22})$$

By (A.17), $f > \rho/(1-\rho)$, which may be combined with the observation that

$$\frac{1}{1+f} - \frac{n-k+1}{n+1} < 0 \iff f > \frac{k}{n-k+1},$$

to deduce that $\eta_I < 0$ for (k, n) sufficiently large, and therefore that

$$\begin{aligned} \bar{\eta}_I^2 &\stackrel{\text{def}}{=} \lim_{n \rightarrow \infty} \eta_I^2 = 2 \left\{ (1-\rho) \ln[(1-\rho)(1+f)] + \rho \ln \left[\frac{\rho(1+f)}{f} \right] \right\} \\ &= 2 [(1-\rho) \ln(1-\rho) + (1-\rho) \ln(1+f) + \rho \ln \rho + \rho \ln(1+f) - \rho \ln f] \\ &= 2 [\ln(1+f) - \rho \ln f - H(\rho)]. \end{aligned} \quad (\text{A.23})$$

Combining (A.20) with a standard exponential tail bound on the complementary error function erfc gives

$$\mathbb{P}(X_n \geq f) \leq \frac{1}{2} e^{-\frac{1}{4}(n+1)\eta_I^2} + S_n(\eta_I), \quad (\text{A.24})$$

to which we can apply (A.16) to obtain

$$\mathbb{P}(X_n \geq f) \leq \mathcal{O}(1) e^{-\frac{1}{4}(n+1)\eta_I^2} \quad \text{for all } n \text{ sufficiently large.}$$

Taking logarithms and letting $n \rightarrow \infty$, we deduce

$$\lim_{n \rightarrow \infty} \frac{1}{n} \ln \mathbb{P}(X_n \geq f) \leq \lim_{n \rightarrow \infty} \frac{1}{n} \ln \mathcal{O}(1) + \lim_{n \rightarrow \infty} \frac{1}{n} \cdot -\frac{1}{4} k \eta_I^2 = -\frac{\rho}{4} \bar{\eta}_I^2,$$

which together with (A.23) proves (A.18). \square

Proof of Lemma 4.5 (Large deviation result for F). Union bounding $\mathbb{P}(X_n^i \geq 1 + f)$ over all $i \in S_n$ gives

$$\mathbb{P}\left\{\bigcup_{i \in S_n} (X_n^i \geq f)\right\} \leq \sum_{i \in S_n} \mathbb{P}(X_n^i \geq f) = |S_n| \cdot \mathbb{P}(X_n^1 \geq f), \quad (\text{A.25})$$

Taking logarithms and limits of the right-hand side of (A.25), using (A.18) and (A.10), we have

$$\lim_{n \rightarrow \infty} \frac{1}{n} \ln [|S_n| \cdot \mathbb{P}(X_n^1 \geq f)] = H(\delta\rho) - \frac{1}{2} [\ln(1+f) - \rho \ln f - H(\rho)],$$

which combines with (A.25) to imply that, for any $\eta > 0$,

$$\frac{1}{n} \ln \mathbb{P}\{\cup_{i \in S_n} (X_n^i \geq f)\} \leq H(\delta\rho) - \frac{1}{2} [\ln(1+f) - \rho \ln f - H(\rho)] + \eta, \quad (\text{A.26})$$

for all n sufficiently large. By the definition of $\mathcal{IF}(\delta, \rho)$ in (4.39), and since the left-hand side of (4.39) on $f > \frac{\rho}{1-\rho}$ is strictly increasing in f , then, for any $\epsilon > 0$, setting $f := f^* = \mathcal{IF}(\delta, \rho) + \epsilon$ and choosing η sufficiently small in (A.26) ensures

$$\frac{1}{n} \ln \mathbb{P}\{\cup_{i \in S_n} (X_n^i \geq f^*)\} \leq -c_I \quad \text{for all } n \text{ sufficiently large,}$$

where c_I is some positive constant, from which it follows that

$$\mathbb{P}\{\cup_{i \in S_n} (X_n^i \geq f^*)\} \leq e^{-c_I \cdot n} \quad \text{for all } n \text{ sufficiently large,}$$

and (4.43) now follows. \square

Appendix B

Proof of Theorem 5.7

We first introduce a set partition definition.

Definition B.1 (Support set partition for N-IHT). Suppose $\delta \in (0, 1]$ and $\rho \in (0, 1/2]$. Given $\zeta > 0$, let us write

$$a^*(\delta, \rho; \zeta) := a(\delta, \rho) + \zeta, \quad (\text{B.1})$$

where $a(\delta, \rho)$ is defined in (5.39), let us write $\{\Gamma_i : i \in S_n\}$ for the set of all possible support sets of cardinality k , and let us disjointly partition $S_n \stackrel{\text{def}}{=} \Theta_n^1 \cup \Theta_n^2$ such that

$$\Theta_n^1 \stackrel{\text{def}}{=} \left\{i \in S_n : \|x_{\Lambda \setminus \Gamma_i}^*\| > \sigma \cdot a^*(\delta, \rho; \zeta)\right\}; \quad \Theta_n^2 \stackrel{\text{def}}{=} \left\{i \in S_n : \|x_{\Lambda \setminus \Gamma_i}^*\| \leq \sigma \cdot a^*(\delta, \rho; \zeta)\right\}. \quad (\text{B.2})$$

The proof of Theorem 5.7 for N-IHT takes broadly the same approach as for the corresponding result for IHT in Section 5.2. However, in order to finally eliminate the dependence upon ϵ in $\underline{\alpha}(\delta, \rho; \epsilon)$, the results corresponding to Lemmas 5.1 and 5.4 for IHT need to be combined together. This is accomplished by Lemma B.1, which establishes that, provided (5.38) holds and ϵ is taken sufficiently small, N-IHT converges to an $\underline{\alpha}(\delta, \rho; \epsilon)$ -stable point on some Γ_i such that $i \in \Theta_n^2$. Lemma B.2 corresponds to Lemma 5.3 for IHT, giving bounds on the approximation error of an $\underline{\alpha}(\delta, \rho; \epsilon)$ -stable point on some Γ_i such that $i \in \Theta_n^2$, for any $\epsilon > 0$. Combining the two lemmas leads us to conclude that N-IHT converges to some limit point with bounded approximation error. We write $NSP_{\underline{\alpha}}$ for the event that there is no $\underline{\alpha}(\delta, \rho; \epsilon)$ -stable point on any Γ_i such that $i \in \Theta_n^1$.

Lemma B.1. Choose $\zeta > 0$. Suppose Assumptions A.2 and A.3 hold, and suppose that (5.38) holds. Then there exists ϵ such that, in the proportional-growth asymptotic, N-IHT converges to an $\underline{\alpha}(\delta, \rho; \epsilon)$ -stable point on some Γ_i such that $i \in \Theta_n^2$, with probability tending to 1 exponentially in n .

Proof. Under Assumption **A.2**, we have by Theorem 3.7 convergence of N-IHT to a $[\kappa(1 + U_{2k})]^{-1}$ -stable point. By Definition 3.1, for any $\alpha_1 < \alpha_2$, the set of α_1 -stable points includes the set of α_2 -stable points, and this observation combines with Lemma 4.6 to imply convergence to an $\underline{\alpha}(\delta, \rho; \epsilon)$ -stable point, where $\underline{\alpha}(\delta, \rho; \epsilon)$ is defined in (5.37), with probability tending to 1 exponentially in n . We now rehearse the argument of Lemma 5.1 to show that, provided ϵ is taken sufficiently small, this stable point must be on Γ_i such that $i \in \Theta_n^2$. For any Γ_i such that $i \in \Theta_n^1$, we have $\Gamma_i \neq \Lambda$, and we may therefore use Theorem 3.2 and Lemma 4.3 with $\Gamma := \Gamma_i$ to deduce that, given some $\epsilon > 0$, a necessary condition for there to be an $\underline{\alpha}(\delta, \rho; \epsilon)$ -stable point on Γ_i is

$$\begin{aligned} & \|x_{\Lambda \setminus \Gamma_i}^*\| \cdot \sqrt{F_{\Gamma_i}} + \sigma \cdot \sqrt{G_{\Gamma_i}} \\ & \geq \underline{\alpha}(\delta, \rho; \epsilon) \left[\left(\frac{n-k}{n} \right) \|x_{\Lambda \setminus \Gamma_i}^*\| \cdot R_{\Gamma_i} - \sigma \cdot \sqrt{\frac{k(n-k)}{n^2} \cdot S_{\Gamma_i} \cdot T_{\Gamma_i}} \right], \end{aligned} \quad (\text{B.3})$$

where

$$\begin{aligned} F_{\Gamma_i} & \sim \frac{k}{n-k+1} F(k, n-k+1); & G_{\Gamma_i} & \sim \frac{k}{n-k+1} F(k, n-k+1); \\ R_{\Gamma_i} & \sim \frac{1}{n-k} \chi_{n-k}^2; & S_{\Gamma_i} & \sim \frac{1}{n-k} \chi_{n-k}^2; & T_{\Gamma_i} & \sim \frac{1}{k} \chi_k^2. \end{aligned}$$

We also have, by (B.2),

$$\sigma \leq \frac{\|x_{\Lambda \setminus \Gamma_i}^*\|}{a^*(\delta, \rho; \zeta)} \quad (\text{B.4})$$

for any Γ_i such that $i \in \Theta_n^1$. Since $\Gamma_i \neq \Lambda$, $\|x_{\Lambda \setminus \Gamma_i}^*\| > 0$, and substitution of (B.4) into (B.3), rearrangement and division by $\|x_{\Lambda \setminus \Gamma_i}^*\|$ yields

$$a^*(\delta, \rho; \zeta) \left[\underline{\alpha}(\delta, \rho; \epsilon) \left(\frac{n-k}{n} \right) \cdot R_{\Gamma_i} - \sqrt{F_{\Gamma_i}} \right] \leq \sqrt{G_{\Gamma_i}} + \underline{\alpha}(\delta, \rho; \epsilon) \sqrt{\frac{k(n-k)}{n^2} \cdot S_{\Gamma_i} \cdot T_{\Gamma_i}},$$

and consequently

$$\begin{aligned} \mathbb{P}(\overline{NSP_{\underline{\alpha}}}) &= \mathbb{P}\{\cup_{i \in \Theta_n^1} (\exists \text{ an } \underline{\alpha}(\delta, \rho; \epsilon)\text{-stable point supported on } \Gamma_i)\} \\ &\leq \mathbb{P}\{\cup_{i \in \Theta_n^1} (\Phi[\rho_n, F_{\Gamma_i}, G_{\Gamma_i}, R_{\Gamma_i}, S_{\Gamma_i}, T_{\Gamma_i}] \geq 0)\}, \end{aligned} \quad (\text{B.5})$$

where we write ρ_n for the sequence of values of the ratio k/n , and where

$$\Phi[\rho, F, G, R, S, T] \stackrel{\text{def}}{=} \sqrt{G} + \underline{\alpha}(\delta, \rho; \epsilon) \sqrt{\rho(1-\rho)(S)(T)} - a^*(\delta, \rho; \zeta) \cdot \left[\underline{\alpha}(\delta, \rho; \epsilon)(1-\rho) \cdot R - \sqrt{F} \right]. \quad (\text{B.6})$$

We now define

$$\begin{aligned} F^* &= G^* \stackrel{\text{def}}{=} \mathcal{IF}(\delta, \rho) + \epsilon; & R^* &\stackrel{\text{def}}{=} 1 - \mathcal{IL}(\delta, \rho, 1-\rho) - \epsilon; \\ S^* &\stackrel{\text{def}}{=} 1 + \mathcal{IU}(\delta, \rho, 1-\rho) + \epsilon; & T^* &\stackrel{\text{def}}{=} 1 + \mathcal{IU}(\delta, \rho, \rho) + \epsilon. \end{aligned} \quad (\text{B.7})$$

Using (B.7), we deduce from (B.5) that

$$\begin{aligned} & \mathbb{P}(\overline{NSP_{\underline{\alpha}}}) \\ & \leq \mathbb{P}\{\cup_{i \in \Theta_n^1} (\Phi[\rho_n, F_{\Gamma_i}, G_{\Gamma_i}, R_{\Gamma_i}, S_{\Gamma_i}, T_{\Gamma_i}] \geq \Phi[\rho_n, F^*, G^*, R^*, S^*, T^*])\} \end{aligned} \quad (\text{B.8})$$

$$+ \mathbb{P}\{\Phi[\rho_n, F^*, G^*, R^*, S^*, T^*] \geq \Phi[\rho, F^*, G^*, R^*, S^*, T^*] + \epsilon\} \quad (\text{B.9})$$

$$+ \mathbb{P}\{\Phi[\rho, F^*, G^*, R^*, S^*, T^*] + \epsilon \geq 0\}, \quad (\text{B.10})$$

since the event in (B.5) lies in the union of the three events in (B.8), (B.9) and (B.10). Now (B.10) is a deterministic event, and $a^*(\delta, \rho; \zeta)$ has been defined in such a way that, for any $\zeta > 0$, provided ϵ is taken sufficiently small, the event has probability 0. This follows from (5.38), (5.39), (B.1), and by the continuity of Φ . The event (B.9) is also deterministic, and by continuity and since $\rho_n \rightarrow \rho$, it follows that there exists some \tilde{n} such that

$$\mathbb{P}\{\Phi[\rho_n, F^*, G^*, R^*, S^*, T^*] \geq \Phi[\rho, F^*, G^*, R^*, S^*, T^*] + \epsilon\} = 0 \quad \text{for all } n \geq \tilde{n}.$$

Taking limits as $n \rightarrow \infty$, the terms (B.9) and (B.10) are zero, leaving only (B.8), and we have

$$\begin{aligned}
& \lim_{n \rightarrow \infty} \mathbb{P}(\overline{NSP_\alpha}) \\
& \leq \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^1} (\Phi[\rho_n, F_{\Gamma_i}, G_{\Gamma_i}, R_{\Gamma_i}, S_{\Gamma_i}, T_{\Gamma_i}] \geq \Phi[\rho_n, F^*, G^*, R^*, S^*, T^*])\} \\
& \leq \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^1} (F_{\Gamma_i} \geq F^*)\} + \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^1} (G_{\Gamma_i} \geq G^*)\} + \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^1} (R_{\Gamma_i} \leq R^*)\} \\
& + \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^1} (S_{\Gamma_i} \geq S^*)\} + \lim_{n \rightarrow \infty} \mathbb{P}\{\cup_{i \in \Theta_n^1} (T_{\Gamma_i} \geq T^*)\}, \tag{B.11}
\end{aligned}$$

where the last line follows from the monotonicity of Φ with respect to F, G, R, S and T . Since $\Theta_n^1 \subseteq S_n$, we may apply Lemmas 4.4 and 4.5 to (B.11), and we deduce $\mathbb{P}(\overline{NSP_\alpha}) \rightarrow 0$ as $n \rightarrow \infty$, exponentially in n , as required. \square

Lemma B.2. *Suppose Assumptions A.2 and A.3 hold, and suppose that (5.38) holds. Given any $\epsilon > 0$, there exists ζ sufficiently small such that, in the proportional-growth asymptotic, any $\underline{\alpha}(\delta, \rho; \epsilon)$ -stable point on Γ_i such that $i \in \Theta_n^2$ satisfies (5.41), with probability tending to 1 exponentially in n .*

Proof. Suppose \bar{x} is a minimum-norm solution on Γ , so that $\bar{x}_\Gamma = A_\Gamma^\dagger b$. Then we may follow the argument of Lemma 5.3 to deduce (5.24), where

$$P_\Gamma \sim \frac{k}{n-k+1} F(k, n-k+1); \quad Q_\Gamma \sim \frac{k}{n-k+1} F(k, n-k+1). \tag{B.12}$$

Combining (5.24) with (B.2), we may further deduce

$$\begin{aligned}
\|\bar{x} - x^*\|^2 & \leq \sigma^2 \left[a^*(\delta, \rho; \zeta) \cdot \sqrt{P_\Gamma} + \sqrt{Q_\Gamma} \right]^2 + [a^*(\delta, \rho; \zeta)]^2 \cdot \sigma^2 \\
& = \sigma^2 \left\{ \left[a^*(\delta, \rho; \zeta) \cdot \sqrt{P_\Gamma} + \sqrt{Q_\Gamma} \right]^2 + [a^*(\delta, \rho; \zeta)]^2 \right\}. \tag{B.13}
\end{aligned}$$

For the sake of brevity, let us define

$$\Psi[P, Q] \stackrel{\text{def}}{=} \sqrt{\left(a^*(\delta, \rho; \zeta) \cdot \sqrt{P} + \sqrt{Q} \right)^2 + a^*(\delta, \rho; \zeta)^2}, \tag{B.14}$$

so that (B.13) may equivalently be written as

$$\|\bar{x} - x^*\| \leq \sigma \cdot \Psi[P_\Gamma, Q_\Gamma]. \tag{B.15}$$

First suppose that $\sigma > 0$. Given $\zeta > 0$, let us define

$$P^* = Q^* \stackrel{\text{def}}{=} \mathcal{IF}(\delta, \rho) + \zeta. \tag{B.16}$$

Now we use (B.15) to perform a union bound over all Γ_i such that $i \in \Theta_n^2$, writing \bar{x}_i for the minimum-norm solution on Γ_i , giving

$$\begin{aligned}
& \mathbb{P}\{\exists \text{ some } \Gamma_i \text{ such that } i \in \Theta_n^2 \text{ and } \|\bar{x}_i - x^*\| > \sigma \cdot \Psi[P^*, Q^*]\} \\
& = \mathbb{P}\left\{ \bigcup_{i \in \Theta_n^2} (\|\bar{x}_i - x^*\| > \sigma \cdot \Psi[P^*, Q^*]) \right\} \tag{B.17}
\end{aligned}$$

$$\leq \mathbb{P}\left\{ \bigcup_{i \in \Theta_n^2} (\|\bar{x}_i - x^*\| > \sigma \cdot \Psi[P_{\Gamma_i}, Q_{\Gamma_i}]) \right\} \tag{B.18}$$

$$\begin{aligned}
& + \mathbb{P}\left\{ \bigcup_{i \in \Theta_n^2} (\sigma \cdot \Psi[P_{\Gamma_i}, Q_{\Gamma_i}] \geq \sigma \cdot \Psi[P^*, Q^*]) \right\}, \tag{B.19}
\end{aligned}$$

since the event in (B.17) lies in the union of the two events in (B.18) and (B.19). It is an immediate consequence of (B.15) that the event in (B.18) has probability 0. Taking limits of (B.19) as $n \rightarrow \infty$, and cancelling σ , we have

$$\begin{aligned} & \lim_{n \rightarrow \infty} \mathbb{P} \{ \exists \text{ some } \Gamma_i \text{ such that } i \in \Theta_n^2 \text{ and } \|\bar{x}_i - x^*\| > \sigma \cdot \Psi[P^*, Q^*] \} \\ & \leq \lim_{n \rightarrow \infty} \mathbb{P} \left\{ \bigcup_{i \in \Theta_n^2} (\Psi[P_{\Gamma_i}, Q_{\Gamma_i}] \geq \Psi[P^*, Q^*]) \right\} \\ & \leq \lim_{n \rightarrow \infty} \mathbb{P} \{ \cup_{i \in \Theta_n^2} (P_{\Gamma_i} \geq P^*) \} + \lim_{n \rightarrow \infty} \mathbb{P} \{ \cup_{i \in \Theta_n^2} (Q_{\Gamma_i} \geq Q^*) \}, \end{aligned} \quad (\text{B.20})$$

where we used the monotonicity of Ψ with respect to P and Q in the last line. Since $\Theta_n^2 \subseteq S_n$, and using (B.12), we may apply Lemma 4.5 to (B.20), yielding that each of the limits in the right-hand side of (B.20) converges to zero exponentially in n , and so finally

$$\lim_{n \rightarrow \infty} \mathbb{P} \{ \exists \text{ some } \Gamma_i \text{ such that } i \in \Theta_n^2 \text{ and } \|\bar{x}_i - x^*\| > \sigma \cdot \Psi[P^*, Q^*] \} = 0,$$

with convergence at a rate exponential in n also by Lemma 4.5. The same result also holds when $\sigma = 0$ by (B.13). Since by Lemma 3.1, any stable point is necessarily a minimum-norm solution, and recalling the definition of $\Psi(P, Q)$ in (5.26), and the definitions of P^*, Q^* in (B.16), we have

$$\lim_{n \rightarrow \infty} \mathbb{P} \left\{ \begin{array}{l} \exists \text{ some } \underline{\alpha}\text{-stable point } \bar{x}_i \text{ on } \Gamma_i \text{ such that } i \in \Theta_n^2 \text{ and} \\ \|\bar{x}_i - x^*\| > \sigma \sqrt{\mathcal{IF}(\delta, \rho) [1 + a(\delta, \rho) + \zeta]^2 + [a(\delta, \rho) + \zeta]^2} \end{array} \right\} = 0, \quad (\text{B.21})$$

with convergence exponential in n . Finally, by continuity,

$$\begin{aligned} \|\bar{x}_i - x^*\| & > \sigma \sqrt{\mathcal{IF}(\delta, \rho) [1 + a(\delta, \rho)]^2 + 1 + [a(\delta, \rho)]^2} \\ \implies \|\bar{x}_i - x^*\| & > \sigma \sqrt{\mathcal{IF}(\delta, \rho) [1 + a(\delta, \rho) + \zeta]^2 + [a(\delta, \rho) + \zeta]^2}, \end{aligned}$$

for some ζ suitably small, and the result now follows from the definition of $\xi(\delta, \rho)$ in (5.40). \square

Proof of Theorem 5.7: By Lemma B.1, there exists $\epsilon > 0$ such that N-IHT converges to an $\underline{\alpha}(\delta, \rho; \epsilon)$ -stable point on some Γ_i such that $i \in \Theta_n^2$, and for this choice of ϵ , we can apply Lemma B.2 to deduce the result.